# Package: mirt (via r-universe)

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**Type** Package

**Title** Multidimensional Item Response Theory

**Description** Analysis of discrete response data using unidimensional and multidimensional item analysis models under the Item Response Theory paradigm (Chalmers (2012) <doi:10.18637/jss.v048.i06>). Exploratory and confirmatory item factor analysis models are estimated with quadrature (EM) or stochastic (MHRM) methods. Confirmatory bi-factor and two-tier models are available for modeling item testlets using dimension reduction EM algorithms, while multiple group analyses and mixed effects designs are included for detecting differential item, bundle, and test functioning, and for modeling item and person covariates. Finally, latent class models such as the DINA, DINO, multidimensional latent class, mixture IRT models, and zero-inflated response models are supported, as well as a wide family of probabilistic unfolding models.

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**Depends** stats, R (>= 4.1.0), stats4, lattice, methods

**Imports** GPArotation, gridExtra, Matrix (>= 1.5-0), Rcpp, mgcv, vegan, Deriv, splines, splines2, pbapply (>= 1.3-0), dcurver, SimDesign

**Suggests** boot, parallel, mirai, latticeExtra, directlabels, shiny, knitr, markdown, Rsolnp, nloptr, sirt, plink, mirtCAT, testthat (>= 3.0.0)

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**LazyLoad** yes

**LazyData** yes

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**License** GPL (>= 3)

**Encoding** UTF-8

**Maintainer** Phil Chalmers <rphilip.chalmers@gmail.com>

**URL** https://philchalmers.github.io/mirt/,
https://github.com/philchalmers/mirt/wiki,
https://groups.google.com/forum/#!forum/mirt-package

**BugReports** https://github.com/philchalmers/mirt/issues?state=open

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**Author** Phil Chalmers [aut, cre] (ORCID:
<https://orcid.org/0000-0001-5332-2810>), Joshua Pritikin
[ctb], Alexander Robitzsch [ctb], Mateusz Zoltak [ctb],
KwonHyun Kim [ctb], Carl F. Falk [ctb], Adam Meade [ctb],
Lennart Schneider [ctb], David King [ctb], Chen-Wei Liu [ctb],
Ogreden Oguzhan [ctb], Samuel Huang [ctb]

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`mirt-package` *Full information maximum likelihood estimation of IRT models.*

---

## Description

Full information maximum likelihood estimation of multidimensional IRT models

## Details

Analysis of dichotomous and polytomous response data using unidimensional and multidimensional latent trait models under the Item Response Theory (IRT) paradigm. Exploratory and confirmatory models can be estimated with quadrature (EM) or stochastic (MHRM) methods. Confirmatory bi-factor and two-tier analyses are available for modeling item testlets. Multiple group analysis and mixed effects designs also are available for detecting differential item and test functioning as well as modeling item and person covariates. Finally, latent class models such as the DINA, DINO, multidimensional latent class, mixture and zero-inflated IRT models, and several other discrete variable models are supported.

Users interested in the most recent version of this package can visit `https://github.com/philchalmers/mirt` and follow the instructions for installing the package from source. Questions regarding the

package can be sent to the mirt-package Google Group, located at https://groups.google.com/forum/#!forum/mirt-package. User contributed files, workshop files, and evaluated help files are also available on the package wiki (https://github.com/philchalmers/mirt/wiki).

## Author(s)

Phil Chalmers <rphilip.chalmers@gmail.com>

## References

Chalmers, R. P. (2012). mirt: A Multidimensional Item Response Theory Package for the R Environment. *Journal of Statistical Software, 48*(6), 1-29. doi:10.18637/jss.v048.i06

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# anova-method *Compare nested models with likelihood-based statistics*

---

## Description

Compare nested models using likelihood ratio test (X2), Akaike Information Criterion (AIC), Bayesian Information Criterion (BIC), Sample-Size Adjusted BIC (SABIC), and Hannan-Quinn (HQ) Criterion. When given a sequence of objects, `anova` tests the models against one another in the order specified. Note that the `object` inputs should be ordered in terms of most constrained model to least constrained.

## Usage

```
## S4 method for signature 'SingleGroupClass'
anova(
  object,
  object2,
  ...,
  bounded = FALSE,
  mix = 0.5,
  frame = 1,
  verbose = FALSE
)
```

## Arguments

| | |
|---|---|
| `object` | an object of class `SingleGroupClass`, `MultipleGroupClass`, or `MixedClass`, reflecting the most constrained model fitted |
| `object2` | a second model estimated from any of the mirt package estimation methods |
| `...` | additional less constrained model objects to be compared sequentially to the previous model |
| `bounded` | logical; are the two models comparing a bounded parameter (e.g., comparing a single 2PL and 3PL model with 1 df)? If `TRUE` then a 50:50 mix of chi-squared distributions is used to obtain the p-value |

| | |
|---|---|
| mix | proportion of chi-squared mixtures. Default is 0.5 |
| frame | (internal parameter not for standard use) |
| verbose | (deprecated argument) |

## Value

a data.frame/mirt_df object

## References

Chalmers, R. P. (2012). mirt: A Multidimensional Item Response Theory Package for the R Environment. *Journal of Statistical Software, 48*(6), 1-29. doi:10.18637/jss.v048.i06

## Examples

```
x <- mirt(Science, 1)
x2 <- mirt(Science, 2)
anova(x, x2)

# compare three models sequentially (X2 not always meaningful)
x3 <- mirt(Science, 1, 'gpcm')
x4 <- mirt(Science, 1, 'nominal')
anova(x, x2, x3, x4)

# in isolation
anova(x)

# with priors on first model
model <- "Theta = 1-4
          PRIOR = (1-4, a1, lnorm, 0, 10)"
xp <- mirt(Science, model)
anova(xp, x2)
anova(xp)

# bounded parameter
dat <- expand.table(LSAT7)
mod <- mirt(dat, 1)
mod2 <- mirt(dat, 1, itemtype = c(rep('2PL', 4), '3PL'))
anova(mod, mod2) #unbounded test
anova(mod, mod2, bounded = TRUE) #bounded

# priors
model <- 'F = 1-5
          PRIOR = (5, g, norm, -1, 1)'
mod1b <- mirt(dat, model, itemtype = c(rep('2PL', 4), '3PL'))
anova(mod1b)

model2 <- 'F = 1-5
          PRIOR = (1-5, g, norm, -1, 1)'
mod2b <- mirt(dat, model2, itemtype = '3PL')
anova(mod1b, mod2b)
```

areainfo *Function to calculate the area under a selection of information curves*

## Description

Compute the area of a test or item information function over a definite integral range.

## Usage

```
areainfo(
  x,
  theta_lim,
  which.items = 1:extract.mirt(x, "nitems"),
  group = NULL,
  ...
)
```

## Arguments

| | |
|---|---|
| x | an object of class 'SingleGroupClass', or an object of class 'MultipleGroupClass' if a suitable group input were supplied |
| theta_lim | range of integration to be computed |
| which.items | an integer vector indicating which items to include in the expected information function. Default uses all possible items |
| group | group argument to pass to extract.group function. Required when the input object is a multiple-group model |
| ... | additional arguments passed to integrate |

## Value

a data.frame with the lower and upper integration range, the information area within the range (Info), the information area over the range -10 to 10 (Total.Info), proportion of total information given the integration range (Info.Proportion), and the number of items included (nitems)

## Author(s)

Phil Chalmers <rphilip.chalmers@gmail.com>

## References

Chalmers, R. P. (2012). mirt: A Multidimensional Item Response Theory Package for the R Environment. *Journal of Statistical Software, 48*(6), 1-29. doi:10.18637/jss.v048.i06

## Examples

```
dat <- expand.table(LSAT7)
mod <- mirt(dat, 1)

areainfo(mod, c(-2,0), which.items = 1) #item 1

areainfo(mod, c(-2,0), which.items = 1:3) #items 1 to 3
areainfo(mod, c(-2,0)) # all items (total test information)

# plot the area
area <- areainfo(mod, c(-2,0))
Theta <- matrix(seq(-3,3, length.out=1000))
info <- testinfo(mod, Theta)
plot(info ~ Theta, type = 'l')

pick <- Theta >= -2 & Theta <=0
polygon(c(-2, Theta[pick], 0), c(0, info[pick], 0), col='lightblue')
text(x = 2, y = 0.5, labels = paste("Total Information:", round(area$TotalInfo, 3),
           "\n\nInformation in (-2, 0):", round(area$Info, 3),
           paste("(", round(100 * area$Proportion, 2), "%)", sep = "")), cex = 1.2)
```

---

# ASVAB *Description of ASVAB data*

---

## Description

Table of counts extracted from Mislvey (1985). Data the 16 possible response patterns observed for four items from the arithmetic reasoning test of the Armed Services Vocational Aptitude Battery (ASVAB), Form 8A, from samples of white males and females and black males and females.

## Author(s)

Phil Chalmers <rphilip.chalmers@gmail.com>

## References

Mislevy, R. J. (1985). Estimation of latent group effects. *Journal of the American Statistical Association, 80*, 993-997.

## Examples

```
data(ASVAB)
datWM <- expand.table(subset(ASVAB, select=c(Item.1:Item.4, White_Male)))
datWF <- expand.table(subset(ASVAB, select=c(Item.1:Item.4, White_Female)))
datBM <- expand.table(subset(ASVAB, select=c(Item.1:Item.4, Black_Male)))
datBF <- expand.table(subset(ASVAB, select=c(Item.1:Item.4, Black_Female)))
```

```
dat <- rbind(datWM, datWF, datBM, datBF)
sex <- rep(c("Male", "Female", "Male", "Female"),
  times=c(nrow(datWM), nrow(datWF), nrow(datBM), nrow(datBF))) |> factor()
color <- rep(c("White", "Black"),
  times=c(nrow(datWM) + nrow(datWF), nrow(datBM) + nrow(datBF))) |> factor()
group <- sex:color

itemstats(dat, group=group)
```

---

`Attitude` *Description of Attitude data*

---

## Description

Table of counts extracted from Andrich (1988). Data the response patterns observed for an eight item survey.

## Details

The items in this survey were:

1. Capital punishment is one of the most hideous practices of our time.
2. The state cannot teach the sacredness of human life by destroying it.
3. Capital punishment is not an effective deterrent to crime.
4. I don't believe in capital punishment but I am not sure it isn't necessary.
5. I think capital punishment is necessary but I wish it were not.
6. Until we find a more civilized way to prevent crime we must have capital punishment.
7. Capital punishment is justified because it does act as a deterrent to crime.
8. Capital punishment gives the criminal what he deserves.

## Author(s)

Phil Chalmers <rphilip.chalmers@gmail.com>

## References

Andrich, D. (1988). The Application of an Unfolding Model of the PIRT Type to the Measurement of Attitude. *Applied Psychological Measurement, 12*, 33-51.

## Examples

```
head(Attitude)
df <- expand.table(Attitude)
itemstats(df)



# estimate SSLM with estimated " latitude of acceptance" (rho)
mod.rho <- mirt(df, 1, itemtype = 'sslm')
coef(mod.rho)
coef(mod.rho, simplify=TRUE)  # slope-intercept-log_rho
coef(mod.rho, simplify=TRUE, IRTpars=TRUE)  # discrimination-difficulty-rho
plot(mod.rho)
plot(mod.rho, type = 'trace')

# without estimating rho, and fixing to rho^2 = 1  (hence,
#   log_rho = -exp(1) = -2.718282 in order to obtain (exp(exp(log_rho))) = 1)
syntax <- "Theta = 1-8
           FIXED = (1-8, log_rho1)
           START = (1-8, log_rho1, -2.71828)"
mod <- mirt(df, syntax, itemtype = 'sslm')  # model found in Andrich (1988)
coef(mod)
coef(mod, simplify=TRUE)  # slope-intercept-log_rho
coef(mod, simplify=TRUE, IRTpars=TRUE)  # discrimination-difficulty-rho
plot(mod)
plot(mod, type = 'trace') # notice that all curves have a fixed height of .5

# goodness of fit (less constrained model fits better)
anova(mod, mod.rho) # original model fits much worse
M2(mod)
M2(mod.rho)
itemfit(mod, p.adjust='fdr')
itemfit(mod.rho, p.adjust='fdr')
```

---

`averageMI` *Collapse values from multiple imputation draws*

---

## Description

This function computes updated parameter and standard error estimates using multiple imputation methodology. Given a set of parameter estimates and their associated standard errors the function returns the weighted average of the overall between and within variability due to the multiple imputations according to Rubin's (1987) methodology.

## Usage

```
averageMI(par, SEpar, as.data.frame = TRUE)
```

## Arguments

| | |
|---|---|
| par | a list containing parameter estimates which were computed the imputed datasets |
| SEpar | a list containing standard errors associated with par |
| as.data.frame | logical; return a data.frame instead of a list? Default is TRUE |

## Value

returns a list or data.frame containing the updated averaged parameter estimates, standard errors, and t-values with the associated degrees of freedom and two tailed p-values

## Author(s)

Phil Chalmers <rphilip.chalmers@gmail.com>

## References

Chalmers, R. P. (2012). mirt: A Multidimensional Item Response Theory Package for the R Environment. *Journal of Statistical Software, 48*(6), 1-29. doi:10.18637/jss.v048.i06

Rubin, D.B. (1987) Multiple Imputation for Nonresponse in Surveys. Wiley & Sons, New York.

## Examples

```
# simulate data
set.seed(1234)
N <- 1000

# covariates
X1 <- rnorm(N); X2 <- rnorm(N)
covdata <- data.frame(X1, X2)
Theta <- matrix(0.5 * X1 + -1 * X2 + rnorm(N, sd = 0.5))

# items and response data
a <- matrix(1, 20); d <- matrix(rnorm(20))
dat <- simdata(a, d, 1000, itemtype = '2PL', Theta=Theta)

mod1 <- mirt(dat, 1, 'Rasch', covdata=covdata, formula = ~ X1 + X2)
coef(mod1, simplify=TRUE)

# draw plausible values for secondary analyses
pv <- fscores(mod1, plausible.draws = 10)
pvmods <- lapply(pv, function(x, covdata) lm(x ~ covdata$X1 + covdata$X2),
                 covdata=covdata)

# compute Rubin's multiple imputation average
so <- lapply(pvmods, summary)
par <- lapply(so, function(x) x$coefficients[, 'Estimate'])
SEpar <- lapply(so, function(x) x$coefficients[, 'Std. Error'])
averageMI(par, SEpar)
```

bfactor *Full-Information Item Bi-factor and Two-Tier Analysis*

## Description

bfactor fits a confirmatory maximum likelihood two-tier/bifactor/testlet model to dichotomous and polytomous data under the item response theory paradigm. The IRT models are fit using a dimensional reduction EM algorithm so that regardless of the number of specific factors estimated the model only uses the number of factors in the second-tier structure plus 1. For the bifactor model the maximum number of dimensions is only 2 since the second-tier only consists of a ubiquitous unidimensional factor. See mirt for appropriate methods to be used on the objects returned from the estimation.

## Usage

```
bfactor(
  data,
  model,
  model2 = paste0("G = 1-", ncol(data)),
  group = NULL,
  quadpts = NULL,
  invariance = "",
  ...
)
```

## Arguments

| | |
|---|---|
| data | a matrix or data.frame that consists of numerically ordered data, organized in the form of integers, with missing data coded as NA |
| model | a numeric vector specifying which factor loads on which item. For example, if for a 4 item test with two specific factors, the first specific factor loads on the first two items and the second specific factor on the last two, then the vector is c(1,1,2,2). For items that should only load on the second-tier factors (have no specific component) NA values may be used as place-holders. These numbers will be translated into a format suitable for mirt.model(), combined with the definition in model2, with the letter 'S' added to the respective factor number<br>Alternatively, input can be specified using the mirt.model syntax with the restriction that each item must load on exactly one specific factor (or no specific factors, if it is only predicted by the general factor specified in model2)<br>IMPORTANT: Additional model information (e.g., keywords such as CONSTRAIN, MEAN, etc) must be specified in the model2 input instead of this specification |
| model2 | a two-tier model specification object defined by mirt.model() or a string to be passed to mirt.model. By default the model will fit a unidimensional model in the second-tier, and therefore be equivalent to the bifactor model |

| | |
|---|---|
| group | a factor variable indicating group membership used for multiple group analyses |
| quadpts | number of quadrature nodes to use after accounting for the reduced number of dimensions. Scheme is the same as the one used in mirt, however it is in regards to the reduced dimensions (e.g., a bifactor model has 2 dimensions to be integrated) |
| invariance | see multipleGroup for details, however, the specific factor variances and means will be constrained according to the dimensional reduction algorithm |
| ... | additional arguments to be passed to the estimation engine. See mirt for more details and examples |

## Details

bfactor follows the item factor analysis strategy explicated by Gibbons and Hedeker (1992), Gibbons et al. (2007), and Cai (2010). Nested models may be compared via an approximate chi-squared difference test or by a reduction in AIC or BIC (accessible via anova). See mirt for more details regarding the IRT estimation approach used in this package.

The two-tier model has a specific block diagonal covariance structure between the primary and secondary latent traits. Namely, the secondary latent traits are assumed to be orthogonal to all traits and have a fixed variance of 1, while the primary traits can be organized to vary and covary with other primary traits in the model.

$$\Sigma_{two-tier} = \begin{pmatrix} G & 0 \\ 0 & diag(S) \end{pmatrix}$$

The bifactor model is a special case of the two-tier model when $G$ above is a 1x1 matrix, and therefore only 1 primary factor is being modeled. Evaluation of the numerical integrals for the two-tier model requires only $ncol(G) + 1$ dimensions for integration since the $S$ second order (or 'specific') factors require only 1 integration grid due to the dimension reduction technique.

Note: for multiple group two-tier analyses only the second-tier means and variances should be freed since the specific factors are not treated independently due to the dimension reduction technique.

## Value

function returns an object of class SingleGroupClass (SingleGroupClass-class) or MultipleGroupClass(MultipleGroupCla class).

## Author(s)

Phil Chalmers <rphilip.chalmers@gmail.com>

## References

Cai, L. (2010). A two-tier full-information item factor analysis model with applications. *Psychometrika, 75*, 581-612.

Chalmers, R. P. (2012). mirt: A Multidimensional Item Response Theory Package for the R Environment. *Journal of Statistical Software, 48*(6), 1-29. doi:10.18637/jss.v048.i06

Bradlow, E.T., Wainer, H., & Wang, X. (1999). A Bayesian random effects model for testlets. *Psychometrika, 64*, 153-168.

Gibbons, R. D., & Hedeker, D. R. (1992). Full-information Item Bi-Factor Analysis. *Psychometrika, 57*, 423-436.

Gibbons, R. D., Darrell, R. B., Hedeker, D., Weiss, D. J., Segawa, E., Bhaumik, D. K., Kupfer, D. J., Frank, E., Grochocinski, V. J., & Stover, A. (2007). Full-Information item bifactor analysis of graded response data. *Applied Psychological Measurement, 31*, 4-19.

Wainer, H., Bradlow, E.T., & Wang, X. (2007). Testlet response theory and its applications. New York, NY: Cambridge University Press.

## See Also

mirt

## Examples

```
### load SAT12 and compute bifactor model with 3 specific factors
data(SAT12)
data <- key2binary(SAT12,
  key = c(1,4,5,2,3,1,2,1,3,1,2,4,2,1,5,3,4,4,1,4,3,3,4,1,3,5,1,3,1,5,4,5))
specific <- c(2,3,2,3,3,2,1,2,1,1,1,3,1,3,1,2,1,1,3,3,1,1,3,1,3,3,1,3,2,3,1,2)
mod1 <- bfactor(data, specific)
summary(mod1)
itemplot(mod1, 18, drop.zeros = TRUE) #drop the zero slopes to allow plotting

# complete factor score predictions (general + specific factors)
eaps <- fscores(mod1)

# factor score predictions for general factors only (more accurate due
# to lower dimensional integration)
eaps_gen <- fscores(mod1, method = 'EAP_general')
head(cbind(eaps_gen, NA, eaps))

# similar EAP estimates, but with respect to sum-scores
fscores(mod1, method = 'EAPsum_2.0', full.scores=FALSE)

# alternative model definition via ?mirt.model syntax
specific2 <- "S1 = 7,9,10,11,13,15,17,18,21,22,24,27,31
              S2 = 1,3,6,8,16,29,32
              S3 = 2,4,5,12,14,19,20,23,25,26,28,30"
mod2 <- bfactor(data, specific2)
anova(mod1, mod2) # same

# also equivalent using item names instead (not run)
specific3 <- "S1 = Item.7, Item.9, Item.10, Item.11, Item.13, Item.15,
                Item.17, Item.18, Item.21, Item.22, Item.24, Item.27, Item.31
              S2 = Item.1, Item.3, Item.6, Item.8, Item.16, Item.29, Item.32
              S3 = Item.2, Item.4, Item.5, Item.12, Item.14, Item.19,
                Item.20, Item.23, Item.25, Item.26, Item.28, Item.30"
# mod3 <- bfactor(data, specific3)
# anova(mod1, mod2, mod3)  # all same
```

```
### Try with fixed guessing parameters added
guess <- rep(.1,32)
mod2 <- bfactor(data, specific, guess = guess)
coef(mod2)
anova(mod1, mod2)

## don't estimate specific factor for item 32
specific[32] <- NA
mod3 <- bfactor(data, specific)
anova(mod3, mod1)

# same, but with syntax (not run)
specific3 <- "S1 = 7,9,10,11,13,15,17,18,21,22,24,27,31
              S2 = 1,3,6,8,16,29
              S3 = 2,4,5,12,14,19,20,23,25,26,28,30"
# mod3b <- bfactor(data, specific3)
# anova(mod3b)


#########
# mixed itemtype example

# simulate data
a <- matrix(c(
1,0.5,NA,
1,0.5,NA,
1,0.5,NA,
1,0.5,NA,
1,0.5,NA,
1,0.5,NA,
1,0.5,NA,
1,NA,0.5,
1,NA,0.5,
1,NA,0.5,
1,NA,0.5,
1,NA,0.5,
1,NA,0.5),ncol=3,byrow=TRUE)

d <- matrix(c(
-1.0,NA,NA,
-1.5,NA,NA,
 1.5,NA,NA,
 0.0,NA,NA,
2.5,1.0,-1,
3.0,2.0,-0.5,
3.0,2.0,-0.5,
3.0,2.0,-0.5,
2.5,1.0,-1,
2.0,0.0,NA,
-1.0,NA,NA,
-1.5,NA,NA,
```

```
 1.5,NA,NA,
 0.0,NA,NA),ncol=3,byrow=TRUE)
items <- rep('2PL', 14)
items[5:10] <- 'graded'

sigma <- diag(3)
dataset <- simdata(a,d,5000,itemtype=items,sigma=sigma)
itemstats(dataset)

specific <- "S1 = 1-7
             S2 = 8-14"
simmod <- bfactor(dataset, specific)
coef(simmod, simplify=TRUE)


#########
# General testlet response model (Wainer, 2007)

# simulate data
set.seed(1234)
a <- matrix(0, 12, 4)
a[,1] <- rlnorm(12, .2, .3)
ind <- 1
for(i in 1:3){
   a[ind:(ind+3),i+1] <- a[ind:(ind+3),1]
   ind <- ind+4
}
print(a)
d <- rnorm(12, 0, .5)
sigma <- diag(c(1, .5, 1, .5))
dataset <- simdata(a,d,2000,itemtype=rep('2PL', 12),sigma=sigma)
itemstats(dataset)

# estimate by applying constraints and freeing the latent variances
specific <- "S1 = 1-4
             S2 = 5-8
             S3 = 9-12"
model <- "G = 1-12
          CONSTRAIN = (1, a1, a2), (2, a1, a2), (3, a1, a2), (4, a1, a2),
            (5, a1, a3), (6, a1, a3), (7, a1, a3), (8, a1, a3),
            (9, a1, a4), (10, a1, a4), (11, a1, a4), (12, a1, a4)
          COV = S1*S1, S2*S2, S3*S3"

simmod <- bfactor(dataset, specific, model)
coef(simmod, simplify=TRUE)

# Constrained testlet model (Bradlow, 1999)
model2 <- "G = 1-12
          CONSTRAIN = (1, a1, a2), (2, a1, a2), (3, a1, a2), (4, a1, a2),
            (5, a1, a3), (6, a1, a3), (7, a1, a3), (8, a1, a3),
            (9, a1, a4), (10, a1, a4), (11, a1, a4), (12, a1, a4),
            (GROUP, COV_22, COV_33, COV_44)
          COV = S1*S1, S2*S2, S3*S3"
```

```
simmod2 <- bfactor(dataset, specific, model2)
coef(simmod2, simplify=TRUE)
anova(simmod2, simmod)


#########
# Two-tier model

# simulate data
set.seed(1234)
a <- matrix(c(
  1,0,0.5,NA,NA,
  1,0,0.5,NA,NA,
  1,0,0.5,NA,NA,
  1,0,0.5,NA,NA,
  1,0,0.5,NA,NA,
  1,0,NA,0.5,NA,
  1,0,NA,0.5,NA,
  1,0,NA,0.5,NA,
  0,1,NA,0.5,NA,
  0,1,NA,0.5,NA,
  0,1,NA,0.5,NA,
  0,1,NA,NA,0.5,
  0,1,NA,NA,0.5,
  0,1,NA,NA,0.5,
  0,1,NA,NA,0.5,
  0,1,NA,NA,0.5),ncol=5,byrow=TRUE)

d <- matrix(rnorm(16))
items <- rep('2PL', 16)

sigma <- diag(5)
sigma[1,2] <- sigma[2,1] <- .4
dataset <- simdata(a,d,2000,itemtype=items,sigma=sigma)
itemstats(dataset)

specific <- "S1 = 1-5
             S2 = 6-11
             S3 = 12-16"
model <- '
    G1 = 1-8
    G2 = 9-16
    COV = G1*G2'

# quadpts dropped for faster estimation, but not as precise
simmod <- bfactor(dataset, specific, model, quadpts = 15, TOL = 1e-3)
coef(simmod, simplify=TRUE)
summary(simmod)
itemfit(simmod, QMC=TRUE)
M2(simmod, QMC=TRUE)
residuals(simmod, QMC=TRUE)
```

```
# EAP predictions for all factors (high dimensional)
eaps_all <- fscores(simmod, QMC=TRUE, quadpts=50000)
head(eaps_all)
maps <- fscores(simmod, method = 'MAP')
head(maps)

######
# Multiple-group bifactor example from Cai, Yang, Hansen (2011)

# Table 1 info from Cai, Yang, and Hansen (2011)
intercept <- c(1, .25, -.25, -1, 1, .25, -.25, -1,1,
               .25, -.25, -1,1, .25, -.25, -1)
theta0 <- c(1, 1.4,1.7,2,
            1.4,1.7,2,1,
            1.7,2,1,1.4,
            2, 1, 1.4, 1.7)
thetan <- c(.8,1.5,1.2,1,
            1,.8,1.5,1.2,
            1.2,1,.8,1.5,
            1.5,1.2,1,.8)
thetaN <- matrix(0, 16, 4)
thetaN[1:4, 1] <- thetan[1:4]
thetaN[1:4+4, 2] <- thetan[1:4+4]
thetaN[1:4+8, 3] <- thetan[1:4+8]
thetaN[1:4+12, 4] <- thetan[1:4+12]

as <- cbind(theta0, thetaN)
as

# data generation for focal group in publication does not have response
# for items 13-16. However, full-data approach presented first
N <- 1000
itemtype <- '2PL'
gmeans <- c(1, -.5, 0, .5, 0)
sigma <- diag(c(.8, 1.2, 1.5, 1, 1))

datG1 <- simdata(as, intercept, N=N, itemtype='2PL')
datG2 <- simdata(as, intercept, N=N, itemtype='2PL',
  mu = gmeans, sigma = sigma)

dat <- rbind(datG1, datG2)
group <- rep(c('G1', 'G2'), each=N)

specific <- "S1 = 1-4
             S2 = 5-8
             S3 = 9-12
             S4 = 13-16"

mod <- bfactor(dat, specific, group=group, SE=TRUE,
  invariance=c('free_means', 'free_vars', colnames(dat)))
coef(mod, simplify=TRUE)

# DIF testing for general dimension only with likelihood ratio tests (not run)
```

```
if(FALSE){
  mirtCluster()  # allocate computing cores
  DIF(mod, which.par=c('a1', 'd'), scheme='drop')
}

## same analysis, however items 13:16 do not exist in the focal group
datG2[,13:16] <- NA    # use this to match publication
dat <- rbind(datG1, datG2)
head(dat)
tail(dat)

# specify mean/cov structure explicitly
model2 <- "G = 1-16
           MEAN [G2] = G, S1, S2, S3
           COV [G2] = G*G, S1*S1, S2*S2, S3*S3"

mod2 <- bfactor(dat, specific, model2, group=group, invariance=colnames(dat))
coef(mod2, simplify=TRUE)
```

---

Bock1997 *Description of Bock 1997 data*

---

## Description

A 3-item tabulated data set extracted from Table 3 in Chapter Two.

## Author(s)

Phil Chalmers <rphilip.chalmers@gmail.com>

## References

Bock, R. D. (1997). The Nominal Categories Model. In van der Linden, W. J. & Hambleton, R. K. *Handbook of modern item response theory*. New York: Springer.

## Examples

```
dat <- expand.table(Bock1997)
head(dat)
itemstats(dat, use_ts=FALSE)

mod <- mirt(dat, 1, 'nominal')

# reproduce table 3 in Bock (1997)
fs <- round(fscores(mod, verbose = FALSE, full.scores = FALSE)[,c('F1','SE_F1')],2)
fttd <- residuals(mod, type = 'exp')
```

```
table <- data.frame(fttd[,-ncol(fttd)], fs)
table

mod <- mirt(dat, 1, 'nominal')
coef(mod)
```

---

`boot.LR` *Parametric bootstrap likelihood-ratio test*

---

## Description

Given two fitted models, compute a parametric bootstrap test to determine whether the less restrictive models fits significantly better than the more restricted model. Note that this hypothesis test also works when prior parameter distributions are included for either model. Function can be run in parallel after using a suitable `mirtCluster` definition.

## Usage

```
boot.LR(mod, mod2, R = 1000, verbose = interactive())
```

## Arguments

| | |
|---|---|
| `mod` | an estimated model object, more constrained than `mod2` |
| `mod2` | an estimated model object |
| `R` | number of parametric bootstraps to use. |
| `verbose` | logical; include additional information in the console? |

## Value

a p-value evaluating whether the more restrictive model fits significantly worse than the less restrictive model

## Author(s)

Phil Chalmers <rphilip.chalmers@gmail.com>

## References

Chalmers, R. P. (2012). mirt: A Multidimensional Item Response Theory Package for the R Environment. *Journal of Statistical Software, 48*(6), 1-29. doi:10.18637/jss.v048.i06

## Examples

```
# standard
dat <- expand.table(LSAT7)
mod1 <- mirt(dat, 1)
mod2 <- mirt(dat, 1, '3PL')

# standard LR test
anova(mod1, mod2)

# bootstrap LR test (run in parallel to save time)
if(interactive()) mirtCluster()
boot.LR(mod1, mod2, R=200)
```

---

`boot.mirt` *Calculate bootstrapped standard errors for estimated models*

---

## Description

Given an internal mirt object estimate the bootstrapped standard errors. It may be beneficial to run the computations using multi-core architecture (e.g., the `parallel` package). Parameters are organized from the freely estimated values in `mod2values(x)` (equality constraints will also be returned in the bootstrapped estimates).

## Usage

```
boot.mirt(x, R = 100, boot.fun = NULL, technical = NULL, ...)
```

## Arguments

| | |
|---|---|
| `x` | an estimated model object |
| `R` | number of draws to use (passed to the `boot()` function) |
| `boot.fun` | a user-defined function used to extract the information from the bootstrap fitted models. Must be of the form `boot.fun(x)`, where `x` is the bootstrap fitted model under investigation, and the return must be a numeric vector. If omitted a default function will be defined internally that returns the estimated parameters from the `mod` object, resulting in bootstrapped parameter estimate results |
| `technical` | technical arguments passed to estimation engine. See `mirt` for details |
| `...` | additional arguments to be passed on to `boot(...)` and mirt's estimation engine |

## Author(s)

Phil Chalmers <rphilip.chalmers@gmail.com>

## References

Chalmers, R. P. (2012). mirt: A Multidimensional Item Response Theory Package for the R Environment. *Journal of Statistical Software, 48*(6), 1-29. doi:10.18637/jss.v048.i06

## Examples

```
# standard
mod <- mirt(Science, 1)
booted <- boot.mirt(mod, R=20)
plot(booted)
booted

## Not run:
#run in parallel using snow back-end using all available cores
mod <- mirt(Science, 1)
booted <- boot.mirt(mod, parallel = 'snow', ncpus = parallel::detectCores())
booted

## End(Not run)

####
# bootstrapped CIs for standardized factor loadings
boot.fun <- function(mod){
  so <- summary(mod, verbose=FALSE)
  as.vector(so$rotF)
}

# test to see if it works before running
boot.fun(mod)

# run
booted.loads <- boot.mirt(mod, boot.fun=boot.fun)
booted.loads
```

---

`coef-method` *Extract raw coefs from model object*

---

## Description

Return a list (or data.frame) of raw item and group level coefficients. Note that while the output to the console is rounded to three digits, the returned list of objects is not. Hence, elements from `cfs <- coef(mod); cfs[[1]]` will contain the non-rounded results (useful for simulations).

## Usage

```
## S4 method for signature 'SingleGroupClass'
coef(
  object,
  CI = 0.95,
  printSE = FALSE,
  rotate = "none",
  Target = NULL,
  IRTpars = FALSE,
  rawug = FALSE,
  as.data.frame = FALSE,
  simplify = FALSE,
  unique = FALSE,
  verbose = TRUE,
  ...
)
```

## Arguments

| | |
|---|---|
| object | an object of class SingleGroupClass, MultipleGroupClass, or MixedClass |
| CI | the amount of converged used to compute confidence intervals; default is 95 percent confidence intervals |
| printSE | logical; print the standard errors instead of the confidence intervals? When IRTpars = TRUE then the delta method will be used to compute the associated standard errors from mirt's default slope-intercept form |
| rotate | see summary method for details. The default rotation is 'none' |
| Target | a dummy variable matrix indicting a target rotation pattern |
| IRTpars | logical; convert slope intercept parameters into traditional IRT parameters? Only applicable to unidimensional models or models with simple structure (i.e., only one non-zero slope). If a suitable ACOV estimate was computed in the fitted model, and printSE = FALSE, then suitable CIs will be included based on the delta method (where applicable) |
| rawug | logical; return the untransformed internal g and u parameters? If FALSE, g and u's are converted with the original format along with delta standard errors |
| as.data.frame | logical; convert list output to a data.frame instead? |
| simplify | logical; if all items have the same parameter names (indicating they are of the same class) then they are collapsed to a matrix, and a list of length 2 is returned containing a matrix of item parameters and group-level estimates |
| unique | return the vector of uniquely estimated parameters |
| verbose | logical; allow information to be printed to the console? |
| ... | additional arguments to be passed |

## References

Chalmers, R. P. (2012). mirt: A Multidimensional Item Response Theory Package for the R Environment. *Journal of Statistical Software, 48*(6), 1-29. doi:10.18637/jss.v048.i06

## See Also

summary-method

## Examples

```
dat <- expand.table(LSAT7)
x <- mirt(dat, 1)
coef(x)
coef(x, IRTpars = TRUE)
coef(x, simplify = TRUE)

#with computed information matrix
x <- mirt(dat, 1, SE = TRUE)
coef(x)
coef(x, printSE = TRUE)
coef(x, as.data.frame = TRUE)

#two factors
x2 <- mirt(Science, 2)
coef(x2)
coef(x2, rotate = 'varimax')
```

---

# createGroup *Create a user defined group-level object with correct generic functions*

---

## Description

Initializes the proper S4 class and methods necessary for mirt functions to use in estimation for defining customized group-level functions. To use the defined objects pass to the `mirt(..., customGroup = OBJECT)` command, and ensure that the class parameters are properly labelled.

## Usage

```
createGroup(
  par,
  est,
  den,
  nfact,
  standardize = FALSE,
  gr = NULL,
  hss = NULL,
  gen = NULL,
  lbound = NULL,
  ubound = NULL,
  derivType = "Richardson"
)
```

## Arguments

| | |
|---|---|
| par | a named vector of the starting values for the parameters |
| est | a logical vector indicating which parameters should be freely estimated by default |
| den | the probability density function given the Theta/ability values. First input contains a vector of all the defined parameters and the second input must be a matrix called Theta. Function also must return a numeric vector object corresponding to the associated densities for each row in the Theta input |
| nfact | number of factors required for the model. E.g., for unidimensional models with only one dimension of integration nfact = 1 |
| standardize | logical; use standardization of the quadrature table method proposed by Woods and Thissen (2006)? If TRUE, the logical elements named 'MEAN_1' and 'COV_11' can be included in the parameter vector, and when these values are set to FALSE in the est input the E-table will be standardized to these fixed values (e.g., par <- c(a1=1, d=0, MEAN_1=0, COV_11=1) with est <- c(TRUE, TRUE, FALSE, FALSE) will standardize the E-table to have a 0 mean and unit variance) |
| gr | gradient function (vector of first derivatives) of the log-likelihood used in estimation. The function must be of the form gr(x, Theta), where x is the object defined by createGroup() and Theta is a matrix of latent trait parameters |
| hss | Hessian function (matrix of second derivatives) of the log-likelihood used in estimation. If not specified a numeric approximation will be used. The input is identical to the gr argument |
| gen | a function used when GenRandomPars = TRUE is passed to the estimation function to generate random starting values. Function must be of the form function(object) ... and must return a vector with properties equivalent to the par object. If NULL, parameters will remain at the defined starting values by default |
| lbound | optional vector indicating the lower bounds of the parameters. If not specified then the bounds will be set to -Inf |
| ubound | optional vector indicating the lower bounds of the parameters. If not specified then the bounds will be set to Inf |
| derivType | if the gr or hss terms are not specified this type will be used to obtain them numerically. Default is 'Richardson' |

## Author(s)

Phil Chalmers <rphilip.chalmers@gmail.com>

## References

Chalmers, R. P. (2012). mirt: A Multidimensional Item Response Theory Package for the R Environment. *Journal of Statistical Software, 48*(6), 1-29. doi:10.18637/jss.v048.i06

## Examples

```
# normal density example, N(mu, sigma^2)
den <- function(obj, Theta) dnorm(Theta, obj@par[1], sqrt(obj@par[2]))
par <- c(mu = 0, sigma2 = .5)
est <- c(FALSE, TRUE)
lbound <- c(-Inf, 0)
grp <- createGroup(par, est, den, nfact = 1, lbound=lbound)

dat <- expand.table(LSAT6)
mod <- mirt(dat, 1, 'Rasch')
modcustom <- mirt(dat, 1, 'Rasch', customGroup=grp)

coef(mod)
coef(modcustom)
```

---

`createItem` *Create a user defined item with correct generic functions*

---

## Description

Initializes the proper S4 class and methods necessary for `mirt` functions to use in estimation. To use the defined objects pass to the `mirt(..., customItems = list())` command, and ensure that the classes are properly labelled and unique in the list. Additionally, the input `mirt(..., customItemsData = list())` can also be included to specify additional item-level information to better recycle custom-item definitions (e.g., for supplying varying Q-matrices), where the `list` input must have the same length as the number of items. For further examples regarding how this function can be used for fitting unfolding-type models see Liu and Chalmers (2018).

## Usage

```
createItem(
  name,
  par,
  est,
  P,
  gr = NULL,
  hss = NULL,
  gen = NULL,
  lbound = NULL,
  ubound = NULL,
  derivType = "Richardson",
  derivType.hss = "Richardson",
  bytecompile = TRUE
)
```

## Arguments

| | |
|---|---|
| name | a character indicating the item class name to be defined |
| par | a named vector of the starting values for the parameters |
| est | a logical vector indicating which parameters should be freely estimated by default |
| P | the probability trace function for all categories (first column is category 1, second category two, etc). First input contains a vector of all the item parameters, the second input must be a matrix called Theta, the third input must be the number of categories called ncat, and (optionally) a fourth argument termed itemdata may be included containing further users specification information. The last optional input is to be utilized within the estimation functions such as mirt via the list input customItemsData to more naturally recycle custom-item definitions. Therefore, these inputs must be of the form<br>function(par, Theta, ncat){...}<br>or<br>function(par, Theta, ncat, itemdata){...}<br>to be valid; however, the names of the arguements is not relavent.<br>Finally, this function must return a matrix object of category probabilities, where the columns represent each respective category |
| gr | gradient function (vector of first derivatives) of the log-likelihood used in estimation. The function must be of the form gr(x, Theta), where x is the object defined by createItem() and Theta is a matrix of latent trait parameters. Tabulated (EM) or raw (MHRM) data are located in the x@dat slot, and are used to form the complete data log-likelihood. If not specified a numeric approximation will be used |
| hss | Hessian function (matrix of second derivatives) of the log-likelihood used in estimation. If not specified a numeric approximation will be used (required for the MH-RM algorithm only). The input is identical to the gr argument |
| gen | a function used when GenRandomPars = TRUE is passed to the estimation function to generate random starting values. Function must be of the form function(object) ... and must return a vector with properties equivalent to the par object. If NULL, parameters will remain at the defined starting values by default |
| lbound | optional vector indicating the lower bounds of the parameters. If not specified then the bounds will be set to -Inf |
| ubound | optional vector indicating the lower bounds of the parameters. If not specified then the bounds will be set to Inf |
| derivType | if the gr term is not specified this type will be used to obtain the gradient numerically or symbolically. Default is the 'Richardson' extrapolation method; see numerical_deriv for details and other options. If 'symbolic' is supplied then the gradient will be computed using a symbolical approach (potentially the most accurate method, though may fail depending on how the P function was defined) |
| derivType.hss | if the hss term is not specified this type will be used to obtain the Hessian numerically. Default is the 'Richardson' extrapolation method; see numerical_deriv |

for details and other options. If 'symbolic' is supplied then the Hessian will be computed using a symbolical approach (potentially the most accurate method, though may fail depending on how the P function was defined)

bytecompile — logical; where applicable, byte compile the functions provided? Default is TRUE to provide

## Details

The summary() function will not return proper standardized loadings since the function is not sure how to handle them (no slopes could be defined at all!). Instead loadings of .001 are filled in as place-holders.

## Author(s)

Phil Chalmers <rphilip.chalmers@gmail.com>

## References

Chalmers, R. P. (2012). mirt: A Multidimensional Item Response Theory Package for the R Environment. *Journal of Statistical Software, 48*(6), 1-29. doi:10.18637/jss.v048.i06

Liu, C.-W. and Chalmers, R. P. (2018). Fitting item response unfolding models to Likert-scale data using mirt in R. *PLoS ONE, 13*, 5. doi:10.1371/journal.pone.0196292

## Examples

```
name <- 'old2PL'
par <- c(a = .5, b = -2)
est <- c(TRUE, TRUE)
P.old2PL <- function(par,Theta, ncat){
     a <- par[1]
     b <- par[2]
     P1 <- 1 / (1 + exp(-1*a*(Theta - b)))
     cbind(1-P1, P1)
}

x <- createItem(name, par=par, est=est, P=P.old2PL)

# So, let's estimate it!
dat <- expand.table(LSAT7)
sv <- mirt(dat, 1, c(rep('2PL',4), 'old2PL'), customItems=list(old2PL=x), pars = 'values')
tail(sv) #looks good
mod <- mirt(dat, 1, c(rep('2PL',4), 'old2PL'), customItems=list(old2PL=x))
coef(mod)
mod2 <- mirt(dat, 1, c(rep('2PL',4), 'old2PL'), customItems=list(old2PL=x), method = 'MHRM')
coef(mod2)

# same definition as above, but using symbolic derivative computations
# (can be more accurate/stable)
xs <- createItem(name, par=par, est=est, P=P.old2PL, derivType = 'symbolic')
```

```
mod <- mirt(dat, 1, c(rep('2PL',4), 'old2PL'), customItems=list(old2PL=xs))
coef(mod, simplify=TRUE)

# several secondary functions supported
M2(mod, calcNull=FALSE)
itemfit(mod)
fscores(mod, full.scores=FALSE)
plot(mod)

# fit the same model, but specify gradient function explicitly (use of a browser() may be helpful)
gr <- function(x, Theta){
     # browser()
     a <- x@par[1]
     b <- x@par[2]
     P <- probtrace(x, Theta)
     PQ <- apply(P, 1, prod)
     r_P <- x@dat / P
     grad <- numeric(2)
     grad[2] <- sum(-a * PQ * (r_P[,2] - r_P[,1]))
     grad[1] <- sum((Theta - b) * PQ * (r_P[,2] - r_P[,1]))

     ## check with internal numerical form to be safe
     # numerical_deriv(x@par[x@est], mirt:::EML, obj=x, Theta=Theta)
     grad
}

x <- createItem(name, par=par, est=est, P=P.old2PL, gr=gr)
mod <- mirt(dat, 1, c(rep('2PL',4), 'old2PL'), customItems=list(old2PL=x))
coef(mod, simplify=TRUE)

### non-linear
name <- 'nonlin'
par <- c(a1 = .5, a2 = .1, d = 0)
est <- c(TRUE, TRUE, TRUE)
P.nonlin <- function(par,Theta, ncat=2){
     a1 <- par[1]
     a2 <- par[2]
     d <- par[3]
     P1 <- 1 / (1 + exp(-1*(a1*Theta + a2*Theta^2 + d)))
     cbind(1-P1, P1)
}

x2 <- createItem(name, par=par, est=est, P=P.nonlin)

mod <- mirt(dat, 1, c(rep('2PL',4), 'nonlin'), customItems=list(nonlin=x2))
coef(mod)

### nominal response model (Bock 1972 version)
Tnom.dev <- function(ncat) {
   T <- matrix(1/ncat, ncat, ncat - 1)
   diag(T[-1, ]) <-  diag(T[-1, ]) - 1
   return(T)
}
```

```
name <- 'nom'
par <- c(alp=c(3,0,-3),gam=rep(.4,3))
est <- rep(TRUE, length(par))
P.nom <- function(par, Theta, ncat){
   alp <- par[1:(ncat-1)]
   gam <- par[ncat:length(par)]
   a <- Tnom.dev(ncat) %*% alp
   c <- Tnom.dev(ncat) %*% gam
   z <- matrix(0, nrow(Theta), ncat)
   for(i in 1:ncat)
       z[,i] <- a[i] * Theta + c[i]
   P <- exp(z) / rowSums(exp(z))
   P
}

nom1 <- createItem(name, par=par, est=est, P=P.nom)
nommod <- mirt(Science, 1, 'nom1', customItems=list(nom1=nom1))
coef(nommod)
Tnom.dev(4) %*% coef(nommod)[[1]][1:3] #a
Tnom.dev(4) %*% coef(nommod)[[1]][4:6] #d
```

---

`deAyala` *Description of deAyala data*

---

## Description

Mathematics data from de Ayala (2009; pg. 14); 5 item dataset in table format.

## Author(s)

Phil Chalmers <rphilip.chalmers@gmail.com>

## References

de Ayala, R. J. (2009). *The theory and practice of item response theory*. Guilford Press.

## Examples

```
dat <- expand.table(deAyala)
head(dat)
itemstats(dat)
```

`DeltaMethod` *Numerical derivative version of delta method*

## Description

Delta method using numerical derivatives (via `numerical_deriv`) for the provided function. Convenient when target transformation function is easier to automate programmatically instead of using explicit formula or math expressions. Can also be useful for checking analytic results.

## Usage

```
DeltaMethod(fn, par, acov, ...)
```

## Arguments

| | |
|---|---|
| `fn` | a function specifying the type of transformation to make for each new parameter of interest. Must be of the form `fn(par, ...)`, or more simply `fn(par)`, and return a numeric vector with one element |
| `par` | numerical vector passed to `fn(par)` (typically vector of MLEs) |
| `acov` | numeric matrix for the ACOV of the MLEs |
| `...` | additional arguments passed to fn |

## Value

returns a list of the transformed parameters, ACOV, and SEs

## Examples

```
# Slightly modified example from ?msm::deltamethod
# Multiple linear regression, E(y) = alpha + beta1 x + beta2 g
x <- 1:100
g <- rep(0:1, each=50)
y <- rnorm(100, 4*x, 5)
toy.lm <- lm(y ~ x + g)
estmean <- coef(toy.lm)
estvar <- vcov(toy.lm)

# Estimate of (1 / (b0 + b1)) and (1 / (b0 + b1 + b2))
1 / (estmean[1] + estmean[2])
1 / (estmean[1] + estmean[2] + estmean[3])

## Not run:
## Approximate standard error
msm::deltamethod (~ 1 / (x1 + x2), estmean, estvar)
msm::deltamethod (~ 1 / (x1 + x2 + x3), estmean, estvar)

## End(Not run)
```

```
# with DeltaMethod
fn <- function(par) 1 / sum(par[1:2])
fn2 <- function(par) 1 / sum(par[1:3])
DeltaMethod(fn, estmean, estvar)$se
DeltaMethod(fn2, estmean, estvar)$se

# index argument for easier flexibility
fn <- function(par, index) 1 / sum(par[index])
DeltaMethod(fn, estmean, estvar, index=1:2)$se
DeltaMethod(fn, estmean, estvar, index=1:3)$se
```

---

DIF *Differential item functioning statistics*

---

## Description

This function runs the Wald and likelihood-ratio approaches for testing differential item functioning (DIF) with two or more groups. This is primarily a convenience wrapper to the `multipleGroup` and `bfactor` functions for performing standard DIF procedures. Independent models can be estimated in parallel by defining a parallel object with `mirtCluster`, which will help to decrease the run time. For best results, the baseline model should contain a set of 'anchor' items and have freely estimated hyper-parameters in the focal groups.

## Usage

```
DIF(
  MGmodel,
  which.par,
  scheme = "add",
  items2test = 1:extract.mirt(MGmodel, "nitems"),
  groups2test = "all",
  seq_stat = "SABIC",
  Wald = FALSE,
  p.adjust = "holm",
  pairwise = FALSE,
  return_models = FALSE,
  return_seq_model = FALSE,
  max_run = Inf,
  plotdif = FALSE,
  type = "trace",
  simplify = TRUE,
  verbose = interactive(),
  ...
)
```

## Arguments

| | |
|---|---|
| MGmodel | an object returned from multipleGroup to be used as the reference model |
| which.par | a character vector containing the parameter names which will be inspected for DIF |
| scheme | type of DIF analysis to perform, either by adding or dropping constraints across groups. These can be:<br>**'add'** parameters in which.par will be constrained each item one at a time for items that are specified in items2test. This is beneficial when examining DIF from a model with parameters freely estimated across groups, and when inspecting differences via the Wald test<br>**'drop'** parameters in which.par will be freely estimated for items that are specified in items2test. This is useful when supplying an overly restrictive model and attempting to detect DIF with a slightly less restrictive model<br>**'add_sequential'** sequentially loop over the items being tested, and at the end of the loop treat DIF tests that satisfy the seq_stat criteria as invariant. The loop is then re-run on the remaining invariant items to determine if they are now displaying DIF in the less constrained model, and when no new invariant item is found the algorithm stops and returns the items that displayed DIF. Note that the DIF statistics are relative to this final, less constrained model which includes the DIF effects<br>**'drop_sequential'** sequentially loop over the items being tested, and at the end of the loop treat items that violate the seq_stat criteria as demonstrating DIF. The loop is then re-run, leaving the items that previously demonstrated DIF as variable across groups, and the remaining test items that previously showed invariance are re-tested. The algorithm stops when no more items showing DIF are found and returns the items that displayed DIF. Note that the DIF statistics are relative to this final, less constrained model which includes the DIF effects |
| items2test | a numeric vector, or character vector containing the item names, indicating which items will be tested for DIF. In models where anchor items are known, omit them from this vector. For example, if items 1 and 2 are anchors in a 10 item test, then items2test = 3:10 would work for testing the remaining items (important to remember when using sequential schemes) |
| groups2test | a character vector indicating which groups to use in the DIF testing investigations. Default is 'all', which uses all group information to perform joint hypothesis tests of DIF (for a two group setup these result in pair-wise tests). For example, if the group names were 'g1', 'g2' and 'g3', and DIF was only to be investigated between group 'g1' and 'g3' then pass groups2test = c('g1', 'g3') |
| seq_stat | select a statistic to test for in the sequential schemes. Potential values are (in descending order of power) 'AIC', 'SABIC', 'HQ', and 'BIC'. If a numeric value is input that ranges between 0 and 1, the 'p' value will be tested (e.g., seq_stat = .05 will test for the difference of p < .05 in the add scheme, or p > .05 in the drop scheme), along with the specified p.adjust input |
| Wald | logical; perform Wald tests for DIF instead of likelihood ratio test? |

| | |
|---|---|
| p.adjust | string to be passed to the p.adjust function to adjust p-values. Adjustments are located in the adj_p element in the returned list. Default uses the Holm-sequential Bonferroni adjustment |
| pairwise | logical; perform pairwise tests between groups when the number of groups is greater than 2? Useful as quickly specified post-hoc tests |
| return_models | logical; return estimated model objects for further analysis? Default is FALSE |
| return_seq_model | logical; on the last iteration of the sequential schemes, return the fitted multiple-group model containing the freely estimated parameters indicative of DIF? This is generally only useful when scheme = 'add_sequential'. Default is FALSE |
| max_run | a number indicating the maximum number of cycles to perform in sequential searches. The default is to perform search until no further DIF is found |
| plotdif | logical; create item plots for items that are displaying DIF according to the seq_stat criteria? Only available for 'add' type schemes |
| type | the type of plot argument passed to plot(). Default is 'trace', though another good option is 'infotrace'. For ease of viewing, the facet_item argument to mirt's plot() function is set to TRUE |
| simplify | logical; simplify the output by returning a data.frame object with the differences between AIC, BIC, etc, as well as the chi-squared test (X2) and associated df and p-values |
| verbose | logical print extra information to the console? |
| ... | additional arguments to be passed to multipleGroup and plot |

## Details

Generally, the pre-computed baseline model should have been configured with two estimation properties: 1) a set of 'anchor' items, where the anchor items have various parameters that have been constrained to be equal across the groups, and 2) contain freely estimated latent mean and variance terms in all but one group (the so-called 'reference' group). These two properties help to fix the metric of the groups so that item parameter estimates do not contain latent distribution characteristics.

## Value

a mirt_df object with the information-based criteria for DIF, though this may be changed to a list output when return_models or simplify are modified. As well, a silent 'DIF_coefficeints' attribute is included to view the item parameter differences between the groups

## Author(s)

Phil Chalmers <rphilip.chalmers@gmail.com>

## References

Chalmers, R. P. (2012). mirt: A Multidimensional Item Response Theory Package for the R Environment. *Journal of Statistical Software, 48*(6), 1-29. doi:10.18637/jss.v048.i06

Chalmers, R. P., Counsell, A., and Flora, D. B. (2016). It might not make a big DIF: Improved Differential Test Functioning statistics that account for sampling variability. *Educational and Psychological Measurement, 76*, 114-140. doi:10.1177/0013164415584576

## See Also

multipleGroup, DRF

## Examples

```
# simulate data where group 2 has a smaller slopes and more extreme intercepts
set.seed(12345)
a1 <- a2 <- matrix(abs(rnorm(15,1,.3)), ncol=1)
d1 <- d2 <- matrix(rnorm(15,0,.7),ncol=1)
a2[1:2, ] <- a1[1:2, ]/3
d1[c(1,3), ] <- d2[c(1,3), ]/4
head(data.frame(a.group1 = a1, a.group2 = a2, d.group1 = d1, d.group2 = d2))
itemtype <- rep('2PL', nrow(a1))
N <- 1000
dataset1 <- simdata(a1, d1, N, itemtype)
dataset2 <- simdata(a2, d2, N, itemtype, mu = .1, sigma = matrix(1.5))
dat <- rbind(dataset1, dataset2)
group <- c(rep('D1', N), rep('D2', N))

#### no anchors, all items tested for DIF by adding item constrains one item at a time.
# define a parallel cluster (optional) to help speed up internal functions
if(interactive()) mirtCluster()

# Information matrix with Oakes' identity (not controlling for latent group differences)
# NOTE: Without properly equating the groups the following example code is not testing for DIF,
     # but instead reflects a combination of DIF + latent-trait distribution effects
model <- multipleGroup(dat, 1, group, SE = TRUE)

# Likelihood-ratio test for DIF (as well as model information)
dif <- DIF(model, c('a1', 'd'))
dif

# function silently includes "DIF_coefficients" attribute to view
# the IRT parameters post-completion
extract.mirt(dif, "DIF_coefficients")

# same as above, but using Wald tests with Benjamini & Hochberg adjustment
DIF(model, c('a1', 'd'), Wald = TRUE, p.adjust = 'fdr')

# equate the groups by assuming the last 5 items have no DIF
itemnames <- colnames(dat)
model <- multipleGroup(dat, 1, group, SE = TRUE,
   invariance = c(itemnames[11:ncol(dat)], 'free_means', 'free_var'))

# test whether adding slopes and intercepts constraints results in DIF. Plot items showing DIF
resulta1d <- DIF(model, c('a1', 'd'), plotdif = TRUE, items2test=1:10)
```

```
resulta1d

# test whether adding only slope constraints results in DIF for all items
DIF(model, 'a1', items2test=1:10)

# Determine whether it's a1 or d parameter causing DIF (could be joint, however)
(a1s <- DIF(model, 'a1', items2test = 1:3))
(ds <- DIF(model, 'd', items2test = 1:3))

### drop down approach (freely estimating parameters across groups) when
### specifying a highly constrained model with estimated latent parameters
model_constrained <- multipleGroup(dat, 1, group,
  invariance = c(colnames(dat), 'free_means', 'free_var'))
dropdown <- DIF(model_constrained, c('a1', 'd'), scheme = 'drop')
dropdown

# View silent "DIF_coefficients" attribute
extract.mirt(dropdown, "DIF_coefficients")

### sequential schemes (add constraints)

### sequential searches using SABIC as the selection criteria
# starting from completely different models
stepup <- DIF(model, c('a1', 'd'), scheme = 'add_sequential',
              items2test=1:10)
stepup

# step down procedure for highly constrained model
stepdown <- DIF(model_constrained, c('a1', 'd'), scheme = 'drop_sequential')
stepdown

# view final MG model (only useful when scheme is 'add_sequential')
updated_mod <- DIF(model, c('a1', 'd'), scheme = 'add_sequential',
               return_seq_model=TRUE)
plot(updated_mod, type='trace')


###################################
# Multi-group example

a1 <- a2 <- a3 <- matrix(abs(rnorm(15,1,.3)), ncol=1)
d1 <- d2 <- d3 <- matrix(rnorm(15,0,.7),ncol=1)
a2[1:2, ] <- a1[1:2, ]/3
d3[c(1,3), ] <- d2[c(1,3), ]/4
head(data.frame(a.group1 = a1, a.group2 = a2, a.group3 = a3,
                d.group1 = d1, d.group2 = d2, d.group3 = d3))
itemtype <- rep('2PL', nrow(a1))
N <- 1000
dataset1 <- simdata(a1, d1, N, itemtype)
dataset2 <- simdata(a2, d2, N, itemtype, mu = .1, sigma = matrix(1.5))
dataset3 <- simdata(a3, d3, N, itemtype, mu = .2)
dat <- rbind(dataset1, dataset2, dataset3)
group <- gl(3, N, labels = c('g1', 'g2', 'g3'))
```

```
# equate the groups by assuming the last 5 items have no DIF
itemnames <- colnames(dat)
model <- multipleGroup(dat, group=group, SE=TRUE,
   invariance = c(itemnames[11:ncol(dat)], 'free_means', 'free_var'))
coef(model, simplify=TRUE)

# omnibus tests
dif <- DIF(model, which.par = c('a1', 'd'), items2test=1:9)
dif

# pairwise post-hoc tests for items flagged via omnibus tests
dif.posthoc <- DIF(model, which.par = c('a1', 'd'), items2test=1:2,
                   pairwise = TRUE)
dif.posthoc

# further probing for df = 1 tests, this time with Wald tests
DIF(model, which.par = c('a1'), items2test=1:2, pairwise = TRUE,
    Wald=TRUE)
DIF(model, which.par = c('d'), items2test=1:2, pairwise = TRUE,
    Wald=TRUE)

##################
# Restricted DIF tests when categories missing across groups

set.seed(1234)
a1 <- a2 <- matrix(abs(rnorm(15,1,.3)), ncol=1)
diffs <- t(apply(matrix(runif(15*4, .3, 1), 15), 1, cumsum))
diffs <- -(diffs - rowMeans(diffs))
d1 <- d2 <- diffs + rnorm(15)
d1[1,] <- d2[1,] <- d2[1,] + 3
d1[1,1] <- d2[1,1] <- 6.5
head(data.frame(a.group1 = a1, a.group2 = a2, d.group1 = d1, d.group2 = d2))
itemtype <- rep('graded', nrow(a1))
N <- 1000
dataset1 <- simdata(a1, d1, N, itemtype=itemtype, mu = -1)
dataset2 <- simdata(a2, d2, N, itemtype=itemtype, mu = 1)
dat <- rbind(dataset1, dataset2)
group <- c(rep('D1', N), rep('D2', N))
itemstats(dat, group = group)

# notice table of counts for item 1, lowest category missing in D2
table(dat[,1], group)

# ensure that missing intercept information borrowed from non-missing group
mod <- multipleGroup(dat, group=group,
                     invariance = c('free_means', 'free_vars', colnames(dat)))
coef(mod, simplify=TRUE)

if(FALSE)
    DIF(mod, which.par = c('a1', 'd1', 'd2', 'd3', 'd4'),
        items2test=1, scheme = 'drop') # Not good ... d1 not identified
```

```
# test item for DIF, however omit d1 as this is not identified
DIF(mod, which.par = c('a1', 'd2', 'd3', 'd4'),
    items2test=1, scheme = 'drop')

# plot resulting DIF test
difmod <- DIF(mod, which.par = c('a1', 'd2', 'd3', 'd4'),
              items2test=1, scheme = 'drop', return_models = TRUE)
itemplot(difmod[[1]], 1)
subset(DRF(difmod[[1]], DIF=TRUE, DIF.cats=TRUE), item == 'Item_1')

itemplot(difmod[[1]], 1, type = 'score')
subset(DRF(difmod[[1]], DIF=TRUE), item == 'Item_1')
```

---

`DiscreteClass-class` *Class "DiscreteClass"*

---

## Description

Defines the object returned from `mdirt`.

## Slots

`Call`: function call

`Data`: list of data, sometimes in different forms

`Options`: list of estimation options

`Fit`: a list of fit information

`Model`: a list of model-based information

`ParObjects`: a list of the S4 objects used during estimation

`OptimInfo`: a list of arguments from the optimization process

`Internals`: a list of internal arguments for secondary computations (inspecting this object is generally not required)

`vcov`: a matrix represented the asymptotic covariance matrix of the parameter estimates

`time`: a data.frame indicating the breakdown of computation times in seconds

## Methods

**print** `signature(x = "DiscreteClass")`

**show** `signature(object = "DiscreteClass")`

**anova** `signature(object = "DiscreteClass")`

**coef** `signature(x = "DiscreteClass")`

**summary** `signature(object = "DiscreteClass")`

**residuals** `signature(object = "DiscreteClass")`

### Author(s)

Phil Chalmers <rphilip.chalmers@gmail.com>

### References

Chalmers, R. P. (2012). mirt: A Multidimensional Item Response Theory Package for the R Environment. *Journal of Statistical Software, 48*(6), 1-29. doi:10.18637/jss.v048.i06

---

## draw_parameters *Draw plausible parameter instantiations from a given model*

---

### Description

Draws plausible parameters from a model using parametric sampling (if the information matrix was computed) or via bootstrap sampling. Primarily for use with the DRF function.

### Usage

```
draw_parameters(
  mod,
  draws,
  method = c("parametric", "boostrap"),
  redraws = 20,
  verbose = FALSE,
  ...
)
```

### Arguments

| | |
|---|---|
| mod | estimated single or multiple-group model |
| draws | number of draws to obtain |
| method | type of plausible values to obtain. Can be 'parametric', for the parametric sampling scheme which uses the estimated information matrix, or 'boostrap' to obtain values from the boot function. Default is 'parametric' |
| redraws | number of redraws to perform when the given parameteric sample does not satisfy the upper and lower parameter bounds. If a valid set cannot be found within this number of draws then an error will be thrown |
| verbose | logical; include additional information in the console? |
| ... | additional arguments to be passed |

### Value

returns a draws x p matrix of plausible parameters, where each row correspeonds to a single set

## Examples

```
set.seed(1234)
n <- 40
N <- 500

# only first 5 items as anchors
model <- 'F = 1-40
          CONSTRAINB = (1-5, a1), (1-5, d)'

a <- matrix(1, n)
d <- matrix(rnorm(n), n)
group <- c(rep('Group_1', N), rep('Group_2', N))

## -------------
# groups completely equal
dat1 <- simdata(a, d, N, itemtype = 'dich')
dat2 <- simdata(a, d, N, itemtype = 'dich')
dat <- rbind(dat1, dat2)
mod <- multipleGroup(dat, model, group=group, SE=TRUE,
                     invariance=c('free_means', 'free_var'))

param_set <- draw_parameters(mod, 100)
head(param_set)
```

---

# DRF *Differential Response Functioning statistics*

---

## Description

Function performs various omnibus differential item (DIF), bundle (DBF), and test (DTF) functioning procedures on an object estimated with `multipleGroup()`. The compensatory and non-compensatory statistics provided are described in Chalmers (2018), which generally can be interpreted as IRT generalizations of the SIBTEST and CSIBTEST statistics. For hypothesis tests, these measures require the ACOV matrix to be computed in the fitted multiple-group model (otherwise, sets of plausible draws from the posterior are explicitly required).

## Usage

```
DRF(
  mod,
  draws = NULL,
  focal_items = 1L:extract.mirt(mod, "nitems"),
  param_set = NULL,
  den.type = "marginal",
  best_fitting = FALSE,
```

```
  CI = 0.95,
  npts = 1000,
  quadpts = NULL,
  theta_lim = c(-6, 6),
  Theta_nodes = NULL,
  plot = FALSE,
  DIF = FALSE,
  DIF.cats = FALSE,
  groups2test = "all",
  pairwise = FALSE,
  simplify = TRUE,
  p.adjust = "none",
  par.strip.text = list(cex = 0.7),
 par.settings = list(strip.background = list(col = "#9ECAE1"), strip.border = list(col =
    "black")),
  auto.key = list(space = "right", points = FALSE, lines = TRUE),
  verbose = interactive(),
  ...
)
```

## Arguments

| | |
|---|---|
| mod | a multipleGroup object which estimated only 2 groups |
| draws | a number indicating how many draws to take to form a suitable multiple imputation or bootstrap estimate of the expected test scores (100 or more). If boot = FALSE, requires an estimated parameter information matrix. Returns a list containing the bootstrap/imputation distribution and null hypothesis test for the sDRF statistics |
| focal_items | a character/numeric vector indicating which items to include in the DRF tests. The default uses all of the items (note that including anchors in the focal items has no effect because they are exactly equal across groups). Selecting fewer items will result in tests of 'differential bundle functioning' |
| param_set | an N x p matrix of parameter values drawn from the posterior (e.g., using the parametric sampling approach, bootstrap, of MCMC). If supplied, then these will be used to compute the DRF measures. Can be much more efficient to precompute these values if DIF, DBF, or DTF are being evaluated within the same model (especially when using the bootstrap method). See draw_parameters |
| den.type | character specifying how the density of the latent traits is computed. Default is 'marginal' to include the proportional information from both groups, 'focal' for just the focal group, and 'reference' for the reference group |
| best_fitting | logical; use the best fitting parametric distribution (Gaussian by default) that was used at the time of model estimation? This will result in much fast computations, however the results are more dependent upon the underlying modelling assumptions. Default is FALSE, which uses the empirical histogram approach |
| CI | range of confidence interval when using draws input |
| npts | number of points to use for plotting. Default is 1000 |

| | |
|---|---|
| quadpts | number of quadrature nodes to use when constructing DRF statistics. Default is extracted from the input model object |
| theta_lim | lower and upper limits of the latent trait (theta) to be evaluated, and is used in conjunction with quadpts and npts |
| Theta_nodes | an optional matrix of Theta values to be evaluated in the draws for the sDRF statistics. However, these values are not averaged across, and instead give the bootstrap confidence intervals at the respective Theta nodes. Useful when following up a large sDRF or uDRF statistic, for example, to determine where the difference between the test curves are large (while still accounting for sampling variability). Returns a matrix with observed variability |
| plot | logical; plot the 'sDRF' functions for the evaluated sDBF or sDTF values across the integration grid or, if DIF = TRUE, the selected items as a faceted plot of individual items? If plausible parameter sets were obtained/supplied then imputed confidence intervals will be included |
| DIF | logical; return a list of item-level imputation properties using the DRF statistics? These can generally be used as a DIF detection method and as a graphical display for understanding DIF within each item |
| DIF.cats | logical; same as DIF = TRUE, however computations will be performed on each item category probability functions rather than the score functions. Only useful for understanding DIF in polytomous items |
| groups2test | when more than 2 groups are being investigated which two groups should be used in the effect size comparisons? |
| pairwise | logical; perform pairwise computations when the applying to multi-group settings |
| simplify | logical; attempt to simplify the output rather than returning larger lists? |
| p.adjust | string to be passed to the p.adjust function to adjust p-values. Adjustments are located in the adj_pvals element in the returned list. Only applicable when DIF = TRUE |
| par.strip.text | plotting argument passed to lattice |
| par.settings | plotting argument passed to lattice |
| auto.key | plotting argument passed to lattice |
| verbose | logical; include additional information in the console? |
| ... | additional arguments to be passed to lattice |

## Details

The effect sizes estimates by the DRF function are

$$sDRF = \int [S(C|\mathbf{\Psi}^{(R)}, \theta)S(C|\mathbf{\Psi}^{(F)}, \theta)]f(\theta)d\theta,$$

$$uDRF = \int |S(C|\mathbf{\Psi}^{(R)}, \theta)S(C|\mathbf{\Psi}^{(F)}, \theta)|f(\theta)d\theta,$$

and

$$dDRF = \sqrt{\int [S(C|\mathbf{\Psi}^{(R)}, \theta)S(C|\mathbf{\Psi}^{(F)}, \theta)]^2 f(\theta)d\theta}$$

where $S(.)$ are the scoring equations used to evaluate the model-implied difference between the focal and reference group. The $f(\theta)$ terms can either be estimated from the posterior via an empirical histogram approach (default), or can use the best fitting prior distribution that is obtain post-convergence (default is a Gaussian distribution). Note that, in comparison to Chalmers (2018), the focal group is the leftmost scoring function while the reference group is the rightmost scoring function. This is largely to keep consistent with similar effect size statistics, such as SIBTEST, DFIT, Wainer's measures of impact, etc, which in general can be seen as special-case estimators of this family.

Finally, for unidimensional models the standardized versions of the above effect sizes are also reported, which are obtained by dividing by the associated (pooled) standard deviation of the item/bundle/test scoring functions given a Gaussian density function with the associated group mean-variance estimates (see `marginal_moments`). These are reported as `DRF*` values in the output to reflect the standardization (in a Cohen's d type metric). Note that this standardization approach reflects a model-based analogues of the `ESSD` and `ETSSD` statistics found in `empirical_ES` (see Meade, 2010).

## Author(s)

Phil Chalmers <rphilip.chalmers@gmail.com>

## References

Chalmers, R. P. (2018). Model-Based Measures for Detecting and Quantifying Response Bias. *Psychometrika, 83*(3), 696-732. doi:10.1007/s1133601896269

## See Also

`multipleGroup`, `DIF`

## Examples

```
set.seed(1234)
n <- 30
N <- 500

# only first 5 items as anchors
model <- 'F = 1-30
          CONSTRAINB = (1-5, a1), (1-5, d)'

a <- matrix(1, n)
d <- matrix(rnorm(n), n)
group <- c(rep('Group_1', N), rep('Group_2', N))

## -------------
# groups completely equal
dat1 <- simdata(a, d, N, itemtype = 'dich')
dat2 <- simdata(a, d, N, itemtype = 'dich')
dat <- rbind(dat1, dat2)
mod <- multipleGroup(dat, model, group=group, SE=TRUE,
                     invariance=c('free_means', 'free_var'))
```

```
plot(mod)
plot(mod, which.items = 6:10) #DBF
plot(mod, type = 'itemscore')
plot(mod, type = 'itemscore', which.items = 10:15)

# empirical histogram approach
DRF(mod)
DRF(mod, focal_items = 6:10) #DBF
DRF(mod, DIF=TRUE)
DRF(mod, DIF=TRUE, focal_items = 10:15)

# Best-fitting Gaussian distributions
DRF(mod, best_fitting=TRUE)
DRF(mod, focal_items = 6:10, best_fitting=TRUE) #DBF
DRF(mod, DIF=TRUE, best_fitting=TRUE)
DRF(mod, DIF=TRUE, focal_items = 10:15, best_fitting=TRUE)

DRF(mod, plot = TRUE)
DRF(mod, focal_items = 6:10, plot = TRUE) #DBF
DRF(mod, DIF=TRUE, plot = TRUE)
DRF(mod, DIF=TRUE, focal_items = 10:15, plot = TRUE)

if(interactive()) mirtCluster()
DRF(mod, draws = 500)
DRF(mod, draws = 500, best_fitting=TRUE)
DRF(mod, draws = 500, plot=TRUE)

# pre-draw parameter set to save computations
#  (more useful when using non-parametric bootstrap)
param_set <- draw_parameters(mod, draws = 500)
DRF(mod, focal_items = 6, param_set=param_set) #DIF test
DRF(mod, DIF=TRUE, param_set=param_set) #DIF test
DRF(mod, focal_items = 6:10, param_set=param_set) #DBF test
DRF(mod, param_set=param_set) #DTF test

DRF(mod, focal_items = 6:10, draws=500) #DBF test
DRF(mod, focal_items = 10:15, draws=500) #DBF test

DIFs <- DRF(mod, draws = 500, DIF=TRUE)
print(DIFs)
DRF(mod, draws = 500, DIF=TRUE, plot=TRUE)

DIFs <- DRF(mod, draws = 500, DIF=TRUE, focal_items = 6:10)
print(DIFs)
DRF(mod, draws = 500, DIF=TRUE, focal_items = 6:10, plot = TRUE)

DRF(mod, DIF=TRUE, focal_items = 6)
DRF(mod, draws=500, DIF=TRUE, focal_items = 6)

# evaluate specific values for sDRF
Theta_nodes <- matrix(seq(-6,6,length.out = 100))

sDTF <- DRF(mod, Theta_nodes=Theta_nodes)
```

```
head(sDTF)
sDTF <- DRF(mod, Theta_nodes=Theta_nodes, draws=200)
head(sDTF)

# sDIF (isolate single item)
sDIF <- DRF(mod, Theta_nodes=Theta_nodes, focal_items=6)
head(sDIF)
sDIF <- DRF(mod, Theta_nodes=Theta_nodes, focal_items = 6, draws=200)
head(sDIF)

## -------------
## random slopes and intercepts for 15 items, and latent mean difference
##    (no systematic DTF should exist, but DIF will be present)
set.seed(1234)
dat1 <- simdata(a, d, N, itemtype = 'dich', mu=.50, sigma=matrix(1.5))
dat2 <- simdata(a + c(numeric(15), rnorm(n-15, 0, .25)),
                d + c(numeric(15), rnorm(n-15, 0, .5)), N, itemtype = 'dich')
dat <- rbind(dat1, dat2)
mod1 <- multipleGroup(dat, 1, group=group)
plot(mod1)
DRF(mod1) #does not account for group differences! Need anchors

mod2 <- multipleGroup(dat, model, group=group, SE=TRUE,
                      invariance=c('free_means', 'free_var'))
plot(mod2)

# significant DIF in multiple items....
# DIF(mod2, which.par=c('a1', 'd'), items2test=16:30)
DRF(mod2)
DRF(mod2, draws=500) #non-sig DTF due to item cancellation

## -------------
## systematic differing slopes and intercepts (clear DTF)
set.seed(1234)
dat1 <- simdata(a, d, N, itemtype = 'dich', mu=.50, sigma=matrix(1.5))
dat2 <- simdata(a + c(numeric(15), rnorm(n-15, 1, .25)),
                d + c(numeric(15), rnorm(n-15, 1, .5)),
                N, itemtype = 'dich')
dat <- rbind(dat1, dat2)
mod3 <- multipleGroup(dat, model, group=group, SE=TRUE,
                      invariance=c('free_means', 'free_var'))
plot(mod3) #visable DTF happening

# DIF(mod3, c('a1', 'd'), items2test=16:30)
DRF(mod3) #unsigned bias. Signed bias (group 2 scores higher on average)
DRF(mod3, draws=500)
DRF(mod3, draws=500, plot=TRUE) #multiple DRF areas along Theta

# plot the DIF
DRF(mod3, draws=500, DIF=TRUE, plot=TRUE)

# evaluate specific values for sDRF
Theta_nodes <- matrix(seq(-6,6,length.out = 100))
```

```
sDTF <- DRF(mod3, Theta_nodes=Theta_nodes, draws=200)
head(sDTF)

# DIF
sDIF <- DRF(mod3, Theta_nodes=Theta_nodes, focal_items = 30, draws=200)
head(sDIF)

## ------------------------------------------------------------------
# polytomous example
# simulate data where group 2 has a different slopes/intercepts
set.seed(4321)
a1 <- a2 <- matrix(rlnorm(20,.2,.3))
a2[c(16:17, 19:20),] <- a1[c(16:17, 19:20),] + c(-.5, -.25, .25, .5)

# for the graded model, ensure that there is enough space between the intercepts,
# otherwise closer categories will not be selected often
diffs <- t(apply(matrix(runif(20*4, .3, 1), 20), 1, cumsum))
diffs <- -(diffs - rowMeans(diffs))
d1 <- d2 <- diffs + rnorm(20)
rownames(d1) <- rownames(d2) <- paste0('Item.', 1:20)
d2[16:20,] <- d1[16:20,] + matrix(c(-.5, -.5, -.5, -.5,
                                    1, 0, 0, -1,
                                    .5, .5, -.5, -.5,
                                    1, .5, 0, -1,
                                    .5, .5, .5, .5), byrow=TRUE, nrow=5)

tail(data.frame(a.group1 = a1, a.group2 = a2), 6)
list(d.group1 = d1[15:20,], d.group2 = d2[15:20,])

itemtype <- rep('graded', nrow(a1))
N <- 600
dataset1 <- simdata(a1, d1, N, itemtype)
dataset2 <- simdata(a2, d2, N, itemtype, mu = -.25, sigma = matrix(1.25))
dat <- rbind(dataset1, dataset2)
group <- c(rep('D1', N), rep('D2', N))

# item 1-10 as anchors
mod <- multipleGroup(dat, group=group, SE=TRUE,
                     invariance=c(colnames(dat)[1:10], 'free_means', 'free_var'))
coef(mod, simplify=TRUE)
plot(mod)
plot(mod, type='itemscore')

# DIF tests vis Wald method
DIF(mod, items2test=11:20,
   which.par=c('a1', paste0('d', 1:4)),
   Wald=TRUE, p.adjust='holm')

DRF(mod)
DRF(mod, DIF=TRUE, focal_items=11:20)
DRF(mod, DIF.cats=TRUE, focal_items=11:20)

## ------------------------------------------------------------------
```

```
### multidimensional DTF

set.seed(1234)
n <- 50
N <- 1000

# only first 5 items as anchors within each dimension
model <- 'F1 = 1-25
          F2 = 26-50
          COV = F1*F2
          CONSTRAINB = (1-5, a1), (1-5, 26-30, d), (26-30, a2)'

a <- matrix(c(rep(1, 25), numeric(50), rep(1, 25)), n)
d <- matrix(rnorm(n), n)
group <- c(rep('Group_1', N), rep('Group_2', N))
Cov <- matrix(c(1, .5, .5, 1.5), 2)
Mean <- c(0, 0.5)

# groups completely equal
dat1 <- simdata(a, d, N, itemtype = 'dich', sigma = cov2cor(Cov))
dat2 <- simdata(a, d, N, itemtype = 'dich', sigma = Cov, mu = Mean)
dat <- rbind(dat1, dat2)
mod <- multipleGroup(dat, model, group=group, SE=TRUE,
                     invariance=c('free_means', 'free_var'))
coef(mod, simplify=TRUE)
plot(mod, degrees = c(45,45))
DRF(mod)

# some intercepts slightly higher in Group 2
d2 <- d
d2[c(10:15, 31:35)] <- d2[c(10:15, 31:35)] + 1
dat1 <- simdata(a, d, N, itemtype = 'dich', sigma = cov2cor(Cov))
dat2 <- simdata(a, d2, N, itemtype = 'dich', sigma = Cov, mu = Mean)
dat <- rbind(dat1, dat2)
mod <- multipleGroup(dat, model, group=group, SE=TRUE,
                     invariance=c('free_means', 'free_var'))
coef(mod, simplify=TRUE)
plot(mod, degrees = c(45,45))

DRF(mod)
DRF(mod, draws = 500)
```

---

DTF *Differential test functioning statistics*

---

## Description

Function performs various omnibus differential test functioning procedures on an object estimated with `multipleGroup()`. If the latent means/covariances are suspected to differ then the input object

should contain a set of 'anchor' items to ensure that only differential test features are being detected rather than group differences. Returns signed (average area above and below) and unsigned (total area) statistics, with descriptives such as the percent average bias between group total scores for each statistic. If a grid of Theta values is passed, these can be evaluated as well to determine specific DTF location effects. For best results, the baseline model should contain a set of 'anchor' items and have freely estimated hyper-parameters in the focal groups. See DIF for details.

## Usage

```
DTF(
  mod,
  draws = NULL,
  CI = 0.95,
  npts = 1000,
  theta_lim = c(-6, 6),
  Theta_nodes = NULL,
  plot = "none",
  auto.key = list(space = "right", points = FALSE, lines = TRUE),
  ...
)
```

## Arguments

| | |
|---|---|
| mod | a multipleGroup object which estimated only 2 groups |
| draws | a number indicating how many draws to take to form a suitable multiple imputation estimate of the expected test scores (usually 100 or more). Returns a list containing the imputation distribution and null hypothesis test for the sDTF statistic |
| CI | range of confidence interval when using draws input |
| npts | number of points to use in the integration. Default is 1000 |
| theta_lim | lower and upper limits of the latent trait (theta) to be evaluated, and is used in conjunction with npts |
| Theta_nodes | an optional matrix of Theta values to be evaluated in the draws for the sDTF statistic. However, these values are not averaged across, and instead give the bootstrap confidence intervals at the respective Theta nodes. Useful when following up a large uDTF/sDTF statistic to determine where the difference between the test curves are large (while still accounting for sampling variability). Returns a matrix with observed variability |
| plot | a character vector indicating which plot to draw. Possible values are 'none', 'func' for the test score functions, and 'sDTF' for the evaluated sDTF values across the integration grid. Each plot is drawn with imputed confidence envelopes |
| auto.key | logical; automatically generate key in lattice plot? |
| ... | additional arguments to be passed to lattice and boot |

## Author(s)

Phil Chalmers <rphilip.chalmers@gmail.com>

## References

Chalmers, R. P. (2012). mirt: A Multidimensional Item Response Theory Package for the R Environment. *Journal of Statistical Software, 48*(6), 1-29. doi:10.18637/jss.v048.i06

Chalmers, R. P., Counsell, A., and Flora, D. B. (2016). It might not make a big DIF: Improved Differential Test Functioning statistics that account for sampling variability. *Educational and Psychological Measurement, 76*, 114-140. doi:10.1177/0013164415584576

## See Also

multipleGroup, DIF

## Examples

```
set.seed(1234)
n <- 30
N <- 500

# only first 5 items as anchors
model <- 'F = 1-30
          CONSTRAINB = (1-5, a1), (1-5, d)'

a <- matrix(1, n)
d <- matrix(rnorm(n), n)
group <- c(rep('Group_1', N), rep('Group_2', N))

## -------------
# groups completely equal
dat1 <- simdata(a, d, N, itemtype = '2PL')
dat2 <- simdata(a, d, N, itemtype = '2PL')
dat <- rbind(dat1, dat2)
mod <- multipleGroup(dat, model, group=group, SE=TRUE,
                     invariance=c('free_means', 'free_var'))
plot(mod)

DTF(mod)
if(interactive()) mirtCluster()
DTF(mod, draws = 1000) #95% C.I. for sDTF containing 0. uDTF is very small
DTF(mod, draws = 1000, plot='sDTF') #sDTF 95% C.I.'s across Theta always include 0

## -------------
## random slopes and intercepts for 15 items, and latent mean difference
##    (no systematic DTF should exist, but DIF will be present)
set.seed(1234)
dat1 <- simdata(a, d, N, itemtype = '2PL', mu=.50, sigma=matrix(1.5))
dat2 <- simdata(a + c(numeric(15), runif(n-15, -.2, .2)),
                d + c(numeric(15), runif(n-15, -.5, .5)), N, itemtype = '2PL')
dat <- rbind(dat1, dat2)
mod1 <- multipleGroup(dat, 1, group=group)
plot(mod1) #does not account for group differences! Need anchors

mod2 <- multipleGroup(dat, model, group=group, SE=TRUE,
```

```
                      invariance=c('free_means', 'free_var'))
plot(mod2)

# significant DIF in multiple items....
# DIF(mod2, which.par=c('a1', 'd'), items2test=16:30)
DTF(mod2)
DTF(mod2, draws=1000) #non-sig DTF due to item cancellation

## -------------
## systematic differing slopes and intercepts (clear DTF)
dat1 <- simdata(a, d, N, itemtype = '2PL', mu=.50, sigma=matrix(1.5))
dat2 <- simdata(a + c(numeric(15), rnorm(n-15, 1, .25)), d + c(numeric(15), rnorm(n-15, 1, .5)),
               N, itemtype = '2PL')
dat <- rbind(dat1, dat2)
mod3 <- multipleGroup(dat, model, group=group, SE=TRUE,
                      invariance=c('free_means', 'free_var'))
plot(mod3) #visable DTF happening

# DIF(mod3, c('a1', 'd'), items2test=16:30)
DTF(mod3) #unsigned bias. Signed bias indicates group 2 scores generally higher on average
DTF(mod3, draws=1000)
DTF(mod3, draws=1000, plot='func')
DTF(mod3, draws=1000, plot='sDTF') #multiple DTF areas along Theta

# evaluate specific values for sDTF
Theta_nodes <- matrix(seq(-6,6,length.out = 100))
sDTF <- DTF(mod3, Theta_nodes=Theta_nodes)
head(sDTF)
sDTF <- DTF(mod3, Theta_nodes=Theta_nodes, draws=100)
head(sDTF)
```

---

`empirical_ES` *Empirical effect sizes based on latent trait estimates*

---

## Description

Computes effect size measures of differential item functioning and differential test/bundle functioning based on expected scores from Meade (2010). Item parameters from both reference and focal group are used in conjunction with focal group empirical theta estimates (and an assumed normally distributed theta) to compute expected scores.

## Usage

```
empirical_ES(
  mod,
  Theta.focal = NULL,
  focal_items = 1L:extract.mirt(mod, "nitems"),
  DIF = TRUE,
```

```
  npts = 61,
  theta_lim = c(-6, 6),
  plot = FALSE,
  type = "b",
  par.strip.text = list(cex = 0.7),
 par.settings = list(strip.background = list(col = "#9ECAE1"), strip.border = list(col =
    "black")),
  ...
)
```

## Arguments

| | |
|---|---|
| mod | a multipleGroup object which estimated only 2 groups. The first group in this object is assumed to be the reference group by default (i.e., ref.group = 1), which conforms to the invariance arguments in multipleGroup |
| Theta.focal | an optional matrix of Theta values from the focal group to be evaluated. If not supplied the default values to fscores will be used in conjunction with the ... arguments passed |
| focal_items | a numeric vector indicating which items to include the tests. The default uses all of the items. Selecting fewer items will result in tests of 'differential bundle functioning' when DIF = FALSE |
| DIF | logical; return a data.frame of item-level imputation properties? If FALSE, only DBF and DTF statistics will be reported |
| npts | number of points to use in the integration. Default is 61 |
| theta_lim | lower and upper limits of the latent trait (theta) to be evaluated, and is used in conjunction with npts |
| plot | logical; plot expected scores of items/test where expected scores are computed using focal group thetas and both focal and reference group item parameters |
| type | type of objects to draw in lattice; default plots both points and lines |
| par.strip.text | plotting argument passed to lattice |
| par.settings | plotting argument passed to lattice |
| ... | additional arguments to be passed to fscores and xyplot |

## DIF

The default DIF = TRUE produces several effect sizes indices at the item level. Signed indices allow DIF favoring the focal group at one point on the theta distribution to cancel DIF favoring the reference group at another point on the theta distribution. Unsigned indices take the absolute value before summing or averaging, thus not allowing cancellation of DIF across theta.

**SIDS** Signed Item Difference in the Sample. The average difference in expected scores across the focal sample using both focal and reference group item parameters.

**UIDS** Unsigned Item Difference in the Sample. Same as SIDS except absolute value of expected scores is taken prior to averaging across the sample.

**D-Max** The maximum difference in expected scores in the sample.

**ESSD** Expected Score Standardized Difference. Cohen's D for difference in expected scores.

**SIDN** Signed Item Difference in a Normal distribution. Identical to SIDS but averaged across a normal distribution rather than the sample.

**UIDN** Unsigned Item Difference in a Normal distribution. Identical to UIDS but averaged across a normal distribution rather than the sample.

### DBF/DTF

`DIF = FALSE` produces a series of test/bundle-level indices that are based on item-level indices.

**STDS** Signed Test Differences in the Sample. The sum of the SIDS across items.

**UTDS** Unsigned Test Differences in the Sample. The sum of the UIDS across items.

**Stark's DTFR** Stark's version of STDS using a normal distribution rather than sample estimated thetas.

**UDTFR** Unsigned Expected Test Scores Differences in the Sample. The difference in observed summed scale scores expected, on average, across a hypothetical focal group with a normally distributed theta, had DF been uniform in nature for all items

**UETSDS** Unsigned Expected Test Score Differences in the Sample. The hypothetical difference expected scale scores that would have been present if scale-level DF had been uniform across respondents (i.e., always favoring the focal group).

**UETSDN** Identical to UETSDS but computed using a normal distribution.

**Test D-Max** Maximum expected test score differences in the sample.

**ETSSD** Expected Test Score Standardized Difference. Cohen's D for expected test scores.

## Author(s)

Adam Meade, with contributions by Phil Chalmers <rphilip.chalmers@gmail.com>

## References

Chalmers, R. P. (2012). mirt: A Multidimensional Item Response Theory Package for the R Environment. *Journal of Statistical Software, 48*(6), 1-29. doi:10.18637/jss.v048.i06

Meade, A. W. (2010). A taxonomy of effect size measures for the differential functioning of items and scales. *Journal of Applied Psychology, 95*, 728-743.

## Examples

```
# no DIF
set.seed(12345)
a <- matrix(abs(rnorm(15,1,.3)), ncol=1)
d <- matrix(rnorm(15,0,.7),ncol=1)
itemtype <- rep('2PL', nrow(a))
N <- 1000
dataset1 <- simdata(a, d, N, itemtype)
dataset2 <- simdata(a, d, N, itemtype, mu = .1, sigma = matrix(1.5))
dat <- rbind(dataset1, dataset2)
```

```
# ensure 'Ref' is the first group (and therefore reference group during estimation)
group <- factor(c(rep('Ref', N), rep('Focal', N)), levels = c('Ref', 'Focal'))

mod <- multipleGroup(dat, 1, group = group,
   invariance = c(colnames(dat)[1:5], 'free_means', 'free_var'))
coef(mod, simplify=TRUE)

empirical_ES(mod)
empirical_ES(mod, DIF=FALSE)
empirical_ES(mod, DIF=FALSE, focal_items = 10:15)

empirical_ES(mod, plot=TRUE)
empirical_ES(mod, plot=TRUE, DIF=FALSE)

###-----------------------------------------------
# DIF
set.seed(12345)
a1 <- a2 <- matrix(abs(rnorm(15,1,.3)), ncol=1)
d1 <- d2 <- matrix(rnorm(15,0,.7),ncol=1)
a2[10:15,] <- a2[10:15,] + rnorm(6, 0, .3)
d2[10:15,] <- d2[10:15,] + rnorm(6, 0, .3)
itemtype <- rep('dich', nrow(a1))
N <- 1000
dataset1 <- simdata(a1, d1, N, itemtype)
dataset2 <- simdata(a2, d2, N, itemtype, mu = .1, sigma = matrix(1.5))
dat <- rbind(dataset1, dataset2)
group <- factor(c(rep('Ref', N), rep('Focal', N)), levels = c('Ref', 'Focal'))

mod <- multipleGroup(dat, 1, group = group,
   invariance = c(colnames(dat)[1:5], 'free_means', 'free_var'))
coef(mod, simplify=TRUE)

empirical_ES(mod)
empirical_ES(mod, DIF = FALSE)
empirical_ES(mod, plot=TRUE)
empirical_ES(mod, plot=TRUE, DIF=FALSE)
```

---

`empirical_plot` *Function to generate empirical unidimensional item and test plots*

---

## Description

Given a dataset containing item responses this function will construct empirical graphics using the observed responses, potentially conditioned on the (reduced) total/composite score. When individual item plots are requested then the total score will be formed without the item of interest (i.e., the total score without that item).

## Usage

```
empirical_plot(
  data,
  which.items = NULL,
  type = "prop",
  smooth = FALSE,
  sort = FALSE,
  formula = resp ~ s(TS, k = 5),
  org.data = NULL,
  ts_fun = rowSums,
  discrim.cut = 0.2,
  main = NULL,
  par.strip.text = list(cex = 0.7),
 par.settings = list(strip.background = list(col = "#9ECAE1"), strip.border = list(col =
    "black")),
  auto.key = list(space = "right", points = FALSE, lines = TRUE),
  ...
)
```

## Arguments

| | |
|---|---|
| data | a data.frame or matrix of item responses (see mirt for typical input) |
| which.items | a numeric vector indicating which items to plot in a faceted image plot. If NULL then empirical test plots will be constructed instead |
| type | character vector specifying type of plot to draw, some of which change as a function of the which.item input. |
| | **'prop' (default)** cumulative total-score proportions. If which.items specified then item-level conditional proportions are instead plotted against the reduced total scores |
| | **'hist'** histogram of total scores |
| | **'freq'** item response frequencies (supports which.item) |
| | **'discrim'** reduced item-total correlations to visualize discrimination effects |
| | **'difficulty'** mean/proportion of each item |
| | **'discrim_diff'** reduced item-total correlation against item difficulty |
| | **'alpha_rm'** effect on coefficient alpha if item were removed |
| | **'boxplot'** conditional boxplots of reduced total scores (supports which.items) |
| | **'bubble'** bivariate frequency bubble plots (requires that which.items is exactly of length two) |
| smooth | logical; include a GAM smoother instead of the raw proportions? Default is FALSE |
| sort | logical; when applicable, sort the items first (e.g., in discrimination plot)? |
| formula | formula used for the GAM smoother |
| org.data | identical to data, but contains the "unscored" response options (e.g., the original coding in a multiple-choice test). Used in various item-level plots for diagnostic purposes, such as in distractor analyses. This will also automatically add size and linetype changes to highlight the detected scored categories |

| | |
|---|---|
| ts_fun | similar to the same argument in itemstats |
| discrim.cut | horizontal cut-off line to use when type = 'discrim'. Default is .2 (to omit, use NA) |
| main | the main title for the plot. If NULL an internal default will be used |
| par.strip.text | plotting argument passed to lattice |
| par.settings | plotting argument passed to lattice |
| auto.key | plotting argument passed to lattice |
| ... | additional arguments to be passed to lattice and coef() |

## Details

Note that some of these plot types should only be used for unidimensional tests with monotonically increasing item response functions. If monotonicity is not true for all items, however, then these plots may serve as a visual diagnostic tool so long as the majority of items are indeed monotonic.

## References

Chalmers, R. P. (2012). mirt: A Multidimensional Item Response Theory Package for the R Environment. *Journal of Statistical Software, 48*(6), 1-29. doi:10.18637/jss.v048.i06

## See Also

itemstats, itemplot, itemGAM

## Examples

```
SAT12[SAT12 == 8] <- NA
data <- key2binary(SAT12,
   key = c(1,4,5,2,3,1,2,1,3,1,2,4,2,1,5,3,4,4,1,4,3,3,4,1,3,5,1,3,1,5,4,5))

# test plot
empirical_plot(data)
empirical_plot(data, type = 'hist')
empirical_plot(data, type = 'hist', ts_fun=rowMeans)
empirical_plot(data, type = 'hist', breaks=20)
empirical_plot(data, type = 'discrim')
empirical_plot(data, type = 'discrim', sort=TRUE)
empirical_plot(data, type = 'difficulty')
empirical_plot(data, type = 'difficulty', sort=TRUE)
empirical_plot(data, type = 'discrim_diff')
empirical_plot(data, type = 'alpha_rm')
empirical_plot(data, type = 'freq')


# items 1, 2 and 5
empirical_plot(data, c(1, 2, 5), type = 'freq')
empirical_plot(data, c(1, 2, 5))
empirical_plot(data, c(1, 2, 5), smooth = TRUE)
```

```
empirical_plot(data, c(1, 2, 5), type = 'boxplot')
empirical_plot(data, c(1, 2), type = 'bubble')
empirical_plot(data, c(1, 5), type = 'bubble')

# replace weird looking items with unscored versions for diagnostics
empirical_plot(data, 32)
data2 <- data
data2[,32] <- SAT12[,32]
empirical_plot(data2, 32)
empirical_plot(data2, 32, smooth = TRUE)

# alternatively, distractor analyses using original dataset
empirical_plot(data, which.items=32, org.data=SAT12)
empirical_plot(data, which.items=32, org.data=SAT12, smooth=TRUE)
empirical_plot(data, which.items=1:12, org.data=SAT12, smooth=TRUE)
empirical_plot(data, which.items=13:32, org.data=SAT12, smooth=TRUE)

# example where item 32 has two correct scoring keys, and item 31
# is incorrectly scored (perhaps miskeyed)
key <- cbind(c(1,4,5,2,3,1,2,1,3,1,2,4,2,1,5,3,4,4,1,4,3,3,4,1,3,5,1,3,1,5,1,5),
             c(rep(NA, 31), 3))
data2 <- key2binary(SAT12, key=key)

# item 30 is correct, but 31 proportions do not climb. Item 32 has two
# highlighted correct answers
empirical_plot(data2, which.items=30:32, org.data=SAT12)
empirical_plot(data2, which.items=30:32, org.data=SAT12, smooth=TRUE)


#################
# polytomous response data
empirical_plot(Science)
empirical_plot(Science, type = 'hist', breaks=20)
empirical_plot(Science, type = 'freq')
empirical_plot(Science, type = 'difficulty')
empirical_plot(Science, type = 'discrim_diff')
empirical_plot(Science, type = 'boxplot')

# item-level
empirical_plot(Science, c(1, 2), type = 'bubble')
empirical_plot(Science, c(1, 3), type = 'bubble')
empirical_plot(Science, which.items = 1:4, type = 'prop')
empirical_plot(Science, which.items = 1:4, type = 'prop', smooth=TRUE)

# last plot very similar to model-based approach (though conditioned
#   on reduced total scores rather than scaled latent trait)
mod <- mirt(Science)
plot(mod, type='trace')

# when missing values present
Science[1:3, 1] <- NA
Science[6:8, 2] <- NA
empirical_plot(Science, type = 'freq')
```

empirical_rxx *Function to calculate the empirical (marginal) reliability*

## Description

Given secondary latent trait estimates and their associated standard errors returned from fscores, compute the empirical reliability.

## Usage

```
empirical_rxx(Theta_SE, T_as_X = FALSE)
```

## Arguments

| | |
|---|---|
| Theta_SE | a matrix of latent trait estimates returned from fscores with the options full.scores = TRUE and full.scores.SE = TRUE |
| T_as_X | logical; should the observed variance be equal to var(X) = var(T) + E(E^2) or var(X) = var(T) when computing empirical reliability estimates? Default (FALSE) uses the former |

## Author(s)

Phil Chalmers <rphilip.chalmers@gmail.com>

## References

Chalmers, R. P. (2012). mirt: A Multidimensional Item Response Theory Package for the R Environment. *Journal of Statistical Software, 48*(6), 1-29. doi:10.18637/jss.v048.i06

## See Also

fscores, marginal_rxx

## Examples

```
dat <- expand.table(deAyala)
itemstats(dat)
mod <- mirt(dat)

theta_se <- fscores(mod, full.scores.SE = TRUE)
empirical_rxx(theta_se)

theta_se <- fscores(mod, full.scores.SE = TRUE, method = 'ML')
```

```
empirical_rxx(theta_se)
empirical_rxx(theta_se, T_as_X = TRUE)
```

---

`estfun.AllModelClass` *Extract Empirical Estimating Functions*

---

## Description

A function for extracting the empirical estimating functions of a fitted `mirt`, `multipleGroup`, `bfactor`, or `mdirt` model. This is the derivative of the log-likelihood with respect to the parameter vector, evaluated at the observed (case-wise) data. In other words, this function returns the case-wise scores, evaluated at the fitted model parameters. Currently, models fitted via the `EM` or `BL` method are supported. For the computations, the internal `Theta` grid of the model is being used which was already used during the estimation of the model itself along with its matching normalized density.

## Usage

```
estfun.AllModelClass(
  x,
  weights = extract.mirt(x, "survey.weights"),
  centering = FALSE
)
```

## Arguments

| | |
|---|---|
| `x` | a fitted model object of class `SingleGroupClass`, `MultipleGroupClass`, or `DiscreteClass` |
| `weights` | by default, the `survey.weights` which were (optionally) specified when fitting the model are included to calculate the scores. If specified by the user, this should be a numeric vector of length equal to the total sample size. Note that if not all cases were weighted equally when fitting the model, the weights must be corrected by taking their square root if the scores are being used to compute the outer product of gradients (OPG) estimate of the variance-covariance matrix (see examples below). |
| `centering` | a boolean variable that allows the centering of the case-wise scores (i.e., setting their expected values to 0). If the case-wise scores were obtained from maximum likelihood estimates, this setting does not affect the result. |

## Value

An n x k matrix corresponding to n observations and k parameters

## Author(s)

Lennart Schneider <lennart.sch@web.de> and Phil Chalmers; centering argument contributed by Rudolf Debelak (<rudolf.debelak@psychologie.uzh.ch>)

## See Also

mirt, multipleGroup, bfactor, mdirt

## Examples

```
# fit a 2PL on the LSAT7 data and get the scores
mod1 <- mirt(expand.table(LSAT7), 1, SE = TRUE, SE.type = "crossprod")
sc1 <- estfun.AllModelClass(mod1)
# get the gradient
colSums(sc1)
# calculate the OPG estimate of the variance-covariance matrix "by hand"
vc1 <- vcov(mod1)
all.equal(crossprod(sc1), chol2inv(chol(vc1)), check.attributes = FALSE)

# Discrete group
modd <- mdirt(expand.table(LSAT7), 2, SE = TRUE, SE.type = "crossprod")
sc1 <- estfun.AllModelClass(modd)
# get the gradient
colSums(sc1)
# calculate the OPG estimate of the variance-covariance matrix "by hand"
vc1 <- vcov(modd)
all.equal(crossprod(sc1), chol2inv(chol(vc1)), check.attributes = FALSE)

# fit a multiple group 2PL and do the same as above
group <- rep(c("G1", "G2"), 500)
mod2 <- multipleGroup(expand.table(LSAT7), 1, group, SE = TRUE,
  SE.type = "crossprod")
sc2 <- estfun.AllModelClass(mod2)
colSums(sc2)
vc2 <- vcov(mod2)
all.equal(crossprod(sc2), chol2inv(chol(vc2)), check.attributes = FALSE)

# fit a bifactor model with 2 specific factors and do the same as above
mod3 <- bfactor(expand.table(LSAT7), c(2, 2, 1, 1, 2), SE = TRUE,
  SE.type = "crossprod")
sc3 <- estfun.AllModelClass(mod3)
colSums(sc3)
vc3 <- vcov(mod3)
all.equal(crossprod(sc3), chol2inv(chol(vc3)), check.attributes = FALSE)

# fit a 2PL not weighting all cases equally
survey.weights <- c(rep(2, sum(LSAT7$freq) / 2), rep(1, sum(LSAT7$freq) / 2))
survey.weights <- survey.weights / sum(survey.weights) * sum(LSAT7$freq)
mod4 <- mirt(expand.table(LSAT7), 1, SE = TRUE, SE.type = "crossprod",
  survey.weights = survey.weights)
sc4 <- estfun.AllModelClass(mod4,
```

```
  weights = extract.mirt(mod4, "survey.weights"))
# get the gradient
colSums(sc4)
# to calculate the OPG estimate of the variance-covariance matrix "by hand",
# the weights must be adjusted by taking their square root
sc4_crp <- estfun.AllModelClass(mod4,
  weights = sqrt(extract.mirt(mod4, "survey.weights")))
vc4 <- vcov(mod4)
all.equal(crossprod(sc4_crp), chol2inv(chol(vc4)), check.attributes = FALSE)
```

---

`expand.table` *Expand summary table of patterns and frequencies*

---

## Description

The `expand.table` function expands a summary table of unique response patterns to a full sized data-set. By default the response frequencies are assumed to be on rightmost column of the input data, though this can be modified.

## Usage

```
expand.table(tabdata, freq = colnames(tabdata)[ncol(tabdata)], sample = FALSE)
```

## Arguments

| | |
|---|---|
| `tabdata` | An object of class `data.frame` or `matrix` with the unique response patterns and the number of frequencies in the rightmost column (though see `freq` for details on how to omit this column) |
| `freq` | either a character vector specifying the column in `tabdata` to be used as the frequency count indicator for each response pattern (defaults to the right-most column) or a integer vector of length `nrow(tabdata)` specifying the frequency counts. When using the latter approach the `tabdata` input should not include any information regarding the counts, and instead should only include the unique response patterns themselves |
| `sample` | logical; randomly switch the rows in the expanded table? This does not change the expanded data, only the row locations |

## Value

Returns a numeric matrix with all the response patterns.

## Author(s)

Phil Chalmers <rphilip.chalmers@gmail.com>

## References

Chalmers, R. P. (2012). mirt: A Multidimensional Item Response Theory Package for the R Environment. *Journal of Statistical Software, 48*(6), 1-29. doi:10.18637/jss.v048.i06

## Examples

```
data(LSAT7)
head(LSAT7) # frequency in right-most column
LSAT7full <- expand.table(LSAT7)
head(LSAT7full)
dim(LSAT7full)

# randomly switch rows in the expanded response table
LSAT7samp <- expand.table(LSAT7, sample = TRUE)
head(LSAT7samp)
colMeans(LSAT7full)
colMeans(LSAT7samp) #equal

#--------


# Generate data from separate response pattern matrix and freq vector
# The following uses Table 2.1 from de Ayala (2009)
f <- c(691,2280,242,235,158,184,1685,1053,134,462,92,65,571,79,87,41,1682,702,
       370,63,626,412,166,52,28,15,2095,1219,500,187,40,3385)

pat <- matrix(c(
   0, 0, 0, 0, 0,
   1, 0, 0, 0, 0,
   0, 1, 0, 0, 0,
   0, 0, 1, 0, 0,
   0, 0, 0, 1, 0,
   0, 0, 0, 0, 1,
   1, 1, 0, 0, 0,
   1, 0, 1, 0, 0,
   0, 1, 1, 0, 0,
   1, 0, 0, 1, 0,
   0, 1, 0, 1, 0,
   0, 0, 1, 1, 0,
   1, 0, 0, 0, 1,
   0, 1, 0, 0, 1,
   0, 0, 1, 0, 1,
   0, 0, 0, 1, 1,
   1, 1, 1, 0, 0,
   1, 1, 0, 1, 0,
   1, 0, 1, 1, 0,
   0, 1, 1, 1, 0,
   1, 1, 0, 0, 1,
   1, 0, 1, 0, 1,
   1, 0, 0, 1, 1,
   0, 1, 1, 0, 1,
   0, 1, 0, 1, 1,
```

```
   0, 0, 1, 1, 1,
   1, 1, 1, 1, 0,
   1, 1, 1, 0, 1,
   1, 1, 0, 1, 1,
   1, 0, 1, 1, 1,
   0, 1, 1, 1, 1,
   1, 1, 1, 1, 1), ncol=5, byrow=TRUE)

colnames(pat) <- paste0('Item.', 1:5)
head(pat)

table2.1 <- expand.table(pat, freq = f)
dim(table2.1)
```

---

`expected.item` *Function to calculate expected value of item*

---

## Description

Given an internal mirt object extracted from an estimated model compute the expected value for an item given the ability parameter(s).

## Usage

```
expected.item(x, Theta, min = 0, include.var = FALSE)
```

## Arguments

| | |
|---|---|
| `x` | an extracted internal mirt object containing item information (see `extract.item`) |
| `Theta` | a vector (unidimensional) or matrix (multidimensional) of latent trait values |
| `min` | a constant value added to the expected values indicating the lowest theoretical category. Default is 0 |
| `include.var` | logical; include the model-implied variance of the expected scores as well? When `TRUE` will return a list containing the expected values (`E`) and variances (`VAR`) |

## Author(s)

Phil Chalmers <rphilip.chalmers@gmail.com>

## References

Chalmers, R. P. (2012). mirt: A Multidimensional Item Response Theory Package for the R Environment. *Journal of Statistical Software, 48*(6), 1-29. doi:10.18637/jss.v048.i06

## See Also

extract.item, expected.test

## Examples

```
mod <- mirt(Science, 1)
extr.2 <- extract.item(mod, 2)
Theta <- matrix(seq(-6,6, length.out=200))
expected <- expected.item(extr.2, Theta, min(Science[,1])) #min() of first item
head(data.frame(expected, Theta=Theta))

expected.item(extr.2, Theta, min(Science[,1]), include.var=TRUE)
```

---

# expected.test *Function to calculate expected test score*

---

## Description

Given an estimated model compute the expected test score. Returns the expected values in the same form as the data used to estimate the model.

## Usage

```
expected.test(
  x,
  Theta,
  group = NULL,
  mins = TRUE,
  individual = FALSE,
  which.items = NULL,
  probs.only = FALSE
)
```

## Arguments

| | |
|---|---|
| x | an estimated mirt object |
| Theta | a matrix of latent trait values (if a vector is supplied, will be coerced to a matrix with one column) |
| group | a number or character signifying which group the item should be extracted from (applies to 'MultipleGroupClass' objects only) |
| mins | logical; include the minimum value constants in the dataset. If FALSE, the expected values for each item are determined from the scoring 0:(ncat-1) |
| individual | logical; return tracelines for individual items? |
| which.items | an integer vector indicating which items to include in the expected test score. Default uses all possible items |
| probs.only | logical; return the probability for each category instead of traceline score functions? Only useful when individual=TRUE |

## References

Chalmers, R. P. (2012). mirt: A Multidimensional Item Response Theory Package for the R Environment. *Journal of Statistical Software, 48*(6), 1-29. doi:10.18637/jss.v048.i06

## See Also

expected.item

## Examples

```
dat <- expand.table(deAyala)
model <- 'F = 1-5
          CONSTRAIN = (1-5, a1)'
mod <- mirt(dat, model)

Theta <- matrix(seq(-6,6,.01))
tscore <- expected.test(mod, Theta)
tail(cbind(Theta, tscore))

# use only first two items (i.e., a bundle)
bscore <- expected.test(mod, Theta, which.items = 1:2)
tail(cbind(Theta, bscore))

# more low-level output (score and probabilty elements)
expected.test(mod, Theta, individual=TRUE)
expected.test(mod, Theta, individual=TRUE, probs.only=TRUE)
```

---

# extract.group *Extract a group from a multiple group mirt object*

---

## Description

Extract a single group from an object defined by multipleGroup, or as a mixture model from mirt.

## Usage

```
extract.group(x, group)
```

## Arguments

| | |
|---|---|
| x | model object of class 'MultipleGroupClass' or 'MixtureClass' |
| group | the name of the group to extract |

## Author(s)

Phil Chalmers <rphilip.chalmers@gmail.com>

## References

Chalmers, R. P. (2012). mirt: A Multidimensional Item Response Theory Package for the R Environment. *Journal of Statistical Software, 48*(6), 1-29. doi:10.18637/jss.v048.i06

## See Also

extract.item, extract.mirt

## Examples

```
set.seed(12345)
a <- matrix(abs(rnorm(15,1,.3)), ncol=1)
d <- matrix(rnorm(15,0,.7),ncol=1)
itemtype <- rep('2PL', nrow(a))
N <- 1000
dataset1 <- simdata(a, d, N, itemtype)
dataset2 <- simdata(a, d, N, itemtype, mu = .1, sigma = matrix(1.5))
dat <- rbind(dataset1, dataset2)
group <- c(rep('D1', N), rep('D2', N))
models <- 'F1 = 1-15'

mod_configural <- multipleGroup(dat, models, group = group)
group.1 <- extract.group(mod_configural, 'D1') #extract first group
summary(group.1)
plot(group.1)
```

---

# extract.item *Extract an item object from* mirt *objects*

---

## Description

Extract the internal mirt objects from any estimated model.

## Usage

```
extract.item(x, item, group = NULL, drop.zeros = FALSE)
```

## Arguments

| | |
|---|---|
| x | mirt model of class 'SingleGroupClass', 'MultipleGroupClass', or 'MixtureClass' |
| item | a number or character signifying which item to extract |
| group | which group the item should be extracted from (applies to 'MultipleGroupClass' and 'MixtureClass' only). Can be a numeric value or the name of the group to be extracted |
| drop.zeros | logical; drop slope values that are numerically close to zero to reduce dimensionality? Useful in objects returned from bfactor or other confirmatory models that contain several zero slopes |

## References

Chalmers, R. P. (2012). mirt: A Multidimensional Item Response Theory Package for the R Environment. *Journal of Statistical Software, 48*(6), 1-29. doi:10.18637/jss.v048.i06

## See Also

`extract.group`, `extract.mirt`

## Examples

```
mod <- mirt(Science, 1)
extr.1 <- extract.item(mod, 1)
```

---

`extract.mirt` *Extract various elements from estimated model objects*

---

## Description

A generic function to extract the internal objects from estimated models.

## Usage

```
extract.mirt(x, what, item = 1, ...)
```

## Arguments

| | |
|---|---|
| `x` | mirt model of class 'SingleGroupClass', 'MultipleGroupClass', 'MixedClass' or 'DiscreteGroupClass' |
| `what` | a string indicating what to extract |
| `item` | if necessary, which item to extract information from. Defaults to 1 if not specified |
| `...` | additional arguments to pass to `summary()` (e.g., for different factor rotations) |

## Details

Objects which can be extracted from mirt objects include:

**logLik** observed log-likelihood

**logPrior** log term contributed by prior parameter distributions

**G2** goodness of fit statistic

**df** degrees of freedom

**p** p-value for G2 statistic

**RMSEA** root mean-square error of approximation based on G2

**CFI** CFI fit statistic

**TLI** TLI fit statistic

**AIC** AIC

**BIC** BIC

**SABIC** sample size adjusted BIC

**HQ** HQ

**LLhistory** EM log-likelihood history

**tabdata** a tabular version of the raw response data input. Frequencies are stored in `freq`

**freq** frequencies associated with `tabdata`

**K** an integer vector indicating the number of unique elements for each item

**mins** an integer vector indicating the lowest category found in the input `data`

**model** input model syntax

**method** estimation method used

**itemtype** a vector of item types for each respective item (e.g., 'graded', '2PL', etc)

**itemnames** a vector of item names from the input data

**factorNames** a vector of factor names from the model definition

**rowID** an integer vector indicating all valid row numbers used in the model estimation (when all cases are used this will be `1:nrow(data)`)

**data** raw input data of item responses

**covdata** raw input data of data used as covariates

**tabdatalong** similar to `tabdata`, however the responses have been transformed into dummy coded variables

**fulldatalong** analogous to `tabdatafull`, but for the raw input data instead of the tabulated frequencies

**EMhistory** if saved, extract the EM iteration history

**F** unrotated factor loadings matrix

**Frot** rotated factor loadings matrix

**exp_resp** expected probability of the unique response patterns

**survey.weights** if supplied, the vector of survey weights used during estimation (NULL if missing)

**converged** a logical value indicating whether the model terminated within the convergence criteria

**iterations** number of iterations it took to reach the convergence criteria

**nestpars** number of freely estimated parameters

**parvec** vector containing uniquely estimated parameters

**vcov** parameter covariance matrix (associated with parvec)

**condnum** the condition number of the Hessian (if computed). Otherwise NA

**constrain** a list of item parameter constraints to indicate which item parameters were equal during estimation

**Prior** prior density distribution for the latent traits

**thetaPosterior** posterior distribution for latent traits when using EM algorithm

**key** if supplied, the data scoring key

**nfact** number of latent traits/factors

**nitems** number of items

**ngroups** number of groups

**groupNames** character vector of unique group names

**group** a character vector indicating the group membership

**invariance** a character vector indicating `invariance` input from `multipleGroup`

**secondordertest** a logical indicating whether the model passed the second-order test based on the Hessian matrix. Indicates whether model is a potential local maximum solution

**SEMconv** logical; check whether the supplemented EM information matrix converged. Will be `NA` if not applicable

**time** estimation time, broken into different sections

## Author(s)

Phil Chalmers <rphilip.chalmers@gmail.com>

## References

Chalmers, R. P. (2012). mirt: A Multidimensional Item Response Theory Package for the R Environment. *Journal of Statistical Software, 48*(6), 1-29. doi:10.18637/jss.v048.i06

## See Also

`extract.group`, `extract.item`, `mod2values`

## Examples

```
mod <- mirt(Science, 1)

extract.mirt(mod, 'logLik')
extract.mirt(mod, 'K')  # unique categories for each item

# multiple group model
grp <- rep(c('G1', 'G2'), each = nrow(Science)/2)
mod2 <- multipleGroup(Science, 1, grp)

grp1 <- extract.group(mod2, 1) #extract single group model
extract.mirt(mod2, 'parvec')
extract.mirt(grp1, 'parvec')
```

`fixedCalib` *Fixed-item calibration method*

## Description

Implements the set of fixed-item calibration methods described by Kim (2006). The initial calibrated model must be fitted via `mirt`, is currently limited to unidimensional models only, and should only be utilized when the new set of responses are obtained from a population with similar distributional characteristics in the latent traits. For more flexible calibration of items, including a fixed-item calibration variant involving anchor items for equating, see `multipleGroup`.

## Usage

```
fixedCalib(
  data,
  model = 1,
  old_mod,
  PAU = "MWU",
  NEMC = "MEM",
  technical = list(),
  ...
)
```

## Arguments

| | |
|---|---|
| `data` | new data to be used for calibration. Note that to be consistent with the `mod` object, observed responses/NA placeholders must be included to link the item names used in the original `mod` definition (i.e., `extract.mirt(mod, what = 'itemnames')`) |
| `model` | type of model to fit for the complete dataset (not that for the fixed items in `old_mod` the factor loadings/constraints specified by the potential `mirt.model` specification is not relevant) |
| `old_mod` | a model of class SingleGroupClass fitted using `mirt` |
| `PAU` | prior ability update (PAU) approach. Supports none (`"NWU"`), one (`"OWU"`), and many (`"MWU"`) |
| `NEMC` | number of EM cycles (NEMC) to use for the to-be-estimated parameters. Supports one (`"OEM"`) and many (`"MEM"`) |
| `technical` | list of technical estimation arguments (see `mirt` for details) |
| `...` | additional arguments to pass to `mirt` |

## References

Kim, S. (2006). A comparative study of IRT fixed parameter calibration methods. *Journal of Educational Measurement, 4*(43), 355-381.

## See Also

`mirt`, `multipleGroup`

## Examples

```
# single factor
set.seed(12345)
J <- 50
a <- matrix(abs(rnorm(J,1,.3)), ncol=1)
d <- matrix(rnorm(J,0,.7),ncol=1)
itemtype <- rep('2PL', nrow(a))

# calibration data theta ~ N(0,1)
N <- 3000
dataset1 <- simdata(a, d, N = N, itemtype=itemtype)

# new data (again, theta ~ N(0,1))
dataset2 <- simdata(a, d, N = 1000, itemtype=itemtype)

# last 40% of experimental items not given to calibration group
#     (unobserved; hence removed)
dataset1 <- dataset1[,-c(J:(J*.6))]
head(dataset1)

#--------------------------------------

# calibrated model from dataset1 only
mod <- mirt(dataset1, model = 1)
coef(mod, simplify=TRUE)

# No Prior Weights Updating and One EM Cycle (NWU-OEM)
NWU_OEM <- fixedCalib(dataset2, model=1, old_mod=mod, PAU='NWU', NEMC='OEM')
coef(NWU_OEM, simplify=TRUE)
data.frame(coef(NWU_OEM, simplify=TRUE)$items[,c('a1','d')],
           pop_a1=a, pop_d=d)
plot(NWU_OEM, type = 'empiricalhist')

# No Prior Weights Updating and Multiple EM Cycles (NWU-MEM)
NWU_MEM <- fixedCalib(dataset2, model = 1, old_mod = mod, PAU = 'NWU')
coef(NWU_MEM, simplify=TRUE)
data.frame(coef(NWU_MEM, simplify=TRUE)$items[,c('a1','d')],
           pop_a1=a, pop_d=d)
plot(NWU_MEM, type = 'empiricalhist')

# One Prior Weights Updating and One EM Cycle (OWU-OEM)
OWU_OEM <- fixedCalib(dataset2, model=1, old_mod=mod, PAU='OWU', NEMC="OEM")
coef(OWU_OEM, simplify=TRUE)
data.frame(coef(OWU_OEM, simplify=TRUE)$items[,c('a1','d')], pop_a1=a, pop_d=d)
plot(OWU_OEM, type = 'empiricalhist')

# One Prior Weights Updating and Multiple EM Cycles (OWU-MEM)
OWU_MEM <- fixedCalib(dataset2, model = 1, old_mod = mod, PAU = 'OWU')
coef(OWU_MEM, simplify=TRUE)
data.frame(coef(OWU_MEM, simplify=TRUE)$items[,c('a1','d')],
           pop_a1=a, pop_d=d)
```

```
plot(OWU_MEM, type = 'empiricalhist')

# Multiple Prior Weights Updating and Multiple EM Cycles (MWU-MEM)
MWU_MEM <- fixedCalib(dataset2, model = 1, old_mod = mod)
coef(MWU_MEM, simplify=TRUE)
data.frame(coef(MWU_MEM, simplify=TRUE)$items[,c('a1','d')],
           pop_a1=a, pop_d=d)
plot(MWU_MEM, type = 'empiricalhist')

# factor scores distribution check
fs <- fscores(MWU_MEM)
hist(fs)
c(mean_calib=mean(fs[1:N, ]), sd_calib=sd(fs[1:N, ]))
c(mean_exper=mean(fs[-c(1:N), ]), sd_exper=sd(fs[-c(1:N), ]))


############################
## Item length constraint example for each participant in the experimental
## items group. In this example, all participants were forced to have a test
## length of J=30, though the item pool had J=50 total items.

# new experimental data (relatively extreme, theta ~ N(.5,1.5))
dataset2 <- simdata(a, d, N = 1000, itemtype=itemtype,
    mu=.5, sigma=matrix(1.5))

# Add missing values to each participant in new dataset where individuals
# were randomly administered 10 experimental items, subject to the constraint
# that each participant received a test with J=30 items.
dataset2 <- t(apply(dataset2, 1, function(x){
   NA_precalib <- sample(1:30, 10)
   NA_experimental <- sample(31:50, 10)
   x[c(NA_precalib, NA_experimental)] <- NA
   x
}))
head(dataset2)

# check that all individuals had 30 items
all(rowSums(!is.na(dataset2)) == 30)

# Multiple Prior Weights Updating and Multiple EM Cycles (MWU-MEM)
MWU_MEM <- fixedCalib(dataset2, model = 1, old_mod = mod)
coef(MWU_MEM, simplify=TRUE)
data.frame(coef(MWU_MEM, simplify=TRUE)$items[,c('a1','d')],
           pop_a1=a, pop_d=d)
plot(MWU_MEM, type = 'empiricalhist')

## factor scores check
fs <- fscores(MWU_MEM)
hist(fs)
c(mean_calib=mean(fs[1:N, ]), sd_calib=sd(fs[1:N, ]))

## shrinkage, but generally different from calibrated sample
c(mean_exper=mean(fs[-c(1:N), ]), sd_exper=sd(fs[-c(1:N), ]))
```

fixef *Compute latent regression fixed effect expected values*

## Description

Create expected values for fixed effects parameters in latent regression models.

## Usage

```
fixef(x)
```

## Arguments

| | |
|---|---|
| x | an estimated model object from the mixedmirt or mirt function |

## Author(s)

Phil Chalmers <rphilip.chalmers@gmail.com>

## References

Chalmers, R. P. (2012). mirt: A Multidimensional Item Response Theory Package for the R Environment. *Journal of Statistical Software, 48*(6), 1-29. doi:10.18637/jss.v048.i06

Chalmers, R. P. (2015). Extended Mixed-Effects Item Response Models with the MH-RM Algorithm. *Journal of Educational Measurement, 52*, 200-222. doi:10.1111/jedm.12072

## See Also

mirt, mixedmirt

## Examples

```
#simulate data
set.seed(1234)
N <- 1000

# covariates
X1 <- rnorm(N); X2 <- rnorm(N)
covdata <- data.frame(X1, X2)
Theta <- matrix(0.5 * X1 + -1 * X2 + rnorm(N, sd = 0.5))

#items and response data
a <- matrix(1, 20); d <- matrix(rnorm(20))
dat <- simdata(a, d, 1000, itemtype = '2PL', Theta=Theta)
```

```
#conditional model using X1 and X2 as predictors of Theta
mod1 <- mirt(dat, 1, 'Rasch', covdata=covdata, formula = ~ X1 + X2)

#latent regression fixed effects (i.e., expected values)
fe <- fixef(mod1)
head(fe)

# with mixedmirt()
mod1b <- mixedmirt(dat, covdata, 1, lr.fixed = ~ X1 + X2, fixed = ~ 0 + items)
fe2 <- fixef(mod1b)
head(fe2)
```

---

`fscores` *Compute factor score estimates (a.k.a, ability estimates, latent trait estimates, etc)*

---

## Description

Computes MAP, EAP, ML (Embretson & Reise, 2000), EAP for sum-scores (Thissen et al., 1995), or WLE (Warm, 1989) factor scores with a multivariate normal prior distribution using equally spaced quadrature. EAP scores for models with more than three factors are generally not recommended since the integration grid becomes very large, resulting in slower estimation and less precision if the `quadpts` are too low. Therefore, MAP scores should be used instead of EAP scores for higher dimensional models. Multiple imputation variants are possible for each estimator if a parameter information matrix was computed, which are useful if the sample size/number of items were small. As well, if the model contained latent regression predictors this information will be used in computing MAP and EAP estimates (for these models, `full.scores=TRUE` will always be used). Finally, plausible value imputation is also available, and will also account for latent regression predictor effects.

## Usage

```
fscores(
  object,
  method = "EAP",
  full.scores = TRUE,
  rotate = "oblimin",
  Target = NULL,
  response.pattern = NULL,
  append_response.pattern = FALSE,
  na.rm = FALSE,
  plausible.draws = 0,
  plausible.type = "normal",
  quadpts = NULL,
  item_weights = rep(1, extract.mirt(object, "nitems")),
```

```
  returnER = FALSE,
  T_as_X = FALSE,
  EAPsum.scores = FALSE,
  return.acov = FALSE,
  mean = NULL,
  cov = NULL,
  covdata = NULL,
  verbose = interactive(),
  full.scores.SE = FALSE,
  theta_lim = c(-6, 6),
  MI = 0,
  use_dentype_estimate = FALSE,
  QMC = FALSE,
  custom_den = NULL,
  custom_theta = NULL,
  expected.info = FALSE,
  min_expected = 1,
  max_theta = 20,
  start = NULL,
  ...
)
```

## Arguments

| | |
|---|---|
| object | a computed model object of class SingleGroupClass, MultipleGroupClass, or DiscreteClass |
| method | type of factor score estimation method. Can be:<br>• "EAP" for the expected a-posteriori (default). For models fit using mdirt this will return the posterior classification probabilities<br>• "EAP_general" for the expected a-posteriori estimates using dimension reduction to obtain estimates of the general factor(s) only (Gibbons, et al., 2007). This is generally more accurate than the 'EAP' approach as the dimensions to integrate are reduced to the number of general factors plus one, and therefore higher number of quadrature nodes can be used (see quadpts for the defaults. For a bifactor model, for example, this method requires exactly 2 dimensions). Model must have been fit using bfactor<br>• "MAP" for the maximum a-posteriori (i.e, Bayes modal)<br>• "ML" for maximum likelihood<br>• "WLE" or "WML" for weighted maximum-likelihood estimation<br>• "EAPsum" and "EAPsum_2.0" for the expected a-posteriori for each sum score (the latter is relevant for bifactor models)<br>• "plausible" for a single plausible value imputation for each case. This is equivalent to setting plausible.draws = 1<br>• "classify" for the posteriori classification probabilities (only applicable when the input model was of class MixtureClass) |
| full.scores | if FALSE then a summary table with factor scores for each unique pattern is displayed as a formatted matrix object. Otherwise, a matrix of factor scores for each response pattern in the data is returned (default) |

rotate prior rotation to be used when estimating the factor scores. See summary-method for details. If the object is not an exploratory model then this argument is ignored

Target target rotation; see summary-method for details

response.pattern an optional argument used to calculate the factor scores and standard errors for a given response vector or matrix/data.frame

append_response.pattern logical; should the inputs from response.pattern also be appended to the factor score output?

na.rm logical; remove rows with any missing values? This is generally not required due to the nature of computing factors scores, however for the "EAPsum" method this may be necessary to ensure that the sum-scores correspond to the same composite score

plausible.draws number of plausible values to draw for future researchers to perform secondary analyses of the latent trait scores. Typically used in conjunction with latent regression predictors (see mirt for details), but can also be generated when no predictor variables were modelled. If plausible.draws is greater than 0 a list of plausible values will be returned

plausible.type type of plausible values to obtain. Can be either 'normal' (default) to use a normal approximation based on the ACOV matrix, or 'MH' to obtain Metropolis-Hastings samples from the posterior (silently passes object to mirt, therefore arguments like technical can be supplied to increase the number of burn-in draws and discarded samples)

quadpts number of quadrature to use per dimension. If not specified, a suitable one will be created which decreases as the number of dimensions increases (and therefore for estimates such as EAP, will be less accurate). This is determined from the switch statement quadpts <- switch(as.character(nfact), '1'=121, '2'=61, '3'=31, '4'=19, '5'=11, '6'=7, 5)

item_weights a user-defined weight vector used in the likelihood expressions to add more/less weight for a given observed response. Default is a vector of 1's, indicating that all the items receive the same weight

returnER logical; return empirical reliability (also known as marginal reliability) estimates as a numeric values?

T_as_X logical; should the observed variance be equal to var(X) = var(T) + E(E^2) or var(X) = var(T) when computing empirical reliability estimates? Default (FALSE) uses the former

EAPsum.scores logical; include the model-implied expected values and variance for the item and total scores when using method = 'EAPsum' or method = 'EAPsum_2.0' with full.scores=FALSE? This information is included in the hidden 'fit' attribute which can be extracted via attr(., 'fit') for later use

return.acov logical; return a list containing covariance matrices instead of factors scores? impute = TRUE not supported with this option

mean a vector for custom latent variable means. If NULL, the default for 'group' values from the computed mirt object will be used

| | |
|---|---|
| cov | a custom matrix of the latent variable covariance matrix. If NULL, the default for 'group' values from the computed mirt object will be used |
| covdata | when latent regression model has been fitted, and the response.pattern input is used to score individuals, then this argument is used to include the latent regression covariate terms for each row vector supplied to response.pattern |
| verbose | logical; print verbose output messages? |
| full.scores.SE | logical; when full.scores == TRUE, also return the standard errors associated with each respondent? Default is FALSE |
| theta_lim | lower and upper range to evaluate latent trait integral for each dimension. If omitted, a range will be generated automatically based on the number of dimensions |
| MI | a number indicating how many multiple imputation draws to perform. Default is 0, indicating that no MI draws will be performed |
| use_dentype_estimate | logical; if the density of the latent trait was estimated in the model (e.g., via Davidian curves or empirical histograms), should this information be used to compute the latent trait estimates? Only applicable for EAP-based estimates (EAP, EAPsum, and plausible) |
| QMC | logical; use quasi-Monte Carlo integration? If quadpts is omitted the default number of nodes is 5000 |
| custom_den | a function used to define the integration density (if required). The NULL default assumes that the multivariate normal distribution with the 'GroupPars' hyper-parameters are used. At the minimum must be of the form:<br>function(Theta, ...)<br>where Theta is a matrix of latent trait values (will be a grid of values if method == 'EAPsum' or method == 'EAP', otherwise Theta will have only 1 row). Additional arguments may included and are caught through the fscores(...) input. The function *must* return a numeric vector of density weights (one for each row in Theta) |
| custom_theta | a matrix of custom integration nodes to use instead of the default, where each column corresponds to the respective dimension in the model |
| expected.info | logical; instead of using the observed information when method = 'ML' or method = 'WLE' use the expected information computed via testinfo? Only currently supported for unidimensional models |
| min_expected | when computing goodness of fit tests when method = 'EAPsum', this value is used to collapse across the conditioned total scores until the expected values are greater than this value. Note that this only affect the goodness of fit tests and not the returned EAP for sum scores table |
| max_theta | the maximum/minimum value any given factor score estimate will achieve using any modal estimator method (e.g., MAP, WLE, ML) |
| start | a matrix of starting values to use for iterative estimation methods. Default will start at a vector of 0's for each response pattern, or will start at the EAP estimates (unidimensional models only). Must be in the form that matches full.scores = FALSE (mostly used in the mirtCAT package) |
| ... | additional arguments to be passed to nlm |

## Details

The function will return either a table with the computed scores and standard errors, the original data matrix with scores appended to the rightmost column, or the scores only. By default the latent means and covariances are determined from the estimated object, though these can be overwritten. Iterative estimation methods can be estimated in parallel to decrease estimation times if a `mirtCluster` object is available.

If the input object is a discrete latent class object estimated from `mdirt` then the returned results will be with respect to the posterior classification for each individual. The method inputs for `'DiscreteClass'` objects may only be `'EAP'`, for posterior classification of each response pattern, `'EAPsum'` and `'EAPsum_2.0'` for posterior classification based on the raw sum-score (`'EAPsum_2.0'` results in marginal estimates of the primary dimensions only using the 2.0 algorithm described by Cai, 2015). For more information on these algorithms refer to the `mirtCAT` package and the associated JSS paper (Chalmers, 2016).

## Author(s)

Phil Chalmers <rphilip.chalmers@gmail.com>

## References

Cai, L. (2015). Lord–Wingersky algorithm version 2.0 for hierarchical item factor models with applications in test scoring, scale alignment, and model fit testing. *Psychometrika, 80*(2), 535–559.

Chalmers, R. P. (2012). mirt: A Multidimensional Item Response Theory Package for the R Environment. *Journal of Statistical Software, 48*(6), 1-29. doi:10.18637/jss.v048.i06

Chalmers, R. P. (2016). Generating Adaptive and Non-Adaptive Test Interfaces for Multidimensional Item Response Theory Applications. *Journal of Statistical Software, 71*(5), 1-39. doi:10.18637/jss.v071.i05

Embretson, S. E. & Reise, S. P. (2000). Item Response Theory for Psychologists. Erlbaum.

Gibbons. R. D., Bock, R. D., Hedeker, D., et al. (2007). Full-Information item bifactor analysis of graded response data. *Applied Psychological Measurement, 31*(1), 4-19.

Thissen, D., Pommerich, M., Billeaud, K., & Williams, V. S. L. (1995). Item Response Theory for Scores on Tests Including Polytomous Items with Ordered Responses. *Applied Psychological Measurement, 19*, 39-49.

Warm, T. A. (1989). Weighted likelihood estimation of ability in item response theory. *Psychometrika, 54*, 427-450.

## See Also

`averageMI`

## Examples

```
mod <- mirt(Science)
tabscores <- fscores(mod, full.scores = FALSE)
head(tabscores)

# convert scores into expected total score information with 95% CIs
```

```
E.total <- expected.test(mod, Theta=tabscores[,'F1'])
E.total_2.5 <- expected.test(mod, Theta=tabscores[,'F1'] +
                                     tabscores[,'SE_F1'] * qnorm(.05/2))
E.total_97.5 <- expected.test(mod, Theta=tabscores[,'F1'] +
                                     tabscores[,'SE_F1'] * qnorm(1-.05/2))

data.frame(Total_score=rowSums(tabscores[,1:4]),
           E.total, E.total_2.5, E.total_97.5) |> head()


fullscores <- fscores(mod)
fullscores_with_SE <- fscores(mod, full.scores.SE=TRUE)
head(fullscores)
head(fullscores_with_SE)

# convert scores into expected total score information with 95% CIs
E.total <- expected.test(mod, Theta=fullscores[,'F1'])
E.total_2.5 <- expected.test(mod, Theta=fullscores_with_SE[,'F1'] +
                                fullscores_with_SE[,'SE_F1'] * qnorm(.05/2))
E.total_97.5 <- expected.test(mod, Theta=fullscores_with_SE[,'F1'] +
                               fullscores_with_SE[,'SE_F1'] * qnorm(1-.05/2))

data.frame(Total_score=rowSums(Science),
           E.total, E.total_2.5, E.total_97.5) |> head()

# change method argument to use MAP estimates
fullscores <- fscores(mod, method='MAP')
head(fullscores)

# calculate MAP for a given response vector
fscores(mod, method='MAP', response.pattern = c(1,2,3,4))
# or matrix
fscores(mod, method='MAP', response.pattern = rbind(c(1,2,3,4), c(2,2,1,3)))

# return only the scores and their SEs
fscores(mod, method='MAP', response.pattern = c(1,2,3,4))

# use custom latent variable properties (diffuse prior for MAP is very close to ML)
fscores(mod, method='MAP', cov = matrix(1000), full.scores = FALSE)
fscores(mod, method='ML', full.scores = FALSE)

# EAPsum table of values based on total scores
(fs <- fscores(mod, method = 'EAPsum', full.scores = FALSE))

# convert expected counts back into marginal probability distribution
within(fs,
   `P(y)` <- expected / sum(observed))

# list of error VCOV matrices for EAPsum (works for other estimators as well)
acovs <- fscores(mod, method = 'EAPsum', full.scores = FALSE, return.acov = TRUE)
acovs

# WLE estimation, run in parallel using available cores
```

```
if(interactive()) mirtCluster()
head(fscores(mod, method='WLE', full.scores = FALSE))

# multiple imputation using 30 draws for EAP scores. Requires information matrix
mod <- mirt(Science, 1, SE=TRUE)
fs <- fscores(mod, MI = 30)
head(fs)

# plausible values for future work
pv <- fscores(mod, plausible.draws = 5)
lapply(pv, function(x) c(mean=mean(x), var=var(x), min=min(x), max=max(x)))

## define a custom_den function (*must* return a numeric vector).
#  EAP with a uniform prior between -3 and 3
fun <- function(Theta, ...) as.numeric(dunif(Theta, min = -3, max = 3))
head(fscores(mod, custom_den = fun))

# compare EAP estimators with same modified prior
fun <- function(Theta, ...) as.numeric(dnorm(Theta, mean=.5))
head(fscores(mod, custom_den = fun))
head(fscores(mod, method = 'EAP', mean=.5))

## Not run:
# custom MAP prior: standard truncated normal between 5 and -2
library(msm)
# need the :: scope for parallel to see the function (not require if no mirtCluster() defined)
fun <- function(Theta, ...) msm::dtnorm(Theta, mean = 0, sd = 1, lower = -2, upper = 5)
head(fscores(mod, custom_den = fun, method = 'MAP', full.scores = FALSE))

## End(Not run)


####################
# scoring via response.pattern input (with latent regression structure)
# simulate data
set.seed(1234)
N <- 1000

# covariates
X1 <- rnorm(N); X2 <- rnorm(N)
covdata <- data.frame(X1, X2)
Theta <- matrix(0.5 * X1 + -1 * X2 + rnorm(N, sd = 0.5))

# items and response data
a <- matrix(1, 20); d <- matrix(rnorm(20))
dat <- simdata(a, d, 1000, itemtype = '2PL', Theta=Theta)

# conditional model using X1 and X2 as predictors of Theta
mod <- mirt(dat, 1, 'Rasch', covdata=covdata, formula = ~ X1 + X2)
coef(mod, simplify=TRUE)

# all EAP estimates that include latent regression information
fs <- fscores(mod, full.scores.SE=TRUE)
```

```
head(fs)

# score only two response patterns
rp <- dat[1:2, ]
cd <- covdata[1:2, ]

fscores(mod, response.pattern=rp, covdata=cd)
fscores(mod, response.pattern=rp[2,], covdata=cd[2,]) # just one pattern
```

---

`gen.difficulty` *Generalized item difficulty summaries*

---

## Description

Function provides the four generalized item difficulty representations for polytomous response models described by Ali, Chang, and Anderson (2015). These estimates are used to gauge how difficult a polytomous item may be.

## Usage

```
gen.difficulty(mod, type = "IRF", interval = c(-30, 30), ...)
```

## Arguments

| | |
|---|---|
| `mod` | a single factor model estimated by `mirt` |
| `type` | type of generalized difficulty parameter to report. Can be `'IRF'` to use the item response function (default), `'mean'` to find the average of the difficulty estimates, `'median'` the median of the difficulty estimates, and `'trimmed'` to find the trimmed mean after removing the first and last difficulty estimates |
| `interval` | interval range to search for `'IRF'` type |
| `...` | additional arguments to pass to `uniroot` |

## Author(s)

Phil Chalmers <rphilip.chalmers@gmail.com>

## References

Ali, U. S., Chang, H.-H., & Anderson, C. J. (2015). *Location indices for ordinal polytomous items based on item response theory* (Research Report No. RR-15-20). Princeton, NJ: Educational Testing Service. http://dx.doi.org/10.1002/ets2.12065

Chalmers, R. P. (2012). mirt: A Multidimensional Item Response Theory Package for the R Environment. *Journal of Statistical Software, 48*(6), 1-29. doi:10.18637/jss.v048.i06

## Examples

```
mod <- mirt(Science, 1)
coef(mod, simplify=TRUE, IRTpars = TRUE)$items

gen.difficulty(mod)
gen.difficulty(mod, type = 'mean')

# also works for dichotomous items (though this is unnecessary)
dat <- expand.table(LSAT7)
mod <- mirt(dat, 1)
coef(mod, simplify=TRUE, IRTpars = TRUE)$items

gen.difficulty(mod)
gen.difficulty(mod, type = 'mean')
```

---

# head.mirt_matrix *Head generic for customized matrix console output*

---

## Description

Provides a nicer output for most printed `matrix` objects defined by functions in `mirt`.

## Usage

```
## S3 method for class 'mirt_matrix'
head(x, digits = 3, ...)
```

## Arguments

| | |
|---|---|
| `x` | object of class `'mirt_matrix'` |
| `digits` | number of digits to round |
| `...` | additional arguments passed to `print(...)` |

---

# imputeMissing *Imputing plausible data for missing values*

---

## Description

Given an estimated model from any of mirt's model fitting functions and an estimate of the latent trait, impute plausible missing data values. Returns the original data in a `data.frame` without any NA values. If a list of `Theta` values is supplied then a list of complete datasets is returned instead.

## Usage

```
imputeMissing(x, Theta, warn = TRUE, ...)
```

## Arguments

| | |
|---|---|
| x | an estimated model x from the mirt package |
| Theta | a matrix containing the estimates of the latent trait scores (e.g., via fscores) |
| warn | logical; print warning messages? |
| ... | additional arguments to pass |

## Author(s)

Phil Chalmers <rphilip.chalmers@gmail.com>

## References

Chalmers, R. P. (2012). mirt: A Multidimensional Item Response Theory Package for the R Environment. *Journal of Statistical Software, 48*(6), 1-29. doi:10.18637/jss.v048.i06

## Examples

```
dat <- expand.table(LSAT7)
(original <- mirt(dat, 1))
NAperson <- sample(1:nrow(dat), 20, replace = TRUE)
NAitem <- sample(1:ncol(dat), 20, replace = TRUE)
for(i in 1:20)
    dat[NAperson[i], NAitem[i]] <- NA
(mod <- mirt(dat, 1))
scores <- fscores(mod, method = 'MAP')

# re-estimate imputed dataset (good to do this multiple times and average over)
fulldata <- imputeMissing(mod, scores)
(fullmod <- mirt(fulldata, 1))

# with multipleGroup
set.seed(1)
group <- sample(c('group1', 'group2'), 1000, TRUE)
mod2 <- multipleGroup(dat, 1, group, TOL=1e-2)
fs <- fscores(mod2)
fulldata2 <- imputeMissing(mod2, fs)
```

itemfit *Item fit statistics*

## Description

Computes item-fit statistics for a variety of unidimensional and multidimensional models. Poorly fitting items should be inspected with the empirical plots/tables for unidimensional models, otherwise itemGAM can be used to diagnose where the functional form of the IRT model was misspecified, or models can be refit using more flexible semi-parametric response models (e.g., itemtype = 'spline'). If the latent trait density was approximated (e.g., Davidian curves, Empirical histograms, etc) then passing use_dentype_estimate = TRUE will use the internally saved quadrature and density components (where applicable). Currently, only S-X2 statistic supported for mixture IRT models. Finally, where applicable the root mean-square error of approximation (RMSEA) is reported to help gauge the magnitude of item misfit.

## Usage

```
itemfit(
  x,
  fit_stats = "S_X2",
  which.items = 1:extract.mirt(x, "nitems"),
  na.rm = FALSE,
  p.adjust = "none",
  group.bins = 10,
  group.size = NA,
  group.fun = mean,
  mincell = 1,
  mincell.X2 = 2,
  return.tables = FALSE,
  pv_draws = 30,
  boot = 1000,
  boot_dfapprox = 200,
  S_X2.plot = NULL,
  S_X2.plot_raw.score = TRUE,
  ETrange = c(-2, 2),
  ETpoints = 11,
  empirical.plot = NULL,
  empirical.CI = 0.95,
  empirical.poly.collapse = FALSE,
  method = "EAP",
  Theta = NULL,
  par.strip.text = list(cex = 0.7),
 par.settings = list(strip.background = list(col = "#9ECAE1"), strip.border = list(col =
    "black")),
  auto.key = list(space = "right", points = FALSE, lines = TRUE),
  ...
)
```

## Arguments

| | |
|---|---|
| x | a computed model object of class SingleGroupClass, MultipleGroupClass, or DiscreteClass |
| fit_stats | a character vector indicating which fit statistics should be computed. Supported inputs are:<br>• 'S_X2' : Orlando and Thissen (2000, 2003) and Kang and Chen's (2007) signed chi-squared test (default)<br>• 'Zh' : Drasgow, Levine, & Williams (1985) Zh<br>• 'X2' : Bock's (1972) chi-squared method. The default inputs compute Yen's (1981) Q1 variant of the X2 statistic (i.e., uses a fixed group.bins = 10). However, Bock's group-size variable median-based method can be computed by passing group.fun = median and modifying the group.size input to the desired number of bins<br>• 'G2' : McKinley & Mills (1985) G2 statistic (similar method to Q1, but with the likelihood-ratio test).<br>• 'PV_Q1' : Chalmers and Ng's (2017) plausible-value variant of the Q1 statistic.<br>• 'PV_Q1*' : Chalmers and Ng's (2017) plausible-value variant of the Q1 statistic that uses parametric bootstrapping to obtain a suitable empirical distribution.<br>• 'X2*' : Stone's (2000) fit statistics that require parametric bootstrapping<br>• 'X2*_df' : Stone's (2000) fit statistics that require parametric bootstrapping to obtain scaled versions of the X2* and degrees of freedom<br>• 'infit' : Compute the infit and outfit statistics<br>Note that 'S_X2' and 'Zh' cannot be computed when there are missing response data (i.e., will require multiple-imputation/row-removal techniques). |
| which.items | an integer vector indicating which items to test for fit. Default tests all possible items |
| na.rm | logical; remove rows with any missing values? This is required for methods such as S-X2 because they require the "EAPsum" method from fscores |
| p.adjust | method to use for adjusting all p-values for each respective item fit statistic (see p.adjust for available options). Default is 'none' |
| group.bins | the number of bins to use for X2 and G2. For example, setting group.bins = 10 will will compute Yen's (1981) Q1 statistic when 'X2' is requested |
| group.size | approximate size of each group to be used in calculating the $\chi^2$ statistic. The default NA disables this command and instead uses the group.bins input to try and construct equally sized bins |
| group.fun | function used when 'X2' or 'G2' are computed. Determines the central tendency measure within each partitioned group. E.g., setting group.fun = median will obtain the median of each respective ability estimate in each subgroup (this is what was used by Bock, 1972) |
| mincell | the minimum expected cell size to be used in the S-X2 computations. Tables will be collapsed across items first if polytomous, and then across scores if necessary |

| | |
|---|---|
| mincell.X2 | the minimum expected cell size to be used in the X2 computations. Tables will be collapsed if polytomous, however if this condition can not be met then the group block will be omitted in the computations |
| return.tables | logical; return tables when investigating 'X2', 'S_X2', and 'X2*'? |
| pv_draws | number of plausible-value draws to obtain for PV_Q1 and PV_Q1* |
| boot | number of parametric bootstrap samples to create for PV_Q1* and X2* |
| boot_dfapprox | number of parametric bootstrap samples to create for the X2*_df statistic to approximate the scaling factor for X2* as well as the scaled degrees of freedom estimates |
| S_X2.plot | argument input is the same as empirical.plot, however the resulting image is constructed according to the S-X2 statistic's conditional sum-score information |
| S_X2.plot_raw.score | logical; use the raw-score information in the plot in stead of the latent trait scale score? Default is FALSE |
| ETrange | range of integration nodes for Stone's X2* statistic |
| ETpoints | number of integration nodes to use for Stone's X2* statistic |
| empirical.plot | a single numeric value or character of the item name indicating which item to plot (via itemplot) and overlay with the empirical $\theta$ groupings (see empirical.CI). Useful for plotting the expected bins based on the 'X2' or 'G2' method |
| empirical.CI | a numeric value indicating the width of the empirical confidence interval ranging between 0 and 1 (default of 0 plots not interval). For example, a 95 interval would be plotted when empirical.CI = .95. Only applicable to dichotomous items |
| empirical.poly.collapse | logical; collapse polytomous item categories to for expected scoring functions for empirical plots? Default is FALSE |
| method | type of factor score estimation method. See fscores for more detail |
| Theta | a matrix of factor scores for each person used for statistics that require empirical estimates. If supplied, arguments typically passed to fscores() will be ignored and these values will be used instead. Also required when estimating statistics with missing data via imputation |
| par.strip.text | plotting argument passed to lattice |
| par.settings | plotting argument passed to lattice |
| auto.key | plotting argument passed to lattice |
| ... | additional arguments to be passed to fscores() and lattice |

## Author(s)

Phil Chalmers <rphilip.chalmers@gmail.com>

## References


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## See Also

personfit, itemGAM

## Examples

```
P <- function(Theta){exp(Theta^2 * 1.2 - 1) / (1 + exp(Theta^2 * 1.2 - 1))}

#make some data
set.seed(1234)
a <- matrix(rlnorm(20, meanlog=0, sdlog = .1),ncol=1)
d <- matrix(rnorm(20),ncol=1)
Theta <- matrix(rnorm(2000))
items <- rep('2PL', 20)
ps <- P(Theta)
baditem <- numeric(2000)
for(i in 1:2000)
   baditem[i] <- sample(c(0,1), 1, prob = c(1-ps[i], ps[i]))
data <- cbind(simdata(a,d, 2000, items, Theta=Theta), baditem=baditem)

x <- mirt(data, 1)
```

```
raschfit <- mirt(data, 1, itemtype='Rasch')
fit <- itemfit(x)
fit

# p-value adjustment
itemfit(x, p.adjust='fdr')

# two different fit stats (with/without p-value adjustment)
itemfit(x, c('S_X2' ,'X2'), p.adjust='fdr')
itemfit(x, c('S_X2' ,'X2'))

# Conditional sum-score plot from S-X2 information
itemfit(x, S_X2.plot = 1) # good fit
itemfit(x, S_X2.plot = 2) # good fit
itemfit(x, S_X2.plot = 21) # bad fit

itemfit(x, 'X2') # just X2
itemfit(x, 'X2', method = 'ML') # X2 with maximum-likelihood estimates for traits
itemfit(x, group.bins=15, empirical.plot = 1, method = 'ML') #empirical item plot with 15 points
itemfit(x, group.bins=15, empirical.plot = 21, method = 'ML')

# PV and X2* statistics (parametric bootstrap stats not run to save time)
itemfit(x, 'PV_Q1')

if(interactive()) mirtCluster() # improve speed of bootstrap samples by running in parallel
# itemfit(x, 'PV_Q1*')
# itemfit(x, 'X2*') # Stone's 1993 statistic
# itemfit(x, 'X2*_df') # Stone's 2000 scaled statistic with df estimate

# empirical tables for X2 statistic
tabs <- itemfit(x, 'X2', return.tables=TRUE, which.items = 1)
tabs

#infit/outfit statistics. method='ML' agrees better with eRm package
itemfit(raschfit, 'infit', method = 'ML') #infit and outfit stats

#same as above, but inputting ML estimates instead (saves time for re-use)
Theta <- fscores(raschfit, method = 'ML')
itemfit(raschfit, 'infit', Theta=Theta)
itemfit(raschfit, empirical.plot=1, Theta=Theta)
itemfit(raschfit, 'X2', return.tables=TRUE, Theta=Theta, which.items=1)

# fit a new more flexible model for the mis-fitting item
itemtype <- c(rep('2PL', 20), 'spline')
x2 <- mirt(data, 1, itemtype=itemtype)
itemfit(x2)
itemplot(x2, 21)
anova(x, x2)

#------------------------------------------------------------

#similar example to Kang and Chen 2007
a <- matrix(c(.8,.4,.7, .8, .4, .7, 1, 1, 1, 1))
```

```
d <- matrix(rep(c(2.0,0.0,-1,-1.5),10), ncol=4, byrow=TRUE)
dat <- simdata(a,d,2000, itemtype = rep('graded', 10))
head(dat)

mod <- mirt(dat, 1)
itemfit(mod)
itemfit(mod, 'X2') # less useful given inflated Type I error rates
itemfit(mod, empirical.plot = 1)
itemfit(mod, empirical.plot = 1, empirical.poly.collapse=TRUE)

# collapsed tables (see mincell.X2) for X2 and G2
itemfit(mod, 'X2', return.tables = TRUE, which.items = 1)

mod2 <- mirt(dat, 1, 'Rasch')
itemfit(mod2, 'infit', method = 'ML')

# massive list of tables for S-X2
tables <- itemfit(mod, return.tables = TRUE)

#observed and expected total score patterns for item 1 (post collapsing)
tables$O[[1]]
tables$E[[1]]

# can also select specific items
# itemfit(mod, return.tables = TRUE, which.items=1)

# fit stats with missing data (run in parallel using all cores)
dat[sample(1:prod(dim(dat)), 100)] <- NA
raschfit <- mirt(dat, 1, itemtype='Rasch')

# use only valid data by removing rows with missing terms
itemfit(raschfit, c('S_X2', 'infit'), na.rm = TRUE)

# note that X2, G2, PV-Q1, and X2* do not require complete datasets
thetas <- fscores(raschfit, method = 'ML') # save for faster computations
itemfit(raschfit, c('X2', 'G2'), Theta=thetas)
itemfit(raschfit, empirical.plot=1, Theta=thetas)
itemfit(raschfit, 'X2', return.tables=TRUE, which.items=1, Theta=thetas)
```

---

`itemGAM` *Parametric smoothed regression lines for item response probability functions*

---

## Description

This function uses a generalized additive model (GAM) to estimate response curves for items that do not seem to fit well in a given model. Using a stable axillary model, traceline functions for poorly

fitting dichotomous or polytomous items can be inspected using point estimates (or plausible values) of the latent trait. Plots of the tracelines and their associated standard errors are available to help interpret the misfit. This function may also be useful when adding new items to an existing, well established set of items, especially when the parametric form of the items under investigation are unknown.

## Usage

```
itemGAM(
  item,
  Theta,
  formula = resp ~ s(Theta, k = 10),
  CI = 0.95,
  theta_lim = c(-3, 3),
  return.models = FALSE,
  ...
)

## S3 method for class 'itemGAM'
plot(
  x,
  y = NULL,
  par.strip.text = list(cex = 0.7),
 par.settings = list(strip.background = list(col = "#9ECAE1"), strip.border = list(col =
    "black")),
  auto.key = list(space = "right", points = FALSE, lines = TRUE),
  ...
)
```

## Arguments

| | |
|---|---|
| item | a single poorly fitting item to be investigated. Can be a vector or matrix |
| Theta | a list or matrix of latent trait estimates typically returned from fscores |
| formula | an R formula to be passed to the gam function. Default fits a spline model with 10 nodes. For multidimensional models, the traits are assigned the names 'Theta1', 'Theta2', ..., 'ThetaN' |
| CI | a number ranging from 0 to 1 indicating the confidence interval range. Default provides the 95 percent interval |
| theta_lim | range of latent trait scores to be evaluated |
| return.models | logical; return a list of GAM models for each category? Useful when the GAMs should be inspected directly, but also when fitting multidimensional models (this is set to TRUE automatically for multidimensional models) |
| ... | additional arguments to be passed to gam or lattice |
| x | an object of class 'itemGAM' |
| y | a NULL value ignored by the plotting function |
| par.strip.text | plotting argument passed to lattice |

| | |
|---|---|
| par.settings | plotting argument passed to lattice |
| auto.key | plotting argument passed to lattice |

## Author(s)

Phil Chalmers <rphilip.chalmers@gmail.com>

## References

Chalmers, R. P. (2012). mirt: A Multidimensional Item Response Theory Package for the R Environment. *Journal of Statistical Software, 48*(6), 1-29. doi:10.18637/jss.v048.i06

## See Also

itemfit

## Examples

```
set.seed(10)
N <- 1000
J <- 30

a <- matrix(1, J)
d <- matrix(rnorm(J))
Theta <- matrix(rnorm(N, 0, 1.5))
dat <- simdata(a, d, N, itemtype = '2PL', Theta=Theta)

# make a bad item
ps <- exp(Theta^2 + Theta) / (1 + exp(Theta^2 + Theta))
item1 <- sapply(ps, function(x) sample(c(0,1), size = 1, prob = c(1-x, x)))

ps2 <- exp(2 * Theta^2 + Theta + .5 * Theta^3) / (1 + exp(2 * Theta^2 + Theta + .5 * Theta^3))
item2 <- sapply(ps2, function(x) sample(c(0,1), size = 1, prob = c(1-x, x)))

# how the actual item looks in the population
plot(Theta, ps, ylim = c(0,1))
plot(Theta, ps2, ylim = c(0,1))

baditems <- cbind(item1, item2)
newdat <- cbind(dat, baditems)

badmod <- mirt(newdat, 1)
itemfit(badmod) #clearly a bad fit for the last two items
mod <- mirt(dat, 1) #fit a model that does not contain the bad items
itemfit(mod)

## Not run:
#### Pure non-parametric way of investigating the items
library(KernSmoothIRT)
ks <- ksIRT(newdat, rep(1, ncol(newdat)), 1)
plot(ks, item=c(1,31,32))
```

```
par(ask=FALSE)

## End(Not run)

# Using point estimates from the model
Theta <- fscores(mod)
IG0 <- itemGAM(dat[,1], Theta) #good item
IG1 <- itemGAM(baditems[,1], Theta)
IG2 <- itemGAM(baditems[,2], Theta)
plot(IG0)
plot(IG1)
plot(IG2)

# same as above, but with plausible values to obtain the standard errors
set.seed(4321)
ThetaPV <- fscores(mod, plausible.draws=10)
IG0 <- itemGAM(dat[,1], ThetaPV) #good item
IG1 <- itemGAM(baditems[,1], ThetaPV)
IG2 <- itemGAM(baditems[,2], ThetaPV)
plot(IG0)
plot(IG1)
plot(IG2)

## for polytomous test items
SAT12[SAT12 == 8] <- NA
dat <- key2binary(SAT12,
                key = c(1,4,5,2,3,1,2,1,3,1,2,4,2,1,5,3,4,4,1,4,3,3,4,1,3,5,1,3,1,5,4,5))
dat <- dat[,-32]
mod <- mirt(dat, 1)

## Not run:
# Kernal smoothing is very sensitive to which category is selected as 'correct'
# 5th category as correct
ks <- ksIRT(cbind(dat, SAT12[,32]), c(rep(1, 31), 5), 1)
plot(ks, items = c(1,2,32))

# 3rd category as correct
ks <- ksIRT(cbind(dat, SAT12[,32]), c(rep(1, 31), 3), 1)
plot(ks, items = c(1,2,32))

## End(Not run)

# splines approach
Theta <- fscores(mod)
IG <- itemGAM(SAT12[,32], Theta)
plot(IG)

set.seed(1423)
ThetaPV <- fscores(mod, plausible.draws=10)
IG2 <- itemGAM(SAT12[,32], ThetaPV)
plot(IG2)

# assuming a simple increasing parametric form (like in a standard IRT model)
```

```
IG3 <- itemGAM(SAT12[,32], Theta, formula = resp ~ Theta)
plot(IG3)
IG3 <- itemGAM(SAT12[,32], ThetaPV, formula = resp ~ Theta)
plot(IG3)

### multidimensional example by returning the GAM objects
mod2 <- mirt(dat, 2)
Theta <- fscores(mod2)
IG4 <- itemGAM(SAT12[,32], Theta, formula = resp ~ s(Theta1, k=10) + s(Theta2, k=10),
   return.models=TRUE)
names(IG4)
plot(IG4[[1L]], main = 'Category 1')
plot(IG4[[2L]], main = 'Category 2')
plot(IG4[[3L]], main = 'Category 3')
```

---

`iteminfo` *Function to calculate item information*

---

## Description

Given an internal mirt item object extracted by using `extract.item`, compute the item information.

## Usage

```
iteminfo(x, Theta, degrees = NULL, total.info = TRUE, multidim_matrix = FALSE)
```

## Arguments

| | |
|---|---|
| `x` | an extracted internal mirt object containing item information (see `extract.item`) |
| `Theta` | a vector (unidimensional) or matrix (multidimensional) of latent trait values |
| `degrees` | a vector of angles in degrees that are between 0 and 90. Only applicable when the input object is multidimensional |
| `total.info` | logical; return the total information curve for the item? If `FALSE`, information curves for each category are returned as a matrix |
| `multidim_matrix` | logical; compute the information matrix for each row in `Theta`? If `Theta` contains more than 1 row then a list of matrices will be returned, otherwise if `Theta` has exactly one row then a matrix will be returned |

## Author(s)

Phil Chalmers <rphilip.chalmers@gmail.com>

## References

Chalmers, R. P. (2012). mirt: A Multidimensional Item Response Theory Package for the R Environment. *Journal of Statistical Software, 48*(6), 1-29. doi:10.18637/jss.v048.i06

## See Also

extract.item

## Examples

```
mod <- mirt(Science, 1)
extr.2 <- extract.item(mod, 2)
Theta <- matrix(seq(-4,4, by = .1))
info.2 <- iteminfo(extr.2, Theta)

#do something with the info?
plot(Theta, info.2, type = 'l', main = 'Item information')



#category information curves
cat.info <- iteminfo(extr.2, Theta, total.info = FALSE)
plot(Theta, cat.info[,1], type = 'l', ylim = c(0, max(cat.info)),
     ylab = 'info', main = 'Category information')
for(i in 2:ncol(cat.info))
   lines(Theta, cat.info[,i], col = i)

## Customized test information plot
T1 <- T2 <- 0
dat <- expand.table(LSAT7)
mod1 <- mirt(dat, 1)
mod2 <- mirt(dat, 1, 'Rasch')
for(i in 1:5){
  T1 <- T1 + iteminfo(extract.item(mod1, i), Theta)
  T2 <- T2 + iteminfo(extract.item(mod2, i), Theta)
}
plot(Theta, T2/T1, type = 'l', ylab = 'Relative Test Information', las = 1)
lines(Theta, T1/T1, col = 'red')

# multidimensional
mod <- mirt(dat, 2, TOL=1e-2)
ii <- extract.item(mod, 1)
Theta <- as.matrix(expand.grid(-4:4, -4:4))

iteminfo(ii, Theta, degrees=c(45,45)) # equal angle
iteminfo(ii, Theta, degrees=c(90,0)) # first dimension only

# information matrices
iteminfo(ii, Theta, multidim_matrix = TRUE)
iteminfo(ii, Theta[1, , drop=FALSE], multidim_matrix = TRUE)
```

---

itemplot *Displays item surface and information plots*

---

## Description

itemplot displays various item based IRT plots, with special options for plotting items that contain several 0 slope parameters. Supports up to three dimensional models.

## Usage

```
itemplot(
  object,
  item,
  type = "trace",
  degrees = 45,
  CE = FALSE,
  CEalpha = 0.05,
  CEdraws = 1000,
  drop.zeros = FALSE,
  theta_lim = c(-6, 6),
  empirical_proportions = FALSE,
  shiny = FALSE,
  rot = list(xaxis = -70, yaxis = 30, zaxis = 10),
  par.strip.text = list(cex = 0.7),
  npts = 200,
 par.settings = list(strip.background = list(col = "#9ECAE1"), strip.border = list(col =
    "black")),
  auto.key = list(space = "right", points = FALSE, lines = TRUE),
  ...
)
```

## Arguments

| | |
|---|---|
| object | a computed model object of class SingleGroupClass or MultipleGroupClass. Input may also be a list for comparing similar item types (e.g., 1PL vs 2PL) |
| item | a single numeric value, or the item name, indicating which item to plot |
| type | plot type to use, information ('info'), standard errors ('SE'), item trace lines ('trace'), cumulative probability plots to indicate thresholds ('threshold'), information and standard errors ('infoSE') or information and trace lines ('infotrace'), category and total information ('infocat'), relative efficiency lines ('RE'), expected score 'score', or information and trace line contours ('infocontour' and 'tracecontour'; not supported for MultipleGroupClass objects) |
| degrees | the degrees argument to be used if there are two or three factors. See iteminfo for more detail. A new vector will be required for three dimensional models to override the default |
| CE | logical; plot confidence envelope? |
| CEalpha | area remaining in the tail for confidence envelope. Default gives 95% confidence region |
| CEdraws | draws number of draws to use for confidence envelope |

| | |
|---|---|
| drop.zeros | logical; drop slope values that are numerically close to zero to reduce dimensionality? Useful in objects returned from bfactor or other confirmatory models that contain several zero slopes |
| theta_lim | lower and upper limits of the latent trait (theta) to be evaluated, and is used in conjunction with npts. Default uses c(-6,6) |
| empirical_proportions | logical; for unidimensional models, include the empirical proportion estimates from the E-table when type = 'trace'? |
| shiny | logical; run interactive display for item plots using the shiny interface. This primarily is an instructive tool for demonstrating how item response curves behave when adjusting their parameters |
| rot | a list of rotation coordinates to be used for 3 dimensional plots |
| par.strip.text | plotting argument passed to lattice |
| npts | number of quadrature points to be used for plotting features. Larger values make plots look smoother |
| par.settings | plotting argument passed to lattice |
| auto.key | plotting argument passed to lattice |
| ... | additional arguments to be passed to lattice and coef() |

## Author(s)

Phil Chalmers <rphilip.chalmers@gmail.com>

## References

Chalmers, R. P. (2012). mirt: A Multidimensional Item Response Theory Package for the R Environment. *Journal of Statistical Software, 48*(6), 1-29. doi:10.18637/jss.v048.i06

## Examples

```
data(LSAT7)
fulldata <- expand.table(LSAT7)
mod1 <- mirt(fulldata,SE=TRUE)
mod2 <- mirt(fulldata, itemtype = 'Rasch')
mod3 <- mirt(fulldata, 2)

itemplot(mod1, 2)
itemplot(mod1, 2, CE = TRUE)
itemplot(mod1, 2, type = 'info')
itemplot(mod1, 2, type = 'info', CE = TRUE)
itemplot(mod1, 2, empirical_proportions=TRUE)
itemplot(mod2, 2, empirical_proportions=TRUE)

mods <- list(twoPL = mod1, onePL = mod2)
itemplot(mods, 1, type = 'RE')

# multidimensional
```

```
itemplot(mod3, 4, type = 'info')
itemplot(mod3, 4, type = 'info',
  col.regions = colorRampPalette(c("white", "red"))(100))
itemplot(mod3, 4, type = 'infocontour')
itemplot(mod3, 4, type = 'tracecontour')

# polytomous items
pmod <- mirt(Science, 1, SE=TRUE)
itemplot(pmod, 3)
itemplot(pmod, 3, type = 'threshold')
itemplot(pmod, 3, CE = TRUE)
itemplot(pmod, 3, type = 'score')
itemplot(pmod, 3, type = 'score', CE = TRUE)
itemplot(pmod, 3, type = 'infotrace')
itemplot(pmod, 3, type = 'infocat')


# use the directlabels package to put labels on tracelines
library(directlabels)
plt <- itemplot(pmod, 3)
direct.label(plt, 'top.points')

# change colour theme of plots
bwtheme <- standard.theme("pdf", color=FALSE)
plot(pmod, type='trace', par.settings=bwtheme)
itemplot(pmod, 1, type = 'trace', par.settings=bwtheme)

# additional modifications can be made via update().
# See ?update.trellis for further documentation
(plt <- itemplot(pmod, 1))
update(plt, ylab = expression(Prob(theta))) # ylab changed

# infoSE plot
itemplot(pmod, 1, type = 'infoSE')

# uncomment to run interactive shiny applet
# itemplot(shiny = TRUE)
```

---

`itemstats` *Generic item summary statistics*

---

## Description

Function to compute generic item summary statistics that do not require prior fitting of IRT models. Contains information about sample sizes (`N`), number of observed categories (`K`), (standardized) coefficient alpha (and alpha if an item is removed; (`alpha_if_rm`)), mean/SD and frequency of total scores, reduced item-total correlations (`cor_if_rm`), average/sd of the correlation between

items, squared multiple correlation (smc), response frequencies, and conditional mean/sd information given the unweighted sum scores. Summary information involving the total scores only included for responses with no missing data to ensure the metric is meaningful, however standardized statistics (e.g., correlations) utilize all possible response information.

## Usage

```
itemstats(
  data,
  group = NULL,
  ts_fun = rowSums,
  use_ts = TRUE,
  itemfreq = "proportions",
  ts.tables = FALSE
)
```

## Arguments

| | |
|---|---|
| data | An object of class data.frame or matrix with the response patterns |
| group | optional grouping variable to condition on when computing summary information |
| ts_fun | function to use to build the linear composite score? Recommend inputs are rowSums (default) and rowMeans, and must include an argument na.rm, though users may define their own as well |
| use_ts | logical; include information that is conditional on a meaningful total score? |
| itemfreq | character vector indicting whether to include item response "proportions" or "counts" for each item. If set to 'none' then this will be omitted |
| ts.tables | logical; include mean/sd summary information pertaining to the unweighted total score? |

## Value

Returns a list containing the summary statistics

## Author(s)

Phil Chalmers <rphilip.chalmers@gmail.com>

## References

Chalmers, R. P. (2012). mirt: A Multidimensional Item Response Theory Package for the R Environment. *Journal of Statistical Software, 48*(6), 1-29. doi:10.18637/jss.v048.i06

## See Also

empirical_plot

## Examples

```
# dichotomous data example
LSAT7full <- expand.table(LSAT7)
head(LSAT7full)
itemstats(LSAT7full)
itemstats(LSAT7full, itemfreq='counts')

# composite score expressed as a mean
itemstats(LSAT7full)$overall
itemstats(LSAT7full, ts_fun=rowMeans)$overall

# behaviour with missing data
LSAT7full[1:5,1] <- NA
itemstats(LSAT7full)

# data with no meaningful total score
head(SAT12)
itemstats(SAT12, use_ts=FALSE)

# extra total scores tables
dat <- key2binary(SAT12,
                  key = c(1,4,5,2,3,1,2,1,3,1,2,4,2,1,
                          5,3,4,4,1,4,3,3,4,1,3,5,1,3,1,5,4,5))
itemstats(dat, ts.tables=TRUE)

# grouping information
group <- gl(2, 300, labels=c('G1', 'G2'))
itemstats(dat, group=group)


#####
# polytomous data example
itemstats(Science)

# polytomous data with missing
newScience <- Science
newScience[1:5,1] <- NA
itemstats(newScience)

# unequal categories
newScience[,1] <- ifelse(Science[,1] == 1, NA, Science[,1])
itemstats(newScience)

merged <- data.frame(LSAT7full[1:392,], Science)
itemstats(merged)
```

---

key2binary *Score a test by converting response patterns to binary data*

---

## Description

The key2binary function will convert response pattern data to a dichotomous format, given a response key.

## Usage

```
key2binary(fulldata, key, score_missing = FALSE)
```

## Arguments

| | |
|---|---|
| fulldata | an object of class data.frame, matrix, or table with the response patterns |
| key | a vector or matrix consisting of the 'correct' response to the items. Each value/row corresponds to each column in fulldata. If the input is a matrix, multiple scoring keys can be supplied for each item. NA values are used to indicate no scoring key (or in the case of a matrix input, no additional scoring keys) |
| score_missing | logical; should missing data elements be returned as incorrect (i.e., 0)? If FALSE, all missing data terms will be kept as missing |

## Value

Returns a numeric matrix with all the response patterns in dichotomous format

## Author(s)

Phil Chalmers <rphilip.chalmers@gmail.com>

## References

Chalmers, R. P. (2012). mirt: A Multidimensional Item Response Theory Package for the R Environment. *Journal of Statistical Software, 48*(6), 1-29. doi:10.18637/jss.v048.i06

## Examples

```
data(SAT12)
head(SAT12)
key <- c(1,4,5,2,3,1,2,1,3,1,2,4,2,1,5,3,4,4,1,4,3,3,4,1,3,5,1,3,1,5,4,5)

dicho.SAT12 <- key2binary(SAT12, key)
head(dicho.SAT12)

# multiple scoring keys
key2 <- cbind(c(1,4,5,2,3,1,2,1,3,1,2,4,2,1,5,3,4,4,1,4,3,3,4,1,3,5,1,3,1,5,4,5),
              c(2,3,NA,1,rep(NA, 28)))
dicho.SAT12 <- key2binary(SAT12, key2)

# keys from raw character responses
resp <- as.data.frame(matrix(c(
  "B","B","D","D","E",
  "B","A","D","D","E",
  "B","A","D","C","E",
```

```
  "D","D","D","C","E",
  "B","C","A","D","A"), ncol=5, byrow=TRUE))

key <- c("B", "D", "D", "C", "E")

d01 <- key2binary(resp, key)
head(d01)

# score/don't score missing values
resp[1,1] <- NA
d01NA <- key2binary(resp, key) # without scoring
d01NA

d01 <- key2binary(resp, key, score_missing = TRUE) # with scoring
d01
```

---

lagrange *Lagrange test for freeing parameters*

---

## Description

Lagrange (i.e., score) test to test whether parameters should be freed from a more constrained baseline model.

## Usage

```
lagrange(mod, parnum, SE.type = "Oakes", type = "Richardson", ...)
```

## Arguments

| | |
|---|---|
| mod | an estimated model |
| parnum | a vector, or list of vectors, containing one or more parameter locations/sets of locations to be tested. See objects returned from mod2values for the locations |
| SE.type | type of information matrix estimator to use. See mirt for further details |
| type | type of numerical algorithm passed to numerical_deriv to obtain the gradient terms |
| ... | additional arguments to pass to mirt |

## Author(s)

Phil Chalmers <rphilip.chalmers@gmail.com>

## References

Chalmers, R. P. (2012). mirt: A Multidimensional Item Response Theory Package for the R Environment. *Journal of Statistical Software, 48*(6), 1-29. doi:10.18637/jss.v048.i06

## See Also

wald

## Examples

```
dat <- expand.table(LSAT7)
mod <- mirt(dat, 1, 'Rasch')
(values <- mod2values(mod))

# test all fixed slopes individually
parnum <- values$parnum[values$name == 'a1']
lagrange(mod, parnum)

# compare to LR test for first two slopes
mod2 <- mirt(dat, 'F = 1-5
                   FREE = (1, a1)', 'Rasch')
coef(mod2, simplify=TRUE)$items
anova(mod, mod2)

mod2 <- mirt(dat, 'F = 1-5
                   FREE = (2, a1)', 'Rasch')
coef(mod2, simplify=TRUE)$items
anova(mod, mod2)

mod2 <- mirt(dat, 'F = 1-5
                   FREE = (3, a1)', 'Rasch')
coef(mod2, simplify=TRUE)$items
anova(mod, mod2)

# test slopes first two slopes and last three slopes jointly
lagrange(mod, list(parnum[1:2], parnum[3:5]))

# test all 5 slopes and first + last jointly
lagrange(mod, list(parnum[1:5], parnum[c(1, 5)]))

######
## DIF using score (Lagrange) test example

n <- 10
N <- 500

# generate 2PL data with DIF on Item_1's intercept in the focal group
a  <- matrix(rep(1, n), ncol = 1)
d1 <- matrix(rnorm(n), ncol = 1)
d2 <- d1
d2[1, 1] <- d1[1, 1] + 1          # shift Item_1's difficulty for the focal group

dat1 <- simdata(a, d1, N, itemtype = '2PL')
dat2 <- simdata(a, d2, N, itemtype = '2PL')
dat  <- rbind(dat1, dat2)
colnames(dat) <- paste0('Item_', 1:n)
```

```
group <- c(rep('Reference', N), rep('Focal', N))

# fully equality-constrained (no-DIF) baseline: all a1/d equated across
# groups, with the latent mean/variance freed in the focal group for
# identification (standard mirt DIF baseline)
mod_baseline <- multipleGroup(dat, model = 1, itemtype = '2PL', group = group,
                              invariance = c(colnames(dat), 'free_means', 'free_var'),
                              SE = TRUE)

values <- mod2values(mod_baseline)
(i1 <- values[values$item == 'Item_1', ])     # inspect parnum for this item

parnum_a1 <- i1$parnum[i1$name == 'a1' & i1$group == 'Focal']
parnum_d  <- i1$parnum[i1$name == 'd'& i1$group == 'Focal']

# joint 2-df score test: does freeing both a1 and d for Item_1 in the
# focal group improve fit relative to the constrained baseline?
lagrange(mod_baseline, list(c(parnum_a1, parnum_d)))

# compare to LR and Wald tests
mod_nest <- multipleGroup(dat, model = 1, itemtype = '2PL', group = group,
                            invariance = c(colnames(dat)[-1], 'free_means', 'free_var'),
                              SE = TRUE)
anova(mod_baseline, mod_nest)
wald(mod_nest)
wald(mod_nest, c('a1.1 = a1.43', 'd.2 = d.44'))

# separate 1-df tests: slope-only DIF vs. intercept-only DIF
lagrange(mod_baseline, list(parnum_a1, parnum_d))
```

---

`likert2int` *Convert ordered Likert-scale responses (character or factors) to integers*

---

## Description

Given a matrix or data.frame object consisting of Likert responses return an object of the same dimensions with integer values.

## Usage

```
likert2int(x, levels = NULL)
```

## Arguments

| | |
|---|---|
| `x` | a matrix of character values or data.frame of character/factor vectors |
| `levels` | a named character vector indicating which integer values should be assigned to which elements. If omitted, the order of the elements will be determined after converting each column in `x` to a factor variable |

## Author(s)

Phil Chalmers <rphilip.chalmers@gmail.com>

## References

Chalmers, R. P. (2012). mirt: A Multidimensional Item Response Theory Package for the R Environment. *Journal of Statistical Software, 48*(6), 1-29. doi:10.18637/jss.v048.i06

## See Also

key2binary, poly2dich

## Examples

```
# simulate data

dat1 <- matrix(sample(c('Disagree', 'Strongly Disagree', 'Agree',
                        'Neutral', 'Strongly Agree'), 1000*5, replace=TRUE),
               nrow=1000, ncol=5)
dat1[2,2] <- dat1[3,3] <- dat1[1,3] <- NA # NAs added for flavour
dat2 <- matrix(sample(c('D', 'SD', 'A', 'N', 'SA'), 1000*5, replace=TRUE),
               nrow=1000, ncol=5)
dat <- cbind(dat1, dat2)

# separately
intdat1 <- likert2int(dat1)
head(dat1)
head(intdat1)

# more useful with explicit levels
lvl1 <- c('Strongly Disagree'=1, 'Disagree'=2, 'Neutral'=3, 'Agree'=4,
          'Strongly Agree'=5)
intdat1 <- likert2int(dat1, levels = lvl1)
head(dat1)
head(intdat1)

# second data
lvl2 <- c('SD'=1, 'D'=2, 'N'=3, 'A'=4, 'SA'=5)
intdat2 <- likert2int(dat2, levels = lvl2)
head(dat2)
head(intdat2)

# full dataset (using both mapping schemes)
intdat <- likert2int(dat, levels = c(lvl1, lvl2))
head(dat)
head(intdat)


#####
# data.frame as input with ordered factors
```

```
dat1 <- data.frame(dat1)
dat2 <- data.frame(dat2)
dat.old <- cbind(dat1, dat2)
colnames(dat.old) <- paste0('Item_', 1:10)
str(dat.old) # factors are leveled alphabetically by default

# create explicit ordering in factor variables
for(i in 1:ncol(dat1))
   levels(dat1[[i]]) <- c('Strongly Disagree', 'Disagree', 'Neutral', 'Agree',
                          'Strongly Agree')

for(i in 1:ncol(dat2))
   levels(dat2[[i]]) <- c('SD', 'D', 'N', 'A', 'SA')

dat <- cbind(dat1, dat2)
colnames(dat) <- colnames(dat.old)
str(dat) # note ordering

intdat <- likert2int(dat)
head(dat)
head(intdat)
```

---

`logLik-method` *Extract log-likelihood*

---

## Description

Extract the observed-data log-likelihood.

## Usage

```
## S4 method for signature 'SingleGroupClass'
logLik(object)
```

## Arguments

| | |
|---|---|
| `object` | an object of class `SingleGroupClass`, `MultipleGroupClass`, or `MixedClass` |

## References

Chalmers, R. P. (2012). mirt: A Multidimensional Item Response Theory Package for the R Environment. *Journal of Statistical Software, 48*(6), 1-29. doi:10.18637/jss.v048.i06

## Examples

```
x <- mirt(Science, 1)
logLik(x)
```

---

LSAT6 *Description of LSAT6 data*

---

## Description

Data from Thissen (1982); contains 5 dichotomously scored items obtained from the Law School Admissions Test, section 6.

## Author(s)

Phil Chalmers <rphilip.chalmers@gmail.com>

## References

Thissen, D. (1982). Marginal maximum likelihood estimation for the one-parameter logistic model. *Psychometrika, 47*, 175-186.

## Examples

```
dat <- expand.table(LSAT6)
head(dat)
itemstats(dat)

model <- 'F = 1-5
         CONSTRAIN = (1-5, a1)'
(mod <- mirt(dat, model))
M2(mod)
itemfit(mod)
coef(mod, simplify=TRUE)

# equivalentely, but with a different parameterization
mod2 <- mirt(dat, 1, itemtype = 'Rasch')
anova(mod, mod2) #equal
M2(mod2)
itemfit(mod2)
coef(mod2, simplify=TRUE)
sqrt(coef(mod2)$GroupPars[2]) #latent SD equal to the slope in mod
```

LSAT7 *Description of LSAT7 data*

## Description

Data from Bock & Lieberman (1970); contains 5 dichotomously scored items obtained from the Law School Admissions Test, section 7.

Data from

## Author(s)

Phil Chalmers <rphilip.chalmers@gmail.com>

## References

Bock, R. D., & Lieberman, M. (1970). Fitting a response model for *n* dichotomously scored items. *Psychometrika, 35*(2), 179-197.

Bock, R. D., & Lieberman, M. (1970). Fitting a response model for *n* dichotomously scored items. *Psychometrika, 35*(2), 179-197.

## Examples

```
dat <- expand.table(LSAT7)
head(dat)
itemstats(dat)

# fit 2PL model for each item
(mod <- mirt(dat))
coef(mod)

# monotonic splines models (see Winsberg, Thissen, and Wainer, 1984)
mod_monospline <- mirt(dat, itemtype = 'monospline')
anova(mod, mod_monospline)
plot(mod_monospline)

# compare item 1 trace-lines
i1 <- extract.item(mod, 1)
i1mono <- extract.item(mod_monospline, 1)
theta <- matrix(seq(-6, 6, length.out=100))
twoPL <- probtrace(i1, theta)[,2]
monospline <- probtrace(i1mono, theta)[,2]

plot(twoPL ~ theta, type = 'l')
lines(monospline ~ theta, col='red')
```

```
dat <- expand.table(LSAT7)
head(dat)
itemstats(dat)

(mod <- mirt(dat, 1))
coef(mod)
```

---

M2 *Compute the M2 model fit statistic*

---

## Description

Computes the M2 (Maydeu-Olivares & Joe, 2006) statistic when all data are dichotomous, the collapsed M2* statistic (collapsing over univariate and bivariate response categories; see Cai and Hansen, 2013), and the hybrid C2 statistic which only collapses only the bivariate moments (Cai and Monro, 2014). The C2 variant is mainly useful when polytomous response models do not have sufficient degrees of freedom to compute M2*. This function also computes associated fit indices that are based on fitting the null model. Supports single and multiple-group models. If the latent trait density was approximated (e.g., Davidian curves, Empirical histograms, etc) then passing `use_dentype_estimate = TRUE` will use the internally saved quadrature and density components (where applicable).

## Usage

```
M2(
  obj,
  type = "M2*",
  calcNull = TRUE,
  quadpts = NULL,
  theta_lim = c(-6, 6),
  CI = 0.9,
  residmat = FALSE,
  QMC = FALSE,
  suppress = 1,
  ...
)
```

## Arguments

| | |
|---|---|
| `obj` | an estimated model object from the mirt package |
| `type` | type of fit statistic to compute. Options are "M2", "M2*" for the univariate and bivariate collapsed version of the M2 statistic ("M2" currently limited to dichotomous response data only), and "C2" for a hybrid between M2 and M2* where only the bivariate moments are collapsed |
| `calcNull` | logical; calculate statistics for the null model as well? Allows for statistics such as the limited information TLI and CFI. Only valid when items all have a suitable null model (e.g., those created via `createItem` will not) |

| | |
|---|---|
| quadpts | number of quadrature points to use during estimation. If NULL, a suitable value will be chosen based on the rubric found in fscores |
| theta_lim | lower and upper range to evaluate latent trait integral for each dimension |
| CI | numeric value from 0 to 1 indicating the range of the confidence interval for RMSEA. Default returns the 90% interval |
| residmat | logical; return the residual matrix used to compute the SRMSR statistic? Only the lower triangle of the residual correlation matrix will be returned (the upper triangle is filled with NA's) |
| QMC | logical; use quasi-Monte Carlo integration? Useful for higher dimensional models. If quadpts not specified, 5000 nodes are used by default |
| suppress | a numeric value indicating which parameter residual dependency combinations to flag as being too high. Absolute values for the standardized residuals greater than this value will be returned, while all values less than this value will be set to NA. Must be used in conjunction with the argument residmat = TRUE |
| ... | additional arguments to pass |

## Value

Returns a data.frame object with the M2-type statistic, along with the degrees of freedom, p-value, RMSEA (with 90% confidence interval), SRMSR for each group, and optionally the TLI and CFI model fit statistics if calcNull = TRUE.

## Author(s)

Phil Chalmers <rphilip.chalmers@gmail.com>

## References

Cai, L. & Hansen, M. (2013). Limited-information goodness-of-fit testing of hierarchical item factor models. *British Journal of Mathematical and Statistical Psychology, 66*, 245-276.

Cai, L. & Monro, S. (2014). *A new statistic for evaluating item response theory models for ordinal data*. National Center for Research on Evaluation, Standards, & Student Testing. Technical Report.

Chalmers, R. P. (2012). mirt: A Multidimensional Item Response Theory Package for the R Environment. *Journal of Statistical Software, 48*(6), 1-29. doi:10.18637/jss.v048.i06

Maydeu-Olivares, A. & Joe, H. (2006). Limited information goodness-of-fit testing in multidimensional contingency tables. *Psychometrika, 71*, 713-732.

## Examples

```
dat <- as.matrix(expand.table(LSAT7))
(mod1 <- mirt(dat, 1))
M2(mod1)
resids <- M2(mod1, residmat=TRUE) #lower triangle of residual correlation matrix
resids
summary(resids[lower.tri(resids)])

# M2 with missing data present
```

```
dat[sample(1:prod(dim(dat)), 250)] <- NA
mod2 <- mirt(dat, 1)
M2(mod2)

# C2 statistic (useful when polytomous IRT models have too few df)
pmod <- mirt(Science, 1)
# This fails with too few df:
# M2(pmod)
# This, however, works:
M2(pmod, type = 'C2')
```

---

`marginal_moments` *Function to calculate the marginal moments for items and bundles fit via IRT models*

---

## Description

Given an estimated model and a prior density function, compute the marginal moments for either and item or a bundle of items. Function returns the first found moments implied by the model and select density function (MEAN, VAR, SKEW, and KURT). Currently limited to unidimensional IRT models.

## Usage

```
marginal_moments(
  mod,
  which.items = NULL,
  group = NULL,
  bundle = TRUE,
  density = NULL,
  Theta_lim = c(-6, 6),
  quadpts = 121,
  ...
)
```

## Arguments

| | |
|---|---|
| `mod` | an object of class `'SingleGroupClass'` or `'MultipleGroupClass'` |
| `which.items` | vector indicating which items to use in the computation of the expected values. Default (`NULL`) uses all available items |
| `group` | optional indicator to return only specific group information for multiple group models. Default compute moments for each group, returning a `list` |
| `bundle` | logical; given `which.items`, should the composite of the response functions be used as a collective bundle? If `TRUE` then `expected.test` will be used, otherwise `expected.item` will be used |

| | |
|---|---|
| density | a density function to use for integration. Default assumes the latent traits are from a normal (Gaussian) distribution. Function definition must be of the form function(quadrature, mean, sd) as the values of the mean/variance are extracted and passed from the supplied model |
| Theta_lim | range of integration grid to use when forming expectations |
| quadpts | number of discrete quadrature to use in the computations |
| ... | additional arguments passed to the density function |

## Author(s)

Phil Chalmers <rphilip.chalmers@gmail.com>

## References

Chalmers, R. P. (2012). mirt: A Multidimensional Item Response Theory Package for the R Environment. *Journal of Statistical Software, 48*(6), 1-29. doi:10.18637/jss.v048.i06

## See Also

expected.item, expected.test

## Examples

```
# single group
dat <- expand.table(deAyala)
mod <- mirt(dat)
TS <- rowSums(dat)

# expected moments of total scores given model
marginal_moments(mod)
c(mean=mean(TS), var=var(TS))  # first two moments of data

# same, however focusing on individual items
marginal_moments(mod, bundle = FALSE)
cbind(mean=colMeans(dat), var=apply(dat, 2, var)) # first two moments of data

############################################
## same as above, however with multiple group model

set.seed(1234)
group <- sample(c('G1', 'G2'), nrow(dat), replace=TRUE)
modMG <- multipleGroup(dat, group=group,
 invariance=c(colnames(dat), 'free_mean', 'free_var'))
coef(modMG, simplify=TRUE)

# expected moments of total scores given model
marginal_moments(modMG)
marginal_moments(modMG, group = 'G1') # specific group only
```

```
# same, however focusing on individual items
marginal_moments(modMG, bundle = FALSE)
```

---

`marginal_rxx` *Function to calculate the marginal reliability*

---

## Description

Given an estimated model and a prior density function, compute the marginal reliability (Thissen and Wainer, 2001). This is only available for unidimensional tests.

## Usage

```
marginal_rxx(mod, density = dnorm, ...)
```

## Arguments

| | |
|---|---|
| `mod` | an object of class `'SingleGroupClass'` |
| `density` | a density function to use for integration. Default assumes the latent traits are from a normal (Gaussian) distribution |
| `...` | additional arguments passed to the density function |

## Author(s)

Phil Chalmers <rphilip.chalmers@gmail.com>

## References

Chalmers, R. P. (2012). mirt: A Multidimensional Item Response Theory Package for the R Environment. *Journal of Statistical Software, 48*(6), 1-29. doi:10.18637/jss.v048.i06

Thissen, D. and Wainer, H. (2001). Test Scoring. Lawrence Erlbaum Associates.

## See Also

`empirical_rxx`, `extract.group`, `testinfo`

## Examples

```
dat <- expand.table(deAyala)
mod <- mirt(dat)

# marginal estimate treating item parameters as known
marginal_rxx(mod)
```

```
# compare to alpha
itemstats(dat)$overall$alpha



# empirical estimate (assuming the same prior)
fscores(mod, returnER = TRUE)

# empirical rxx the alternative way, given theta scores and SEs
fs <- fscores(mod, full.scores.SE=TRUE)
head(fs)
empirical_rxx(fs)

##############
# example demonstrating correlation attenuation

theta <- rnorm(1000)
X <- theta + rnorm(1000, sd=2)
cor(X, theta)    # correlation without measurement error (what you want)

# measured with a 10 item GRM test
nitems <- 10
a <- matrix(rlnorm(nitems,.2,.3))
diffs <- t(apply(matrix(runif(nitems*4, .3, 1), nitems), 1, cumsum))
diffs <- -(diffs - rowMeans(diffs))
d <- diffs + rnorm(nitems)
dat <- simdata(a, d, itemtype = 'graded', Theta=matrix(theta))

# correlation with total score (attenuated)
cor(rowSums(dat), X)

# fit single group model
mod <- mirt(dat)

# EAP correlation (also attenuated)
fs <- fscores(mod)
cor(fs, X)

# correction for attenuation, r_x.theta = r_x.theta.hat / sqrt(rxx_theta.hat)
(rxx <- marginal_rxx(mod))  # alternatively, could use empirical_rxx()
cor(fs, X) / sqrt(rxx)  # correction estimate
cor(X, theta)           # compare to true correlation
```

---

MDIFF *Compute multidimensional difficulty index*

---

## Description

Returns a matrix containing the MDIFF values (Reckase, 2009). Only supported for items of class 'dich' and 'graded'.

## Usage

```
MDIFF(x, which.items = NULL, group = NULL)
```

## Arguments

| | |
|---|---|
| x | an object of class 'SingleGroupClass', or an object of class 'MultipleGroupClass' if a suitable group input were supplied |
| which.items | a vector indicating which items to select. If NULL is used (the default) then MDISC will be computed for all items |
| group | group argument to pass to extract.group function. Required when the input object is a multiple-group model |

## Author(s)

Phil Chalmers <rphilip.chalmers@gmail.com>

## References

Chalmers, R. P. (2012). mirt: A Multidimensional Item Response Theory Package for the R Environment. *Journal of Statistical Software, 48*(6), 1-29. doi:10.18637/jss.v048.i06

Reckase, M. D. (2009). Multidimensional Item Response Theory. Springer.

## See Also

extract.group, MDISC

## Examples

```
mod <- mirt(Science, 2)
MDIFF(mod)

mod <- mirt(expand.table(LSAT7), 2)
MDIFF(mod)
```

mdirt *Multidimensional discrete item response theory*

## Description

mdirt fits a variety of item response models with discrete latent variables. These include, but are not limited to, latent class analysis, multidimensional latent class models, multidimensional discrete latent class models, DINA/DINO models, grade of measurement models, C-RUM, and so on. If response models are not defined explicitly then customized models can defined using the createItem function.

## Usage

```
mdirt(
  data,
  model,
  customTheta = NULL,
  structure = NULL,
  item.Q = NULL,
  nruns = 1,
  method = "EM",
  covdata = NULL,
  formula = NULL,
  itemtype = "lca",
  optimizer = "nlminb",
  return_max = TRUE,
  group = NULL,
  GenRandomPars = FALSE,
  verbose = interactive(),
  pars = NULL,
  technical = list(),
  ...
)
```

## Arguments

| | |
|---|---|
| data | a matrix or data.frame that consists of numerically ordered data, organized in the form of integers, with missing data coded as NA |
| model | number of mutually exclusive classes to fit, or alternatively a more specific mirt.model definition (which reflects the so-called Q-matrix). Note that when using a mirt.model, the order with which the syntax factors/attributes are defined are associated with the columns in the customTheta input |
| customTheta | input passed to technical = list(customTheta = ...), but is included directly in this function for convenience. This input is most interesting for discrete latent models because it allows customized patterns of latent classes (i.e., defines the possible combinations of the latent attribute profile). The default |

builds the pattern `customTheta = diag(model)`, which is the typical pattern for the traditional latent class analysis whereby class membership mutually distinct and exhaustive. See `thetaComb` for a quick method to generate a matrix with all possible combinations

`structure` an R formula allowing the profile probability patterns (i.e., the structural component of the model) to be fitted according to a log-linear model. When `NULL`, all profile probabilities (except one) will be estimated. Use of this input requires that the `customTheta` input is supplied, and that the column names in this matrix match the names found within this formula

`item.Q` a list of item-level Q-matrices indicating how the respective categories should be modeled by the underlying attributes. Each matrix must represent a $K_i \times A$ matrix, where $K_i$ represents the number of categories for the ith item, and $A$ is the number of attributes included in the `Theta` matrix; otherwise, a value of`NULL` will default to a matrix consisting of 1's for each $K_i \times A$ element except for the first row, which contains only 0's for proper identification. Incidentally, the first row of each matrix `must` contain only 0's so that the first category represents the reference category for identification

`nruns` a numeric value indicating how many times the model should be fit to the data when using random starting values. If greater than 1, `GenRandomPars` is set to `TRUE` by default. Using this returns a list of fitted model objects, where the model with the highest log-likelihood should generally be selected as the model best associated with the MLE. Note that if a `mirtCluster` was defined earlier then the runs will be run in parallel

`method` estimation method. Can be 'EM' or 'BL' (see `mirt` for more details)

`covdata` a data.frame of data used for latent regression models

`formula` an R formula (or list of formulas) indicating how the latent traits can be regressed using external covariates in `covdata`. If a named list of formulas is supplied (where the names correspond to the latent trait/attribute names in `model`) then specific regression effects can be estimated for each factor. Supplying a single formula will estimate the regression parameters for all latent variables by default

`itemtype` a vector indicating the itemtype associated with each item. For discrete models this is limited to only 'lca' or items defined using a `createItem` definition

`optimizer` optimizer used for the M-step, set to `'nlminb'` by default. See `mirt` for more details

`return_max` logical; when `nruns > 1`, return the model that has the most optimal maximum likelihood criteria? If FALSE, returns a list of all the estimated objects

`group` a factor variable indicating group membership used for multiple group analyses

`GenRandomPars` logical; use random starting values

`verbose` logical; turn on messages to the R console

`pars` used for modifying starting values; see `mirt` for details

`technical` list of lower-level inputs. See `mirt` for details

`...` additional arguments to be passed to the estimation engine. See `mirt` for more details and examples

## Details

Posterior classification accuracy for each response pattern may be obtained via the `fscores` function. The `summary()` function will display the category probability values given the class membership, which can also be displayed graphically with `plot()`, while `coef()` displays the raw coefficient values (and their standard errors, if estimated). Finally, `anova()` is used to compare nested models, while `M2` and `itemfit` may be used for model fitting purposes.

## 'lca' model definition

The latent class IRT model with two latent classes has the form

$$P(x = k|\theta_1, \theta_2, a1, a2) = \frac{exp(a1\theta_1 + a2\theta_2)}{\sum_j^K exp(a1\theta_1 + a2\theta_2)}$$

where the $\theta$ values generally take on discrete points (such as 0 or 1). For proper identification, the first category slope parameters ($a1$ and $a2$) are never freely estimated. Alternatively, supplying a different grid of $\theta$ values will allow the estimation of similar models (multidimensional discrete models, grade of membership, etc.). See the examples below.

When the `item.Q` for is utilized, the above equation can be understood as

$$P(x = k|\theta_1, \theta_2, a1, a2) = \frac{exp(a1\theta_1 Q_{j1} + a2\theta_2 Q_{j2})}{\sum_j^K exp(a1\theta_1 Q_{j1} + a2\theta_2 Q_{j2})}$$

where by construction `Q` is a $K_i \times A$ matrix indicating whether the category should be modeled according to the latent class structure. For the standard latent class model, the Q-matrix has as many rows as categories, as many columns as the number of classes/attributes modeled, and consist of 0's in the first row and 1's elsewhere. This of course can be over-written by passing an alternative `item.Q` definition for each respective item.

## Author(s)

Phil Chalmers <rphilip.chalmers@gmail.com>

## References

Chalmers, R. P. (2012). mirt: A Multidimensional Item Response Theory Package for the R Environment. *Journal of Statistical Software, 48*(6), 1-29.

Proctor, C. H. (1970). A probabilistic formulation and statistical analysis for Guttman scaling. *Psychometrika, 35*, 73-78. doi:10.18637/jss.v048.i06

## See Also

`thetaComb`, `fscores`, `mirt.model`, `M2`, `itemfit`, `boot.mirt`, `mirtCluster`, `wald`, `coef-method`, `summary-method`, `anova-method`, `residuals-method`

## Examples

```
# LSAT6 dataset
dat <- expand.table(LSAT6)

# fit with 2-3 latent classes
(mod2 <- mdirt(dat, 2))

(mod3 <- mdirt(dat, 3))
summary(mod2)
residuals(mod2)
residuals(mod2, type = 'exp')
anova(mod2, mod3)
M2(mod2)
itemfit(mod2)

# generate classification plots
plot(mod2)
plot(mod2, facet_items = FALSE)
plot(mod2, profile = TRUE)

# available for polytomous data
mod <- mdirt(Science, 2)
summary(mod)
plot(mod)
plot(mod, profile=TRUE)

# classification based on response patterns
fscores(mod2, full.scores = FALSE)

# classify individuals either with the largest posterior probability.....
fs <- fscores(mod2)
head(fs)
classes <- 1:2
class_max <- classes[apply(apply(fs, 1, max) == fs, 1, which)]
table(class_max)

# ... or by probability sampling (i.e., plausible value draws)
class_prob <- apply(fs, 1, function(x) sample(1:2, 1, prob=x))
table(class_prob)

# plausible value imputations for stochastic classification in both classes
pvs <- fscores(mod2, plausible.draws=10)
tabs <- lapply(pvs, function(x) apply(x, 2, table))
tabs[[1]]


# fit with random starting points (run in parallel to save time)
if(interactive()) mirtCluster()
mod <- mdirt(dat, 2, nruns=10)

#--------------------------
# Grade of measurement model
```

```
# define a custom Theta grid for including a 'fuzzy' class membership
(Theta <- matrix(c(1, 0, .5, .5, 0, 1), nrow=3 , ncol=2, byrow=TRUE))
(mod_gom <- mdirt(dat, 2, customTheta = Theta))
summary(mod_gom)

#-----------------
# Multidimensional discrete latent class model

dat <- key2binary(SAT12,
     key = c(1,4,5,2,3,1,2,1,3,1,2,4,2,1,5,3,4,4,1,4,3,3,4,1,3,5,1,3,1,5,4,5))

# define Theta grid for three latent classes
(Theta <- thetaComb(0:1, 3))
(mod_discrete <- mdirt(dat, 3, customTheta = Theta))
summary(mod_discrete)

# Located latent class model
model <- mirt.model('C1 = 1-32
                     C2 = 1-32
                     C3 = 1-32
                     CONSTRAIN = (1-32, a1), (1-32, a2), (1-32, a3)')
(mod_located <- mdirt(dat, model, customTheta = diag(3)))
summary(mod_located)

#-----------------
### DINA model example
# generate some suitable data for a two dimensional DINA application
#     (first columns are intercepts)
set.seed(1)
Theta <- expand.table(matrix(c(1,0,0,0,
                               1,1,0,0,
                               1,0,1,0,
                               1,1,1,1), 4, 4, byrow=TRUE),
                      freq = c(200,200,100,500))
a <- matrix(c(rnorm(15, -1.5, .5), rlnorm(5, .2, .3), numeric(15), rlnorm(5, .2, .3),
              numeric(15), rlnorm(5, .2, .3)), 15, 4)

guess <- plogis(a[11:15,1]) # population guess
slip <- 1 - plogis(rowSums(a[11:15,])) # population slip

dat <- simdata(a, Theta=Theta, itemtype = 'lca')

# first column is the intercept, 2nd and 3rd are attributes
theta <- cbind(1, thetaComb(0:1, 2))
theta <- cbind(theta, theta[,2] * theta[,3]) #DINA interaction of main attributes
model <- mirt.model('Intercept = 1-15
                     A1 = 1-5
                     A2 = 6-10
                     A1A2 = 11-15')

# last 5 items are DINA (first 10 are unidimensional C-RUMs)
DINA <- mdirt(dat, model, customTheta = theta)
```

```
coef(DINA, simplify=TRUE)
summary(DINA)
M2(DINA) # fits well (as it should)

cfs <- coef(DINA, simplify=TRUE)$items[11:15,]
cbind(guess, estguess = plogis(cfs[,1]))
cbind(slip, estslip = 1 - plogis(rowSums(cfs)))


### DINO model example
theta <- cbind(1, thetaComb(0:1, 2))
# define theta matrix with negative interaction term
(theta <- cbind(theta, -theta[,2] * theta[,3]))

model <- mirt.model('Intercept = 1-15
                     A1 = 1-5, 11-15
                     A2 = 6-15
                     Yoshi = 11-15
                     CONSTRAIN = (11,a2,a3,a4), (12,a2,a3,a4), (13,a2,a3,a4),
                                 (14,a2,a3,a4), (15,a2,a3,a4)')

# last five items are DINOs (first 10 are unidimensional C-RUMs)
DINO <- mdirt(dat, model, customTheta = theta)
coef(DINO, simplify=TRUE)
summary(DINO)
M2(DINO) #doesn't fit as well, because not the generating model

## C-RUM (analogous to MIRT model)
theta <- cbind(1, thetaComb(0:1, 2))
model <- mirt.model('Intercept = 1-15
                     A1 = 1-5, 11-15
                     A2 = 6-15')

CRUM <- mdirt(dat, model, customTheta = theta)
coef(CRUM, simplify=TRUE)
summary(CRUM)

# good fit, but over-saturated (main effects for items 11-15 can be set to 0)
M2(CRUM)

#------------------
# multidimensional latent class model

dat <- key2binary(SAT12,
     key = c(1,4,5,2,3,1,2,1,3,1,2,4,2,1,5,3,4,4,1,4,3,3,4,1,3,5,1,3,1,5,4,5))

# 5 latent classes within 2 different sets of items
model <- mirt.model('C1 = 1-16
                     C2 = 1-16
                     C3 = 1-16
                     C4 = 1-16
                     C5 = 1-16
                     C6 = 17-32
```

```
                    C7 = 17-32
                    C8 = 17-32
                    C9 = 17-32
                    C10 = 17-32
                  CONSTRAIN = (1-16, a1), (1-16, a2), (1-16, a3), (1-16, a4), (1-16, a5),
                      (17-32, a6), (17-32, a7), (17-32, a8), (17-32, a9), (17-32, a10)')

theta <- diag(10) # defined explicitly. Otherwise, this profile is assumed
mod <- mdirt(dat, model, customTheta = theta)
coef(mod, simplify=TRUE)
summary(mod)

#------------------
# multiple group with constrained group probabilities
 dat <- key2binary(SAT12,
   key = c(1,4,5,2,3,1,2,1,3,1,2,4,2,1,5,3,4,4,1,4,3,3,4,1,3,5,1,3,1,5,4,5))
group <- rep(c('G1', 'G2'), each = nrow(SAT12)/2)
Theta <- diag(2)

# the latent class parameters are technically located in the (nitems + 1) location
model <- mirt.model('A1 = 1-32
                    A2 = 1-32
                    CONSTRAINB = (33, c1)')
mod <- mdirt(dat, model, group = group, customTheta = Theta)
coef(mod, simplify=TRUE)
summary(mod)


#------------------
# Probabilistic Guttman Model (Proctor, 1970)

# example analysis can also be found in the sirt package (see ?prob.guttman)
data(data.read, package = 'sirt')
head(data.read)

Theta <- matrix(c(1,0,0,0,
                  1,1,0,0,
                  1,1,1,0,
                  1,1,1,1), 4, byrow=TRUE)

model <- mirt.model("INTERCEPT = 1-12
                    C1 = 1,7,9,11
                    C2 = 2,5,8,10,12
                    C3 = 3,4,6")

mod <- mdirt(data.read, model, customTheta=Theta)
summary(mod)

M2(mod)
itemfit(mod)
```

MDISC *Compute multidimensional discrimination index*

## Description

Returns a vector containing the MDISC values for each item in the model input object (Reckase, 2009).

## Usage

```
MDISC(x, group = NULL)
```

## Arguments

| | |
|---|---|
| x | an object of class 'SingleGroupClass', or an object of class 'MultipleGroupClass' if a suitable group input were supplied |
| group | group argument to pass to extract.group function. Required when the input object is a multiple-group model |

## Author(s)

Phil Chalmers <rphilip.chalmers@gmail.com>

## References

Chalmers, R. P. (2012). mirt: A Multidimensional Item Response Theory Package for the R Environment. *Journal of Statistical Software, 48*(6), 1-29. doi:10.18637/jss.v048.i06

Reckase, M. D. (2009). Multidimensional Item Response Theory. Springer.

## See Also

extract.group

## Examples

```
mod <- mirt(Science, 2)
MDISC(mod)
```

mirt *Full-Information Item Factor Analysis (Multidimensional Item Response Theory)*

## Description

mirt fits a maximum likelihood (or maximum a posteriori) factor analysis model to any mixture of dichotomous and polytomous data under the multidimensional item response theory paradigm using either Cai's (2010) Metropolis-Hastings Robbins-Monro (MHRM) algorithm, with an EM algorithm approach outlined by Bock and Aitkin (1981) using rectangular or quasi-Monte Carlo integration grids, or with the stochastic EM (i.e., the first two stages of the MH-RM algorithm). Unidimensional and multidimensional dominance/compensatory response models or unfolding/pairwise comparison models can be specified independently via the itemtype argument.

## Usage

```
mirt(
  data,
  model = 1,
  itemtype = NULL,
  guess = 0,
  upper = 1,
  SE = FALSE,
  covdata = NULL,
  formula = NULL,
  itemdesign = NULL,
  item.formula = NULL,
  SE.type = "Oakes",
  method = "EM",
  optimizer = NULL,
  dentype = "Gaussian",
  pars = NULL,
  constrain = NULL,
  calcNull = FALSE,
  draws = 5000,
  survey.weights = NULL,
  quadpts = NULL,
  TOL = NULL,
  gpcm_mats = list(),
  grsm.block = NULL,
  rsm.block = NULL,
  monopoly.k = 1L,
  key = NULL,
  large = FALSE,
  GenRandomPars = FALSE,
  accelerate = "Ramsay",
  verbose = interactive(),
```

```
  solnp_args = list(),
  nloptr_args = list(),
  spline_args = list(),
  control = list(),
  technical = list(),
  ...
)
```

## Arguments

| | |
|---|---|
| data | a matrix or data.frame that consists of numerically ordered data, organized in the form of integers, with missing data coded as NA (to convert from an ordered factor data.frame see data.matrix) |
| model | a string to be passed (or an object returned from) mirt.model, declaring how the IRT model is to be estimated (loadings, constraints, priors, etc). For exploratory IRT models, a single numeric value indicating the number of factors to extract is also supported. Default is 1, indicating that a unidimensional model will be fit unless otherwise specified |
| itemtype | type of items to be modeled, declared as either a) a single value to be recycled for each item, b) a vector for each respective item, or c) if applicable, a matrix with columns equal to the number of items and rows equal to the number of latent classes. The NULL default assumes that the items follow a graded or 2PL structure, however they may be changed to the following: |

- 'Rasch' - Rasch/partial credit model by constraining slopes to 1 and freely estimating the variance parameters (alternatively, can be specified by applying equality constraints to the slope parameters in 'gpcm' and '2PL'; Rasch, 1960)
- '1PL', '2PL', '3PL', '3PLu', and '4PL' - 1-4 parameter logistic model, where 3PL estimates the lower asymptote only while 3PLu estimates the upper asymptote only (Lord and Novick, 1968; Lord, 1980). Note that specifying '1PL' will not automatically estimate the variance of the latent trait compared to the 'Rasch' type
- '5PL' - 5 parameter logistic model to estimate asymmetric logistic response curves. Currently restricted to unidimensional models
- 'CLL' - complementary log-log link model. Currently restricted to unidimensional models
- 'ULL' - unipolar log-logistic model (Lucke, 2015). Note the use of this itemtype will automatically use a log-normal distribution for the latent traits
- 'graded' - graded response model (Samejima, 1969)
- 'PCgraded' - partially non-compensatory graded response model (Chalmers, 2020)
- 'grsm' - graded ratings scale model in the classical IRT parameterization (restricted to unidimensional models; Muraki, 1992)
- 'gpcm' and 'gpcmIRT' - generalized partial credit model in the slope-intercept and classical parameterization. 'gpcmIRT' is restricted to unidimensional models. Note that optional scoring matrices for 'gpcm' are available with the gpcm_mats input (Muraki, 1992)

- 'rsm' - Rasch rating scale model using the 'gpcmIRT' structure (unidimensional only; Andrich, 1978)
- 'nominal' - nominal response model (Bock, 1972)
- 'ideal' - dichotomous ideal point model (Maydeu-Olivares, 2006)
- 'ggum' - generalized graded unfolding model (Roberts, Donoghue, & Laughlin, 2000) and its multidimensional extension
- 'hcm' and 'ghcm' - (generalized) hyperbolic cosine model (Andrich and Luo, 1993; Andrich, 1996) for dichotomous or ordered polytomous items (see Luo, 2001)
- 'alm' and 'galm' - (generalized) absolute logistic model (Luo and Andrich, 2005) for dichotomous or ordered polytomous items
- 'sslm' and 'gsslm' - (generalized) simple squared logistic model (Andrich, 1988) for dichotomous or ordered polytomous items (see Luo, 2001)
- 'paralla' and 'gparalla' - (generalized) parallellogram analysis model (Hoijtink, 1990) for dichotomous or ordered polytomous items (see Luo, 2001)
- 'sequential' - multidimensional sequential response model (Tutz, 1990) in slope-intercept form
- 'Tutz' - same as the 'sequential' itemtype, except the slopes are fixed to 1 and the latent variance terms are freely estimated (similar to the 'Rasch' itemtype input)
- 'PC1PL', 'PC2PL', and 'PC3PL' - 1-3 parameter partially compensatory model. Note that constraining the slopes to be equal across items will also reduce the model to Embretson's (a.k.a. Whitely's) multicomponent model (1980), while for 'PC1PL' the slopes are fixed to 1 while the latent trait variance terms are estimated
- '2PLNRM', '3PLNRM', '3PLuNRM', and '4PLNRM' - 2-4 parameter nested logistic model, where 3PLNRM estimates the lower asymptote only while 3PLuNRM estimates the upper asymptote only (Suh and Bolt, 2010)
- 'spline' - spline response model with the bs (default) or the ns function
- 'monospline' - monotonic spline response model with a constrained version of the I-spline basis from iSpline (Ramsay and Winsberg, 1991; Winsberg, Thissen, and Wainer, 1984)
- 'monopoly' - monotonic polynomial model for unidimensional tests for dichotomous and polytomous response data (Falk and Cai, 2016)

Additionally, user defined item classes can also be defined using the createItem function

guess — fixed pseudo-guessing parameters. Can be entered as a single value to assign a global guessing parameter or may be entered as a numeric vector corresponding to each item

upper — fixed upper bound parameters for 4-PL model. Can be entered as a single value to assign a global guessing parameter or may be entered as a numeric vector corresponding to each item

SE — logical; estimate the standard errors by computing the parameter information matrix? See SE.type for the type of estimates available

covdata — a data.frame of data used for latent regression models

formula — an R formula (or list of formulas) indicating how the latent traits can be regressed using external covariates in covdata. If a named list of formulas is supplied (where the names correspond to the latent trait names in model) then specific regression effects can be estimated for each factor. Supplying a single formula will estimate the regression parameters for all latent traits by default

itemdesign — a data.frame with rows equal to the number of items and columns containing any item-design effects. If items should be included in the design structure (i.e., should be left in their canonical structure) then fewer rows can be used, however the rownames must be defined and matched with colnames in the data input. The item design matrix is constructed with the use of item.formula. Providing this input will fix the associated 'd' intercepts to 0, where applicable

item.formula — an R formula used to specify any intercept decomposition (e.g., the LLTM; Fischer, 1983). Note that only the right-hand side of the formula is required for compensatory models.

For non-compensatory itemtypes (e.g., 'PC1PL') the formula must include the name of the latent trait in the left hand side of the expression to indicate which of the trait specification should have their intercepts decomposed (see MLTM; Embretson, 1984)

SE.type — type of estimation method to use for calculating the parameter information matrix for computing standard errors and wald tests. Can be:

- 'Richardson', 'forward', or 'central' for the numerical Richardson, forward difference, and central difference evaluation of observed Hessian matrix
- 'crossprod' and 'Louis' for standard error computations based on the variance of the Fisher scores as well as Louis' (1982) exact computation of the observed information matrix. Note that Louis' estimates can take a long time to obtain for large sample sizes and long tests
- 'sandwich' for the sandwich covariance estimate based on the 'crossprod' and 'Oakes' estimates (see Chalmers, 2018, for details)
- 'sandwich.Louis' for the sandwich covariance estimate based on the 'crossprod' and 'Louis' estimates
- 'Oakes' for Oakes' (1999) method using a central difference approximation (see Chalmers, 2018, for details)
- 'SEM' for the supplemented EM (disables the accelerate option automatically; EM only)
- 'Fisher' for the expected information, 'complete' for information based on the complete-data Hessian used in EM algorithm
- 'MHRM' and 'FMHRM' for stochastic approximations of observed information matrix based on the Robbins-Monro filter or a fixed number of MHRM draws without the RM filter. These are the only options supported when method = 'MHRM'
- 'numerical' to obtain the numerical estimate from a call to optim when method = 'BL'

Note that both the 'SEM' method becomes very sensitive if the ML solution has has not been reached with sufficient precision, and may be further sensitive

if the history of the EM cycles is not stable/sufficient for convergence of the respective estimates. Increasing the number of iterations (increasing NCYCLES and decreasing TOL, see below) will help to improve the accuracy, and can be run in parallel if a mirtCluster object has been defined (this will be used for Oakes' method as well). Additionally, inspecting the symmetry of the ACOV matrix for convergence issues by passing technical = list(symmetric = FALSE) can be helpful to determine if a sufficient solution has been reached

method — a character object specifying the estimation algorithm to be used. The default is 'EM', for the standard EM algorithm with fixed quadrature, 'QMCEM' for quasi-Monte Carlo EM estimation, or 'MCEM' for Monte Carlo EM estimation. The option 'MHRM' may also be passed to use the MH-RM algorithm, 'SEM' for the Stochastic EM algorithm (first two stages of the MH-RM stage using an optimizer other than a single Newton-Raphson iteration), and 'BL' for the Bock and Lieberman approach (generally not recommended for longer tests).

The 'EM' is generally effective with 1-3 factors, but methods such as the 'QMCEM', 'MCEM', 'SEM', or 'MHRM' should be used when the dimensions are 3 or more. Note that when the optimizer is stochastic the associated SE.type is automatically changed to SE.type = 'MHRM' by default to avoid the use of quadrature

optimizer — a character indicating which numerical optimizer to use. By default, the EM algorithm will use the 'BFGS' when there are no upper and lower bounds box-constraints and 'nlminb' when there are.

Other options include the Newton-Raphson ('NR'), which can be more efficient than the 'BFGS' but not as stable for more complex IRT models (such as the nominal or nested logit models) and the related 'NR1' which is also the Newton-Raphson but consists of only 1 update that has been coupled with RM Hessian (only applicable when the MH-RM algorithm is used). The MH-RM algorithm uses the 'NR1' by default, though currently the 'BFGS', 'L-BFGS-B', and 'NR' are also supported with this method (with fewer iterations by default) to emulate stochastic EM updates. As well, the 'Nelder-Mead' and 'SANN' estimators are available, but their routine use generally is not required or recommended.

Additionally, estimation subroutines from the Rsolnp and nloptr packages are available by passing the arguments 'solnp' and 'nloptr', respectively. This should be used in conjunction with the solnp_args and nloptr_args specified below. If equality constraints were specified in the model definition only the parameter with the lowest parnum in the pars = 'values' data.frame is used in the estimation vector passed to the objective function, and group hyper-parameters are omitted. Equality an inequality functions should be of the form function(p, optim_args), where optim_args is a list of internally parameters that largely can be ignored when defining constraints (though use of browser() here may be helpful)

dentype — type of density form to use for the latent trait parameters. Current options include

- 'Gaussian' (default) assumes a multivariate Gaussian distribution with an associated mean vector and variance-covariance matrix
- 'empiricalhist' or 'EH' estimates latent distribution using an empirical histogram described by Bock and Aitkin (1981). Only applicable for unidimensional models estimated with the EM algorithm. For this option, the

number of cycles, TOL, and quadpts are adjusted accommodate for less precision during estimation (namely: `TOL = 3e-5`, `NCYCLES = 2000`, `quadpts = 121`)

- `'empiricalhist_Woods'` or `'EHW'` estimates latent distribution using an empirical histogram described by Bock and Aitkin (1981), with the same specifications as in `dentype = 'empiricalhist'`, but with the extrapolation-interpolation method described by Woods (2007). NOTE: to improve stability in the presence of extreme response styles (i.e., all highest or lowest in each item) the `technical` option `zeroExtreme = TRUE` may be required to down-weight the contribution of these problematic patterns
- `'Davidian-#'` estimates semi-parametric Davidian curves described by Woods and Lin (2009), where the # placeholder represents the number of Davidian parameters to estimate (e.g., `'Davidian-6'` will estimate 6 smoothing parameters). By default, the number of `quadpts` is increased to 121, and this method is only applicable for unidimensional models estimated with the EM algorithm

Note that when `itemtype = 'ULL'` then a log-normal(0,1) density is used to support the unipolar scaling

`pars` a data.frame with the structure of how the starting values, parameter numbers, estimation logical values, etc, are defined. The user may observe how the model defines the values by using `pars = 'values'`, and this object can in turn be modified and input back into the estimation with `pars = mymodifiedpars`

`constrain` a list of user declared equality constraints. To see how to define the parameters correctly use `pars = 'values'` initially to see how the parameters are labeled. To constrain parameters to be equal create a list with separate concatenated vectors signifying which parameters to constrain. For example, to set parameters 1 and 5 equal, and also set parameters 2, 6, and 10 equal use `constrain = list(c(1,5), c(2,6,10))`. Constraints can also be specified using the `mirt.model` syntax (recommended)

`calcNull` logical; calculate the Null model for additional fit statistics (e.g., TLI)? Only applicable if the data contains no NA's and the data is not overly sparse

`draws` the number of Monte Carlo draws to estimate the log-likelihood for the MH-RM algorithm. Default is 5000

`survey.weights` a optional numeric vector of survey weights to apply for each case in the data (EM estimation only). If not specified, all cases are weighted equally (the standard IRT approach). The sum of the `survey.weights` must equal the total sample size for proper weighting to be applied

`quadpts` number of quadrature points per dimension (must be larger than 2). By default the number of quadrature uses the following scheme: `switch(as.character(nfact), '1'=61, '2'=31, '3'=15, '4'=9, '5'=7, 3)`. However, if the method input is set to `'QMCEM'` and this argument is left blank then the default number of quasi-Monte Carlo integration nodes will be set to 5000 in total

`TOL` convergence threshold for EM or MH-RM; defaults are .0001 and .001. If `SE.type = 'SEM'` and this value is not specified, the default is set to `1e-5`. To evaluate the model using only the starting values pass `TOL = NaN`, and to evaluate the starting values without the log-likelihood pass `TOL = NA`

gpcm_mats
: a list of matrices specifying how the scoring coefficients in the (generalized) partial credit model should be constructed. If omitted, the standard gpcm format will be used (i.e., `seq(0, k, by = 1)` for each trait). This input should be used if traits should be scored different for each category (e.g., `matrix(c(0:3, 1,0,0,0), 4, 2)` for a two-dimensional model where the first trait is scored like a gpcm, but the second trait is only positively indicated when the first category is selected). Can be used when `itemtypes` are `'gpcm'` or `'Rasch'`, but only when the respective element in `gpcm_mats` is not `NULL`

grsm.block
: an optional numeric vector indicating where the blocking should occur when using the grsm, NA represents items that do not belong to the grsm block (other items that may be estimated in the test data). For example, to specify two blocks of 3 with a 2PL item for the last item: `grsm.block = c(rep(1,3), rep(2,3), NA)`. If NULL the all items are assumed to be within the same group and therefore have the same number of item categories

rsm.block
: same as `grsm.block`, but for `'rsm'` blocks

monopoly.k
: a vector of values (or a single value to repeated for each item) which indicate the degree of the monotone polynomial fitted, where the monotone polynomial corresponds to `monopoly.k * 2 + 1` (e.g., `monopoly.k = 2` fits a 5th degree polynomial). Default is `monopoly.k = 1`, which fits a 3rd degree polynomial

key
: a numeric vector of the response scoring key. Required when using nested logit item types, and must be the same length as the number of items used. Items that are not nested logit will ignore this vector, so use `NA` in item locations that are not applicable

large
: a `logical` indicating whether unique response patterns should be obtained prior to performing the estimation so as to avoid repeating computations on identical patterns. The default `TRUE` provides the correct degrees of freedom for the model since all unique patterns are tallied (typically only affects goodness of fit statistics such as G2, but also will influence nested model comparison methods such as `anova(mod1, mod2)`), while `FALSE` will use the number of rows in `data` as a placeholder for the total degrees of freedom. As such, model objects should only be compared if all flags were set to `TRUE` or all were set to `FALSE`

  Alternatively, if the collapse table of frequencies is desired for the purpose of saving computations (i.e., only computing the collapsed frequencies for the data onte-time) then a character vector can be passed with the arguement `large = 'return'` to return a list of all the desired table information used by `mirt`. This list object can then be reused by passing it back into the `large` argument to avoid re-tallying the data again (again, useful when the dataset are very large and computing the tabulated data is computationally burdensome). This strategy is shown below:

  **Compute organized data** e.g., `internaldat <- mirt(Science, 1, large = 'return')`

  **Pass the organized data to all estimation functions** e.g., `mod <- mirt(Science, 1, large = internaldat)`

GenRandomPars
: logical; generate random starting values prior to optimization instead of using the fixed internal starting values?

accelerate
: a character vector indicating the type of acceleration to use. Default is `'Ramsay'`, but may also be `'squarem'` for the SQUAREM procedure (specifically, the

gSqS3 approach) described in Varadhan and Roldand (2008). To disable the acceleration, pass 'none'

verbose — logical; print observed- (EM) or complete-data (MHRM) log-likelihood after each iteration cycle? Default is TRUE

solnp_args — a list of arguments to be passed to the solnp::solnp() function for equality constraints, inequality constraints, etc

nloptr_args — a list of arguments to be passed to the nloptr::nloptr() function for equality constraints, inequality constraints, etc

spline_args — a named list of lists containing information to be passed to the bs (default) ns, and iSpline for each spline/monospline itemtype. Each element must refer to the name of the itemtype with the spline, while the internal list names refer to the arguments which are passed. For example, if item 2 were called 'read2', and item 5 were called 'read5', both of which were of itemtype 'spline' but item 5 should use the ns form, then a modified list for each input might be of the form:

spline_args = list(read2 = list(degree = 4), read5 = list(fun = 'ns', knots = c(-2, 2)))

This code input changes the bs() splines function to have a degree = 4 input, while the second element changes to the ns() function with knots set a c(-2, 2)

control — a list passed to the respective optimizers (i.e., optim(), nlminb(), etc). Additional arguments have been included for the 'NR' optimizer: 'tol' for the convergence tolerance in the M-step (default is TOL/1000), while the default number of iterations for the Newton-Raphson optimizer is 50 (modified with the 'maxit' control input)

technical — a list containing lower level technical parameters for estimation. May be:

**NCYCLES** maximum number of EM or MH-RM cycles; defaults are 500 and 2000

**MAXQUAD** maximum number of quadrature, which you can increase if you have more than 4GB or RAM on your PC; default 20000

**theta_lim** range of integration grid for each dimension; default is c(-6, 6). Note that when itemtype = 'ULL' a log-normal distribution is used and the range is change to c(.01, and 6^2), where the second term is the square of the theta_lim input instead

**set.seed** seed number used during estimation. Default is 12345

**SEtol** standard error tolerance criteria for the S-EM and MHRM computation of the information matrix. Default is 1e-3

**symmetric** logical; force S-EM/Oakes information matrix estimates to be symmetric? Default is TRUE so that computation of standard errors are more stable. Setting this to FALSE can help to detect solutions that have not reached the ML estimate

**SEM_window** ratio of values used to define the S-EM window based on the observed likelihood differences across EM iterations. The default is c(0, 1 - SEtol), which provides nearly the very full S-EM window (i.e., nearly all EM cycles used). To use the a smaller SEM window change the window to to something like c(.9, .999) to start at a point farther into the EM history

**warn** logical; include warning messages during estimation? Default is TRUE

**message** logical; include general messages during estimation? Default is TRUE

**customK** a numeric vector used to explicitly declare the number of response categories for each item. This should only be used when constructing mirt model for reasons other than parameter estimation (such as to obtain factor scores), and requires that the input data all have 0 as the lowest category. The format is the same as the `extract.mirt(mod, 'K')` slot in all converged models

**customPriorFun** a custom function used to determine the normalized density for integration in the EM algorithm. Must be of the form `function(Theta, Etable){...}`, and return a numeric vector with the same length as number of rows in `Theta`. The `Etable` input contains the aggregated table generated from the current E-step computations. For proper integration, the returned vector should sum to 1 (i.e., normalized). Note that if using the `Etable` it will be NULL on the first call, therefore the prior will have to deal with this issue accordingly

**zeroExtreme** logical; assign extreme response patterns a `survey.weight` of 0 (formally equivalent to removing these data vectors during estimation)? When `dentype = 'EHW'`, where Woods' extrapolation is utilized, this option may be required if the extrapolation causes expected densities to tend towards positive or negative infinity. The default is `FALSE`

**customTheta** a custom `Theta` grid, in matrix form, used for integration. If not defined, the grid is determined internally based on the number of `quadpts`

**fixedTheta** a `matrix` of latent trait values taken to be fixed and known. This will perform a single M-step optimization to obtain item parameter estimates, holding constant the elements in `fixedTheta`, using the `'MHRM'` engine with the BFGS/L-BFGS-B algorithm. Matrix input must have as many rows as there are rows in `data`

**nconstrain** same specification as the `constrain` list argument, however imposes a negative equality constraint instead (e.g., $a12 = -a21$, which is specified as `nconstrain = list(c(12, 21))`). Note that each specification in the list must be of length 2, where the second element is taken to be -1 times the first element

**delta** the deviation term used in numerical estimates when computing the ACOV matrix with the 'forward' or 'central' numerical approaches, as well as Oakes' method with the Richardson extrapolation. Default is 1e-5

**parallel** logical; use the parallel cluster defined by `mirtCluster`? Default is TRUE

**storeEMhistory** logical; store the iteration history when using the EM algorithm? Default is FALSE. When TRUE, use `extract.mirt` to extract

**internal_constraints** logical; include the internal constraints when using certain IRT models (e.g., 'grsm' itemtype). Disable this if you want to use special optimizers such as the solnp. Default is `TRUE`

**gain** a vector of two values specifying the numerator and exponent values for the RM gain function $(val1/cycle)^val2$. Default is `c(0.10, 0.75)`

**BURNIN** number of burn in cycles (stage 1) in MH-RM; default is 150

**SEMCYCLES** number of SEM cycles (stage 2) in MH-RM; default is 100

**MHDRAWS** number of Metropolis-Hasting draws to use in the MH-RM at each iteration; default is 5

**MHcand** a vector of values used to tune the MH sampler. Larger values will cause the acceptance ratio to decrease. One value is required for each group in unconditional item factor analysis (`mixedmirt()` requires additional values for random effect). If null, these values are determined internally, attempting to tune the acceptance of the draws to be between .1 and .4

**MHRM_SE_draws** number of fixed draws to use when `SE=TRUE` and `SE.type = 'FMHRM'` and the maximum number of draws when `SE.type = 'MHRM'`. Default is 2000

**MCEM_draws** a function used to determine the number of quadrature points to draw for the `'MCEM'` method. Must include one argument which indicates the iteration number of the EM cycle. Default is `function(cycles) 500 + (cycles - 1)*2`, which starts the number of draws at 500 and increases by 2 after each full EM iteration

**info_if_converged** logical; compute the information matrix when using the MH-RM algorithm only if the model converged within a suitable number of iterations? Default is `TRUE`

**logLik_if_converged** logical; compute the observed log-likelihood when using the MH-RM algorithm only if the model converged within a suitable number of iterations? Default is `TRUE`

**keep_vcov_PD** logical; attempt to keep the variance-covariance matrix of the latent traits positive definite during estimation in the EM algorithm? This generally improves the convergence properties when the traits are highly correlated. Default is `TRUE`

... additional arguments to be passed

## Details

Models containing 'explanatory' person or item level predictors can only be included by using the `mixedmirt` function, though latent regression models can be fit using the `formula` input in this function. Tests that form a two-tier or bi-factor structure should be estimated with the `bfactor` function, which uses a dimension reduction EM algorithm for modeling item parcels. Multiple group analyses (useful for DIF and DTF testing) are also available using the `multipleGroup` function.

## Value

function returns an object of class `SingleGroupClass` (SingleGroupClass-class)

## Confirmatory and Exploratory IRT

Specification of the confirmatory item factor analysis model follows many of the rules in the structural equation modeling framework for confirmatory factor analysis. The variances of the latent factors are automatically fixed to 1 to help facilitate model identification. All parameters may be fixed to constant values or set equal to other parameters using the appropriate declarations. Confirmatory models may also contain 'explanatory' person or item level predictors, though including predictors is currently limited to the `mixedmirt` function.

When specifying a single number greater than 1 as the `model` input to mirt an exploratory IRT model will be estimated. Rotation and target matrix options are available if they are passed to generic functions such as `summary-method` and `fscores`. Factor means and variances are fixed to ensure proper identification.

If the model is an exploratory item factor analysis estimation will begin by computing a matrix of quasi-polychoric correlations. A factor analysis with `nfact` is then extracted and item parameters are estimated by $a_{ij} = f_{ij}/u_j$, where $f_{ij}$ is the factor loading for the *j*th item on the *i*th factor, and $u_j$ is the square root of the factor uniqueness, $\sqrt{1 - h_j^2}$. The initial intercept parameters are determined by calculating the inverse normal of the item facility (i.e., item easiness), $q_j$, to obtain $d_j = q_j/u_j$. A similar implementation is also used for obtaining initial values for polytomous items.

### A note on upper and lower bound parameters

Internally the $g$ and $u$ parameters are transformed using a logit transformation ($log(x/(1-x))$), and can be reversed by using $1/(1 + exp(-x))$ following convergence. This also applies when computing confidence intervals for these parameters, and is done so automatically if `coef(mod, rawug = FALSE)`.

As such, when applying prior distributions to these parameters it is recommended to use a prior that ranges from negative infinity to positive infinity, such as the normally distributed prior via the `'norm'` input (see `mirt.model`).

### Convergence for quadrature methods

Unrestricted full-information factor analysis is known to have problems with convergence, and some items may need to be constrained or removed entirely to allow for an acceptable solution. As a general rule dichotomous items with means greater than .95, or items that are only .05 greater than the guessing parameter, should be considered for removal from the analysis or treated with prior parameter distributions. The same type of reasoning is applicable when including upper bound parameters as well. For polytomous items, if categories are rarely endorsed then this will cause similar issues. Also, increasing the number of quadrature points per dimension, or using the quasi-Monte Carlo integration method, may help to stabilize the estimation process in higher dimensions. Finally, solutions that are not well defined also will have difficulty converging, and can indicate that the model has been misspecified (e.g., extracting too many dimensions).

### Convergence for MH-RM method

For the MH-RM algorithm, when the number of iterations grows very high (e.g., greater than 1500) or when `Max Change = .2500` values are repeatedly printed to the console too often (indicating that the parameters were being constrained since they are naturally moving in steps greater than 0.25) then the model may either be ill defined or have a very flat likelihood surface, and genuine maximum-likelihood parameter estimates may be difficult to find. Standard errors are computed following the model convergence by passing `SE = TRUE`, to perform an addition MH-RM stage but treating the maximum-likelihood estimates as fixed points.

### Additional helper functions

Additional functions are available in the package which can be useful pre- and post-estimation. These are:

`mirt.model` Define the IRT model specification use special syntax. Useful for defining between and within group parameter constraints, prior parameter distributions, and specifying the slope coefficients for each factor

`coef-method` Extract raw coefficients from the model, along with their standard errors and confidence intervals

`summary-method` Extract standardized loadings from model. Accepts a `rotate` argument for exploratory item response model

`anova-method` Compare nested models using likelihood ratio statistics as well as information criteria such as the AIC and BIC

`residuals-method` Compute pairwise residuals between each item using methods such as the LD statistic (Chen & Thissen, 1997), as well as response pattern residuals

`plot-method` Plot various types of test level plots including the test score and information functions and more

`itemplot` Plot various types of item level plots, including the score, standard error, and information functions, and more

`createItem` Create a customized `itemtype` that does not currently exist in the package

`imputeMissing` Impute missing data given some computed Theta matrix

`fscores` Find predicted scores for the latent traits using estimation methods such as EAP, MAP, ML, WLE, and EAPsum

`wald` Compute Wald statistics follow the convergence of a model with a suitable information matrix

`M2` Limited information goodness of fit test statistic based to determine how well the model fits the data

`itemfit` **and** `personfit` Goodness of fit statistics at the item and person levels, such as the S-X2, infit, outfit, and more

`boot.mirt` Compute estimated parameter confidence intervals via the bootstrap methods

`mirtCluster` Define a cluster for the package functions to use for capitalizing on multi-core architecture to utilize available CPUs when possible. Will help to decrease estimation times for tasks that can be run in parallel

## IRT Models

The parameter labels use the follow convention, here using two factors and $K$ as the total number of categories (using $k$ for specific category instances).

**Rasch** Only one intercept estimated, and the latent variance of $\theta$ is freely estimated. If the data have more than two categories then a partial credit model is used instead (see 'gpcm' below).

$$P(x = 1|\theta, d) = \frac{1}{1 + exp(-(\theta + d))}$$

**1-4PL** Depending on the model $u$ may be equal to 1 (e.g., 3PL), $g$ may be equal to 0 (e.g., 2PL), or the as may be fixed to 1 (e.g., 1PL).

$$P(x = 1|\theta, \psi) = g + \frac{(u - g)}{1 + exp(-(a_1 * \theta_1 + a_2 * \theta_2 + d))}$$

**5PL** Currently restricted to unidimensional models

$$P(x = 1|\theta, \psi) = g + \frac{(u - g)}{1 + exp(-(a_1 * \theta_1 + d))^S}$$

where $S$ allows for asymmetry in the response function and is transformation constrained to be greater than 0 (i.e., `log(S)` is estimated rather than `S`)

**CLL** Complementary log-log model (see Shim, Bonifay, and Wiedermann, 2022)

$$P(x = 1|\theta, b) = 1 - exp(-exp(\theta - b))$$

Currently restricted to unidimensional dichotomous data.

**graded** The graded model consists of sequential 2PL models,

$$P(x = k|\theta, \psi) = P(x \geq k|\theta, \phi) - P(x \geq k + 1|\theta, \phi)$$

Note that $P(x \geq 1|\theta, \phi) = 1$ while $P(x \geq K + 1|\theta, \phi) = 0$

**ULL** The unipolar log-logistic model (ULL; Lucke, 2015) is defined the same as the graded response model, however

$$P(x \leq k|\theta, \psi) = \frac{\lambda_k \theta^\eta}{1 + \lambda_k \theta^\eta}$$

. Internally the $\lambda$ parameters are exponentiated to keep them positive, and should therefore the reported estimates should be interpreted in log units

**grsm** A more constrained version of the graded model where graded spacing is equal across item blocks and only adjusted by a single 'difficulty' parameter (c) while the latent variance of $\theta$ is freely estimated (see Muraki, 1990 for this exact form). This is restricted to unidimensional models only.

**gpcm/nominal** For the gpcm the $d$ values are treated as fixed and ordered values from $0 : (K - 1)$ (in the nominal model $d_0$ is also set to 0). Additionally, for identification in the nominal model $ak_0 = 0$, $ak_{(K-1)} = (K - 1)$.

$$P(x = k|\theta, \psi) = \frac{exp(ak_{k-1} * (a_1 * \theta_1 + a_2 * \theta_2) + d_{k-1})}{\sum_{k=1}^{K} exp(ak_{k-1} * (a_1 * \theta_1 + a_2 * \theta_2) + d_{k-1})}$$

For the partial credit model (when `itemtype = 'Rasch'`; unidimensional only) the above model is further constrained so that $ak = (0, 1, \ldots, K - 1)$, $a_1 = 1$, and the latent variance of $\theta_1$ is freely estimated. Alternatively, the partial credit model can be obtained by containing all the slope parameters in the gpcms to be equal. More specific scoring function may be included by passing a suitable list or matrices to the `gpcm_mats` input argument.

In the nominal model this parametrization helps to identify the empirical ordering of the categories by inspecting the $ak$ values. Larger values indicate that the item category is more positively related to the latent trait(s) being measured. For instance, if an item was truly ordinal (such as a Likert scale), and had 4 response categories, we would expect to see $ak_0 < ak_1 < ak_2 < ak_3$ following estimation. If on the other hand $ak_0 > ak_1$ then it would appear that the second category is less related to to the trait than the first, and therefore the second category should be understood as the 'lowest score'.

NOTE: The nominal model can become numerical unstable if poor choices for the high and low values are chosen, resulting in `ak` values greater than `abs(10)` or more. It is recommended to choose high and low anchors that cause the estimated parameters to fall between 0 and $K - 1$ either by theoretical means or by re-estimating the model with better values following convergence.

**gpcmIRT and rsm** The gpcmIRT model is the classical generalized partial credit model for unidimensional response data. It will obtain the same fit as the `gpcm` presented above, however the parameterization allows for the Rasch/generalized rating scale model as a special case.

E.g., for a K = 4 category response model,

$$P(x = 0|\theta, \psi) = exp(0)/G$$
$$P(x = 1|\theta, \psi) = exp(a(\theta - b1) + c)/G$$
$$P(x = 2|\theta, \psi) = exp(a(2\theta - b1 - b2) + 2c)/G$$
$$P(x = 3|\theta, \psi) = exp(a(3\theta - b1 - b2 - b3) + 3c)/G$$

where

$$G = exp(0) + exp(a(\theta - b1) + c) + exp(a(2\theta - b1 - b2) + 2c) + exp(a(3\theta - b1 - b2 - b3) + 3c)$$

Here $a$ is the slope parameter, the $b$ parameters are the threshold values for each adjacent category, and $c$ is the so-called difficulty parameter when a rating scale model is fitted (otherwise, $c = 0$ and it drops out of the computations).

The gpcmIRT can be constrained to the partial credit IRT model by either constraining all the slopes to be equal, or setting the slopes to 1 and freeing the latent variance parameter.

Finally, the rsm is a more constrained version of the (generalized) partial credit model where the spacing is equal across item blocks and only adjusted by a single 'difficulty' parameter (c). Note that this is analogous to the relationship between the graded model and the grsm (with an additional constraint regarding the fixed discrimination parameters).

**sequential/Tutz** The multidimensional sequential response model has the form

$$P(x = k|\theta, \psi) = \prod(1 - F(a_1\theta_1 + a_2\theta_2 + d_{sk}))F(a_1\theta_1 + a_2\theta_2 + d_{jk})$$

where $F(\cdot)$ is the cumulative logistic function. The Tutz variant of this model (Tutz, 1990) (via `itemtype = 'Tutz'`) assumes that the slope terms are all equal to 1 and the latent variance terms are estimated (i.e., is a Rasch variant).

**ideal** The ideal point model has the form, with the upper bound constraint on $d$ set to 0:

$$P(x = 1|\theta, \psi) = exp(-0.5 * (a_1 * \theta_1 + a_2 * \theta_2 + d)^2)$$

**partcomp** Partially non-compensatory models consist of the product of 2PL probability curves.

$$P(x = 1|\theta, \psi) = g + (1 - g)(\frac{1}{1 + exp(-(a_1 * \theta_1 + d_1))}_1^c * \frac{1}{1 + exp(-(a_2 * \theta_2 + d_2))}_2^c)$$

where $c_1$ and $c_2$ are binary indicator variables reflecting whether the item should include the select compensatory component (1) or not (0). Note that constraining the slopes to be equal across items will reduce the model to Embretson's (Whitely's) multicomponent model (1980).

**PCgraded** Partially non-compensatory models generalization for graded response models as described in Chalmers (2020). Note that for numerical stability reasons the intercepts per dimension are expressed in a deviation form to satisfy the ordered-intercept constraint during optimization. The setup is generally the same as `partcomp` in terms of model declaration, while in the `K=2` case this reduces to the partially non-compensatory 2PL as a special case.

**2-4PLNRM** Nested logistic curves for modeling distractor items. Requires a scoring key. The model is broken into two components for the probability of endorsement. For successful endorsement the probability trace is the 1-4PL model, while for unsuccessful endorsement:

$$P(x = 0|\theta, \psi) = (1 - P_{1-4PL}(x = 1|\theta, \psi)) * P_{nominal}(x = k|\theta, \psi)$$

which is the product of the complement of the dichotomous trace line with the nominal response model. In the nominal model, the slope parameters defined above are constrained to be 1's, while the last value of the $ak$ is freely estimated.

**ggum** The (multidimensional) generalized graded unfolding model is a class of ideal point models useful for ordinal response data. The form is

$$P(z = k|\theta, \psi) = \frac{exp\left[\left(z\sqrt{\sum_{d=1}^{D} a_{id}^2(\theta_{jd} - b_{id})^2}\right) + \sum_{k=0}^{z} \psi_{ik}\right] + exp\left[\left((M - z)\sqrt{\sum_{d=1}^{D} a_{id}^2(\theta_{jd} - b_{id})^2}\right) + \sum_{k=0}^{z} \psi_{ik}\right]}{\sum_{w=0}^{C}\left(exp\left[\left(w\sqrt{\sum_{d=1}^{D} a_{id}^2(\theta_{jd} - b_{id})^2}\right) + \sum_{k=0}^{z} \psi_{ik}\right] + exp\left[\left((M - w)\sqrt{\sum_{d=1}^{D} a_{id}^2(\theta_{jd} - b\right.\right.\right.}$$

where $\theta_{jd}$ is the location of the $j$th individual on the $d$th dimension, $b_{id}$ is the difficulty location of the $i$th item on the $d$th dimension, $a_{id}$ is the discrimination of the $j$th individual on the $d$th dimension (where the discrimination values are constrained to be positive), $\psi_{ik}$ is the $k$th subjective response category threshold for the $i$th item, assumed to be symmetric about the item and constant across dimensions, where $\psi_{ik} = \sum_{d=1}^{D} a_{id}t_{ik}$ $z = 1, 2, \ldots, C$ (where $C$ is the number of categories minus 1), and $M = 2C + 1$.

**(g)hcm, (g)alm, (g)sslm, and (g)paralla** Following Luo (2001), this family of response models can be characterized under the same ordinal response functioning structure, differing only in their linking functions ($\psi(x)$). For example, for a two-dimensional model the equation used is

$$p_k = \frac{\psi(\rho_k)}{\psi(\rho_k) + \psi(a_1\theta_1 + a_2\theta_2 + d)}$$

which is expressed in slope-intercept form to accommodate multidimensionality. The "generalized" versions of this family estimate the slope and `rho` parameters, which allow each item to differ in the steepness of the unfolding model functions; otherwise, slopes are fixed to the value of 1, though the `rho` parameters must be set to 0 manually. For ordered polytomous items the response function follows the Guttman-scaling logic

$$P_1 = q_1 \cdot q_2 \cdot q_3 \cdots q_k$$

$$P_2 = p_1 \cdot q_2 \cdot q_3 \cdots q_k$$

$$P_3 = p_1 \cdot p_2 \cdot q_3 \cdots q_k$$

$$\cdots$$

$$P_K = p_1 \cdot p_2 \cdot p_3 \cdots p_k$$

Note that for estimation purposes the `rho` parameters are expressed in log units so that they remain positive during estimation. Hence, the parameterization used herein is

$$p_k = \frac{\psi(exp(\rho_k^*))}{\psi(exp(\rho_k^*))) + \psi(a_1\theta_1 + a_2\theta_2 + d)}$$

where $\rho_k^*$ is in natural log units. Currently supported models in this family are the:

- (generalized) hyperbolic cosine model ($\psi(x) = cosh(x)$),
- (generalized) absolute logistic model ($\psi(x) = exp(|x|)$),
- (generalized) simple squared logistic model ($\psi(x) = exp(x^2)$), and
- (generalized) parallellogram analysis model ($\psi(x) = x^2$), respectively.

all of which are available for dichotomous and ordered polytomous response option items.

**spline** Spline response models attempt to model the response curves uses non-linear and potentially non-monotonic patterns. The form is

$$P(x = 1|\theta, \eta) = \frac{1}{1 + exp(-(\eta_1 * X_1 + \eta_2 * X_2 + \cdots + \eta_n * X_n))}$$

where the $X_n$ are from the spline design matrix $X$ organized from the grid of $\theta$ values. B-splines with a natural or polynomial basis are supported, and the `intercept` input is set to `TRUE` by default.

**monospline** The structure of the monotone spline is the same as the `'spline'` type, however is built from a constrained version of an I-spline from `iSpline`, which are cumulative sums of B-splines. In this case, the intercept is left as an unconstrained parameter to be estimated, however all other terms are constrained to be positive. For this model to work correctly `intercept` must always be `TRUE`.

**monopoly** Monotone polynomial model for polytomous response data of the form

$$P(x = k|\theta, \psi) = \frac{exp(\sum_1^k (m^*(\psi) + \xi_{c-1})}{\sum_1^C exp(\sum_1^K (m^*(\psi) + \xi_{c-1}))}$$

where $m^*(\psi)$ is the monotone polynomial function without the intercept.

## HTML help files, exercises, and examples

To access examples, vignettes, and exercise files that have been generated with knitr please visit https://github.com/philchalmers/mirt/wiki.

## Author(s)

Phil Chalmers <rphilip.chalmers@gmail.com>

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## See Also

`bfactor`, `multipleGroup`, `mixedmirt`, `expand.table`, `key2binary`, `mod2values`, `extract.item`, `iteminfo`, `testinfo`, `probtrace`, `simdata`, `averageMI`, `fixef`, `extract.mirt`, `itemstats`

## Examples

```
# load LSAT section 7 data and compute 1 and 2 factor models
data <- expand.table(LSAT7)
itemstats(data)

(mod1 <- mirt(data, 1))
```

```
coef(mod1)
summary(mod1)
plot(mod1)
plot(mod1, type = 'trace')


(mod2 <- mirt(data, 1, SE = TRUE)) #standard errors via the Oakes method
(mod2 <- mirt(data, 1, SE = TRUE, SE.type = 'SEM')) #standard errors with SEM method
coef(mod2)
(mod3 <- mirt(data, 1, SE = TRUE, SE.type = 'Richardson')) #with numerical Richardson method
residuals(mod1)
plot(mod1) #test score function
plot(mod1, type = 'trace') #trace lines
plot(mod2, type = 'info') #test information
plot(mod2, MI=200) #expected total score with 95% confidence intervals

# estimated 3PL model for item 5 only
(mod1.3PL <- mirt(data, 1, itemtype = c('2PL', '2PL', '2PL', '2PL', '3PL')))
coef(mod1.3PL)

# internally g and u pars are stored as logits, so usually a good idea to include normal prior
#  to help stabilize the parameters. For a value around .182 use a mean
#  of -1.5 (since 1 / (1 + exp(-(-1.5))) == .182)
model <- 'F = 1-5
         PRIOR = (5, g, norm, -1.5, 3)'
mod1.3PL.norm <- mirt(data, model, itemtype = c('2PL', '2PL', '2PL', '2PL', '3PL'))
coef(mod1.3PL.norm)
#limited information fit statistics
M2(mod1.3PL.norm)

# unidimensional ideal point model
idealpt <- mirt(data, 1, itemtype = 'ideal')
plot(idealpt, type = 'trace', facet_items = TRUE)
plot(idealpt, type = 'trace', facet_items = FALSE)

# two factors (exploratory)
mod2 <- mirt(data, 2)
coef(mod2)
summary(mod2, rotate = 'oblimin') #oblimin rotation
residuals(mod2)
plot(mod2)
plot(mod2, rotate = 'oblimin')

anova(mod1, mod2) #compare the two models
scoresfull <- fscores(mod2) #factor scores for each response pattern
head(scoresfull)
scorestable <- fscores(mod2, full.scores = FALSE) #save factor score table
head(scorestable)

# confirmatory (as an example, model is not identified since you need 3 items per factor)
# Two ways to define a confirmatory model: with mirt.model, or with a string

# these model definitions are equivalent
```

```
cmodel <- mirt.model('
   F1 = 1,4,5
   F2 = 2,3')
cmodel2 <- 'F1 = 1,4,5
            F2 = 2,3'

cmod <- mirt(data, cmodel)
# cmod <- mirt(data, cmodel2) # same as above
coef(cmod)
anova(cmod, mod2)
# check if identified by computing information matrix
(cmod <- mirt(data, cmodel, SE = TRUE))

############
# data from the 'ltm' package in numeric format
itemstats(Science)

pmod1 <- mirt(Science, 1)
plot(pmod1)
plot(pmod1, type = 'trace')
plot(pmod1, type = 'itemscore')
summary(pmod1)

# Constrain all slopes to be equal with the constrain = list() input or mirt.model() syntax
# first obtain parameter index
values <- mirt(Science,1, pars = 'values')
values #note that slopes are numbered 1,5,9,13, or index with values$parnum[values$name == 'a1']
(pmod1_equalslopes <- mirt(Science, 1, constrain = list(c(1,5,9,13))))
coef(pmod1_equalslopes)

# using mirt.model syntax, constrain all item slopes to be equal
model <- 'F = 1-4
          CONSTRAIN = (1-4, a1)'
(pmod1_equalslopes <- mirt(Science, model))
coef(pmod1_equalslopes)

coef(pmod1_equalslopes)
anova(pmod1_equalslopes, pmod1) #significantly worse fit with almost all criteria

pmod2 <- mirt(Science, 2)
summary(pmod2)
plot(pmod2, rotate = 'oblimin')
itemplot(pmod2, 1, rotate = 'oblimin')
anova(pmod1, pmod2)

# unidimensional fit with a generalized partial credit and nominal model
(gpcmod <- mirt(Science, 1, 'gpcm'))
coef(gpcmod)

# for the nominal model the lowest and highest categories are assumed to be the
#  theoretically lowest and highest categories that related to the latent trait(s)
(nomod <- mirt(Science, 1, 'nominal'))
coef(nomod) #ordering of ak values suggest that the items are indeed ordinal
```

```
anova(gpcmod, nomod)
itemplot(nomod, 3)

# generalized graded unfolding model
(ggum <- mirt(Science, 1, 'ggum'))
coef(ggum, simplify=TRUE)
plot(ggum)
plot(ggum, type = 'trace')
plot(ggum, type = 'itemscore')

# monotonic polyomial models
(monopoly <- mirt(Science, 1, 'monopoly'))
coef(monopoly, simplify=TRUE)
plot(monopoly)
plot(monopoly, type = 'trace')
plot(monopoly, type = 'itemscore')

# unipolar IRT model
unimod <- mirt(Science, itemtype = 'ULL')
coef(unimod, simplify=TRUE)
plot(unimod)
plot(unimod, type = 'trace')
itemplot(unimod, 1)

# following use the correct log-normal density for latent trait
itemfit(unimod)
M2(unimod, type = 'C2')
fs <- fscores(unimod)
hist(fs, 20)
fscores(unimod, method = 'EAPsum', full.scores = FALSE)

## example applying survey weights.
# weight the first half of the cases to be more representative of population
survey.weights <- c(rep(2, nrow(Science)/2), rep(1, nrow(Science)/2))
survey.weights <- survey.weights/sum(survey.weights) * nrow(Science)
unweighted <- mirt(Science, 1)
weighted <- mirt(Science, 1, survey.weights=survey.weights)

############
# empirical dimensionality testing that includes 'guessing'

data(SAT12)
data <- key2binary(SAT12,
  key = c(1,4,5,2,3,1,2,1,3,1,2,4,2,1,5,3,4,4,1,4,3,3,4,1,3,5,1,3,1,5,4,5))
itemstats(data)

mod1 <- mirt(data, 1)
extract.mirt(mod1, 'time') #time elapsed for each estimation component

# optionally use Newton-Raphson for (generally) faster convergence in the M-step's
mod1 <- mirt(data, 1, optimizer = 'NR')
extract.mirt(mod1, 'time')
```

```
mod2 <- mirt(data, 2, optimizer = 'NR')
# difficulty converging with reduced quadpts, reduce TOL
mod3 <- mirt(data, 3, TOL = .001, optimizer = 'NR')
anova(mod1,mod2)
anova(mod2, mod3) #negative AIC, 2 factors probably best

# same as above, but using the QMCEM method for generally better accuracy in mod3
mod3 <- mirt(data, 3, method = 'QMCEM', TOL = .001, optimizer = 'NR')
anova(mod2, mod3)

# with fixed guessing parameters
mod1g <- mirt(data, 1, guess = .1)
coef(mod1g)

############
# graded rating scale example

# make some data
set.seed(1234)
a <- matrix(rep(1, 10))
d <- matrix(c(1,0.5,-.5,-1), 10, 4, byrow = TRUE)
c <- seq(-1, 1, length.out=10)
data <- simdata(a, d + c, 2000, itemtype = rep('graded',10))
itemstats(data)

mod1 <- mirt(data, 1)
mod2 <- mirt(data, 1, itemtype = 'grsm')
coef(mod2)
anova(mod2, mod1) #not sig, mod2 should be preferred
itemplot(mod2, 1)
itemplot(mod2, 5)
itemplot(mod2, 10)

############
# 2PL nominal response model example (Suh and Bolt, 2010)
data(SAT12)
SAT12[SAT12 == 8] <- NA #set 8 as a missing value
head(SAT12)

# correct answer key
key <- c(1,4,5,2,3,1,2,1,3,1,2,4,2,1,5,3,4,4,1,4,3,3,4,1,3,5,1,3,1,5,4,5)
scoredSAT12 <- key2binary(SAT12, key)
mod0 <- mirt(scoredSAT12, 1)

# for first 5 items use 2PLNRM and nominal
scoredSAT12[,1:5] <- as.matrix(SAT12[,1:5])
mod1 <- mirt(scoredSAT12, 1, c(rep('nominal',5),rep('2PL', 27)))
mod2 <- mirt(scoredSAT12, 1, c(rep('2PLNRM',5),rep('2PL', 27)), key=key)
coef(mod0)$Item.1
coef(mod1)$Item.1
coef(mod2)$Item.1
itemplot(mod0, 1)
itemplot(mod1, 1)
```

```
itemplot(mod2, 1)

# compare added information from distractors
Theta <- matrix(seq(-4,4,.01))
par(mfrow = c(2,3))
for(i in 1:5){
    info <- iteminfo(extract.item(mod0,i), Theta)
    info2 <- iteminfo(extract.item(mod2,i), Theta)
   plot(Theta, info2, type = 'l', main = paste('Information for item', i), ylab = 'Information')
    lines(Theta, info, col = 'red')
}
par(mfrow = c(1,1))

# test information
plot(Theta, testinfo(mod2, Theta), type = 'l', main = 'Test information', ylab = 'Information')
lines(Theta, testinfo(mod0, Theta), col = 'red')

###########
# using the MH-RM algorithm
data(LSAT7)
fulldata <- expand.table(LSAT7)
(mod1 <- mirt(fulldata, 1, method = 'MHRM'))

# Confirmatory models

# simulate data
a <- matrix(c(
1.5,NA,
0.5,NA,
1.0,NA,
1.0,0.5,
 NA,1.5,
 NA,0.5,
 NA,1.0,
 NA,1.0),ncol=2,byrow=TRUE)

d <- matrix(c(
-1.0,NA,NA,
-1.5,NA,NA,
 1.5,NA,NA,
 0.0,NA,NA,
3.0,2.0,-0.5,
2.5,1.0,-1,
2.0,0.0,NA,
1.0,NA,NA),ncol=3,byrow=TRUE)

sigma <- diag(2)
sigma[1,2] <- sigma[2,1] <- .4
items <- c(rep('2PL',4), rep('graded',3), '2PL')
dataset <- simdata(a,d,2000,items,sigma)

# analyses
# CIFA for 2 factor crossed structure
```

```
model.1 <- '
  F1 = 1-4
  F2 = 4-8
  COV = F1*F2'

# compute model, and use parallel computation of the log-likelihood
if(interactive()) mirtCluster()
mod1 <- mirt(dataset, model.1, method = 'MHRM')
coef(mod1)
summary(mod1)
residuals(mod1)

#####
# bifactor
model.3 <- '
  G = 1-8
  F1 = 1-4
  F2 = 5-8'

mod3 <- mirt(dataset,model.3, method = 'MHRM')
coef(mod3)
summary(mod3)
residuals(mod3)
anova(mod1,mod3)

#####
# polynomial/combinations
data(SAT12)
data <- key2binary(SAT12,
               key = c(1,4,5,2,3,1,2,1,3,1,2,4,2,1,5,3,4,4,1,4,3,3,4,1,3,5,1,3,1,5,4,5))

model.quad <- '
       F1 = 1-32
  (F1*F1) = 1-32'


model.combo <- '
       F1 = 1-16
       F2 = 17-32
  (F1*F2) = 1-8'

(mod.quad <- mirt(data, model.quad))
summary(mod.quad)
(mod.combo <- mirt(data, model.combo))
anova(mod.combo, mod.quad)

# non-linear item and test plots
plot(mod.quad)
plot(mod.combo, type = 'SE')
itemplot(mod.quad, 1, type = 'score')
itemplot(mod.combo, 2, type = 'score')
itemplot(mod.combo, 2, type = 'infocontour')
```

```
## empirical histogram examples (normal, skew and bimodality)
# make some data
set.seed(1234)
a <- matrix(rlnorm(50, .2, .2))
d <- matrix(rnorm(50))
ThetaNormal <- matrix(rnorm(2000))
ThetaBimodal <- scale(matrix(c(rnorm(1000, -2), rnorm(1000,2)))) #bimodal
ThetaSkew <- scale(matrix(rchisq(2000, 3))) #positive skew
datNormal <- simdata(a, d, 2000, itemtype = '2PL', Theta=ThetaNormal)
datBimodal <- simdata(a, d, 2000, itemtype = '2PL', Theta=ThetaBimodal)
datSkew <- simdata(a, d, 2000, itemtype = '2PL', Theta=ThetaSkew)

normal <- mirt(datNormal, 1, dentype = "empiricalhist")
plot(normal, type = 'empiricalhist')
histogram(ThetaNormal, breaks=30)

bimodal <- mirt(datBimodal, 1, dentype = "empiricalhist")
plot(bimodal, type = 'empiricalhist')
histogram(ThetaBimodal, breaks=30)

skew <- mirt(datSkew, 1, dentype = "empiricalhist")
plot(skew, type = 'empiricalhist')
histogram(ThetaSkew, breaks=30)

#####
# non-linear parameter constraints with Rsolnp package (nloptr supported as well):
# Find Rasch model subject to the constraint that the intercepts sum to 0

dat <- expand.table(LSAT6)
itemstats(dat)

# free latent mean and variance terms
model <- 'Theta = 1-5
          MEAN = Theta
          COV = Theta*Theta'

# view how vector of parameters is organized internally
sv <- mirt(dat, model, itemtype = 'Rasch', pars = 'values')
sv[sv$est, ]

# constraint: create function for solnp to compute constraint, and declare value in eqB
eqfun <- function(p, optim_args) sum(p[1:5]) #could use browser() here, if it helps
LB <- c(rep(-15, 6), 1e-4) # more reasonable lower bound for variance term

mod <- mirt(dat, model, sv=sv, itemtype = 'Rasch', optimizer = 'solnp',
   solnp_args=list(eqfun=eqfun, eqB=0, LB=LB))
print(mod)
coef(mod)
(ds <- sapply(coef(mod)[1:5], function(x) x[,'d']))
sum(ds)

# same likelihood location as: mirt(dat, 1, itemtype = 'Rasch')
```

```
#######
# Latent regression Rasch model, followed by factor score regression

# Example assumes that exogenous variables relate to latent trait,
# and latent trait related to endogenous variables not included in the
# measurement model. Equates to the following structure
#
# Exogenous (X3 not related):
# Theta ~ beta*X1 + beta*X2 + 0*X3 + e
#
# Endogenous:
# Y0 ~ 0*Theta + e      (no relationship)
# Y1 ~ gamma*Theta + e
# Y2 ~ gamma*Theta + gamma*Theta^2 + e

# generate suitable data
set.seed(1234)
N <- 1000

# covariates (exogeneous variables for Theta)
X1 <- rnorm(N); X2 <- rnorm(N)
covdata <- data.frame(X1, X2, X3 = rnorm(N))
Theta <- matrix(0.2 * X1  -0.3 * X2 + rnorm(N, sd=sqrt(1 - .2^2 - .3^2)))
var(Theta)

# relation to endogenous variables, Y1 and Y2, not included in measurement model
Y0 <- rnorm(N)   # no relationship
Y1 <- .5 * Theta + rnorm(N)
Y2 <- .5 * Theta + .25 * Theta^2 + rnorm(N)

# true relationship between Y and Theta
lm(Y0 ~ Theta) |> summary()
lm(Y1 ~ Theta) |> summary()
lm(Y2 ~ Theta + I(Theta^2)) |> summary()

# items and response data
a <- matrix(1, 20); d <- matrix(rnorm(20))
dat <- simdata(a, d, N, itemtype = '2PL', Theta=Theta)

# unconditional Rasch model
mod0 <- mirt(dat, 1, 'Rasch', SE=TRUE)
coef(mod0, printSE=TRUE)

#######
# For exogenous predictor variables (e.g., Theta ~ X1 + X2 + e)
#   include directly in model, and perform LR/Wald tests
#######
#
# conditional model using X1, X2, and X3 (bad) as predictors of Theta
mod1 <- mirt(dat, 1, 'Rasch', covdata=covdata, formula = ~ X1 + X2 + X3, SE=TRUE)
coef(mod1, printSE=TRUE)
```

```
coef(mod1, simplify=TRUE)
anova(mod0, mod1)  # jointly significant predictors of theta

# large sample z-ratios and p-values via Wald tests
cfs <- coef(mod1, printSE=TRUE)
(z <- cfs$lr.betas[[1]] / cfs$lr.betas[[2]])
ps <- round(pnorm(abs(z[,1]), lower.tail=FALSE)*2, 3)
names(ps) <- paste0('p.', names(ps))
ps

# drop X3 predictor for nested LR comparison
mod1b <- mirt(dat, 1, 'Rasch', covdata=covdata, formula = ~ X1 + X2)
anova(mod1b, mod1)  # same information as Wald test

#####
# For exogenous Y's, treat scores as in factor score regression.
#  However, update the VCOV/SEs and R^2 as the regression
#  assumes the scores in the IV side are error-free (too optimistic)
#####

# EAP estimates and SEs
fs <- fscores(mod1b, full.scores.SE=TRUE)
theta <- fs[,1]

# relationship with endogenous Y variable
summary(lmmod0 <- lm(Y0 ~ theta))
summary(lmmod1 <- lm(Y1 ~ theta))
summary(lmmod2 <- lm(Y2 ~ theta + I(theta^2)))

# improve SEs/R^2 in regression model by accounting for unreliability
(rxx <- empirical_rxx(fs))

# updates most lm() output, but not everything (do not use for other purposes!)
update.summary <- function(lmmod, rxx){
  so <- summary(lmmod)
  so$r.squared <- so$r.squared / rxx
  so$adj.r.squared <- so$adj.r.squared / rxx
  so$cov.unscaled  <- vcov(lmmod) / rxx
  so$coefficients[,2] <- sqrt(diag(so$cov.unscaled))
  so$coefficients[,3] <- so$coefficients[,1] / so$coefficients[,2]
  so$coefficients[,4] <- pt(abs(so$coefficients[,3]), df=lmmod$df.residual,
    lower.tail = FALSE) * 2
  so
}

update.summary(lmmod0, rxx)
update.summary(lmmod1, rxx)
update.summary(lmmod2, rxx)


############
# Example using Gauss-Hermite quadrature with custom input functions
```

```
## Not run:
library(fastGHQuad)
data(SAT12)
data <- key2binary(SAT12,
               key = c(1,4,5,2,3,1,2,1,3,1,2,4,2,1,5,3,4,4,1,4,3,3,4,1,3,5,1,3,1,5,4,5))
GH <- gaussHermiteData(50)
Theta <- matrix(GH$x)

# This prior works for uni- and multi-dimensional models
prior <- function(Theta, Etable){
    P <- grid <- GH$w / sqrt(pi)
    if(ncol(Theta) > 1)
        for(i in 2:ncol(Theta))
            P <- expand.grid(P, grid)
     if(!is.vector(P)) P <- apply(P, 1, prod)
     P
}

GHmod1 <- mirt(data, 1, optimizer = 'NR',
              technical = list(customTheta = Theta, customPriorFun = prior))
coef(GHmod1, simplify=TRUE)

Theta2 <- as.matrix(expand.grid(Theta, Theta))
GHmod2 <- mirt(data, 2, optimizer = 'NR', TOL = .0002,
              technical = list(customTheta = Theta2, customPriorFun = prior))
summary(GHmod2, suppress=.2)

## End(Not run)

############
# Davidian curve example

dat <- key2binary(SAT12,
               key = c(1,4,5,2,3,1,2,1,3,1,2,4,2,1,5,3,4,4,1,4,3,3,4,1,3,5,1,3,1,5,4,5))
dav <- mirt(dat, 1, dentype = 'Davidian-4') # use four smoothing parameters
plot(dav, type = 'Davidian') # shape of latent trait distribution
coef(dav, simplify=TRUE)

fs <- fscores(dav) # assume normal prior
fs2 <- fscores(dav, use_dentype_estimate=TRUE) # use Davidian estimated prior shape
head(cbind(fs, fs2))

itemfit(dav) # assume normal prior
itemfit(dav, use_dentype_estimate=TRUE) # use Davidian estimated prior shape

############
# Unfolding models

# polytomous hyperbolic cosine model with
#  estimated latitude of acceptance (rho parameters)
mod <- mirt(Science, model=1, itemtype = 'hcm')
coef(mod, simplify=TRUE)$items
coef(mod, simplify=TRUE, IRTpars=TRUE)$items
```

```
plot(mod)
plot(mod, type = 'trace')
plot(mod, type = 'itemscore')

# EAP estimates
fs <- fscores(mod)
head(fs)

itemfit(mod)
M2(mod, type = 'C2')

############
# 5PL and restricted 5PL example
dat <- expand.table(LSAT7)

mod2PL <- mirt(dat)
mod2PL

# Following does not converge without including strong priors
# mod5PL <- mirt(dat, itemtype = '5PL')
# mod5PL

# restricted version of 5PL (asymmetric 2PL)
model <- 'Theta = 1-5
          FIXED = (1-5, g), (1-5, u)'

mod2PL_asym <- mirt(dat, model=model, itemtype = '5PL')
mod2PL_asym
coef(mod2PL_asym, simplify=TRUE)
coef(mod2PL_asym, simplify=TRUE, IRTpars=TRUE)

# no big difference statistically or visually
anova(mod2PL, mod2PL_asym)
plot(mod2PL, type = 'trace')
plot(mod2PL_asym, type = 'trace')


###################
# LLTM example

a <- matrix(rep(1,30))
d <- rep(c(1,0, -1),each = 10)  # first easy, then medium, last difficult
dat <- simdata(a, d, 1000, itemtype = '2PL')

# unconditional model for intercept comparisons
mod <- mirt(dat, itemtype = 'Rasch')
coef(mod, simplify=TRUE)

# Suppose that the first 10 items were suspected to be easy, followed by 10 medium difficulty items,
# then finally the last 10 items are difficult,
# and we wish to test this item structure hypothesis (more intercept designs are possible
# by including more columns).
```

```
itemdesign <- data.frame(difficulty =
   factor(c(rep('easy', 10), rep('medium', 10), rep('hard', 10))))
rownames(itemdesign) <- colnames(dat)
itemdesign

# LLTM with mirt()
lltm <- mirt(dat, itemtype = 'Rasch', SE=TRUE,
   item.formula = ~ 0 + difficulty, itemdesign=itemdesign)
coef(lltm, simplify=TRUE)
coef(lltm, printSE=TRUE)
anova(lltm, mod)  # models fit effectively the same; hence, intercept variability well captured

# additional information for LLTM
plot(lltm)
plot(lltm, type = 'trace')
itemplot(lltm, item=1)
itemfit(lltm)
head(fscores(lltm))  #EAP estimates
fscores(lltm, method='EAPsum', full.scores=FALSE)
M2(lltm) # goodness of fit
head(personfit(lltm))
residuals(lltm)

# intercept across items also possible by removing ~ 0 portion, just interpreted differently
lltm.int <- mirt(dat, itemtype = 'Rasch',
   item.formula = ~ difficulty, itemdesign=itemdesign)
anova(lltm, lltm.int) # same
coef(lltm.int, simplify=TRUE)

# using unconditional modeling for first four items
itemdesign.sub <- itemdesign[5:nrow(itemdesign), , drop=FALSE]
itemdesign.sub    # note that rownames are required in this case
lltm.4 <- mirt(dat, itemtype = 'Rasch',
   item.formula = ~ 0 + difficulty, itemdesign=itemdesign.sub)
coef(lltm.4, simplify=TRUE) # first four items are the standard Rasch
anova(lltm, lltm.4) # similar fit, hence more constrained model preferred

# LLTM with mixedmirt() (more flexible in general, but slower)
LLTM <- mixedmirt(dat, model=1, fixed = ~ 0 + difficulty,
                  itemdesign=itemdesign, SE=FALSE)
summary(LLTM)
coef(LLTM)

# LLTM with random error estimate (not supported with mirt() )
LLTM.e <- mixedmirt(dat, model=1, fixed = ~ 0 + difficulty,
                  random = ~ 1|items, itemdesign=itemdesign, SE=FALSE)
coef(LLTM.e)


###################
# General MLTM example (Embretson, 1984)

set.seed(42)
```

```
as <- matrix(rep(1,60), ncol=2)
as[11:18,1] <- as[1:9,2] <- 0
d1 <- rep(c(3,1),each = 6)  # first easy, then medium, last difficult for first trait
d2 <- rep(c(0,1,2),times = 4)    # difficult to easy
d <- rnorm(18)
ds <- rbind(cbind(d1=NA, d2=d), cbind(d1, d2))
(pars <- data.frame(a=as, d=ds))
dat <- simdata(as, ds, 2500,
  itemtype = c(rep('dich', 18), rep('partcomp', 12)))
itemstats(dat)

# unconditional model
syntax <- "theta1 = 1-9, 19-30
           theta2 = 10-30
           COV = theta1*theta2"
itemtype <- c(rep('Rasch', 18), rep('PC1PL', 12))
mod <- mirt(dat, syntax, itemtype=itemtype)
coef(mod, simplify=TRUE)
data.frame(est=coef(mod, simplify=TRUE)$items, pop=data.frame(a=as, d=ds))
itemplot(mod, 1)
itemplot(mod, 30)

# MLTM design only for PC1PL items
itemdesign <- data.frame(t1_difficulty= factor(d1, labels=c('medium', 'easy')),
                         t2_difficulty=factor(d2, labels=c('hard', 'medium', 'easy')))
rownames(itemdesign) <- colnames(dat)[19:30]
itemdesign

# fit MLTM design, leaving first 18 items as 'Rasch' type
mltm <- mirt(dat, syntax, itemtype=itemtype, itemdesign=itemdesign,
             item.formula = list(theta1 ~ 0 + t1_difficulty,
                                 theta2 ~ 0 + t2_difficulty), SE=TRUE)
coef(mltm, simplify=TRUE)
coef(mltm, printSE=TRUE)
anova(mltm, mod) # similar fit; hence more constrained version preferred
M2(mltm) # goodness of fit
head(personfit(mltm))
residuals(mltm)

# EAP estimates
fscores(mltm) |> head()
```

---

`mirt.model` *Specify model information*

---

## Description

The `mirt.model` function scans/reads user input to specify the confirmatory model. Item locations must be used in the specifications if no `itemnames` argument is supplied. This is called implicitly by estimation functions when a string is passed to the `model` argument.

## Usage

```
mirt.model(
  input = NULL,
  itemnames = NULL,
  file = "",
  COV = NULL,
  quiet = TRUE,
  ...
)
```

## Arguments

| | |
|---|---|
| `input` | input for writing out the model syntax. Can either be a string declaration of class character or the so-called Q-matrix or class `matrix` that specifies the model either with integer or logical values. If the Q-matrix method is chosen covariances terms can be specified with the `COV` input |
| `itemnames` | a character vector or factor indicating the item names. If a data.frame or matrix object is supplied the names will be extracted using `colnames(itemnames)`. Supplying this input allows the syntax to be specified with the raw item names rather than item locations |
| `file` | a input specifying an external file that declares the input. |
| `COV` | a symmetric, logical matrix used to declare which covariance terms are estimated |
| `quiet` | logical argument passed to `scan()` to suppress console read message |
| `...` | additional arguments for `scan()` |

## Details

Factors are first named and then specify which numerical items they affect (i.e., where the slope is not equal to 0), separated either by commas or by - to indicate a range of items. Products between factors may be specified by enclosing the left hand term within brackets. To finish the declaration of a model simply enter a blank line with only a carriage return (i.e., the 'enter' or 'return' key), or instead read in an input version of the model syntax. The associated slopes throughout the package label these coefficients as `a1, a2, ..., ak`, where the associated number is assigned according to the respective order of the defined factors.

For example, if the syntax were

`"G = 1-10 F = 1-5 A = 6-10"`

then the `G` factor would be assigned the slopes `a1` for each item, `F` assigned the slopes `a2`, and `A` assigned the slopes `a3`. The same principle applies to the `bfactor` function whereby the slopes are automatically included for the specific factors after the general factor structure has been assigned.

There is an optional keyword for specifying the correlation between relationships between factors called COV, and non-linear factor products can be included by enclosing the product combination on the left hand side of the declaration (e.g., (F1*F1) would create a quadratic factor for F1).

The keywords CONSTRAIN, CONSTRAINB, PRIOR, FIXED, FREE, START, UBOUND, LBOUND can be applied to specific sub-groups in multiple-group models by included square brackets before the = sign, where groups are separated by commas. For example, to apply within-group equality constraints to a group called "male", then specifying:

```
CONSTRAIN [male] = (1-5, a1)
```

is appropriate, while specifying the same constraints to the sub-groups "male" and "female" would appear as

```
CONSTRAIN [male, female] = (1-5, a1)
```

For all other groups in the multi-group model, these within-group equality constraints would not appear. Therefore, these bracketed group specifications are useful when modifying priors, starting values, between/within group equality constraints, and so on when the specifications for each sub-group may differ.

Additionally, the use of negations can be used to omit specific groups in the constraint specifications by prefixing the string with a - operator, such as the following which applies between-group constraints to all groups except "Group2" and "Group3":

```
CONSTRAINB [-Group2, -Group3] = (1-5, a1)
```

Finally, the keyword GROUP can be used to specify the group-level hyper-parameter terms, such as the means and variance of the default Gaussian distribution. For example, to set the starting value of the variance parameter (COV_11) to 1.5:

```
START = (GROUP, COV_11, 1.5)
```

**COV** Specify the relationship between the latent factors. Estimating a correlation between factors is declared by joining the two factors with an asterisk (e.g., F1*F2), or with an asterisk between three or more factors to estimate all the possible correlations (e.g., F1*F2*F3). Specifications with the same factor (e.g., F1*F1) will free the variance of said factor instead

**MEAN** A comma separated list specifying which latent factor means to freely estimate. E.g., MEAN = F1, F2 will free the latent means for factors F1 and F2

**CONSTRAIN** A bracketed, comma separated list specifying equality constrains between items. The input format is CONSTRAIN = (items, ..., parameterName(s)), (items, ..., parameterName).

For example, in a single group 10-item dichotomous tests, using the default 2PL model, the first and last 5 item slopes (a1) can be constrained to be equal by using CONSTRAIN = (1-5, a1), (6-10, a1), or some combination such as CONSTRAIN = (1-3,4,5,a1), (6,7,8-10,a1).

When constraining parameters to be equal across items with different parameter names, a balanced bracketed vector must be supplied. E.g., setting the first slope for item 1 equal to the second slope in item 3 would be CONSTRAIN = (1, 3, a1, a2)

**CONSTRAINB** A bracketed, comma separate list specifying equality constrains between groups. The input format is CONSTRAINB = (items, ..., parameterName), (items, ..., parameterName).

For example, in a two group 10-item dichotomous tests, using the default 2PL model, the first 5 item slopes (a1) can be constrained to be equal across both groups by using CONSTRAINB = (1-5, a1), or some combination such as CONSTRAINB = (1-3,4,5,a1)

**PRIOR** A bracketed, comma separate list specifying prior parameter distributions. The input format is `PRIOR = (items, ..., parameterName, priorType, val1, val2), (items, ..., parameterName, priorType, val1, val2)`. For example, in a single group 10-item dichotomous tests, using the default 2PL model, defining a normal prior of N(0,2) for the first 5 item intercepts (d) can be defined by `PRIOR = (1-5, d, norm, 0, 2)`

Currently supported priors are of the form: `(items, norm, mean, sd)` for the normal/Gaussian, `(items, lnorm, log_mean, log_sd)` for log-normal, `(items, beta, alpha, beta)` for beta, and `(items, expbeta, alpha, beta)` for the beta distribution after applying the function `plogis` to the input value (note, this is specifically for applying a beta prior to the lower-bound parameters in 3/4PL models). Note that the beta prior specification focuses on the mode of the beta distribution centered around the value $\frac{\alpha}{(\alpha+\beta)}$ rather than the mean; for the mean of the beta distribution specify `alpha + 1` and `beta + 1` instead

**LBOUND** A bracketed, comma separate list specifying lower bounds for estimated parameters (used in optimizers such as `L-BFGS-B` and `nlminb`). The input format is `LBOUND = (items, ..., parameterName, value), (items, ..., parameterName, value)`.

For example, in a single group 10-item dichotomous tests, using the 3PL model and setting lower bounds for the 'g' parameters for the first 5 items to 0.2 is accomplished with `LBOUND = (1-5, g, 0.2)`

**UBOUND** same as LBOUND, but specifying upper bounds in estimated parameters

**START** A bracketed, comma separate list specifying the starting values for individual parameters. The input is of the form `(items, ..., parameterName, value)`. For instance, setting the 10th and 12th to 15th item slope parameters (a1) to 1.0 is specified with `START = (10, 12-15, a1, 1.0)`

For more hands on control of the starting values pass the argument `pars = 'values'` through whatever estimation function is being used

**FIXED** A bracketed, comma separate list specifying which parameters should be fixed at their starting values (i.e., not freely estimated). The input is of the form `(items, ..., parameterName)`. For instance, fixing the 10th and 12th to 15th item slope parameters (a1) is accomplished with `FIXED = (10, 12-15, a1)`

For more hands on control of the estimated values pass the argument `pars = 'values'` through whatever estimation function is being used

**FREE** Equivalent to the `FIXED` input, except that parameters are freely estimated instead of fixed at their starting value

**NEXPLORE** Number of exploratory factors to extract. Usually this is not required because passing a numeric value to the `model` argument in the estimation function will generate an exploratory factor analysis model, however if different start values, priors, lower and upper bounds, etc, are desired then this input can be used

## Value

Returns a model specification object to be used in `mirt`, `bfactor`, `multipleGroup`, or `mixedmirt`

## Author(s)

Phil Chalmers <rphilip.chalmers@gmail.com> and Alexander Robitzsch

## References

Chalmers, R. P. (2012). mirt: A Multidimensional Item Response Theory Package for the R Environment. *Journal of Statistical Software, 48*(6), 1-29. doi:10.18637/jss.v048.i06

## Examples

```
# interactively through the console (not run)
#model <- mirt.model()
#  F1 = 1,2,3,4-10
#  F2 = 10-20
#  (F1*F2) = 1,2,3,4-10
#  COV = F1*F2


# Or alternatively with a string input
s <- 'F1 = 1,2,3,4-10
      F2 = 10-20
      (F1*F2) = 1,2,3,4-10
      COV = F1*F2'
model <- mirt.model(s)

# strings can also be passed to the estimation functions directly,
#   which silently calls mirt.model(). E.g., using the string above:
# mod <- mirt(data, s)


# Q-matrix specification
Q <- matrix(c(1,1,1,0,0,0,0,0,0,1,1,1), ncol=2, dimnames = list(NULL, c('Factor1', 'Factor2')))
COV <- matrix(c(FALSE, TRUE, TRUE, FALSE), 2)
model <- mirt.model(Q, COV=COV)

## constrain various items slopes and all intercepts in single group model to be equal,
#   and use a log-normal prior for all the slopes
s <- 'F = 1-10
      CONSTRAIN = (1-3, 5, 6, a1), (1-10, d)
      PRIOR = (1-10, a1, lnorm, .2, .2)'
model <- mirt.model(s)


## constrain various items slopes and intercepts across groups for use in multipleGroup(),
#  and constrain first two slopes within 'group1' to be equal
s <- 'F = 1-10
      CONSTRAIN = (1-2, a1)
      CONSTRAINB = (1-3, 5, 6, a1), (1-10, d)'
model <- mirt.model(s)


## specify model using raw item names
data(data.read, package = 'sirt')
dat <- data.read
```

```
# syntax with variable names
mirtsyn2 <- "
       F1 = A1,B2,B3,C4
       F2 = A1-A4,C2,C4
       MEAN = F1
       COV = F1*F1, F1*F2
       CONSTRAIN=(A2-A4,a2),(A3,C2,d)
       PRIOR = (C3,A2-A4,a2,lnorm, .2, .2),(B3,d,norm,0,.0001)"
# create a mirt model
mirtmodel <- mirt.model(mirtsyn2, itemnames=dat)
# or equivalently:
# mirtmodel <- mirt.model(mirtsyn2, itemnames=colnames(dat))

# mod <- mirt(dat , mirtmodel)

# using sprintf() to functionally fill in information (useful for long tests
# or more complex specifications)
nitems <- 100
s <- sprintf('F = 1-%i
      CONSTRAIN = (%s, a1)
      CONSTRAINB = (%s, a1), (1-%i, d)',
      nitems, "1,2,4,50,100",
      paste0(1:45, collapse=','),
      nitems)
cat(s)
model <- mirt.model(s)
```

---

`mirtCluster` *Define a parallel cluster object to be used in internal functions*

---

## Description

This function defines a object that is placed in a relevant internal environment defined in mirt. Internal functions such as `calcLogLik`, `fscores`, etc, will utilize this object automatically to capitalize on parallel processing architecture. The object defined is a call from either `mirai::make_cluster()` (the default) or `parallel::makeCluster()`. Note that if you are defining other parallel objects (for simulation designs, for example) it is not recommended to define a mirtCluster.

## Usage

```
mirtCluster(spec, omp_threads, remove = FALSE, use_mirai = TRUE, ...)
```

## Arguments

| | |
|---|---|
| `spec` | input that is passed to `parallel::makeCluster()`. If no input is given the maximum number of available local cores minus 1 will be used. Setting this |

| | |
|---|---|
| | to NULL will skip a new definition (allows omp_threads to be used independently) |
| omp_threads | number of OpenMP threads to use for OpenMP-backed E-step and item trace computations. Not used when argument input is missing |
| remove | logical; remove previously defined mirtCluster()? |
| use_mirai | logical; use the mirai cluster definition instead of parallel? |
| ... | additional arguments to pass to parallel::makeCluster |

## Author(s)

Phil Chalmers <rphilip.chalmers@gmail.com>

## References

Chalmers, R. P. (2012). mirt: A Multidimensional Item Response Theory Package for the R Environment. *Journal of Statistical Software, 48*(6), 1-29. doi:10.18637/jss.v048.i06

## Examples

```
## Not run:
if(interactive()){
  # use all available cores
  mirtCluster()
  mirtCluster(remove = TRUE)

  # make 4 cores available for parallel computing
  mirtCluster(4)
  mirtCluster(remove = TRUE)

  # create 3 core architecture in R, and 4 thread architecture with OpenMP
  mirtCluster(spec = 3, omp_threads = 4)

  # leave previous multicore objects, but change omp_threads
  mirtCluster(spec = NULL, omp_threads = 2)
}


## End(Not run)
```

---

# MixedClass-class *Class "MixedClass"*

---

## Description

Defines the object returned from mixedmirt.

### Slots

`Call`: function call

`Data`: list of data, sometimes in different forms

`Options`: list of estimation options

`Fit`: a list of fit information

`Model`: a list of model-based information

`ParObjects`: a list of the S4 objects used during estimation

`OptimInfo`: a list of arguments from the optimization process

`Internals`: a list of internal arguments for secondary computations (inspecting this object is generally not required)

`vcov`: a matrix represented the asymptotic covariance matrix of the parameter estimates

`time`: a data.frame indicating the breakdown of computation times in seconds

### Methods

**coef** `signature(object = "MixedClass")`

**print** `signature(x = "MixedClass")`

**residuals** `signature(object = "MixedClass")`

**show** `signature(object = "MixedClass")`

**summary** `signature(object = "MixedClass")`

**logLik** `signature(object = "MixedClass")`

**anova** `signature(object = "MixedClass")`

### Author(s)

Phil Chalmers <rphilip.chalmers@gmail.com>

### References

Chalmers, R. P. (2012). mirt: A Multidimensional Item Response Theory Package for the R Environment. *Journal of Statistical Software, 48*(6), 1-29. doi:10.18637/jss.v048.i06

---

## `mixedmirt` *Mixed effects modeling for MIRT models*

---

### Description

`mixedmirt` fits MIRT models using FIML estimation to dichotomous and polytomous IRT models conditional on fixed and random effect of person and item level covariates. This can also be understood as 'explanatory IRT' if only fixed effects are modeled, or multilevel/mixed IRT if random and fixed effects are included. The method uses the MH-RM algorithm exclusively. Additionally, computation of the log-likelihood can be sped up by using parallel estimation via `mirtCluster`.

## Usage

```
mixedmirt(
  data,
  covdata = NULL,
  model = 1,
  fixed = ~1,
  random = NULL,
  itemtype = "Rasch",
  lr.fixed = ~1,
  lr.random = NULL,
  itemdesign = NULL,
  constrain = NULL,
  pars = NULL,
  return.design = FALSE,
  SE = TRUE,
  internal_constraints = TRUE,
  technical = list(SEtol = 1e-04),
  ...
)
```

## Arguments

| | |
|---|---|
| data | a matrix or data.frame that consists of numerically ordered data, organized in the form of integers, with missing data coded as NA |
| covdata | a data.frame that consists of the nrow(data) by K 'person level' fixed and random predictors. If missing data are present in this object then the observations from covdata and data will be removed row-wise via the rowSums(is.na(covdata)) > 0 |
| model | an object returned from, or a string to be passed to, mirt.model() to declare how the IRT model is to be estimated. See mirt.model and mirt for more detail |
| fixed | a right sided R formula for specifying the fixed effect (aka 'explanatory') predictors from covdata and itemdesign. To estimate the intercepts for each item the keyword items is reserved and automatically added to the itemdesign input. If any polytomous items are being model the items are argument is not valid since all intercept parameters are freely estimated and identified with the parameterizations found in mirt, and the first column in the fixed design matrix (commonly the intercept or a reference group) is omitted |
| random | a right sided formula or list of formulas containing crossed random effects of the form v1 + ... v_n \| G, where G is the grouping variable and v_n are random numeric predictors within each group. If no intercept value is specified then by default the correlations between the v's and G are estimated, but can be suppressed by including the ~ -1 + ... or 0 constant. G may contain interaction terms, such as group:items to include cross or person-level interactions effects |
| itemtype | same as itemtype in mirt, except when the fixed or random inputs are used does not support the following item types: c('PC2PL', 'PC3PL', '2PLNRM', '3PLNRM', '3PLuNRM', '4PLNRM') |

`lr.fixed` an R formula (or list of formulas) to specify regression effects in the latent variables from the variables in `covdata`. This is used to construct models such as the so-called 'latent regression model' to explain person-level ability/trait differences. If a named list of formulas is supplied (where the names correspond to the latent trait names in `model`) then specific regression effects can be estimated for each factor. Supplying a single formula will estimate the regression parameters for all latent traits by default.

`lr.random` a list of random effect terms for modeling variability in the latent trait scores, where the syntax uses the same style as in the `random` argument. Useful for building so-called 'multilevel IRT' models which are non-Rasch (multilevel Rasch models do not technically require these because they can be built using the `fixed` and `random` inputs alone)

`itemdesign` a `data.frame` object used to create a design matrix for the items, where each `nrow(itemdesign) == nitems` and the number of columns is equal to the number of fixed effect predictors (i.e., item intercepts). By default an `items` variable is reserved for modeling the item intercept parameters

`constrain` a list indicating parameter equality constrains. See `mirt` for more detail

`pars` used for parameter starting values. See `mirt` for more detail

`return.design` logical; return the design matrices before they have (potentially) been reassigned?

`SE` logical; compute the standard errors by approximating the information matrix using the MHRM algorithm? Default is TRUE

`internal_constraints` logical; use the internally defined constraints for constraining effects across persons and items? Default is TRUE. Setting this to FALSE runs the risk of underidentification

`technical` the technical list passed to the MH-RM estimation engine, with the SEtol default increased to .0001. Additionally, the argument `RANDSTART` is available to indicate at which iteration (during the burn-in stage) the additional random effect variables should begin to be approximated (i.e., elements in `lr.random` and `random`). The default for `RANDSTART` is to start at iteration 100, and when random effects are included the default number of burn-in iterations is increased from 150 to 200. See `mirt` for further details

`...` additional arguments to be passed to the MH-RM estimation engine. See `mirt` for more details and examples

## Details

For dichotomous response models, `mixedmirt` follows the general form

$$P(x = 1|\theta, \psi) = g + \frac{(u - g)}{1 + exp(-1 * [\theta a + X\beta + Z\delta])}$$

where X is a design matrix with associated $\beta$ fixed effect intercept coefficients, and Z is a design matrix with associated $\delta$ random effects for the intercepts. For simplicity and easier interpretation, the unique item intercept values typically found in $X\beta$ are extracted and reassigned within

mirt's 'intercept' parameters (e.g., 'd'). To observe how the design matrices are structured prior to reassignment and estimation pass the argument return.design = TRUE.

Polytomous IRT models follow a similar format except the item intercepts are automatically estimated internally, rendering the items argument in the fixed formula redundant and therefore must be omitted from the specification. If there are a mixture of dichotomous and polytomous items the intercepts for the dichotomous models are also estimated for consistency.

The decomposition of the $\theta$ parameters is also possible to form latent regression and multilevel IRT models by using the lr.fixed and lr.random inputs. These effects decompose $\theta$ such that

$$\theta = V\Gamma + W\zeta + \epsilon$$

where V and W are fixed and random effects design matrices for the associated coefficients.

To simulate expected a posteriori predictions for the random effect terms use the randef function.

## Value

function returns an object of class MixedClass (MixedClass-class).

## Author(s)

Phil Chalmers <rphilip.chalmers@gmail.com>

## References

Chalmers, R. P. (2012). mirt: A Multidimensional Item Response Theory Package for the R Environment. *Journal of Statistical Software, 48*(6), 1-29. doi:10.18637/jss.v048.i06

Chalmers, R. P. (2015). Extended Mixed-Effects Item Response Models with the MH-RM Algorithm. *Journal of Educational Measurement, 52*, 200-222. doi:10.1111/jedm.12072

## See Also

mirt, randef, fixef, boot.mirt

## Examples

```
# make some data
set.seed(1234)
N <- 750
a <- matrix(rlnorm(10,.3,1),10,1)
d <- matrix(rnorm(10), 10)
Theta <- matrix(sort(rnorm(N)))
pseudoIQ <- Theta * 5 + 100  + rnorm(N, 0 , 5)
pseudoIQ <- (pseudoIQ - mean(pseudoIQ))/10  #rescale variable for numerical stability
group <- factor(rep(c('G1','G2','G3'), each = N/3))
data <- simdata(a,d,N, itemtype = rep('2PL',10), Theta=Theta)
covdata <- data.frame(group, pseudoIQ)
```

```
itemstats(data)

# use parallel computing
if(interactive()) mirtCluster()

# specify IRT model
model <- 'Theta = 1-10'

# model with no person predictors
mod0 <- mirt(data, model, itemtype = 'Rasch')

# group as a fixed effect predictor (aka, uniform dif)
mod1 <- mixedmirt(data, covdata, model, fixed = ~ 0 + group + items)
anova(mod0, mod1)
summary(mod1)
coef(mod1)

## Not run:
# same model as above in lme4
wide <- data.frame(id=1:nrow(data),data,covdata)
long <- reshape2::melt(wide, id.vars = c('id', 'group', 'pseudoIQ'))
library(lme4)
lmod0 <- glmer(value ~ 0 + variable + (1|id), long, family = binomial)
lmod1 <- glmer(value ~ 0 + group + variable + (1|id), long, family = binomial)
anova(lmod0, lmod1)

## End(Not run)

# model using 2PL items instead of Rasch
mod1b <- mixedmirt(data, covdata, model, fixed = ~ 0 + group + items, itemtype = '2PL')
anova(mod1, mod1b) #better with 2PL models using all criteria (as expected, given simdata pars)

# continuous predictor with group
mod2 <- mixedmirt(data, covdata, model, fixed = ~ 0 + group + items + pseudoIQ)
summary(mod2)
anova(mod1b, mod2)

# view fixed design matrix with and without unique item level intercepts
withint <- mixedmirt(data, covdata, model, fixed = ~ 0 + items + group, return.design = TRUE)
withoutint <- mixedmirt(data, covdata, model, fixed = ~ 0 + group, return.design = TRUE)

# notice that in result above, the intercepts 'items1 to items 10' were reassigned to 'd'
head(withint$X)
tail(withint$X)
head(withoutint$X) # no intercepts design here to be reassigned into item intercepts
tail(withoutint$X)

####################################################
### random effects
# make the number of groups much larger
covdata$group <- factor(rep(paste0('G',1:50), each = N/50))

# random groups
```

```
rmod1 <- mixedmirt(data, covdata, 1, fixed = ~ 0 + items, random = ~ 1|group)
summary(rmod1)
coef(rmod1)

# random groups and random items
rmod2 <- mixedmirt(data, covdata, 1, random = list(~ 1|group, ~ 1|items))
summary(rmod2)
eff <- randef(rmod2) #estimate random effects

# random slopes with fixed intercepts (suppressed correlation)
rmod3 <- mixedmirt(data, covdata, 1, fixed = ~ 0 + items, random = ~ -1 + pseudoIQ|group)
summary(rmod3)
eff <- randef(rmod3)
str(eff)

###################################################
## LLTM, and 2PL version of LLTM
data(SAT12)
data <- key2binary(SAT12,
               key = c(1,4,5,2,3,1,2,1,3,1,2,4,2,1,5,3,4,4,1,4,3,3,4,1,3,5,1,3,1,5,4,5))
model <- 'Theta = 1-32'

# for unconditional intercept comparison
mod <- mirt(data, model, itemtype='Rasch')
coef(mod, simplify=TRUE)

# Suppose that the first 16 items were suspected to be easier than the last 16 items,
#   and we wish to test this item structure hypothesis (more intercept designs are possible
#   by including more columns).
itemdesign <- data.frame(itemorder = factor(c(rep('easier', 16), rep('harder', 16))))
rownames(itemdesign) <- colnames(data)
itemdesign

# notice that the 'fixed = ~ ... + items' argument is omitted
LLTM <- mixedmirt(data, model = model, fixed = ~ 0 + itemorder, itemdesign = itemdesign,
   SE = TRUE) # SE argument ensures that the information matrix is computed accurately
summary(LLTM)
coef(LLTM)
wald(LLTM)
L <- matrix(c(-1, 1, 0), 1)
wald(LLTM, L) #first half different from second

# compare to items with estimated slopes (2PL)
twoPL <- mixedmirt(data, model = model, fixed = ~ 0 + itemorder, itemtype = '2PL',
                   itemdesign = itemdesign)
# twoPL not mixing too well (AR should be between .2 and .5), decrease MHcand
twoPL <- mixedmirt(data, model = model, fixed = ~ 0 + itemorder, itemtype = '2PL',
                  itemdesign = itemdesign, technical = list(MHcand = 0.8))
anova(twoPL, LLTM) #much better fit
summary(twoPL)
coef(twoPL)

wald(twoPL)
```

```
L <- matrix(0, 1, 34)
L[1, 1] <- 1
L[1, 2] <- -1
wald(twoPL, L) # n.s., which is the correct conclusion. Rasch approach gave wrong inference

## LLTM with item error term
LLTMwithError <- mixedmirt(data, model = model, fixed = ~ 0 + itemorder, random = ~ 1|items,
    itemdesign = itemdesign)
summary(LLTMwithError)
# large item level variance after itemorder is regressed; not a great predictor of item difficulty
coef(LLTMwithError)

#################################################
### Polytomous example

# make an arbitrary group difference
covdat <- data.frame(group = rep(c('m', 'f'), nrow(Science)/2))

# partial credit model
mod <- mixedmirt(Science, covdat, model=1, fixed = ~ 0 + group)
coef(mod)

# gpcm to estimate slopes
mod2 <- mixedmirt(Science, covdat, model=1, fixed = ~ 0 + group,
                  itemtype = 'gpcm')
summary(mod2)
anova(mod, mod2)

# graded model
mod3 <- mixedmirt(Science, covdat, model=1, fixed = ~ 0 + group,
                  itemtype = 'graded')
coef(mod3)


#################################################
# latent regression with Rasch and 2PL models

set.seed(1)
n <- 300
a <- matrix(1, 10)
d <- matrix(rnorm(10))
Theta <- matrix(c(rnorm(n, 0), rnorm(n, 1), rnorm(n, 2)))
covdata <- data.frame(group=rep(c('g1','g2','g3'), each=n))
dat <- simdata(a, d, N=n*3, Theta=Theta, itemtype = '2PL')
itemstats(dat)

# had we known the latent abilities, we could have computed the regression coefs
summary(lm(Theta ~ covdata$group))

# but all we have is observed test data. Latent regression helps to recover these coefs
# Rasch model approach (and mirt equivalent)
rmod0 <- mirt(dat, 1, 'Rasch') # unconditional
```

```
# these two models are equivalent
rmod1a <- mirt(dat, 1, 'Rasch', covdata = covdata, formula = ~ group)
rmod1b <- mixedmirt(dat, covdata, 1, fixed = ~ 0 + items + group)
anova(rmod0, rmod1b)
coef(rmod1a, simplify=TRUE)
summary(rmod1b)

# 2PL, requires different input to allow Theta variance to remain fixed
mod0 <- mirt(dat, 1) # unconditional
mod1a <- mirt(dat, 1, covdata = covdata, formula = ~ group, itemtype = '2PL')
mod1b <- mixedmirt(dat, covdata, 1, fixed = ~ 0 + items, lr.fixed = ~group, itemtype = '2PL')
anova(mod0, mod1b)
coef(mod1a)$lr.betas
summary(mod1b)

# specifying specific regression effects is accomplished by passing a list of formula
model <- 'F1 = 1-5
          F2 = 6-10'
covdata$contvar <- rnorm(nrow(covdata))
mod2 <- mirt(dat, model, itemtype = 'Rasch', covdata=covdata,
        formula = list(F1 = ~ group + contvar, F2 = ~ group))
coef(mod2)[11:12]
mod2b <- mixedmirt(dat, covdata, model, fixed = ~ 0 + items,
        lr.fixed = list(F1 = ~ group + contvar, F2 = ~ group))
summary(mod2b)

####################################################
## Simulated Multilevel Rasch Model

set.seed(1)
N <- 2000
a <- matrix(rep(1,10),10,1)
d <- matrix(rnorm(10))
cluster = 100
random_intercept = rnorm(cluster,0,1)
Theta = numeric()
for (i in 1:cluster)
    Theta <- c(Theta, rnorm(N/cluster,0,1) + random_intercept[i])

group = factor(rep(paste0('G',1:cluster), each = N/cluster))
covdata <- data.frame(group)
dat <- simdata(a,d,N, itemtype = rep('2PL',10), Theta=matrix(Theta))
itemstats(dat)

# null model
mod1 <- mixedmirt(dat, covdata, 1, fixed = ~ 0 + items, random = ~ 1|group)
summary(mod1)

# include level 2 predictor for 'group' variance
covdata$group_pred <- rep(random_intercept, each = N/cluster)
mod2 <- mixedmirt(dat, covdata, 1, fixed = ~ 0 + items + group_pred, random = ~ 1|group)

# including group means predicts nearly all variability in 'group'
```

```
summary(mod2)
anova(mod1, mod2)

# can also be fit for Rasch/non-Rasch models with the lr.random input
mod1b <- mixedmirt(dat, covdata, 1, fixed = ~ 0 + items, lr.random = ~ 1|group)
summary(mod1b)

mod2b <- mixedmirt(dat, covdata, 1, fixed = ~ 0 + items + group_pred, lr.random = ~ 1|group)
summary(mod2b)
anova(mod1b, mod2b)

mod3 <- mixedmirt(dat, covdata, 1, fixed = ~ 0 + items, lr.random = ~ 1|group, itemtype = '2PL')
summary(mod3)
anova(mod1b, mod3)

head(cbind(randef(mod3)$group, random_intercept))
```

---

## MixtureClass-class *Class "MixtureClass"*

---

### Description

Defines the object returned from `multipleGroup` when estimated with mixture distributions.

### Slots

`Call`: function call

`Data`: list of data, sometimes in different forms

`Options`: list of estimation options

`Fit`: a list of fit information

`Model`: a list of model-based information

`ParObjects`: a list of the S4 objects used during estimation

`OptimInfo`: a list of arguments from the optimization process

`Internals`: a list of internal arguments for secondary computations (inspecting this object is generally not required)

`vcov`: a matrix represented the asymptotic covariance matrix of the parameter estimates

`time`: a data.frame indicating the breakdown of computation times in seconds

### Methods

**coef** `signature(object = "MixtureClass")`

**print** `signature(x = "MixtureClass")`

**show** `signature(object = "MixtureClass")`

**anova** `signature(object = "MixtureClass")`

### Author(s)

Phil Chalmers <rphilip.chalmers@gmail.com>

### References

Chalmers, R. P. (2012). mirt: A Multidimensional Item Response Theory Package for the R Environment. *Journal of Statistical Software, 48*(6), 1-29. doi:10.18637/jss.v048.i06

---

## mod2values *Convert an estimated mirt model to a data.frame*

---

### Description

Given an estimated model from any of mirt's model fitting functions this function will convert the model parameters into the design data frame of starting values and other parameter characteristics (similar to using the `pars = 'values'` for obtaining starting values).

### Usage

```
mod2values(x)
```

### Arguments

| | |
|---|---|
| x | an estimated model x from the mirt package |

### Author(s)

Phil Chalmers <rphilip.chalmers@gmail.com>

### References

Chalmers, R. P. (2012). mirt: A Multidimensional Item Response Theory Package for the R Environment. *Journal of Statistical Software, 48*(6), 1-29. doi:10.18637/jss.v048.i06

### See Also

extract.mirt

### Examples

```
dat <- expand.table(LSAT7)
mod <- mirt(dat, "F=1-5
                  CONSTRAIN=(1-5, a1)")
values <- mod2values(mod)
values

# use the converted values as starting values in a new model, and reduce TOL
```

```
mod2 <- mirt(dat, 1, pars = values, TOL=1e-5)
coef(mod2, simplify=TRUE)

# use parameters on different dataset
mod3 <- mirt(expand.table(LSAT6), pars=values)
coef(mod3, simplify=TRUE)

# supports differing itemtypes on second model
sv <- mirt(Science, itemtype=c('graded', rep('gpcm', 3)), pars='values')
mod3 <- mirt(Science, pars = sv)  # itemtype omitted
coef(mod3, simplify=TRUE)$items
extract.mirt(mod3, 'itemtype')
```

---

`multipleGroup` *Multiple Group Estimation*

---

## Description

`multipleGroup` performs a full-information maximum-likelihood multiple group analysis for any combination of dichotomous and polytomous data under the item response theory paradigm using either Cai's (2010) Metropolis-Hastings Robbins-Monro (MHRM) algorithm or with an EM algorithm approach. This function may be used for detecting differential item functioning (DIF), thought the `DIF` function may provide a more convenient approach. If the grouping variable is not specified then the `dentype` input can be modified to fit mixture models to estimate any latent group components.

## Usage

```
multipleGroup(
  data,
  model = 1,
  group,
  itemtype = NULL,
  invariance = "",
  method = "EM",
  dentype = "Gaussian",
  itemdesign = NULL,
  item.formula = NULL,
  nruns = 1,
  return_max = TRUE,
  GenRandomPars = FALSE,
  verbose = interactive(),
  ...
)
```

## Arguments

| | |
|---|---|
| data | a matrix or data.frame that consists of numerically ordered data, organized in the form of integers, with missing data coded as NA |
| model | string to be passed to, or a model object returned from, mirt.model declaring how the global model is to be estimated (useful to apply constraints here) |
| group | a character or factor vector indicating group membership. If a character vector is supplied this will be automatically transformed into a factor variable. As well, the first level of the (factorized) grouping variable will be treated as the "reference" group |
| itemtype | can be same type of input as is documented in mirt, however may also be a ngroups by nitems matrix specifying the type of IRT models for each group, respectively. Rows of this input correspond to the levels of the group input. For mixture models the rows correspond to the respective mixture grouping variables to be constructed, and the IRT models should be within these mixtures |
| invariance | a character vector containing the following possible options:<br>'free_mean' **or** 'free_means' freely estimate all latent means in all focal groups (reference group constrained to a vector of 0's). To estimate all means use 'free_all_means'<br>'free_var' **or** 'free_vars' freely estimate all latent variances in focal groups (reference group variances all constrained to 1). To estimate all variances use either To estimate all variances use 'free_all_vars'<br>'slopes' to constrain all the slopes to be equal across all groups<br>'intercepts' to constrain all the intercepts to be equal across all groups, note for nominal models this also includes the category specific slope parameters<br>Additionally, specifying specific item name bundles (from colnames(data)) will constrain all freely estimated parameters in each item to be equal across groups. This is useful for selecting 'anchor' items for vertical and horizontal scaling, and for detecting differential item functioning (DIF) across groups |
| method | a character object that is either 'EM', 'QMCEM', or 'MHRM' (default is 'EM'). See mirt for details |
| dentype | type of density form to use for the latent trait parameters. Current options include all of the methods described in mirt, as well as<br>• 'mixture-#' estimates mixtures of Gaussian distributions, where the # placeholder represents the number of potential grouping variables (e.g., 'mixture-3' will estimate 3 underlying classes). Each class is assigned the group name MIXTURE_#, where # is the class number.<br>Note that internally the mixture coefficients are stored as log values where the first mixture group coefficient is fixed at 0. Additionally, it is recommended to use the nruns argument as mixture IRT models are known to contain local maximums |
| itemdesign | see mirt for details |
| item.formula | see mirt for details |

| | |
|---|---|
| nruns | a numeric value indicating how many times the model should be fit to the data when using random starting values, which is particularly useful when evaluating mixture IRT Models. If greater than 1, GenRandomPars is set to TRUE by default. Using this returns a list of fitted model objects, where the model with the highest log-likelihood should generally be selected as the model best associated with the MLE (this is done automatically if return_max = TRUE). Note that if a mirtCluster was defined earlier then the runs will be run in parallel |
| return_max | logical; when nruns > 1, return the model that has the most optimal maximum likelihood criteria? If FALSE, returns a list of all the estimated objects |
| GenRandomPars | see mirt for details |
| verbose | see mirt for details |
| ... | additional arguments to be passed to the estimation engine. See mirt for details and examples |

## Details

By default the estimation in multipleGroup assumes that the models are maximally independent, and therefore could initially be performed by sub-setting the data and running identical models with mirt and aggregating the results (e.g., log-likelihood). However, constrains may be automatically imposed across groups by invoking various invariance keywords. Users may also supply a list of parameter equality constraints to by constrain argument, of define equality constraints using the mirt.model syntax (recommended).

## Value

function returns an object of class MultipleGroupClass (MultipleGroupClass-class).

## Author(s)

Phil Chalmers <rphilip.chalmers@gmail.com>

## References

Chalmers, R. P. (2012). mirt: A Multidimensional Item Response Theory Package for the R Environment. *Journal of Statistical Software, 48*(6), 1-29. doi:10.18637/jss.v048.i06

Magnus, B. E. and Garnier-Villarreal (2022). A multidimensional zero-inflated graded response model for ordinal symptom data. *Psychological Methods*, 27, 261-279.

Wall, M., M., Park, J., Y., and Moustaki I. (2015). IRT modeling in the presence of zero-inflation with application to psychiatric disorder severity. *Applied Psychological Measurement* 39: 583-597.

## See Also

mirt, DIF, extract.group, DRF

## Examples

```
# single factor
set.seed(12345)
a <- matrix(abs(rnorm(15,1,.3)), ncol=1)
d <- matrix(rnorm(15,0,.7),ncol=1)
itemtype <- rep('2PL', nrow(a))
N <- 1000
dataset1 <- simdata(a, d, N, itemtype)
dataset2 <- simdata(a, d, N, itemtype, mu = .1, sigma = matrix(1.5))
dat <- rbind(dataset1, dataset2)
group <- c(rep('D1', N), rep('D2', N))

# marginal information
itemstats(dat)

# conditional information
itemstats(dat, group=group)

mod_configural <- multipleGroup(dat, 1, group = group) #completely separate analyses
# limited information fit statistics
M2(mod_configural)

mod_metric <- multipleGroup(dat, 1, group = group,
                            invariance=c('slopes', 'free_var')) #equal slopes
# equal intercepts, free variance and means
mod_scalar2 <- multipleGroup(dat, 1, group = group,
                           invariance=c('slopes', 'intercepts', 'free_var','free_means'))
mod_scalar1 <- multipleGroup(dat, 1, group = group,  #fixed means
                             invariance=c('slopes', 'intercepts', 'free_var'))
mod_fullconstrain <- multipleGroup(dat, 1, group = group,
                             invariance=c('slopes', 'intercepts'))
extract.mirt(mod_fullconstrain, 'time') #time of estimation components

# optionally use Newton-Raphson for (generally) faster convergence in the
#  M-step's, though occasionally less stable
mod_fullconstrain <- multipleGroup(dat, 1, group = group, optimizer = 'NR',
                             invariance=c('slopes', 'intercepts'))
extract.mirt(mod_fullconstrain, 'time') #time of estimation components

summary(mod_scalar2)
coef(mod_scalar2, simplify=TRUE)
residuals(mod_scalar2)
plot(mod_configural)
plot(mod_configural, type = 'info')
plot(mod_configural, type = 'trace')
plot(mod_configural, type = 'trace', which.items = 1:4)
itemplot(mod_configural, 2)
itemplot(mod_configural, 2, type = 'RE')

anova(mod_metric, mod_configural) #equal slopes
anova(mod_scalar2, mod_metric) #equal intercepts, free variance and mean
```

```
anova(mod_scalar1, mod_scalar2) #fix mean
anova(mod_fullconstrain, mod_scalar1) #fix variance

# compared all at once (in order of most constrained to least)
anova(mod_fullconstrain, mod_scalar2, mod_configural)


# test whether first 6 slopes should be equal across groups
values <- multipleGroup(dat, 1, group = group, pars = 'values')
values
constrain <- list(c(1, 63), c(5,67), c(9,71), c(13,75), c(17,79), c(21,83))
equalslopes <- multipleGroup(dat, 1, group = group, constrain = constrain)
anova(equalslopes, mod_configural)

# same as above, but using mirt.model syntax
newmodel <- '
    F = 1-15
    CONSTRAINB = (1-6, a1)'
equalslopes <- multipleGroup(dat, newmodel, group = group)
coef(equalslopes, simplify=TRUE)

############
# vertical scaling (i.e., equating when groups answer items others do not)
dat2 <- dat
dat2[group == 'D1', 1:2] <- dat2[group != 'D1', 14:15] <- NA
head(dat2)
tail(dat2)

# items with missing responses need to be constrained across groups for identification
nms <- colnames(dat2)
mod <- multipleGroup(dat2, 1, group, invariance = nms[c(1:2, 14:15)])

# this will throw an error without proper constraints (SEs cannot be computed either)
# mod <- multipleGroup(dat2, 1, group)

# model still does not have anchors, therefore need to add a few (here use items 3-5)
mod_anchor <- multipleGroup(dat2, 1, group,
                            invariance = c(nms[c(1:5, 14:15)], 'free_means', 'free_var'))
coef(mod_anchor, simplify=TRUE)

# check if identified by computing information matrix
mod_anchor <- multipleGroup(dat2, 1, group, pars = mod2values(mod_anchor), TOL=NaN, SE=TRUE,
                            invariance = c(nms[c(1:5, 14:15)], 'free_means', 'free_var'))
mod_anchor
coef(mod_anchor)
coef(mod_anchor, printSE=TRUE)


#############
# DIF test for each item (using all other items as anchors)
itemnames <- colnames(dat)
refmodel <- multipleGroup(dat, 1, group = group, SE=TRUE,
                          invariance=c('free_means', 'free_var', itemnames))
```

```
# loop over items (in practice, run in parallel to increase speed). May be better to use ?DIF
estmodels <- vector('list', ncol(dat))
for(i in 1:ncol(dat))
    estmodels[[i]] <- multipleGroup(dat, 1, group = group, verbose = FALSE,
                            invariance=c('free_means', 'free_var', itemnames[-i]))
anova(refmodel, estmodels[[1]])
(anovas <- lapply(estmodels, function(x, refmodel) anova(refmodel, x),
   refmodel=refmodel))

# family-wise error control
p <- do.call(rbind, lapply(anovas, function(x) x[2, 'p']))
p.adjust(p, method = 'BH')

# same as above, except only test if slopes vary (1 df)
# constrain all intercepts
estmodels <- vector('list', ncol(dat))
for(i in 1:ncol(dat))
    estmodels[[i]] <- multipleGroup(dat, 1, group = group, verbose = FALSE,
                            invariance=c('free_means', 'free_var', 'intercepts',
                            itemnames[-i]))

(anovas <- lapply(estmodels, function(x, refmodel) anova(refmodel, x),
   refmodel=refmodel))

# quickly test with Wald test using DIF()
mod_configural2 <- multipleGroup(dat, 1, group = group, SE=TRUE)
DIF(mod_configural2, which.par = c('a1', 'd'), Wald=TRUE, p.adjust = 'fdr')



##############
# Three group model where the latent variable parameters are constrained to
# be equal in the focal groups

set.seed(12345)
a <- matrix(abs(rnorm(15,1,.3)), ncol=1)
d <- matrix(rnorm(15,0,.7),ncol=1)
itemtype <- rep('2PL', nrow(a))
N <- 1000
dataset1 <- simdata(a, d, N, itemtype)
dataset2 <- simdata(a, d, N, itemtype, mu = .1, sigma = matrix(1.5))
dataset3 <- simdata(a, d, N, itemtype, mu = .1, sigma = matrix(1.5))
dat <- rbind(dataset1, dataset2, dataset3)
group <- rep(c('D1', 'D2', 'D3'), each=N)

# marginal information
itemstats(dat)

# conditional information
itemstats(dat, group=group)

model <- 'F1 = 1-15
```

```
          FREE[D2, D3] = (GROUP, MEAN_1), (GROUP, COV_11)
          CONSTRAINB[D2,D3] = (GROUP, MEAN_1), (GROUP, COV_11)'

mod <- multipleGroup(dat, model, group = group, invariance = colnames(dat))
coef(mod, simplify=TRUE)

#############
# Testing main effects in multiple independent grouping variables
set.seed(1234)
a <- matrix(abs(rnorm(15,1,.3)), ncol=1)
d <- matrix(rnorm(15,0,.7),ncol=1)
itemtype <- rep('2PL', nrow(a))
N <- 500

# generated data have interaction effect for latent means, as well as a
# main effect across D but no main effect across G
d11 <- simdata(a, d, N, itemtype, mu = 0)
d12 <- simdata(a, d, N, itemtype, mu = 0)
d13 <- simdata(a, d, N, itemtype, mu = 0)
d21 <- simdata(a, d, N, itemtype, mu = 1/2)
d22 <- simdata(a, d, N, itemtype, mu = 1/2)
d23 <- simdata(a, d, N, itemtype, mu = -1)
dat <- do.call(rbind, list(d11, d12, d13, d21, d22, d23))
group <- rep(c('G1.D1', 'G1.D2', 'G1.D3', 'G2.D1', 'G2.D2', 'G2.D3'), each=N)
table(group)

## Not run:
# in practice, group would be organized in a data.frame as follows
df <- data.frame(group)
dfw <- tidyr::separate_wider_delim(df, group, delim='.', names = c('G', 'D'))
head(dfw)

# for use with multipleGroup() combine into a single long group
group <- with(dfw, factor(G):factor(D))

# conditional information
itemstats(dat, group=group)

mod <- multipleGroup(dat, group = group, SE=TRUE,
                     invariance = c(colnames(dat), 'free_mean', 'free_var'))
coef(mod, simplify=TRUE)
sapply(coef(mod, simplify=TRUE), \(x) unname(x$means)) # mean estimates
wald(mod) # view parameter names for later testing

# test for main effect over G group (manually compute marginal mean)
wald(mod, "0 + MEAN_1.123 + MEAN_1.185 = MEAN_1.247 + MEAN_1.309 + MEAN_1.371")

# test for main effect over D group  (manually compute marginal means)
wald(mod, c("0 + MEAN_1.247 = MEAN_1.123 + MEAN_1.309",
            "0 + MEAN_1.247 = MEAN_1.185 + MEAN_1.371"))

# post-hoc tests (better practice would include p.adjust() )
wald(mod, "0 + MEAN_1.247 = MEAN_1.123 + MEAN_1.309") # D1 vs D2
```

```
wald(mod, "0 + MEAN_1.247 = MEAN_1.185 + MEAN_1.371") # D1 vs D3
wald(mod, "MEAN_1.123 + MEAN_1.309 = MEAN_1.185 + MEAN_1.371") # D2 vs D3

## End(Not run)

#############
# multiple factors

a <- matrix(c(abs(rnorm(5,1,.3)), rep(0,15),abs(rnorm(5,1,.3)),
     rep(0,15),abs(rnorm(5,1,.3))), 15, 3)
d <- matrix(rnorm(15,0,.7),ncol=1)
mu <- c(-.4, -.7, .1)
sigma <- matrix(c(1.21,.297,1.232,.297,.81,.252,1.232,.252,1.96),3,3)
itemtype <- rep('2PL', nrow(a))
N <- 1000
dataset1 <- simdata(a, d, N, itemtype)
dataset2 <- simdata(a, d, N, itemtype, mu = mu, sigma = sigma)
dat <- rbind(dataset1, dataset2)
group <- c(rep('D1', N), rep('D2', N))

# group models
model <- '
   F1 = 1-5
   F2 = 6-10
   F3 = 11-15'

# define mirt cluster to use parallel architecture
if(interactive()) mirtCluster()

# EM approach (not as accurate with 3 factors, but generally good for quick model comparisons)
mod_configural <- multipleGroup(dat, model, group = group) #completely separate analyses
mod_metric <- multipleGroup(dat, model, group = group,
                            invariance=c('slopes', 'free_var')) #equal slopes
mod_fullconstrain <- multipleGroup(dat, model, group = group, #equal means, slopes, intercepts
                             invariance=c('slopes', 'intercepts'))

anova(mod_metric, mod_configural)
anova(mod_fullconstrain, mod_metric)

# same as above, but with MHRM (generally  more accurate with 3+ factors, but slower)
mod_configural <- multipleGroup(dat, model, group = group, method = 'MHRM')
mod_metric <- multipleGroup(dat, model, group = group,
                            invariance=c('slopes', 'free_var'), method = 'MHRM')
mod_fullconstrain <- multipleGroup(dat, model, group = group, method = 'MHRM',
                             invariance=c('slopes', 'intercepts'))

anova(mod_metric, mod_configural)
anova(mod_fullconstrain, mod_metric)

############
# polytomous item example
set.seed(12345)
a <- matrix(abs(rnorm(15,1,.3)), ncol=1)
```

```
d <- matrix(rnorm(15,0,.7),ncol=1)
d <- cbind(d, d-1, d-2)
itemtype <- rep('graded', nrow(a))
N <- 1000
dataset1 <- simdata(a, d, N, itemtype)
dataset2 <- simdata(a, d, N, itemtype, mu = .1, sigma = matrix(1.5))
dat <- rbind(dataset1, dataset2)
group <- c(rep('D1', N), rep('D2', N))
model <- 'F1 = 1-15'

mod_configural <- multipleGroup(dat, model, group = group)
plot(mod_configural)
plot(mod_configural, type = 'SE')
plot(mod_configural, type = 'gen.difficulty')
itemplot(mod_configural, 1)
itemplot(mod_configural, 1, type = 'info')
plot(mod_configural, type = 'trace') # messy, score function typically better
plot(mod_configural, type = 'itemscore')

fs <- fscores(mod_configural, full.scores = FALSE)
head(fs[["D1"]])
fscores(mod_configural, method = 'EAPsum', full.scores = FALSE)

# constrain slopes within each group to be equal (but not across groups)
model2 <- 'F1 = 1-15
           CONSTRAIN = (1-15, a1)'
mod_configural2 <- multipleGroup(dat, model2, group = group)
plot(mod_configural2, type = 'SE')
plot(mod_configural2, type = 'RE')
itemplot(mod_configural2, 10)

#############
## empirical histogram example (normal and bimodal groups)
set.seed(1234)
a <- matrix(rlnorm(50, .2, .2))
d <- matrix(rnorm(50))
ThetaNormal <- matrix(rnorm(2000))
ThetaBimodal <- scale(matrix(c(rnorm(1000, -2), rnorm(1000,2)))) #bimodal
Theta <- rbind(ThetaNormal, ThetaBimodal)
dat <- simdata(a, d, 4000, itemtype = '2PL', Theta=Theta)
group <- rep(c('G1', 'G2'), each=2000)

EH <- multipleGroup(dat, 1, group=group, dentype="empiricalhist", invariance = colnames(dat))
coef(EH, simplify=TRUE)
plot(EH, type = 'empiricalhist', npts = 60)

# DIF test for item 1
EH1 <- multipleGroup(dat, 1, group=group, dentype="empiricalhist", invariance = colnames(dat)[-1])
anova(EH, EH1)

#--------------------------------
# Mixture model (no prior group variable specified)
```

```
set.seed(12345)
nitems <- 20
a1 <- matrix(.75, ncol=1, nrow=nitems)
a2 <- matrix(1.25, ncol=1, nrow=nitems)
d1 <- matrix(rnorm(nitems,0,1),ncol=1)
d2 <- matrix(rnorm(nitems,0,1),ncol=1)
itemtype <- rep('2PL', nrow(a1))
N1 <- 500
N2 <- N1*2 # second class twice as large

dataset1 <- simdata(a1, d1, N1, itemtype)
dataset2 <- simdata(a2, d2, N2, itemtype)
dat <- rbind(dataset1, dataset2)
# group <- c(rep('D1', N1), rep('D2', N2))

# Mixture Rasch model (Rost, 1990)
models <- 'F1 = 1-20
           CONSTRAIN = (1-20, a1)'
mod_mix <- multipleGroup(dat, models, dentype = 'mixture-2', GenRandomPars = TRUE)
coef(mod_mix, simplify=TRUE)
summary(mod_mix)
plot(mod_mix)
plot(mod_mix, type = 'trace')
plot(mod_mix, type = 'gen.difficulty')
itemplot(mod_mix, 1, type = 'info')

head(fscores(mod_mix)) # theta estimates
head(fscores(mod_mix, method = 'classify')) # classification probability
itemfit(mod_mix)

# Above works fine, but its generally a good idea to evaluate models
# with multiple random starting values in case local maximums are an issue.
# To do this use the argument "nruns", which returns the best model
if(interactive()) mirtCluster()
mod_mix <- multipleGroup(dat, models, dentype = 'mixture-2', nruns=5)
mod_mix

# For obtaining isolated estimates within each mixture, use extract.group()
#   to construct single-group extractions of the mixtures
mix1 <- extract.group(mod_mix, group = "MIXTURE_1")
mix2 <- extract.group(mod_mix, group = "MIXTURE_2")

# EAP estimates per mixture group, ignoring the original mixture structure.
#   Used to demonstrate the behaviour of how the individuals would have
#   been scored if they (deterministically) belonged to one class
data.frame(EAP_mix1=unname(fscores(mix1)),
           EAP_mix2=unname(fscores(mix2)),
           EAP=unname(fscores(mod_mix))) |> head()

############
# Mixture 2PL model
mod_mix2 <- multipleGroup(dat, 1, dentype = 'mixture-2', nruns=5)
anova(mod_mix, mod_mix2)
```

```
coef(mod_mix2, simplify=TRUE)
itemfit(mod_mix2)

# Compare to single group
mod <- mirt(dat)
anova(mod, mod_mix2)

########################################
# Zero-inflated 2PL IRT model (Wall, Park, and Moustaki, 2015)

n <- 1000
nitems <- 20

a <- rep(2, nitems)
d <- rep(c(-2,-1,0,1,2), each=nitems/5)
zi_p <- 0.2 # Proportion of people in zero class

theta <- rnorm(n, 0, 1)
zeros <- matrix(0, n*zi_p, nitems)
nonzeros <- simdata(a, d, n*(1-zi_p), itemtype = '2PL',
                    Theta = as.matrix(theta[1:(n*(1-zi_p))]))
data <- rbind(nonzeros, zeros)

# define class with extreme theta but fixed item parameters
zi2PL <- "F = 1-20
          START [MIXTURE_1] = (GROUP, MEAN_1, -100), (GROUP, COV_11, .00001),
                              (1-20, a1, 1.0), (1-20, d, 0)
          FIXED [MIXTURE_1] = (GROUP, MEAN_1), (GROUP, COV_11),
                              (1-20, a1), (1-20, d)"

# define custom Theta integration grid that contains extreme theta + normal grid
technical <- list(customTheta = matrix(c(-100, seq(-6,6,length.out=61))))

# fit ZIM-IRT
zi2PL.fit <- multipleGroup(data, zi2PL, dentype = 'mixture-2', technical=technical)
coef(zi2PL.fit, simplify=TRUE)

# classification estimates
pi_hat <- fscores(zi2PL.fit, method = 'classify')
head(pi_hat)
tail(pi_hat)

# EAP estimates (not useful for zip class)
fs <- fscores(zi2PL.fit)
head(fs)
tail(fs)

########################################
# Zero-inflated graded response model (Magnus and Garnier-Villarreal, 2022)

n <- 1000
nitems <- 20
```

```
a <- matrix(rlnorm(20,.2,.3))

# for the graded model, ensure that there is enough space between the intercepts,
# otherwise closer categories will not be selected often (minimum distance of 0.3 here)
diffs <- t(apply(matrix(runif(20*4, .3, 1), 20), 1, cumsum))
diffs <- -(diffs - rowMeans(diffs))
d <- diffs + rnorm(20)

zi_p <- 0.2 # Proportion of people in zero/lowest category class

theta <- rnorm(n, 0, 1)
zeros <- matrix(0, n*zi_p, nitems)
nonzeros <- simdata(a, d, n*(1-zi_p), itemtype = 'graded',
                    Theta = as.matrix(theta[1:(n*(1-zi_p))]))
data <- rbind(nonzeros, zeros)

# intercepts will be labelled as d1 through d4
apply(data, 2, table)

# ignoring zero inflation (bad idea)
modGRM <- mirt(data)
coef(modGRM, simplify=TRUE)

# Define class with extreme theta but fixed item parameters
#   For GRM in zero-inflated class the intercept values are arbitrary
#   as the model forces the responses all into the first category (hence,
#   spacing arbitrarily set to 1)
ziGRM <- "F = 1-20
          START [MIXTURE_1] = (GROUP, MEAN_1, -100), (GROUP, COV_11, .00001),
                              (1-20, a1, 1.0),
                             (1-20, d1, 2), (1-20, d2, 1), (1-20, d3, 0), (1-20, d4, -1)
          FIXED [MIXTURE_1] = (GROUP, MEAN_1), (GROUP, COV_11),
                              (1-20, a1),
                              (1-20, d1), (1-20, d2), (1-20, d3), (1-20, d4)"

# define custom Theta integration grid that contains extreme theta + normal grid
technical <- list(customTheta = matrix(c(-100, seq(-6,6,length.out=61))))

# fit zero-inflated GRM
ziGRM.fit <- multipleGroup(data, ziGRM, dentype = 'mixture-2', technical=technical)
coef(ziGRM.fit, simplify=TRUE)

# classification estimates
pi_hat <- fscores(ziGRM.fit, method = 'classify')
head(pi_hat)
tail(pi_hat)

# EAP estimates (not useful for zip class)
fs <- fscores(ziGRM.fit)
head(fs)
tail(fs)
```

MultipleGroupClass-class
*Class "MultipleGroupClass"*

## Description

Defines the object returned from `multipleGroup`.

## Slots

`Call`: function call

`Data`: list of data, sometimes in different forms

`Options`: list of estimation options

`Fit`: a list of fit information

`Model`: a list of model-based information

`ParObjects`: a list of the S4 objects used during estimation

`OptimInfo`: a list of arguments from the optimization process

`Internals`: a list of internal arguments for secondary computations (inspecting this object is generally not required)

`vcov`: a matrix represented the asymptotic covariance matrix of the parameter estimates

`time`: a data.frame indicating the breakdown of computation times in seconds

## Methods

**coef** `signature(object = "MultipleGroupClass")`

**print** `signature(x = "MultipleGroupClass")`

**show** `signature(object = "MultipleGroupClass")`

**anova** `signature(object = "MultipleGroupClass")`

## Author(s)

Phil Chalmers <rphilip.chalmers@gmail.com>

## References

Chalmers, R. P. (2012). mirt: A Multidimensional Item Response Theory Package for the R Environment. *Journal of Statistical Software, 48*(6), 1-29. doi:10.18637/jss.v048.i06

numerical_deriv *Compute numerical derivatives*

## Description

Compute numerical derivatives using forward/backward difference, central difference, or Richardson extrapolation.

## Usage

```
numerical_deriv(
  par,
  f,
  ...,
  delta = 1e-05,
  gradient = TRUE,
  type = "Richardson"
)
```

## Arguments

| | |
|---|---|
| par | a vector of parameters to find partial derivative at |
| f | the objective function being evaluated |
| ... | additional arguments to be passed to f |
| delta | the term used to perturb the f function. Default is 1e-5 |
| gradient | logical; compute the gradient terms? If FALSE then the Hessian is computed instead |
| type | type of difference to compute. Can be either 'forward' for the forward difference, 'central' for the central difference, or 'Richardson' for the Richardson extrapolation (default). Backward difference is achieved by supplying a negative delta value with 'forward'. When type = 'Richardson', the default value of delta is increased to delta * 1000 for the Hessian and delta * 10 for the gradient to provide a reasonable perturbation starting location (each delta is halved at each iteration). |

## Author(s)

Phil Chalmers <rphilip.chalmers@gmail.com>

## Examples

```
f <- function(x) 3*x[1]^3 - 4*x[2]^2
par <- c(3,8)

# grad = 9 * x^2 , -8 * y
```

```
(actual <- c(9 * par[1]^2, -8 * par[2]))
numerical_deriv(par, f, type = 'forward')
numerical_deriv(par, f, type = 'central')
numerical_deriv(par, f, type = 'Richardson') # default

# Hessian = h11 -> 18 * x, h22 -> -8, h12 -> h21 -> 0
(actual <- matrix(c(18 * par[1], 0, 0, -8), 2, 2))
numerical_deriv(par, f, type = 'forward', gradient = FALSE)
numerical_deriv(par, f, type = 'central', gradient = FALSE)
numerical_deriv(par, f, type = 'Richardson', gradient = FALSE) # default
```

---

`personfit` *Person fit statistics*

---

## Description

`personfit` calculates the Zh values from Drasgow, Levine and Williams (1985) for unidimensional and multidimensional models, as well as the infit and outfit statistics. The returned object is a `data.frame` consisting either of the tabulated data or full data with the statistics appended to the rightmost columns.

## Usage

```
personfit(
  x,
  method = "EAP",
  Theta = NULL,
  stats.only = TRUE,
  return.resids = FALSE,
  ...
)
```

## Arguments

| | |
|---|---|
| `x` | a computed model object of class `SingleGroupClass` or `MultipleGroupClass` |
| `method` | type of factor score estimation method. See `fscores` for more detail |
| `Theta` | a matrix of factor scores used for statistics that require empirical estimates. If supplied, arguments typically passed to `fscores()` will be ignored and these values will be used instead |
| `stats.only` | logical; return only the person fit statistics without their associated response pattern? |
| `return.resids` | logical; return the standardized and unstandardized N by J matrices of person and item residuals? If `TRUE` will return a named list of each residual type |
| `...` | additional arguments to be passed to `fscores()` |

## Author(s)

Phil Chalmers <rphilip.chalmers@gmail.com>

## References

Chalmers, R. P. (2012). mirt: A Multidimensional Item Response Theory Package for the R Environment. *Journal of Statistical Software, 48*(6), 1-29. doi:10.18637/jss.v048.i06

Drasgow, F., Levine, M. V., & Williams, E. A. (1985). Appropriateness measurement with polychotomous item response models and standardized indices. *British Journal of Mathematical and Statistical Psychology, 38*, 67-86.

Reise, S. P. (1990). A comparison of item- and person-fit methods of assessing model-data fit in IRT. *Applied Psychological Measurement, 14*, 127-137.

Wright B. D. & Masters, G. N. (1982). *Rating scale analysis*. MESA Press.

## See Also

itemfit

## Examples

```
#make some data
set.seed(1)
a <- matrix(rlnorm(20),ncol=1)
d <- matrix(rnorm(20),ncol=1)
items <- rep('2PL', 20)
data <- simdata(a,d, 2000, items)

# first observation responds 1 for most difficult, 0 for easiest
data[1,] <- ifelse(d > 0, 0, 1)

# second observations answers first half as 1 second half as 0
data[2,] <- rep(1:0, each = 10)

x <- mirt(data, 1)
fit <- personfit(x)
head(fit)

# raw/standardized residuals
resid_list <- personfit(x, return.resids=TRUE)
head(resid_list$resid) # unstandardized
head(resid_list$std.resid) # standardized (approximate z-scores)

# with missing data
data[3, c(1,3,5,7)] <- NA
x.miss <- mirt(data, 1)
fit.miss <- personfit(x.miss)
head(fit.miss)
head(personfit(x.miss, return.resids=TRUE))
```

```
#using precomputed Theta
Theta <- fscores(x, method = 'MAP', full.scores = TRUE)
head(personfit(x, Theta=Theta))

# multiple group Rasch model example
set.seed(12345)
a <- matrix(rep(1, 16), ncol=1)
d <- matrix(rnorm(16,0,.7),ncol=1)
itemtype <- rep('dich', nrow(a))
N <- 1000
dataset1 <- simdata(a, d, N, itemtype)
dataset2 <- simdata(a, d, N, itemtype, sigma = matrix(1.5))
dat <- rbind(dataset1, dataset2)

# first observation responds 1 for most difficult, 0 for easiest
dat[1,] <- ifelse(d > 0, 0, 1)

group <- c(rep('D1', N), rep('D2', N))
models <- 'F1 = 1-16'
mod_Rasch <- multipleGroup(dat, models, itemtype = 'Rasch', group = group)
coef(mod_Rasch, simplify=TRUE)
pf <- personfit(mod_Rasch, method='MAP')
head(pf)
```

---

`pirt` *Projective IRT (PIRT) models*

---

## Description

Computes the projective IRT model parameters either using the logistic kernel approximation approaches (Stucky et al. 2013; Ip 2010/Doebler and Doebler, 2022) or the MML-PIRT approach by Chalmers et al. (2026). These return information pertaining to a lower-dimensional (often unidimensional) IRT model that has marginalized one or more latent traits. When the method is the MML-PIRT a working `mirt` object will be returned, otherwise if the target is the logistic kernel approximations then a list of the PIRT estimates (and potentially their delta-method SEs) are returned.

## Usage

```
pirt(
  mod,
  model = 1,
  project = 1,
  itemtype = extract.mirt(mod, "itemtype"),
  SE = TRUE,
  estimator = "MML",
```

```
  invariance = "",
  IRTpars = FALSE,
  shortform = NULL,
  ...
)
```

## Arguments

| | |
|---|---|
| mod | fitted model from mirt |
| model | type of PIRT model to fit to marginalized table |
| project | which dimension to project to in the stored E-table. Default projects all dimensions to the first dimension (project = 1), as this is the assumed primary dimension in, for instance, bfactor() |
| itemtype | type of model to fit to the projected marginalized factor. Defaults to the same type as the parent object mod. See mirt for details |
| SE | logical; compute the ACOV matrix to obtain standard errors (SE)? |
| estimator | type of PIRT estimator to use. Default is 'MML' to fit and return the MML-PIRT variable described by Chalmers et al. (in review). 'FA' will use Stucky et al.'s (2012) factor analysis logistic approximation to project to the first dimension, and 'LKA' will use the use logistic kernel approximation formula presented in Ip (2010)/Doebler and Doebler(2020), however these only project to the first dimension? The latter two estimators are also only applicable when input model is from the M*PL and MGRM family (see for polytomous MGRM derivation), and only return lists of the resulting parameters and SE estimates<br>Note that for methods other than 'MML' the SEs will be returned only if the parent model as a suitably estimated ACOV matrix to perform the delta method (e.g., bfactor(..., SE=TRUE) |
| invariance | type of group invariance to specify (see multipleGroup). Included here as the multiple-group PIRT approach needs to also calibrate the scale and location of the reference group, which is done silently internally whenever 'free_means' or 'free_vars' are specified |
| IRTpars | logical; report classical IRT parameterization and SEs? Only used when FAapprox = TRUE |
| shortform | character vector of item names, or numeric vector of item locations, used create a short-form by specifying which items to extract from the PIRT model. For instance, selecting one item from each specific factor has the property of identical marginal and conditional bifactors (see Stucky, Thissen, and Elden, 2013). Only supported when using the MML-PIRT approach as this returns the associated model object |
| ... | extra information passed to the estimation engine |

## Details

Logistic kernel approximation approaches are limited only to the M2PL (and related models, like the M4PL) and MGRM family of ordinal models, while MML-PIRT can be applied to any MIRT model supported by the package.

Standard error are computed when `FAapprox = TRUE` via the delta method only when the parent model estimated the ACOV matrix, while the MML-PIRT approach does not require this prerequisite as the ACOV can be computed directly. By default, MML-PIRT computes the associated ACOV matrix using the Oakes' identity given then marginalized E-table.

## References

Chalmers, R. P., Falk, C. F., and Reise, S. P. (2026). A General Approach for Estimating Projective IRT Models. Applied Psychological Measurement. DOI: 10.1177/01466216261474955

Doebler, A., and Doebler, P. (2022). Rotate and Project: Measurement of the Intended Concept with Unidimensional Item Response Theory. from Multidimensional Ordinal Items. Multivariate Behavioral Research, 57 (1), 40-56.

Ip, H. (2010). Empirically indistinguishable multidimensional IRT and locally dependent unidimensional item response models. British Journal of Mathematical and Statistical Psychology, 63(2),395-416.

Stucky, B. D., Thissen, D., & Orlando E., M. (2012). Using Logistic Approximations of Marginal Trace Lines to Develop Short Assessments. Applied Psychological Measurement, 37(1), 41-57.

## Examples

```
# Table 3 in Ip (2010)
as <- cbind(rep(c(2.63, 1.63, 1.38), each=5),
            c(rep(3.07, 5), numeric(10)),
            c(numeric(5), rep(1.36, 5), numeric(5)),
            c(numeric(10), rep(.69,5)))
as

set.seed(8675309)
d <- round(rnorm(15, 0, sd=1.5), 2)
nitems <- length(d)
nfact <- ncol(as)
dat <- simdata(as, d, 10000, itemtype='2PL')
itemstats(dat)


# bifactor model, storing E-table
mod <- bfactor(dat,
               "S1 = 1-5
                S2 = 6-10
                S3 = 11-15")
summary(mod)
coef(mod, simplify=TRUE)$items

# PIRT model via MML-PIRT
pmod <- pirt(mod)
coef(pmod)
coef(pmod, printSE=TRUE)
coef(pmod, simplify=TRUE)$items
```

```
# Logistic approximations (parameters only)
pirt(mod, estimator = 'FA')
pirt(mod, estimator = 'LKA')  # effectively the same

# plots
plot(pmod)
plot(pmod, type = 'info')
plot(pmod, type = 'itemscore')
plot(pmod, type = 'infotrace')
itemplot(pmod, 1, type = 'info', CE=TRUE)

# standardized loadings
summary(pmod)

# factor scores (EAP). Not optimal as the estimates/SEs ignore
#   the marginal components
fscores(pmod, method = 'EAPsum', full.scores=FALSE) # EAPs for sum-scores

# EAPs using Lord-Wingersky 2.0 approach (recommended if using sum-scores)
fscores(mod, method = 'EAPsum_2.0', full.scores=FALSE)

# EAP scores from full response data (not generally recommended)
fscores(pmod, full.scores=FALSE) |> head()
fs <- fscores(pmod)

# compare to bifactor scores (marginal vs conditional)
bfs <- fscores(mod, method = 'MAP')
cor(bfs[,1], fs)

# item fit (silly, but possible)
itemfit(pmod)

# three item shortform by taking the highest PIRT discrimination/slope terms
# per specific factor to ensure that the marginal and joint likelihood
# functions are comparable (see Stucky et al. 2013 for details)
slopes <- coef(pmod, simplify=TRUE)$items[,'a1']
slopes
namemax <- \(slp) names(slp)[which.max(slp)]
shortform <- c(namemax(slopes[1:5]),
               namemax(slopes[6:10]), namemax(slopes[11:15]))
pirt_short <- pirt(mod, shortform=shortform)
coef(pirt_short, simplify=TRUE)$items

# EAP-sum for short form
fscores(pirt_short, method='EAPsum', full.scores=FALSE)

############################
# same example, but with polytomous data

# Table 3 in Ip (2010)
as <- cbind(rep(c(2.63, 1.63, 1.38), each=5),
            c(rep(3.07, 5), numeric(10)),
            c(numeric(5), rep(1.36, 5), numeric(5)),
```

```
            c(numeric(10), rep(.69,5)))
as

set.seed(8675309)
diffs <- t(apply(matrix(runif(20*4, .5, 1), 20), 1, cumsum))
diffs <- -(diffs - rowMeans(diffs))
d <- diffs + rnorm(20)
nitems <- nrow(d)
nfact <- ncol(as)
dat <- simdata(as, d, 10000, itemtype='graded')
itemstats(dat)

# bifactor model, storing E-table
mod <- bfactor(dat,
               "S1 = 1-5
                S2 = 6-10
                S3 = 11-15")
summary(mod)
coef(mod, simplify=TRUE)$items

# PIRT model via MML-PIRT
pmod <- pirt(mod)
coef(pmod)
coef(pmod, printSE=TRUE)
coef(pmod, simplify=TRUE)$items

# Logistic approximations (parameters only)
pirt(mod, estimator = 'FA')
pirt(mod, estimator = 'LKA')  # see Doebler and Doebler (2020)

# plots
plot(pmod)
plot(pmod, type = 'info')
plot(pmod, type = 'trace')
plot(pmod, type = 'infotrace')
itemplot(pmod, 1, type = 'info', CE=TRUE)

# standardized loadings
summary(pmod)

# factor scores (EAP)
fscores(pmod, method = 'EAPsum', full.scores=FALSE) # EAPs for sum-scores

# EAP scores from full response data
fs <- fscores(pmod)

# compare to bifactor scores (marginal vs conditional)
bfs <- fscores(mod, method = 'MAP')
cor(bfs[,1], fs)


############################
# Two tier example projected to simple structure
```

```
# simulate data
set.seed(1234)
a <- matrix(c(
    0,1,0.5,NA,NA,
    0,1,0.5,NA,NA,
    0,1,0.5,NA,NA,
    0,1,0.5,NA,NA,
    0,1,0.5,NA,NA,
    0,1,NA,0.5,NA,
    0,1,NA,0.5,NA,
    0,1,NA,0.5,NA,
    1,0,NA,0.5,NA,
    1,0,NA,0.5,NA,
    1,0,NA,0.5,NA,
    1,0,NA,NA,0.5,
    1,0,NA,NA,0.5,
    1,0,NA,NA,0.5,
    1,0,NA,NA,0.5,
    1,0,NA,NA,0.5),ncol=5,byrow=TRUE)

d <- matrix(rnorm(16))
items <- rep('2PL', 16)

sigma <- diag(5)
sigma[1,2] <- sigma[2,1] <- .4
dataset <- simdata(a,d,2000,itemtype=items,sigma=sigma)
itemstats(dataset)

specific <- "
S1 = 1-5
S2 = 6-11
S3 = 12-16"

model <- '
    G1 = 1-8
    G2 = 9-16
    COV = G1*G2'

# quadpts dropped for faster estimation, but not as precise
simmod <- bfactor(dataset, specific, model, quadpts = 15, TOL = 1e-3)
coef(simmod, simplify=TRUE)

# simple structure projection
pirt2 <- pirt(simmod, model=model, project=1:2)
coef(pirt2, simplify=TRUE)
```

---

## pirt_DIF *Simulated datasets for PIRT-DIF*

---

## Description

Three associated datasets for PIRT-DIF, stored as a list (more information to come).

## Author(s)

Phil Chalmers <rphilip.chalmers@gmail.com>

## Examples

```
data(pirt_DIF)

# dataset 1
dat1 <- pirt_DIF$dat1
group <- dat1$group
dat <- dat1[,-1]
itemstats(dat, group=group)
```

---

# PLCI.mirt *Compute profiled-likelihood (or posterior) confidence intervals*

---

## Description

Computes profiled-likelihood based confidence intervals. Supports the inclusion of equality constraints. Object returns the confidence intervals and whether the respective interval could be found.

## Usage

```
PLCI.mirt(
  mod,
  parnum = NULL,
  alpha = 0.05,
  search_bound = TRUE,
  step = 0.5,
  lower = TRUE,
  upper = TRUE,
  inf2val = 30,
  NealeMiller = FALSE,
  verbose = interactive(),
  ...
)
```

## Arguments

| | |
|---|---|
| mod | a converged mirt model |
| parnum | a numeric vector indicating which parameters to estimate. Use mod2values to determine parameter numbers. If NULL, all possible parameters are used |
| alpha | two-tailed alpha critical level |
| search_bound | logical; use a fixed grid of values around the ML estimate to determine more suitable optimization bounds? Using this has much better behaviour than setting fixed upper/lower bound values and searching from more extreme ends |
| step | magnitude of steps used when search_bound is TRUE. Smaller values create more points to search a suitable bound for (up to the lower bound value visible with mod2values). When upper/lower bounds are detected this value will be adjusted accordingly |
| lower | logical; search for the lower CI? |
| upper | logical; search for the upper CI? |
| inf2val | a numeric used to change parameter bounds which are infinity to a finite number. Decreasing this too much may not allow a suitable bound to be located. Default is 30 |
| NealeMiller | logical; use the Neale and Miller (1997) approximation? Default is FALSE |
| verbose | logical; include additional information in the console? |
| ... | additional arguments to pass to the estimation functions |

## Author(s)

Phil Chalmers <rphilip.chalmers@gmail.com>

## References

Chalmers, R. P. (2012). mirt: A Multidimensional Item Response Theory Package for the R Environment. *Journal of Statistical Software, 48*(6), 1-29. doi:10.18637/jss.v048.i06

Chalmers, R. P., Pek, J., & Liu, Y. (2017). Profile-likelihood Confidence Intervals in Item Response Theory Models. *Multivariate Behavioral Research, 52*, 533-550. doi:10.1080/00273171.2017.1329082

Neale, M. C. & Miller, M. B. (1997). The use of likelihood-based confidence intervals in genetic models. *Behavior Genetics, 27*, 113-120.

## See Also

boot.mirt

## Examples

```
if(interactive()) mirtCluster() #use all available cores to estimate CI's in parallel
dat <- expand.table(LSAT7)
mod <- mirt(dat, 1)

result <- PLCI.mirt(mod)
```

```
result


mod2 <- mirt(Science, 1)
result2 <- PLCI.mirt(mod2)
result2

# only estimate CI's slopes
sv <- mod2values(mod2)
parnum <- sv$parnum[sv$name == 'a1']
result3 <- PLCI.mirt(mod2, parnum)
result3
```

---

`plot,MultipleGroupClass,missing-method`
*Plot various test-implied functions from models*

---

## Description

Plot various test implied response functions from models estimated in the mirt package.

## Usage

```
## S4 method for signature 'MultipleGroupClass,missing'
plot(
  x,
  y,
  type = "score",
  npts = 200,
  drop2 = TRUE,
  degrees = 45,
  which.items = 1:extract.mirt(x, "nitems"),
  rot = list(xaxis = -70, yaxis = 30, zaxis = 10),
  facet_items = TRUE,
  theta_lim = c(-6, 6),
  par.strip.text = list(cex = 0.7),
  gen.diff_type = "IRF",
 par.settings = list(strip.background = list(col = "#9ECAE1"), strip.border = list(col =
    "black")),
  auto.key = list(space = "right", points = FALSE, lines = TRUE),
  ...
)

## S4 method for signature 'SingleGroupClass,missing'
plot(
  x,
```

```
  y,
  type = "score",
  npts = 200,
  drop2 = TRUE,
  degrees = 45,
  theta_lim = c(-6, 6),
  which.items = 1:extract.mirt(x, "nitems"),
  gen.diff_type = "IRF",
  MI = 0,
  CI = 0.95,
  rot = list(xaxis = -70, yaxis = 30, zaxis = 10),
  facet_items = TRUE,
  main = NULL,
  drape = TRUE,
  colorkey = TRUE,
  ehist.cut = 1e-10,
  add.ylab2 = TRUE,
  par.strip.text = list(cex = 0.7),
 par.settings = list(strip.background = list(col = "#9ECAE1"), strip.border = list(col =
    "black")),
  auto.key = list(space = "right", points = FALSE, lines = TRUE),
  profile = FALSE,
  ...
)
```

## Arguments

| | |
|---|---|
| x | an object of class SingleGroupClass, MultipleGroupClass, or DiscreteClass |
| y | an arbitrary missing argument required for R CMD check |
| type | type of plot to view. Can be<br>'info' test information function<br>'rxx' for the reliability function<br>'infocontour' for the test information contours<br>'SE' for the test standard error function<br>'infotrace' item information traceline plots<br>'infoSE' a combined test information and standard error plot<br>'trace' item probability traceline plots<br>'itemscore' item scoring traceline plots<br>'score' expected total score surface<br>'scorecontour' expected total score contour plot<br>'posteriorTheta' posterior for the latent trait distribution<br>'gen.difficulty' plots items by generalized difficulty estimates (see gen.difficulty)<br>'EAPsum' compares sum-scores to the expected values based on the EAP for sum-scores method (see fscores)<br>Note that if dentype = 'empiricalhist' was used in estimation then the type 'empiricalhist' also will be available to generate the empirical histogram |

| | |
|---|---|
| | plot, and if dentype = 'Davidian-#' was used then the type 'Davidian' will also be available to generate the curve estimates at the quadrature nodes used during estimation |
| npts | number of quadrature points to be used for plotting features. Larger values make plots look smoother |
| drop2 | logical; where appropriate, for dichotomous response items drop the lowest category and provide information pertaining only to the second response option? |
| degrees | numeric value ranging from 0 to 90 used in plot to compute angle for information-based plots with respect to the first dimension. If a vector is used then a bubble plot is created with the summed information across the angles specified (e.g., degrees = seq(0, 90, by=10)) |
| which.items | numeric vector indicating which items to be used when plotting. Default is to use all available items |
| rot | allows rotation of the 3D graphics |
| facet_items | logical; apply grid of plots across items? If FALSE, items will be placed in one plot for each group |
| theta_lim | lower and upper limits of the latent trait (theta) to be evaluated, and is used in conjunction with npts |
| par.strip.text | plotting argument passed to lattice |
| gen.diff_type | argument passed to type in gen.difficulty |
| par.settings | plotting argument passed to lattice |
| auto.key | plotting argument passed to lattice |
| ... | additional arguments to be passed to lattice |
| MI | a single number indicating how many imputations to draw to form bootstrapped confidence intervals for the selected test statistic. If greater than 0 a plot will be drawn with a shaded region for the interval |
| CI | a number from 0 to 1 indicating the confidence interval to select when MI input is used. Default uses the 95% confidence (CI = .95) |
| main | argument passed to lattice. Default generated automatically |
| drape | logical argument passed to lattice. Default generated automatically |
| colorkey | logical argument passed to lattice. Default generated automatically |
| ehist.cut | a probability value indicating a threshold for excluding cases in empirical histogram plots. Values larger than the default will include more points in the tails of the plot, potentially squishing the 'meat' of the plot to take up less area than visually desired |
| add.ylab2 | logical argument passed to lattice. Default generated automatically |
| profile | logical; provide a profile plot of response probabilities (objects returned from mdirt only) |

## References

Chalmers, R. P. (2012). mirt: A Multidimensional Item Response Theory Package for the R Environment. *Journal of Statistical Software, 48*(6), 1-29. doi:10.18637/jss.v048.i06

## Examples

```
x <- mirt(Science, SE=TRUE)
plot(x)
plot(x, type = 'info')
plot(x, type = 'infotrace')
plot(x, type = 'infotrace', facet_items = FALSE)
plot(x, type = 'infoSE')
plot(x, type = 'rxx')
plot(x, type = 'posteriorTheta')
plot(x, type = 'gen.difficulty')

# confidence interval plots when information matrix computed
plot(x)
plot(x, MI=100)
plot(x, type='info', MI=100)
plot(x, type='SE', MI=100)
plot(x, type='rxx', MI=100)

# use the directlabels package to put labels on tracelines
library(directlabels)
plt <- plot(x, type = 'trace')
direct.label(plt, 'top.points')

# additional modifications can be made via update().
# See ?update.trellis for further documentation
plt
update(plt, ylab = expression(Prob(theta)),
            main = "Item Traceline Functions") # ylab/main changed

set.seed(1234)
group <- sample(c('g1','g2'), nrow(Science), TRUE)
x2 <- multipleGroup(Science, 1, group)
plot(x2)
plot(x2, type = 'trace')
plot(x2, type = 'trace', which.items = 1:2)
plot(x2, type = 'itemscore', which.items = 1:2)
plot(x2, type = 'trace', which.items = 1, facet_items = FALSE) #facet by group
plot(x2, type = 'info')
plot(x2, type = 'gen.difficulty')

x3 <- mirt(Science, 2)
plot(x3, type = 'info')
plot(x3, type = 'SE', theta_lim = c(-3,3))
```

---

`poly2dich` *Change polytomous items to dichotomous item format*

---

## Description

Transforms a matrix of items into a new matrix where the select polytomous items have been converted into comparable dichotomous items with the same information.

## Usage

```
poly2dich(data, which.items = 1:ncol(data), sep = "_cat.")
```

## Arguments

| | |
|---|---|
| `data` | an object of class `data.frame` or `matrix` |
| `which.items` | a vector indicating which items should be transformed into the dichotomous form. Default uses all input items |
| `sep` | character vector pattern to append to each item name in `data` |

## Value

Returns an integer matrix

## Author(s)

Phil Chalmers <rphilip.chalmers@gmail.com>

## References

Chalmers, R. P. (2012). mirt: A Multidimensional Item Response Theory Package for the R Environment. *Journal of Statistical Software, 48*(6), 1-29. doi:10.18637/jss.v048.i06

## Examples

```
data(Science)

head(Science)
newScience <- poly2dich(Science)
head(newScience)

newScience2 <- poly2dich(Science, which.items = 2)
head(newScience2)
```

## print-method *Print the model objects*

### Description

Print model object summaries to the console.

### Usage

```
## S4 method for signature 'SingleGroupClass'
print(x)
```

### Arguments

| | |
|---|---|
| x | an object of class SingleGroupClass, MultipleGroupClass, or MixedClass |

### References

Chalmers, R. P. (2012). mirt: A Multidimensional Item Response Theory Package for the R Environment. *Journal of Statistical Software, 48*(6), 1-29. doi:10.18637/jss.v048.i06

### Examples

```
x <- mirt(Science, 1)
print(x)
```

## print.mirt_df *Print generic for customized data.frame console output*

### Description

Provides a nicer output for most printed data.frame objects defined by functions in mirt.

### Usage

```
## S3 method for class 'mirt_df'
print(x, digits = 3, ...)
```

### Arguments

| | |
|---|---|
| x | object of class 'mirt_df' |
| digits | number of digits to round |
| ... | additional arguments passed to print(...) |

## print.mirt_list *Print generic for customized list console output*

### Description

Provides a nicer output for most printed `list` objects defined by functions in `mirt`.

### Usage

```
## S3 method for class 'mirt_list'
print(x, digits = 3, ...)
```

### Arguments

| | |
|---|---|
| `x` | object of class `'mirt_list'` |
| `digits` | number of digits to round |
| `...` | additional arguments passed to `print(...)` |

## print.mirt_matrix *Print generic for customized matrix console output*

### Description

Provides a nicer output for most printed `matrix` objects defined by functions in `mirt`.

### Usage

```
## S3 method for class 'mirt_matrix'
print(x, digits = 3, ...)

## S3 method for class 'matrix'
print(x, ...)
```

### Arguments

| | |
|---|---|
| `x` | object of class `'mirt_matrix'` |
| `digits` | number of digits to round |
| `...` | additional arguments passed to `print(...)` |

probtrace *Function to calculate probability trace lines*

## Description

Given an internal mirt object extracted from an estimated model, or the single-group estimated model itself, compute the probability trace lines for all categories.

## Usage

```
probtrace(x, Theta)
```

## Arguments

| | |
|---|---|
| x | either an extracted internal mirt object containing item information (see extract.item) or a model of class SingleGroupClass typically returned by the function mirt or bfactor |
| Theta | a vector (unidimensional) or matrix (unidimensional/multidimensional) of latent trait values |

## Author(s)

Phil Chalmers <rphilip.chalmers@gmail.com>

## References

Chalmers, R. P. (2012). mirt: A Multidimensional Item Response Theory Package for the R Environment. *Journal of Statistical Software, 48*(6), 1-29. doi:10.18637/jss.v048.i06

## See Also

extract.item, extract.group

## Examples

```
mod <- mirt(Science, 1)

# single item probabilty tracelines for Item 2
extr.2 <- extract.item(mod, 2)
Theta <- matrix(seq(-4,4, by = .1))
traceline <- probtrace(extr.2, Theta)
head(data.frame(traceline, Theta=Theta))

# probability tracelines for all items in test
tracelines <- probtrace(mod, Theta)
head(tracelines)
```

randef *Compute posterior estimates of random effect*

## Description

Stochastically compute random effects for `MixedClass` objects with Metropolis-Hastings samplers and averaging over the draws to obtain expected a posteriori predictions. Returns a list of the estimated effects.

## Usage

```
randef(x, ndraws = 1000, thin = 10, return.draws = FALSE)
```

## Arguments

| | |
|---|---|
| `x` | an estimated model object from the `mixedmirt` function |
| `ndraws` | total number of draws to perform. Default is 1000 |
| `thin` | amount of thinning to apply. Default is to use every 10th draw |
| `return.draws` | logical; return a list containing the thinned draws of the posterior? |

## Author(s)

Phil Chalmers <rphilip.chalmers@gmail.com>

## References

Chalmers, R. P. (2012). mirt: A Multidimensional Item Response Theory Package for the R Environment. *Journal of Statistical Software, 48*(6), 1-29.

Chalmers, R. P. (2015). Extended Mixed-Effects Item Response Models with the MH-RM Algorithm. *Journal of Educational Measurement, 52*, 200-222. doi:10.1111/jedm.12072 doi:10.18637/jss.v048.i06

## Examples

```
# make an arbitrary groups
covdat <- data.frame(group = rep(paste0('group', 1:49), each=nrow(Science)/49))

# partial credit model
mod <- mixedmirt(Science, covdat, model=1, random = ~ 1|group)
summary(mod)

effects <- randef(mod, ndraws = 2000, thin = 20)
head(effects$Theta)
head(effects$group)

# lr.random input
mod2 <- mixedmirt(Science, covdat, model=1, lr.random = ~ 1|group)
```

```
summary(mod2)

effects <- randef(mod2, ndraws = 2000)
head(effects$Theta)
head(effects$group)
```

---

RCI *Model-based Reliable Change Index*

---

## Description

Computes an IRT version of the "reliable change index" (RCI) proposed by Jacobson and Traux (1991) but modified to use IRT information about scores and measurement error (see Jabrayilov, Emons, and Sijtsma (2016)). Main benefit of the IRT approach is the inclusion of response pattern information in the pre/post data score estimates, as well as conditional standard error of measurement information. Models can be specified as separate unidimensional IRT models fitted via `mirt` (or extracted from `multipleGroup` via `extract.group`), or a two-dimensional model where the latent traits correspond to the two test administrations.

## Usage

```
RCI(
  mod_pre,
  predat,
  postdat,
  mod_post = mod_pre,
  cutoffs = NULL,
  SEM.pre = NULL,
  SEM.post = NULL,
  Fisher = FALSE,
  zero_cor = TRUE,
  expected.scores = FALSE,
  shiny = FALSE,
  main = "Test Scores",
  ...
)
```

## Arguments

| | |
|---|---|
| `mod_pre` | single-group model fitted by `mirt`. If not supplied the information will be extracted from the data input objects to compute the classical test theory version of the RCI statistics |
| `predat` | a vector (if one individual) or matrix/data.frame of response data to be scored, where each individuals' responses are included in exactly one row.<br>If total score information only is to be used instead of the complete response matrix then a `matrix` object with exactly `ncol = 1` columns should be provided |

| | |
|---|---|
| postdat | same as predat, but with respect to the post/follow-up measurement. Ignored when a two-dimensional IRT model is included.<br>For the original RCI approach, a matrix containing the complete responses, or a matrix with one column containing only the total scores, can be provided |
| mod_post | (optional) IRT model for post-test if different from pre-test; otherwise, the pre-test model will be used. Ignored when a two-dimensional model IRT is included |
| cutoffs | optional vector of length 2 indicating the type of cut-offs to report (e.g., c(-1.96, 1.96) reflects the 95 percent z-score type cut-off) |
| SEM.pre | standard error of measurement for the pretest. This can be used instead of rxx.pre and SD.pre |
| SEM.post | (optional) standard error of measurement for the post-test. Using this will create a pooled version of the SEM; otherwise, SEM.post = SEM.pre |
| Fisher | logical; use the Fisher/expected information function to compute the SE terms? If FALSE the SE information will be extracted from the select fscores method (default). Only applicable for unidimensional models |
| zero_cor | logical; when the supplied mod_pre is a two-factor model should the covariance/correlation between the latent traits be forced to be 0? |
| expected.scores | logical; when using IRT scoring methods, should the expected total scores be reported instead of the scaled (theta) scores returned from fscores? When set to TRUE the factor score estimates are passed to expected.test, and delta method standard error for the difference between teh scores are reported, however the original z and p-value information will be unchanged |
| shiny | logical; launch an interactive shiny applications for real-time scoring of supplied total-scores or response vectors? Only requires mod_pre and (optional) mod_post inputs |
| main | main label to use when shiny=TRUE |
| ... | additional arguments passed to fscores |

## Author(s)

Phil Chalmers <rphilip.chalmers@gmail.com>

## References

Chalmers, R. P. (2012). mirt: A Multidimensional Item Response Theory Package for the R Environment. *Journal of Statistical Software, 48*(6), 1-29. doi:10.18637/jss.v048.i06

Jacobson, N. S., & Truax, P. (1991). Clinical significance: A statistical approach to defining meaningful change in psychotherapy research. Journal of Consulting and Clinical Psychology, 59, 12-19.

Jabrayilov, R. , Emons, W. H. M., & Sijtsma, K. (2016). Comparison of Classical Test Theory and Item Response Theory in Individual Change Assessment. *Applied Psychological Measurement, 40* (8), 559-572.

## Examples

```
# simulate some data
N <- 1000
J <- 20     # number of items
a <- matrix(rlnorm(J,.2,.3))
d <- rnorm(J)

theta <- matrix(rnorm(N))
dat_pre <- simdata(a, d, itemtype = '2PL', Theta = theta)

# first 3 cases decrease by 1/2
theta2 <- theta - c(1/2, 1/2, 1/2, numeric(N-3))
dat_post <- simdata(a, d, itemtype = '2PL', Theta = theta2)

mod <- mirt(dat_pre)

# all changes using fitted model from pre data
RCI(mod, predat=dat_pre, postdat=dat_post)

# reported all expected change estimates in the original test-score metric
RCI(mod, predat=dat_pre, postdat=dat_post, expected.scores=TRUE)

# single response pattern change using EAP information
RCI(mod, predat=dat_pre[1,], postdat=dat_post[1,])

# WLE estimator with Fisher information for SE (see Jabrayilov et al. 2016)
RCI(mod, predat = dat_pre[1,], postdat = dat_post[1,],
    method = 'WLE', Fisher = TRUE)

# multiple respondents
RCI(mod, predat = dat_pre[1:6,], postdat = dat_post[1:6,])

# include large-sample z-type cutoffs
RCI(mod, predat = dat_pre[1:6,], postdat = dat_post[1:6,],
    cutoffs = c(-1.96, 1.96))

######
# CTT version by omitting IRT model
    # Requires either sample or population SEM's as input
(istats <- itemstats(dat_pre)$overall)
SEM.alpha <- istats$SEM.alpha    # SEM estimate of dat_pre

# assumes SEM.post = SEM.pre
RCI(predat = dat_pre, postdat = dat_post, SEM.pre=SEM.alpha)

# include cutoffs
RCI(predat = dat_pre, postdat = dat_post, SEM.pre=SEM.alpha,
    cutoffs=c(-1.96, 1.96))

# allows SEM.post != SEM.pre
```

```
(istats.post <- itemstats(dat_post)$overall)
SEM.alpha.post <- istats.post$SEM.alpha

RCI(predat = dat_pre, postdat = dat_post,
   SEM.pre=SEM.alpha, SEM.post=SEM.alpha.post)

# Supplying only the total scores
TS_pre <- matrix(rowSums(dat_pre))
TS_post <- matrix(rowSums(dat_post))
RCI(predat = TS_pre, postdat = TS_post,
   SEM.pre=SEM.alpha, SEM.post=SEM.alpha.post)

######

# interactive shiny interfaces for live scoring
mod_pre <- mirt(Science)

# (optional) setup mod_post to have medium effect size change (d = 0.5)
sv <- mod2values(mod_pre)
sv$value[sv$name == 'MEAN_1'] <- 0.5
mod_post <- mirt(Science, pars=sv, TOL=NA)

# only use pre-test model for scoring
if(interactive()){
    RCI(mod_pre=mod_pre, shiny=TRUE)

    # use both pre-test and post-test models for including empirical priors
    RCI(mod_pre=mod_pre, mod_post=mod_post, shiny=TRUE,
        main='Perceptions of Science and Technology')
 }


############################
# Example where individuals take completely different item set pre-post
#   but prior calibration has been performed to equate the items

dat <- key2binary(SAT12,
  key = c(1,4,5,2,3,1,2,1,3,1,2,4,2,1,5,3,4,4,1,4,3,3,4,1,3,5,1,3,1,5,4,5))

mod <- mirt(dat)

# with N=5 individuals under investigation
predat <- postdat <- dat[1:5,]
predat[, 17:32] <- NA
postdat[, 1:16] <- NA

head(predat)
head(postdat)

RCI(mod, predat, postdat)

# expected scores /32 (even though only /16 items answered)
RCI(mod, predat, postdat, expected.scores=TRUE)
```

```
######
# Two-dimensional IRT model for each time point, 20 items (no DIF)

J <- 20
N <- 500
slopes <- rlnorm(J, .2, .2)
a <- matrix(c(slopes, numeric(J*2), slopes),J*2)
ints <- rnorm(J)
d <- matrix(c(ints, ints), ncol=1)
data.frame(a=a, d=d)

# mean effects across time
mu <- c(0, -1/2)
sigma <- matrix(c(1, .7, .7, 1), 2,2)

dat <- simdata(a, d, N, mu=mu, sigma=sigma, itemtype = '2PL')
colnames(dat) <- c(paste0('Item.pre_', 1:20), paste0('Item.post_', 1:20))
itemstats(dat[,1:20])$overall
itemstats(dat[,21:40])$overall

# build equality constraints across time points
constr <- NULL
for(i in (1:J)){
    constr <- c(constr, paste0("(", i, ',', i+J, ",a1,a2)"))
    constr <- c(constr, paste0("(", i, ',', i+J, ",d)"))
}
constr <- paste0(constr, collapse=',')

# define model where item parameters constrained over time, and
# latent trait has potential scale-location changes (e.g., regression to the
# mean effects)
model <- sprintf("
                  thetapre = 1-%i,
                  thetapost = %i-%i,
                  COV = thetapre*thetapost, thetapost*thetapost
                  MEAN = thetapost
                  CONSTRAIN = %s", J, J+1, 2*J, constr)
cat(model)

# fit the model to calibration data
mod <- mirt(dat, model = model, SE=TRUE)
coef(mod, printSE=TRUE)
coef(mod, simplify=TRUE)
summary(mod)

# test data
Theta <- cbind(c(0, 1, 2), c(0,1,2))
nochange <- simdata(a, d, itemtype = '2PL', Theta = Theta)
change <- simdata(a, d, itemtype = '2PL', Theta = Theta +
                      cbind(0, c(-2, -2, -2)))

# total score differences
```

```
data.frame(pre=rowSums(nochange[,1:J]),
           post=rowSums(nochange[,1:J + J]))
data.frame(pre=rowSums(change[,1:J]),
           post=rowSums(change[,1:J + J]))


RCI(mod, predat = nochange)
RCI(mod, predat = change)

# expected total-score metric reported instead
RCI(mod, predat = nochange, expected.scores=TRUE)
RCI(mod, predat = change, expected.scores=TRUE)

############
# Similar example using the GRM instead, however
#  accounting for missing categories pre-post test (requires
#  the customK technical and strong equality constraints
#  to control the missing parameters)

J <- 20
N <- 500
slopes <- rlnorm(J, .2, .2)
a <- matrix(c(slopes, numeric(J*2), slopes),J*2)
ints <- rnorm(J)
diffs <- t(apply(matrix(runif(20*4, .3, 1), 20), 1, cumsum))
diffs <- -(diffs - rowMeans(diffs))
d <- diffs + rnorm(20)
d <- rbind(d,d)
data.frame(a=a, d=d)

# mean effects across time
mu <- c(0, -1/2)
sigma <- matrix(c(1, .7, .7, 1), 2,2)

dat <- simdata(a, d, N, mu=mu, sigma=sigma, itemtype = 'graded')
colnames(dat) <- c(paste0('Item.pre_', 1:20), paste0('Item.post_', 1:20))
itemstats(dat[,1:20])$overall
itemstats(dat[,21:40])$overall

# add missing response elements to dataset
datM <- dat
datM[datM[,1] == 4, 1] <- 3
datM[datM[,2] == 1, 2] <- 0
datM[datM[,3] == 0, 3] <- 1
itemstats(datM[,1:20])

# build equality constraints across time points
constr <- NULL
for(i in (1:J)){
    # constrain slopes across items/dimensions
    constr <- c(constr, paste0("(", i, ',', i+J, ",a1,a2)"))
    # constrain intercepts to be equal (required if any category missing)
    constr <- c(constr, paste0(paste0("(", i, ',', i+J, ",d", 1:4, ")")))
```

```
}
constr <- paste0(constr, collapse=',')

# define model where item parameters constrained over time, and
# latent trait has potential scale-location changes (e.g., regression to the
# mean effects)
model <- sprintf("
                  thetapre = 1-%i,
                  thetapost = %i-%i,
                  COV = thetapre*thetapost, thetapost*thetapost
                  MEAN = thetapost
                  CONSTRAIN = %s", J, J+1, 2*J, constr)
cat(model)

# fit on original data (no missing responses)
mod <- mirt(dat, model = model, itemtype='graded', SE=TRUE)
coef(mod, printSE=TRUE)
coef(mod, simplify=TRUE)
summary(mod)

# fit on items with missing responses requires customK
modM <- mirt(datM, model = model, itemtype='graded', SE=TRUE,
             technical = list(customK = rep(5, J*2)))

# test data
Theta <- cbind(c(0, 1, 2), c(0,1,2))
nochange <- simdata(a, d, itemtype = 'graded', Theta = Theta)
change <- simdata(a, d, itemtype = 'graded', Theta = Theta +
                      cbind(0, c(-2, -2, -2)))

# total score differences
data.frame(pre=rowSums(nochange[,1:J]),
           post=rowSums(nochange[,1:J + J]))
data.frame(pre=rowSums(change[,1:J]),
           post=rowSums(change[,1:J + J]))


RCI(modM, predat = nochange)
RCI(modM, predat = change)

# expected total-score metric reported instead
RCI(modM, predat = nochange, expected.scores=TRUE)
RCI(modM, predat = change, expected.scores=TRUE)
```

---

`read.mirt` *Translate mirt parameters into suitable structure for plink package*

---

## Description

This function exports item parameters from the `mirt` package to the `plink` package.

## Usage

```
read.mirt(x, as.irt.pars = TRUE, ...)
```

## Arguments

| | |
|---|---|
| x | a single object (or list of objects) returned from mirt, bfactor, or a single object returned by multipleGroup |
| as.irt.pars | if TRUE, the parameters will be output as an irt.pars object |
| ... | additional arguments to be passed to coef() |

## Author(s)

Phil Chalmers <rphilip.chalmers@gmail.com>

## Examples

```
## unidimensional
library(plink)

data <- expand.table(LSAT7)
(mod1 <- mirt(data, 1))
plinkpars <- read.mirt(mod1)
plot(plinkpars)
plot(mod1, type = 'trace')

# graded
mod2 <- mirt(Science, 1)
plinkpars <- read.mirt(mod2)
plot(plinkpars)
plot(mod2, type = 'trace')

# gpcm
mod3 <- mirt(Science, 1, itemtype = 'gpcm')
plinkpars <- read.mirt(mod3)
plot(plinkpars)
plot(mod3, type = 'trace')

# nominal
mod4 <- mirt(Science, 1, itemtype = 'nominal')
plinkpars <- read.mirt(mod4)
plot(plinkpars)
plot(mod4, type = 'trace')

## multidimensional

data <- expand.table(LSAT7)
(mod1 <- mirt(data, 2))
plinkpars <- read.mirt(mod1)
plinkpars
```

```
plot(plinkpars)
plot(mod1, type = 'trace')

cmod <- mirt.model('
   F1 = 1,4,5
   F2 = 2-4')
model <- mirt(data, cmod)
plot(read.mirt(model))
itemplot(model, 1)

# graded
mod2 <- mirt(Science, 2)
plinkpars <- read.mirt(mod2)
plinkpars
plot(plinkpars)
plot(mod2, type = 'trace')

### multiple group equating example
set.seed(1234)
dat <- expand.table(LSAT7)
group <- sample(c('g1', 'g2'), nrow(dat), TRUE)
dat1 <- dat[group == 'g1', ]
dat2 <- dat[group == 'g2', ]
mod1 <- mirt(dat1, 1)
mod2 <- mirt(dat2, 1)

# convert and combine pars
plinkMG <- read.mirt(list(g1=mod1, g2=mod2))

# equivalently:
# mod <- multipleGroup(dat, 1, group)
# plinkMG <- read.mirt(mod)

combine <- matrix(1:5, 5, 2)
comb <- combine.pars(plinkMG, combine, grp.names=unique(group))
out <- plink(comb, rescale="SL")
equate(out)
equate(out, method = 'OSE')
```

---

`remap.distance` *Remap item categories to have integer distances of 1*

---

## Description

The mirt package's estimation setup requires that all item responses have spaces equal to 1 (e.g., a Likert scale scored from 1 through 5). In the event that categories are missing the categories must be re-coded. This function is automatically called by the package estimation functions (e.g., `mirt`), however for convince this function has been extracted for users to better understand the remapping consequences.

### Usage

```
remap.distance(data, message = TRUE)
```

### Arguments

| | |
|---|---|
| data | the response data to remap as a data.frame or matrix |
| message | logical; print message information pertaining to which items were remapped? |

### Author(s)

Phil Chalmers <rphilip.chalmers@gmail.com>

### References

Chalmers, R. P. (2012). mirt: A Multidimensional Item Response Theory Package for the R Environment. *Journal of Statistical Software, 48*(6), 1-29. doi:10.18637/jss.v048.i06

### Examples

```
# category 2 for item 1 missing
dat <- Science
dat[,1] <- ifelse(Science[,1] == 2, 1, Science[,1])
apply(dat, 2, table)

# mirt() automatically remaps categories
mod <- mirt(dat, 1)
coef(mod, simplify=TRUE)

# this is the transformed data used by mirt()
remap_dat <- remap.distance(dat)
apply(remap_dat, 2, table)
```

---

## residuals-method *Compute model residuals*

---

### Description

Return model implied residuals for linear dependencies between items or at the person level. If the latent trait density was approximated (e.g., Davidian curves, Empirical histograms, etc) then passing use_dentype_estimate = TRUE will use the internally saved quadrature and density components (where applicable).

## Usage

```
## S4 method for signature 'SingleGroupClass'
residuals(
  object,
  type = "LD",
  p.adjust = "none",
  df.p = FALSE,
  approx.z = FALSE,
  full.scores = FALSE,
  QMC = FALSE,
  printvalue = NULL,
  tables = FALSE,
  verbose = TRUE,
  Theta = NULL,
  suppress = NA,
  theta_lim = c(-6, 6),
  quadpts = NULL,
  fold = TRUE,
  upper = TRUE,
  technical = list(),
  ...
)
```

## Arguments

| | |
|---|---|
| object | an object of class SingleGroupClass or MultipleGroupClass. Bifactor models are automatically detected and utilized for better accuracy |
| type | type of residuals to be displayed. Can be either 'LD' or 'LDG2' for a local dependence matrix based on the X2 or G2 statistics (Chen & Thissen, 1997), 'Q3' for the statistic proposed by Yen (1984), 'JSI' for the jack-knife statistic proposed Edwards et al. (2018), 'exp' for the expected values for the frequencies of every response pattern, and 'expfull' for the expected values for every theoretically observable response pattern. For the 'LD' and 'LDG2' types, the upper diagonal elements represent the standardized residuals in the form of signed Cramers V coefficients |
| p.adjust | method to use for adjusting all p-values (see p.adjust for available options). Default is 'none' |
| df.p | logical; print the degrees of freedom and p-values? |
| approx.z | logical; transform $\chi^2(df)$ information from LD tests into approximate z-ratios instead using the transformation $z = \sqrt{2 * \chi^2} - \sqrt{2 * df - 1}$? |
| full.scores | logical; compute relevant statistics for each subject in the original data? |
| QMC | logical; use quasi-Monte Carlo integration? If quadpts is omitted the default number of nodes is 5000 |
| printvalue | a numeric value to be specified when using the res='exp' option. Only prints patterns that have standardized residuals greater than abs(printvalue). The default (NULL) prints all response patterns |

| | |
|---|---|
| tables | logical; for LD type, return the observed, expected, and standardized residual tables for each item combination? |
| verbose | logical; allow information to be printed to the console? |
| Theta | a matrix of factor scores used for statistics that require empirical estimates (i.e., Q3). If supplied, arguments typically passed to fscores() will be ignored and these values will be used instead |
| suppress | a numeric value indicating which parameter local dependency combinations to flag as being too high (for LD, LDG2, and Q3 the standardize correlations are used; for JSI, the z-ratios are used). Absolute values for the standardized estimates greater than this value will be returned, while all values less than this value will be set to missing |
| theta_lim | range for the integration grid |
| quadpts | number of quadrature nodes to use. The default is extracted from model (if available) or generated automatically if not available |
| fold | logical; apply the sum 'folding' described by Edwards et al. (2018) for the JSI statistic? |
| upper | logical; which portion of the matrix (upper versus lower triangle) should the suppress argument be applied to? |
| technical | list of technical arguments when models are re-estimated (see mirt for details) |
| ... | additional arguments to be passed to fscores() |

## References

Chalmers, R. P. (2012). mirt: A Multidimensional Item Response Theory Package for the R Environment. *Journal of Statistical Software, 48*(6), 1-29. doi:10.18637/jss.v048.i06

Chen, W. H. & Thissen, D. (1997). Local dependence indices for item pairs using item response theory. *Journal of Educational and Behavioral Statistics, 22*, 265-289.

Edwards, M. C., Houts, C. R. & Cai, L. (2018). A Diagnostic Procedure to Detect Departures From Local Independence in Item Response Theory Models. *Psychological Methods, 23*, 138-149.

Yen, W. (1984). Effects of local item dependence on the fit and equating performance of the three parameter logistic model. *Applied Psychological Measurement, 8*, 125-145.

## Examples

```
x <- mirt(Science, 1)
residuals(x)
residuals(x, tables = TRUE)
residuals(x, type = 'exp')
residuals(x, suppress = .15)
residuals(x, df.p = TRUE)
residuals(x, df.p = TRUE, p.adjust = 'fdr') # apply FWE control

# Pearson's X2 estimate for goodness-of-fit
full_table <- residuals(x, type = 'expfull')
```

```
head(full_table)
X2 <- with(full_table, sum((freq - exp)^2 / exp))
df <- nrow(full_table) - extract.mirt(x, 'nestpars') - 1
p <- pchisq(X2, df = df, lower.tail=FALSE)
data.frame(X2, df, p, row.names='Pearson-X2')

# above FOG test as a function
PearsonX2 <- function(x){
   full_table <- residuals(x, type = 'expfull')
   X2 <- with(full_table, sum((freq - exp)^2 / exp))
   df <- nrow(full_table) - extract.mirt(x, 'nestpars') - 1
   p <- pchisq(X2, df = df, lower.tail=FALSE)
   data.frame(X2, df, p, row.names='Pearson-X2')
}
PearsonX2(x)


# extract results manually
out <- residuals(x, df.p = TRUE, verbose=FALSE)
str(out)
out$df.p[1,2]

# with and without supplied factor scores
Theta <- fscores(x)
residuals(x, type = 'Q3', Theta=Theta)
residuals(x, type = 'Q3', method = 'ML')

# Edwards et al. (2018) JSI statistic
N <- 250
a <- rnorm(10, 1.7, 0.3)
d <- rnorm(10)
dat <- simdata(a, d, N=250, itemtype = '2PL')

mod <- mirt(dat, 1)
residuals(mod, type = 'JSI')
residuals(mod, type = 'JSI', fold=FALSE) # unfolded

# LD between items 1-2
aLD <- numeric(10)
aLD[1:2] <- rnorm(2, 2.55, 0.15)
a2 <- cbind(a, aLD)
dat <- simdata(a2, d, N=250, itemtype = '2PL')

mod <- mirt(dat, 1)

# JSI executed in parallel over multiple cores
if(interactive()) mirtCluster()
residuals(mod, type = 'JSI')
```

`reverse.score` *Reverse score one or more items from a response matrix*

## Description

Reverse score specific items given empirical range or specific scoring range.

## Usage

```
reverse.score(data, which, range = NULL, append = ".RS")
```

## Arguments

| | |
|---|---|
| `data` | an object of class `data.frame`, `matrix`, or `table` with the response patterns |
| `which` | names of items or column integer location in `data` that should be rescored. If missing the all columns in `data` will be reverse scored |
| `range` | (optional) a named `list` to specify the low and high score ranges. Specified names must match the names found in `data`, and each element of this list should contain only two values. If items specified in `which` are omitted from this list then the empirical min/max information will be used instead |
| `append` | character vector indicating what to append to the item names that have been rescored |

## Value

returns the original `data` object with the specified items reverse scored replacing the original scoring scheme

## Author(s)

Phil Chalmers <rphilip.chalmers@gmail.com>

## References

Chalmers, R. P. (2012). mirt: A Multidimensional Item Response Theory Package for the R Environment. *Journal of Statistical Software, 48*(6), 1-29. doi:10.18637/jss.v048.i06

## Examples

```
a <- rlnorm(20)
a[c(1,5,10)] <- -a[c(1,5,10)]
diffs <- t(apply(matrix(runif(20*4, .3, 1), 20), 1, cumsum))
diffs <- -(diffs - rowMeans(diffs))
d <- diffs + rnorm(20)
dat <- simdata(a,d,itemtype='graded', N=300)
head(dat)
```

```
# fitted model has negative slopes due to flipped scoring
mod <- mirt(dat)
coef(mod, simplify=TRUE)$items
plot(mod, type = 'itemscore')


# reverse the scoring for items 1, 5, and 10 only using empirical min/max
revdat <- reverse.score(dat, c('Item_1', 'Item_5', 'Item_10'))
head(revdat)

# compare
apply(dat[,c(1,5,10)], 2, table)
apply(revdat[,c(1,5,10)], 2, table)


# slopes all positive now
mod2 <- mirt(revdat)
coef(mod2, simplify=TRUE)$items
plot(mod2, type = 'itemscore')


# use different empirical scoring information due to options not used
  # 0 score not observed for item 1, though should have been rescored to a 4
dat[dat[,1] == 0, 1] <- 1
table(dat[,1])

# 4 score not observed for item 5, though should have been rescored to a 0
dat[dat[,5] == 4, 5] <- 3
table(dat[,5])

# specify theoretical scoring values in the range list
revdat2 <- reverse.score(dat, c('Item_1', 'Item_5', 'Item_10'),
                              range = list(Item_1 = c(0,4), Item_5 = c(0,4)))
head(revdat2)
table(dat[,1])
table(revdat2[,1])

table(dat[,5])
table(revdat2[,5])
```

---

RMSD_DIF *RMSD effect size statistic to quantify category-level DIF*

---

## Description

This function computes a set of RMSD "badness-of-fit" statistics when investing DIF across a set of grouping variables. In a first step, a (potentially highly constrained) multiple group model is fitted, while in a second step the item (and person) parameters are estimated based on all examines across

all groups. Category level DIF is assessed based on how well the pseudo-table of counts match the (constrained) probability functions implied by the original multiple group model (while also weighing across the implied density function of the latent traits). If the RSMD fit is poor, indicating non-ignorable DIF, then the multiple-group model should be adjusted to better account for the large response bias due to using a pooled model. See Lee and von Davier (2020) and Buchholz and Hartig (2019) for details.

## Usage

```
RMSD_DIF(pooled_mod, flag = 0, probfun = TRUE, dentype = "norm")
```

## Arguments

| | |
|---|---|
| `pooled_mod` | a multiple-group model (used to compute the model-implied probability in the goodness-of-fit test) or a single-group model for computing RMSD goodness-of-fit values relative to the model-implied information |
| `flag` | a numeric value used as a cut-off to help flag larger RMSD values (e.g., `flag = .03` will highlight only categories with RMSD values greater than .03) |
| `probfun` | logical; use probability functions to compute RMSD? If FALSE, the expected score functions will be integrated instead, which may be useful for collapsing across the categories in polytomous items |
| `dentype` | density to use for the latent trait. Can be `'norm'` to use a normal Gaussian density where the mean/variance are extracted from the model object(default), `'snorm'` for a standard normal distribution, or `'empirical'` to use the density estimate obtained via the E-table |

## Details

This function is generally designed to work with multiple-group models, however if a single-group model is passed this will return the RMSD information regarding the observed vs expected response functions.

## Author(s)

Phil Chalmers <rphilip.chalmers@gmail.com>

## References

Buchholz, J., and Hartig, J. (2019). Comparing Attitudes Across Groups: An IRT-Based Item-Fit Statistic for the Analysis of Measurement Invariance. *Applied Psychological Measurement, 43*(3), 241-250. doi:10.1177/0146621617748323

Lee, S. S., and von Davier, M. (2020). Improving measurement properties of the PISA home possessions scale through partial invariance modeling. *Psychological test and assessment modeling*, 62(1):55-83.

## See Also

`DIF`, `DRF`, `multipleGroup`, `empirical_ES`

## Examples

```
#----- generate some data
set.seed(12345)
a <- a2 <- matrix(abs(rnorm(15,1,.3)), ncol=1)
d <- d2 <- matrix(rnorm(15,0,.7),ncol=1)

# item 1 has DIF
d2[1] <- d[1] - .5
a2[1] <- a[1] + 1

itemtype <- rep('2PL', nrow(a))
N <- 1000
dataset1 <- simdata(a, d, N, itemtype)
dataset2 <- simdata(a2, d2, N, itemtype)
dat <- rbind(dataset1, dataset2)
group <- c(rep('D1', N), rep('D2', N))

#-----

# fully pooled model
pooled_mod <- multipleGroup(dat, 1, group=group,
   invariance = c(colnames(dat), 'free_mean', 'free_var'))
coef(pooled_mod, simplify=TRUE)

RMSD_DIF(pooled_mod)
RMSD_DIF(pooled_mod, dentype = 'empirical')
RMSD_DIF(pooled_mod, flag = .03)

# more freely estimated model (item 1 has 2 parameters estimated)
MGmod <- multipleGroup(dat, 1, group=group,
                       invariance = c(colnames(dat)[-1], 'free_mean', 'free_var'))
coef(MGmod, simplify=TRUE)

# RMSD in item.1 now reduced (MG model accounts for DIF)
RMSD_DIF(MGmod)
RMSD_DIF(MGmod, flag = .03)

##################
# NA placeholders included when groups do not respond to specific items

a1 <- a2 <- rlnorm(20)
d <- d2 <- rnorm(20)
# item 5 contains DIF
a2[5] <- a1[5] + 1
d2[5] <- d[5] - 1/2
g <- rbeta(20, 5, 17)

dat1 <- simdata(a1, d, guess = g, N=1000, itemtype = '3PL')
dat1[, 11:13] <- NA  # items 11:13 items NA for g1
dat2 <- simdata(a2, d2, guess = g, N=1000, itemtype = '3PL',
```

```
   mu=1/4, sigma=matrix(.75))
dat2[,1:3] <- NA # items 1:3 items NA for g2
dat <- rbind(dat1, dat2)
group <- c(rep('g1', 1000), rep('g2', 1000))

mod <- multipleGroup(dat, "Theta = 1-20
                           PRIOR = (1-20, g, norm, -1, 0.5)",
                     group=group, itemtype='3PL',
                     invariance = c(colnames(dat), 'free_mean', 'free_var'))
coef(mod, simplify = TRUE)

RMSD_DIF(mod)
RMSD_DIF(mod, flag = .03)

#################
# Single group model (GOF only)
mod <- mirt(dat)
RMSD_DIF(mod)
RMSD_DIF(mod, probfun=FALSE)

# polytomous
mod2 <- mirt(Science)
RMSD_DIF(mod2)
RMSD_DIF(mod2, probfun=FALSE)

#################
# polytomous DIF example
set.seed(12345)
a <- a2 <- matrix(rlnorm(20,.2,.3))

# for the graded model, ensure that there is enough space between the intercepts,
# otherwise closer categories will not be selected often (minimum distance of 0.3 here)
diffs <- t(apply(matrix(runif(20*4, .3, 1), 20), 1, cumsum))
diffs <- -(diffs - rowMeans(diffs))
d <- d2 <- diffs + rnorm(20)

# item 1 has slope + dif for first intercept parameter
d2[1] <- d[1] - .5
a2[1] <- a[1] + 1

itemtype <- rep('graded', nrow(a))
N <- 1000
dataset1 <- simdata(a, d, N, itemtype)
dataset2 <- simdata(a2, d2, N, itemtype)
dat <- rbind(dataset1, dataset2)
group <- c(rep('D1', N), rep('D2', N))

#-----

# fully pooled model
pooled_mod <- multipleGroup(dat, 1, group=group,
         invariance = c(colnames(dat), 'free_mean', 'free_var'))
coef(pooled_mod, simplify=TRUE)
```

```
# Item_1 fits poorly in several categories (RMSD > .05)
RMSD_DIF(pooled_mod)
RMSD_DIF(pooled_mod, flag = .05)
RMSD_DIF(pooled_mod, flag = .1, probfun = FALSE) # use expected score function

# more freely estimated model (item 1 has more parameters estimated)
MGmod <- multipleGroup(dat, 1, group=group,
                       invariance = c(colnames(dat)[-1], 'free_mean', 'free_var'))
coef(MGmod, simplify=TRUE)

# RMSDs in Item_1 now reduced (MG model better accounts for DIF)
RMSD_DIF(MGmod)
RMSD_DIF(MGmod, flag = .05)
RMSD_DIF(MGmod, probfun = FALSE, flag = .1) # use expected score function
```

---

SAT12 *Description of SAT12 data*

---

## Description

Data obtained from the TESTFACT (Woods et al., 2003) manual, with 32 response pattern scored items for a grade 12 science assessment test (SAT) measuring topics of chemistry, biology, and physics. The scoring key for these data is [1, 4, 5, 2, 3, 1, 2, 1, 3, 1, 2, 4, 2, 1, 5, 3, 4, 4, 1, 4, 3, 3, 4, 1, 3, 5, 1, 3, 1, 5, 4, 5], respectively. However, careful analysis using the nominal response model suggests that the scoring key for item 32 may be incorrect, and should be changed from 5 to 3.

## Author(s)

Phil Chalmers <rphilip.chalmers@gmail.com>

## References

Wood, R., Wilson, D. T., Gibbons, R. D., Schilling, S. G., Muraki, E., & Bock, R. D. (2003). TESTFACT 4 for Windows: Test Scoring, Item Statistics, and Full-information Item Factor Analysis [Computer software]. Lincolnwood, IL: Scientific Software International.

## Examples

```
itemstats(SAT12, use_ts = FALSE)

# score the data (missing scored as 0)
head(SAT12)
dat <- key2binary(SAT12,
```

```
    key = c(1,4,5,2,3,1,2,1,3,1,2,4,2,1,5,3,4,4,1,4,3,3,4,1,3,5,1,3,1,5,4,5))
head(dat)
itemstats(dat)

# score the data, missing (value of 8) treated as NA
SAT12missing <- SAT12
SAT12missing[SAT12missing == 8] <- NA
dat <- key2binary(SAT12missing,
    key = c(1,4,5,2,3,1,2,1,3,1,2,4,2,1,5,3,4,4,1,4,3,3,4,1,3,5,1,3,1,5,4,5))
head(dat)

# potentially better scoring for item 32 (based on nominal model finding)
dat <- key2binary(SAT12,
    key = c(1,4,5,2,3,1,2,1,3,1,2,4,2,1,5,3,4,4,1,4,3,3,4,1,3,5,1,3,1,5,4,3))

# fit 2PL model to each item
mod <- mirt(dat)
plot(mod)
itemfit(mod)

# fit monotonic spline model to items 4 through 6 to improve fit
itemtype <- rep('2PL', 32)
itemtype[4:6] <- 'monospline'
mod_mspline <- mirt(dat, itemtype = itemtype)
anova(mod, mod_mspline)

plot(mod, which.items=4:6, type = 'trace')
plot(mod_mspline, which.items=4:6, type = 'trace')
```

---

`Science` *Description of Science data*

---

## Description

A 4-item data set borrowed from `ltm` package in R, first example of the `grm()` function. See more complete documentation therein, as well as Karlheinz and Melich (1992).

## Author(s)

Phil Chalmers <rphilip.chalmers@gmail.com>

## References

Karlheinz, R. and Melich, A. (1992). Euro-Barometer 38.1: *Consumer Protection and Perceptions of Science and Technology*. INRA (Europe), Brussels. [computer file]

## Examples

```
itemstats(Science)

mod <- mirt(Science, 1)
plot(mod, type = 'trace')
```

---

`secondOrderTest` *Second-order test of convergence*

---

## Description

Test whether terminated estimation criteria for a given model passes the second order test by checking the positive definiteness of the resulting Hessian matrix. This function, which accepts the symmetric Hessian/information matrix as the input, returns `TRUE` if the matrix is positive definite and `FALSE` otherwise.

## Usage

```
secondOrderTest(mat, ..., method = "eigen")
```

## Arguments

| | |
|---|---|
| `mat` | symmetric matrix to test for positive definiteness (typically the Hessian at the highest point of model estimator, such as MLE or MAP) |
| `...` | arguments passed to either `eigen`, `chol`, or `'det'` for the positiveness of the eigen values, positiveness of leading minors via the Cholesky decomposition, or evaluation of whether the determinant is greater than 0 |
| `method` | method to use to test positive definiteness. Default is `'eigen'` |

## Value

a matrix with all possible combinations

## Author(s)

Phil Chalmers <rphilip.chalmers@gmail.com>

## References

Chalmers, R. P. (2012). mirt: A Multidimensional Item Response Theory Package for the R Environment. *Journal of Statistical Software, 48*(6), 1-29. doi:10.18637/jss.v048.i06

## Examples

```
# PD matrix
mod <- mirt(Science, 1, SE=TRUE)
info <- solve(vcov(mod))   ## observed information
secondOrderTest(info)
secondOrderTest(info, method = 'chol')
secondOrderTest(info, method = 'det')

# non-PD matrix
mat <- matrix(c(1,0,0,0,1,1,0,1,1), ncol=3)
mat
secondOrderTest(mat)
secondOrderTest(mat, method = 'chol')
secondOrderTest(mat, method = 'det')
```

---

# show-method *Show model object*

---

## Description

Print model object summaries to the console.

## Usage

```
## S4 method for signature 'SingleGroupClass'
show(object)
```

## Arguments

| | |
|---|---|
| `object` | an object of class `SingleGroupClass`, `MultipleGroupClass`, or `MixedClass` |

## References

Chalmers, R. P. (2012). mirt: A Multidimensional Item Response Theory Package for the R Environment. *Journal of Statistical Software, 48*(6), 1-29. doi:10.18637/jss.v048.i06

## Examples

```
x <- mirt(Science, 1)
show(x)
```

SIBTEST *(Generalized) Simultaneous Item Bias Test (SIBTEST)*

## Description

Classical test theory approach to detecting unidirectional and bidirectional (with one crossing location) DIF. This family of statistics is intended for unidimensional tests, and applies a regression-corrected matched-total score approach to quantify the response bias between two or more groups. Can be used for DIF, DBF, and DTF testing with two or more discrete groups.

## Usage

```
SIBTEST(
  dat,
  group,
  suspect_set,
  match_set,
  focal_name = unique(group)[2],
  guess_correction = 0,
  Jmin = 5,
  na.rm = FALSE,
  randomize = FALSE,
  C = cbind(1, -diag(length(unique(group)) - 1L)),
  pairwise = FALSE,
  DIF = FALSE,
  p.adjust.method = "none",
  permute = 1000,
  pk_focal = FALSE,
  correction = TRUE,
  remove_cross = FALSE,
  details = FALSE,
  plot = "none",
  ...
)
```

## Arguments

| | |
|---|---|
| dat | integer-based dataset to be tested, containing dichotomous or polytomous responses |
| group | a (factor) vector indicating group membership with the same length as the number of rows in dat |
| suspect_set | an integer vector indicating which items to inspect with SIBTEST. Including only one value will perform a DIF test, while including more than one will perform a simultaneous bundle test (DBF); including all non-matched items will perform DTF. If missing, a simultaneous test using all the items not listed in match_set will be used (i.e., DTF) |

`match_set` an integer vector indicating which items to use as the items which are matched (i.e., contain no DIF). These are analogous to 'anchor' items in the likelihood method to locate DIF. If missing, all items other than the items found in the `suspect_set` will be used

`focal_name` name of the focal group; e.g., `'focal'`. If not specified then one will be selected automatically using `unique(group)[2]`

`guess_correction`
a vector of numbers from 0 to 1 indicating how much to correct the items for guessing. It's length should be the same as ncol(dat)

`Jmin` the minimum number of observations required when splitting the data into focal and reference groups conditioned on the matched set

`na.rm` logical; remove rows in `dat` with any missing values? If `TRUE`, rows with missing data will be removed, as well as the corresponding elements in the `group` input

`randomize` logical; perform the crossing test for non-compensatory bias using Li and Stout's (1996) permutation approach? Default is `FALSE`, which uses the ad-hoc mixed degrees of freedom method suggested by Chalmers (2018)

`C` a contrast matrix to use for pooled testing with more than two groups. Default uses an effects coding approach, where the last group (last column of the matrix) is treated as the reference group, and each column is associated with the respective name via `unique(group)` (i.e., the first column is the coefficient for `unique(group)[1]`, second column for `unique(group)[2]`, and so on)

`pairwise` logical; perform pairwise comparisons in multi-group applications?

`DIF` logical; should the elements in `suspect_set` be treated one at a time to test for DIF? Use of this logical will treat all other items as part of the `match_set` unless this input is provided explicitly. Default is `FALSE` to allow DBF and DTF tests

`p.adjust.method`
a character input dictating which `method` to use in `p.adjust`. when studying more than two groups. Default does not present any p-value adjustments

`permute` number of permutations to perform when `randomize = TRUE`. Default is 1000

`pk_focal` logical; using the group weights from the focal group instead of the total sample? Default is FALSE as per Shealy and Stout's recommendation

`correction` logical; apply the composite correction for the difference between focal composite scores using the true-score regression technique? Default is `TRUE`, reflecting Shealy and Stout's linear extrapolation method

`remove_cross` logical; remove the subtest information associated with the approximate crossing location? If TRUE this reflects the CSIBTEST definition of Li and Stout (1996); if FALSE, this reflects the version of CSIBTEST utilized by Chalmers (2018). Only applicable in two-group settings (in multi-group this is fixed to FALSE)

`details` logical; return a data.frame containing the details required to compute SIBTEST?

`plot` a character input indicating the type of plot to construct. Options are `'none'` (default), `'observed'` for the scaled focal subtest scores against the matched subtest scores, `'weights'` for the proportion weights used (i.e., the proportion of observations at each matched score), `'difference'` for the difference between the scaled focal subtest scores against the matched subtest scores, and

'wdifference' for the conditional differences multiplied by each respective weight. Note that the last plot reflects the components used in SIBTEST, and therefore the sum of these plotted observations will equal the beta coefficient for SIBTEST

... additional plotting arguments to be passed

## Details

SIBTEST is similar to the Mantel-Haenszel approach for detecting DIF but uses a regression correction based on the KR-20/coefficient alpha reliability index to correct the observed differences when the latent trait distributions are not equal. Function supports the standard SIBTEST for dichotomous and polytomous data (compensatory) and supports crossing DIF testing (i.e., non-compensatory/non-uniform) using the asymptotic sampling distribution version of the Crossing-SIBTEST (CSIBTEST) statistic described by Chalmers (2018) and the permutation method described by Li and Stout (1996). This function also supports the multi-group generalizations (GSIBTEST and GCSIBTEST) proposed by Chalmers and Zheng (2023), where users may specify alternative contrast matrices to evaluate specific comparisons between groups as well as perform joint hypothesis tests.

## Value

a data.frame type object containing the SIBTEST results. Note that the beta coefficient for (G)CSIBTEST are reported as absolute values to reflect the sum of the respective area information above and below the estimated crossing locations

## Author(s)

Phil Chalmers <rphilip.chalmers@gmail.com>

## References

Chalmers, R. P. (2018). Improving the Crossing-SIBTEST statistic for detecting non-uniform DIF. *Psychometrika, 83*, 2, 376-386.

Chalmers, R. P. (2012). mirt: A Multidimensional Item Response Theory Package for the R Environment. *Journal of Statistical Software, 48*(6), 1-29. doi:10.18637/jss.v048.i06

Chalmers, R. P. & Zheng, G. (2023). Multi-group Generalizations of SIBTEST and Crossing-SIBTEST. *Applied Measurement in Education, 36*(2), 171-191, doi:10.1080/08957347.2023.2201703.

Chang, H. H., Mazzeo, J. & Roussos, L. (1996). DIF for Polytomously Scored Items: An Adaptation of the SIBTEST Procedure. *Journal of Educational Measurement, 33*, 333-353.

Li, H.-H. & Stout, W. (1996). A new procedure for detection of crossing DIF. *Psychometrika, 61*, 647-677.

Shealy, R. & Stout, W. (1993). A model-based standardization approach that separates true bias/DIF from group ability differences and detect test bias/DTF as well as item bias/DIF. *Psychometrika, 58*, 159-194.

## Examples

```
set.seed(1234)
n <- 30
N <- 500
a <- matrix(1, n)
d <- matrix(rnorm(n), n)
group <- c(rep('reference', N), rep('focal', N*2))

## -------------
# groups completely equal
dat1 <- simdata(a, d, N, itemtype = 'dich')
dat2 <- simdata(a, d, N*2, itemtype = 'dich')
dat <- rbind(dat1, dat2)

# DIF (all other items as anchors)
SIBTEST(dat, group, suspect_set = 6)

# Some plots depicting the above tests
SIBTEST(dat, group, suspect_set = 6, plot = 'observed')
SIBTEST(dat, group, suspect_set = 6, plot = 'weights')
SIBTEST(dat, group, suspect_set = 6, plot = 'wdifference')

# Include CSIBTEST with randomization method
SIBTEST(dat, group, suspect_set = 6, randomize = TRUE)

# remove crossing-location (identical to Li and Stout 1996 definition of CSIBTEST)
SIBTEST(dat, group, suspect_set = 6, randomize = TRUE, remove_cross=TRUE)

# DIF (specific anchors)
SIBTEST(dat, group, match_set = 1:5, suspect_set = 6)
SIBTEST(dat, group, match_set = 1:5, suspect_set = 6, randomize=TRUE)

# DBF (all and specific anchors, respectively)
SIBTEST(dat, group, suspect_set = 11:30)
SIBTEST(dat, group, match_set = 1:5, suspect_set = 11:30)

# DTF
SIBTEST(dat, group, suspect_set = 11:30)
SIBTEST(dat, group, match_set = 1:10) #equivalent

# different hyper pars
dat1 <- simdata(a, d, N, itemtype = 'dich')
dat2 <- simdata(a, d, N*2, itemtype = 'dich', mu = .5, sigma = matrix(1.5))
dat <- rbind(dat1, dat2)
SIBTEST(dat, group, 6:30)
SIBTEST(dat, group, 11:30)

# DIF testing with anchors 1 through 5
SIBTEST(dat, group, 6, match_set = 1:5)
SIBTEST(dat, group, 7, match_set = 1:5)
```

```
SIBTEST(dat, group, 8, match_set = 1:5)

# DIF testing with all other items as anchors
SIBTEST(dat, group, 6)
SIBTEST(dat, group, 7)
SIBTEST(dat, group, 8)

## -------------
## systematic differing slopes and intercepts (clear DTF)
dat1 <- simdata(a, d, N, itemtype = 'dich')
dat2 <- simdata(a + c(numeric(15), rnorm(n-15, 1, .25)), d + c(numeric(15), rnorm(n-15, 1, 1)),
  N*2, itemtype = 'dich')
dat <- rbind(dat1, dat2)
SIBTEST(dat, group, 6:30)
SIBTEST(dat, group, 11:30)

# Some plots depicting the above tests
SIBTEST(dat, group, suspect_set = 11:30, plot = 'observed')
SIBTEST(dat, group, suspect_set = 11:30, plot = 'weights')
SIBTEST(dat, group, suspect_set = 11:30, plot = 'wdifference')

# DIF testing using valid anchors
SIBTEST(dat, group, suspect_set = 6, match_set = 1:5)
SIBTEST(dat, group, suspect_set = 7, match_set = 1:5)
SIBTEST(dat, group, suspect_set = 30, match_set = 1:5)

# test DIF using specific match_set
SIBTEST(dat, group, suspect_set = 6:30, match_set = 1:5, DIF=TRUE)

# test DIF using all-other-as-anchors method (not typically recommended)
SIBTEST(dat, group, suspect_set = 1:30, DIF=TRUE)

# randomization method is fairly poor when smaller matched-set used
SIBTEST(dat, group, suspect_set = 30, match_set = 1:5, randomize=TRUE)
SIBTEST(dat, group, suspect_set = 30, randomize=TRUE)

## ---------------------------------
# three group SIBTEST test
set.seed(1234)
n <- 30
N <- 1000
a <- matrix(1, n)
d <- matrix(rnorm(n), n)
group <- c(rep('group1', N), rep('group2', N), rep('group3', N))

# groups completely equal
dat1 <- simdata(a, d, N, itemtype = 'dich')
dat2 <- simdata(a, d, N, itemtype = 'dich')
dat3 <- simdata(a, d, N, itemtype = 'dich')
dat <- rbind(dat1, dat2, dat3)

# omnibus test using effects-coding contrast matrix (default)
SIBTEST(dat, group, suspect_set = 6)
```

```
SIBTEST(dat, group, suspect_set = 6, randomize=TRUE)

# explicit contrasts
SIBTEST(dat, group, suspect_set = 6, randomize=TRUE,
        C = matrix(c(1,-1,0), 1))

# test all items for DIF
SIBTEST(dat, group, suspect_set = 1:ncol(dat), DIF=TRUE)
SIBTEST(dat, group, suspect_set = 16:ncol(dat), DIF=TRUE,
        match_set = 1:15) # specific anchors

# post-hoc between two groups only
pick <- group %in% c('group1', 'group2')
SIBTEST(subset(dat, pick), group[pick], suspect_set = 1:ncol(dat), DIF=TRUE)

# post-hoc pairwise comparison for all groups
SIBTEST(dat, group, suspect_set = 1:ncol(dat), DIF=TRUE, pairwise = TRUE)

## systematic differing slopes and intercepts
dat2 <- simdata(a + c(numeric(15), .5,.5,.5,.5, numeric(10)),
        d + c(numeric(15), 0,.6,.7,.8,.9, numeric(10)),
        N, itemtype = 'dich')
dat <- rbind(dat1, dat2, dat3)

SIBTEST(dat, group, suspect_set = 16)
SIBTEST(dat, group, suspect_set = 16, randomize=TRUE)

SIBTEST(dat, group, suspect_set = 19)
SIBTEST(dat, group, suspect_set = 19, randomize=TRUE)

SIBTEST(dat, group, suspect_set = c(16, 19), DIF=TRUE)
SIBTEST(dat, group, suspect_set = c(16, 19), DIF=TRUE, pairwise=TRUE)
```

---

`simdata` *Simulate response patterns*

---

## Description

Simulates response patterns for compensatory and noncompensatory MIRT models from multivariate normally distributed factor ($\theta$) scores, or from a user input matrix of $\theta$'s.

## Usage

```
simdata(
  a,
  d,
  N,
```

```
  itemtype,
  sigma = NULL,
  mu = NULL,
  guess = 0,
  upper = 1,
  rho = NULL,
  nominal = NULL,
  t = NULL,
  Theta = NULL,
  gpcm_mats = list(),
  returnList = FALSE,
  model = NULL,
  equal.K = TRUE,
  which.items = NULL,
  mins = 0,
  lca_cats = NULL,
  prob.list = NULL
)
```

## Arguments

| | |
|---|---|
| a | a matrix/vector of slope parameters. If slopes are to be constrained to zero then use NA or simply set them equal to 0 |
| d | a matrix/vector of intercepts. The matrix should have as many columns as the item with the largest number of categories, and filled empty locations with NA. When a vector is used the test is assumed to consist only of dichotomous items (because only one intercept per item is provided). When `itemtype = 'lca'` intercepts will not be used |
| N | sample size |
| itemtype | a character vector of length `nrow(a)` (or 1, if all the item types are the same) specifying the type of items to simulate. Inputs can either be the same as the inputs found in the `itemtype` argument in `mirt` or the internal classes defined by the package. Typical `itemtype` inputs that are passed to `mirt` are used then these will be converted into the respective internal classes automatically.<br><br>If the internal class of the object is specified instead, the inputs can be `'dich'`, `'graded'`, `'gpcm'`, `'sequential'`, `'nominal'`, `'nestlogit'`, `'partcomp'`, `'gumm'`, `'lca'`, and all the models under the Luo (2001) family (see `mirt`), for dichotomous, graded, generalized partial credit, sequential, nominal, nested logit, partially compensatory, generalized graded unfolding model, latent class analysis model, and ordered unfolding models. Note that for the gpcm, nominal, and nested logit models there should be as many parameters as desired categories, however to parametrized them for meaningful interpretation the first category intercept should equal 0 for these models (second column for `'nestlogit'`, since first column is for the correct item traceline). For nested logit models the 'correct' category is always the lowest category (i.e., == 1). It may be helpful to use `mod2values` on data-sets that have already been estimated to understand the itemtypes more intimately |

`sigma` a covariance matrix of the underlying distribution. Default is the identity matrix. Used when `Theta` is not supplied

`mu` a mean vector of the underlying distribution. Default is a vector of zeros. Used when `Theta` is not supplied

`guess` a vector of guessing parameters for each item; only applicable for dichotomous items. Must be either a scalar value that will affect all of the dichotomous items, or a vector with as many values as to be simulated items

`upper` same as `guess`, but for upper bound parameters

`rho` a matrix of `rho` values to be used for the Lui (2001) family of ordered unfolding models (see `mirt`) to control the latitude of acceptance. All values must be positive

`nominal` a matrix of specific item category slopes for nominal models. Should be the dimensions as the intercept specification with one less column, with `NA` in locations where not applicable. Note that during estimation the first slope will be constrained to 0 and the last will be constrained to the number of categories minus 1, so it is best to set these as the values for the first and last categories as well

`t` matrix of t-values for the 'ggum' itemtype, where each row corresponds to a given item. Also determines the number of categories, where `NA` can be used for non-applicable categories

`Theta` a user specified matrix of the underlying ability parameters, where `nrow(Theta) == N` and `ncol(Theta) == ncol(a)`. When this is supplied the `N` input is not required

`gpcm_mats` a list of matrices specifying the scoring scheme for generalized partial credit models (see `mirt` for details)

`returnList` logical; return a list containing the data, item objects defined by `mirt` containing the population parameters and item structure, and the latent trait matrix `Theta`? Default is FALSE

`model` a single group object, typically returned by functions such as `mirt` or `bfactor`. Supplying this will render all other parameter elements (excluding the `Theta`, `N`, `mu`, and `sigma` inputs) redundant (unless explicitly provided). This input can therefore be used to create parametric bootstrap data whereby plausible data implied by the estimated model can be generated and evaluated

`equal.K` logical; when a `model` input is supplied, should the generated data contain the same number of categories as the original data indicated by `extract.mirt(model, 'K')`? Default is TRUE, which will redrawn data until this condition is satisfied

`which.items` an integer vector used to indicate which items to simulate when a `model` input is included. Default simulates all items

`mins` an integer vector (or single value to be used for each item) indicating what the lowest category should be. If `model` is supplied then this will be extracted from `slot(mod, 'Data')$mins`, otherwise the default is 0

`lca_cats` a vector indicating how many categories each lca item should have. If not supplied then it is assumed that 2 categories should be generated for each item

`prob.list` an optional list containing matrix/data.frames of probabilities values for each category to be simulated. This is useful when creating customized probability functions to be sampled from

## Details

Returns a data matrix simulated from the parameters, or a list containing the data, item objects, and Theta matrix.

## Author(s)

Phil Chalmers <rphilip.chalmers@gmail.com>

## References

Chalmers, R. P. (2012). mirt: A Multidimensional Item Response Theory Package for the R Environment. *Journal of Statistical Software, 48*(6), 1-29. doi:10.18637/jss.v048.i06

Reckase, M. D. (2009). *Multidimensional Item Response Theory*. New York: Springer.

## Examples

```
### Parameters from Reckase (2009), p. 153

set.seed(1234)

a <- matrix(c(
 .7471, .0250, .1428,
 .4595, .0097, .0692,
 .8613, .0067, .4040,
1.0141, .0080, .0470,
 .5521, .0204, .1482,
1.3547, .0064, .5362,
1.3761, .0861, .4676,
 .8525, .0383, .2574,
1.0113, .0055, .2024,
 .9212, .0119, .3044,
 .0026, .0119, .8036,
 .0008, .1905,1.1945,
 .0575, .0853, .7077,
 .0182, .3307,2.1414,
 .0256, .0478, .8551,
 .0246, .1496, .9348,
 .0262, .2872,1.3561,
 .0038, .2229, .8993,
 .0039, .4720, .7318,
 .0068, .0949, .6416,
 .3073, .9704, .0031,
 .1819, .4980, .0020,
 .4115,1.1136, .2008,
 .1536,1.7251, .0345,
 .1530, .6688, .0020,
 .2890,1.2419, .0220,
 .1341,1.4882, .0050,
 .0524, .4754, .0012,
 .2139, .4612, .0063,
 .1761,1.1200, .0870),30,3,byrow=TRUE)*1.702
```

```
d <- matrix(c(.1826,-.1924,-.4656,-.4336,-.4428,-.5845,-1.0403,
  .6431,.0122,.0912,.8082,-.1867,.4533,-1.8398,.4139,
  -.3004,-.1824,.5125,1.1342,.0230,.6172,-.1955,-.3668,
  -1.7590,-.2434,.4925,-.3410,.2896,.006,.0329),ncol=1)*1.702

mu <- c(-.4, -.7, .1)
sigma <- matrix(c(1.21,.297,1.232,.297,.81,.252,1.232,.252,1.96),3,3)

dataset1 <- simdata(a, d, 2000, itemtype = '2PL')
dataset2 <- simdata(a, d, 2000, itemtype = '2PL', mu = mu, sigma = sigma)

#mod <- mirt(dataset1, 3, method = 'MHRM')
#coef(mod)



### Unidimensional graded response model with 5 categories each

a <- matrix(rlnorm(20,.2,.3))

# for the graded model, ensure that there is enough space between the intercepts,
# otherwise closer categories will not be selected often (minimum distance of 0.3 here)
diffs <- t(apply(matrix(runif(20*4, .3, 1), 20), 1, cumsum))
diffs <- -(diffs - rowMeans(diffs))
d <- diffs + rnorm(20)

dat <- simdata(a, d, 500, itemtype = 'graded')
# mod <- mirt(dat, 1)

### An example of a mixed item, bifactor loadings pattern with correlated specific factors

a <- matrix(c(
.8,.4,NA,
.4,.4,NA,
.7,.4,NA,
.8,NA,.4,
.4,NA,.4,
.7,NA,.4),ncol=3,byrow=TRUE)

d <- matrix(c(
-1.0,NA,NA,
 1.5,NA,NA,
 0.0,NA,NA,
0.0,-1.0,1.5,  #the first 0 here is the recommended constraint for nominal
0.0,1.0,-1, #the first 0 here is the recommended constraint for gpcm
2.0,0.0,NA),ncol=3,byrow=TRUE)

nominal <- matrix(NA, nrow(d), ncol(d))
# the first 0 and last (ncat - 1) = 2 values are the recommended constraints
nominal[4, ] <- c(0,1.2,2)

sigma <- diag(3)
```

```
sigma[2,3] <- sigma[3,2] <- .25
items <- c('2PL','2PL','2PL','nominal','gpcm','graded')

dataset <- simdata(a,d,2000,items,sigma=sigma,nominal=nominal)

#mod <- bfactor(dataset, c(1,1,1,2,2,2), itemtype=c(rep('2PL', 3), 'nominal', 'gpcm','graded'))
#coef(mod)

#### Convert standardized factor loadings to slopes

F2a <- function(F, D=1.702){
    h2 <- rowSums(F^2)
    a <- (F / sqrt(1 - h2)) * D
    a
}

(F <- matrix(c(rep(.7, 5), rep(.5,5))))
(a <- F2a(F))
d <- rnorm(10)

dat <- simdata(a, d, 5000, itemtype = '2PL')
mod <- mirt(dat, 1)
coef(mod, simplify=TRUE)$items
summary(mod)

mod2 <- mirt(dat, 'F1 = 1-10
                   CONSTRAIN = (1-5, a1), (6-10, a1)')
summary(mod2)
anova(mod2, mod)

#### Convert classical 3PL paramerization into slope-intercept form
nitems <- 50
as <- rlnorm(nitems, .2, .2)
bs <- rnorm(nitems, 0, 1)
gs <- rbeta(nitems, 5, 17)

# convert first item (only intercepts differ in resulting transformation)
traditional2mirt(c('a'=as[1], 'b'=bs[1], 'g'=gs[1], 'u'=1), cls='3PL')

# convert all difficulties to intercepts
ds <- numeric(nitems)
for(i in 1:nitems)
   ds[i] <- traditional2mirt(c('a'=as[i], 'b'=bs[i], 'g'=gs[i], 'u'=1),
                             cls='3PL')[2]

dat <- simdata(as, ds, N=5000, guess=gs, itemtype = '3PL')

# estimate with beta prior for guessing parameters
# mod <- mirt(dat, model="Theta = 1-50
#                         PRIOR = (1-50, g, expbeta, 5, 17)", itemtype = '3PL')
# coef(mod, simplify=TRUE, IRTpars=TRUE)$items
# data.frame(as, bs, gs, us=1)
```

```
#### Unidimensional nonlinear factor pattern

theta <- rnorm(2000)
Theta <- cbind(theta,theta^2)

a <- matrix(c(
.8,.4,
.4,.4,
.7,.4,
.8,NA,
.4,NA,
.7,NA),ncol=2,byrow=TRUE)
d <- matrix(rnorm(6))
itemtype <- rep('2PL',6)

nonlindata <- simdata(a=a, d=d, itemtype=itemtype, Theta=Theta)

#model <- '
#F1 = 1-6
#(F1 * F1) = 1-3'
#mod <- mirt(nonlindata, model)
#coef(mod)

#### 2PLNRM model for item 4 (with 4 categories), 2PL otherwise

a <- matrix(rlnorm(4,0,.2))

# first column of item 4 is the intercept for the correct category of 2PL model,
#    otherwise nominal model configuration
d <- matrix(c(
-1.0,NA,NA,NA,
 1.5,NA,NA,NA,
 0.0,NA,NA,NA,
 1, 0.0,-0.5,0.5),ncol=4,byrow=TRUE)

nominal <- matrix(NA, nrow(d), ncol(d))
nominal[4, ] <- c(NA,0,.5,.6)

items <- c(rep('2PL',3),'nestlogit')

dataset <- simdata(a,d,2000,items,nominal=nominal)

#mod <- mirt(dataset, 1, itemtype = c('2PL', '2PL', '2PL', '2PLNRM'), key=c(NA,NA,NA,0))
#coef(mod)
#itemplot(mod,4)

# return list of simulation parameters
listobj <- simdata(a,d,2000,items,nominal=nominal, returnList=TRUE)
str(listobj)

# generate dataset from converged model
mod <- mirt(Science, 1, itemtype = c(rep('gpcm', 3), 'nominal'))
```

```
sim <- simdata(model=mod, N=1000)
head(sim)

Theta <- matrix(rnorm(100))
sim <- simdata(model=mod, Theta=Theta)
head(sim)

# alternatively, define a suitable object with functions from the mirtCAT package
# help(generate.mirt_object)
library(mirtCAT)

nitems <- 50
a1 <- rlnorm(nitems, .2,.2)
d <- rnorm(nitems)
g <- rbeta(nitems, 20, 80)
pars <- data.frame(a1=a1, d=d, g=g)
head(pars)

obj <- generate.mirt_object(pars, '3PL')
dat <- simdata(N=200, model=obj)

#### 10 item GGUMs test with 4 categories each
a <- rlnorm(10, .2, .2)
b <- rnorm(10) #passed to d= input, but used as the b parameters
diffs <- t(apply(matrix(runif(10*3, .3, 1), 10), 1, cumsum))
t <- -(diffs - rowMeans(diffs))

dat <- simdata(a, b, 1000, 'ggum', t=t)
apply(dat, 2, table)
# mod <- mirt(dat, 1, 'ggum')
# coef(mod)

### 10 items with the hyperbolic cosine model with differing category lengths
a <- matrix(1, 10)
d <- rnorm(10)
rho <- matrix(1:2, nrow=10, ncol=2, byrow=TRUE)
rho[1:2,2] <- NA   # first two items have K=2 categories

dat <- simdata(a, d, 1000, 'hcm', rho=rho)
itemstats(dat)
# mod <- mirt(dat, 1, 'hcm')
# list(est=coef(mod, simplify=TRUE)$items, pop=cbind(a, d, log(rho)))


######
# prob.list example

# custom probability function that returns a matrix
fun <- function(a, b, theta){
    P <- 1 / (1 + exp(-a * (theta-b)))
    cbind(1-P, P)
}
```

```
set.seed(1)
theta <- matrix(rnorm(100))
prob.list <- list()
nitems <- 5
a <- rlnorm(nitems, .2, .2); b <- rnorm(nitems, 0, 1/2)
for(i in 1:nitems) prob.list[[i]] <- fun(a[i], b[i], theta)
str(prob.list)

dat <- simdata(prob.list=prob.list)
head(dat)

# prob.list input is useful when defining custom items as well
name <- 'old2PL'
par <- c(a = .5, b = -2)
est <- c(TRUE, TRUE)
P.old2PL <- function(par,Theta, ncat){
     a <- par[1]
     b <- par[2]
     P1 <- 1 / (1 + exp(-1*a*(Theta - b)))
     cbind(1-P1, P1)
}

x <- createItem(name, par=par, est=est, P=P.old2PL)

prob.list[[1]] <- x@P(x@par, theta)
```

---

`SingleGroupClass-class` *Class "SingleGroupClass"*

---

## Description

Defines the object returned from `mirt` when model is exploratory.

## Slots

`Call`: function call

`Data`: list of data, sometimes in different forms

`Options`: list of estimation options

`Fit`: a list of fit information

`Model`: a list of model-based information

`ParObjects`: a list of the S4 objects used during estimation

`OptimInfo`: a list of arguments from the optimization process

`Internals`: a list of internal arguments for secondary computations (inspecting this object is generally not required)

`vcov`: a matrix represented the asymptotic covariance matrix of the parameter estimates

`time`: a data.frame indicating the breakdown of computation times in seconds

## Methods

**anova** `signature(object = "SingleGroupClass")`

**coef** `signature(object = "SingleGroupClass")`

**plot** `signature(x = "SingleGroupClass", y = "missing")`

**print** `signature(x = "SingleGroupClass")`

**residuals** `signature(object = "SingleGroupClass")`

**show** `signature(object = "SingleGroupClass")`

**summary** `signature(object = "SingleGroupClass")`

## Author(s)

Phil Chalmers <rphilip.chalmers@gmail.com>

## References

Chalmers, R. P. (2012). mirt: A Multidimensional Item Response Theory Package for the R Environment. *Journal of Statistical Software, 48*(6), 1-29. doi:10.18637/jss.v048.i06

---

# SLF *Social Life Feelings Data*

---

## Description

A 5-item data set analyzed by Bartholomew (1998). Data contains dichotomous responses (endorsement vs non-endorsement) from 1490 German respondents to five statements on perceptions of social life.

## Author(s)

Phil Chalmers <rphilip.chalmers@gmail.com>

## References

Bartholomew, D., J. (1998). Scaling unobservable constructs in social science. Journal of the Royal Statistical Society - Series C, 47, 1-13.

## Examples

```
# tabular format
data(SLF)
SLF

# full dataset
full <- expand.table(SLF)
itemstats(full)

mod <- mirt(full)
plot(mod, type = 'trace')
```

---

summary-method *Summary of model object*

---

## Description

Transforms coefficients into a standardized factor loading's metric. For MixedClass objects, the fixed and random coefficients are printed. Note that while the output to the console is rounded to three digits, the returned list of objects is not. For simulations, use output <- summary(mod, verbose = FALSE) to suppress the console messages.

## Usage

```
## S4 method for signature 'SingleGroupClass'
summary(
  object,
  SE = TRUE,
  rotate = "oblimin",
  Target = NULL,
  suppress = 0,
  suppress.cor = 0,
  verbose = TRUE,
  ...
)
```

## Arguments

| | |
|---|---|
| object | an object of class SingleGroupClass, MultipleGroupClass, or MixedClass |
| SE | logical; include the standard errors for the standardized loadings? Requires the initial model to have included and estimated of the asymptotic covariance matrix (via, for instance, mirt(..., SE = TRUE)). If TRUE SEs are computed using the delta method |

| | |
|---|---|
| rotate | a string indicating which rotation to use for exploratory models, primarily from the GPArotation package (see documentation therein).<br>Rotations currently supported are: 'promax', 'oblimin', 'varimax', 'quartimin', 'targetT', 'targetQ', 'pstT', 'pstQ', 'oblimax', 'entropy', 'quartimax', 'simplimax', 'bentlerT', 'bentlerQ', 'tandemI', 'tandemII', 'geominT', 'geominQ', 'cfT', 'cfQ', 'infomaxT', 'infomaxQ', 'mccammon', 'bifactorT', 'bifactorQ'.<br>For models that are not exploratory this input will automatically be set to 'none' |
| Target | a dummy variable matrix indicting a target rotation pattern. This is required for rotations such as 'targetT', 'targetQ', 'pstT', and 'pstQ' |
| suppress | a numeric value indicating which (possibly rotated) factor loadings should be suppressed. Typical values are around .3 in most statistical software. Default is 0 for no suppression |
| suppress.cor | same as suppress, but for the correlation matrix output |
| verbose | logical; allow information to be printed to the console? |
| ... | additional arguments to be passed |

## References

Chalmers, R. P. (2012). mirt: A Multidimensional Item Response Theory Package for the R Environment. *Journal of Statistical Software, 48*(6), 1-29. doi:10.18637/jss.v048.i06

## See Also

coef-method

## Examples

```
x <- mirt(Science, 2)
summary(x)
summary(x, rotate = 'varimax')
```

---

# tail.mirt_matrix *Tail generic for customized matrix console output*

---

## Description

Provides a nicer output for most printed matrix objects defined by functions in mirt.

## Usage

```
## S3 method for class 'mirt_matrix'
tail(x, digits = 3, ...)
```

### Arguments

| | |
|---|---|
| x | object of class 'mirt_matrix' |
| digits | number of digits to round |
| ... | additional arguments passed to print(...) |

---

## testinfo *Function to calculate test information*

---

### Description

Given an estimated model compute the test information.

### Usage

```
testinfo(
  x,
  Theta,
  degrees = NULL,
  group = NULL,
  individual = FALSE,
  which.items = 1:extract.mirt(x, "nitems")
)
```

### Arguments

| | |
|---|---|
| x | an object of class 'SingleGroupClass', or an object of class 'MultipleGroupClass' if a suitable group input were supplied |
| Theta | a matrix of latent trait values |
| degrees | a vector of angles in degrees that are between 0 and 90. Only applicable when the input object is multidimensional |
| group | group argument to pass to extract.group function. Required when the input object is a multiple-group model |
| individual | logical; return a data.frame of information traceline for each item? |
| which.items | an integer vector indicating which items to include in the expected information function. Default uses all possible items |

### Author(s)

Phil Chalmers <rphilip.chalmers@gmail.com>

### References

Chalmers, R. P. (2012). mirt: A Multidimensional Item Response Theory Package for the R Environment. *Journal of Statistical Software, 48*(6), 1-29. doi:10.18637/jss.v048.i06

## Examples

```
dat <- expand.table(deAyala)
(mirt(dat, 1, '2PL', pars = 'values'))
mod <- mirt(dat, 1, '2PL', constrain = list(c(1,5,9,13,17)))

Theta <- matrix(seq(-4,4,.01))
tinfo <- testinfo(mod, Theta)
plot(Theta, tinfo, type = 'l')



# compare information loss between two tests
tinfo_smaller <- testinfo(mod, Theta, which.items = 3:5)

# removed item informations
plot(Theta, iteminfo(extract.item(mod, 1), Theta), type = 'l')
plot(Theta, iteminfo(extract.item(mod, 2), Theta), type = 'l')

# most loss of info around -1 when removing items 1 and 2; expected given item info functions
plot(Theta, tinfo_smaller - tinfo, type = 'l')
```

---

`thetaComb` *Create all possible combinations of vector input*

---

## Description

This function constructs all possible k-way combinations of an input vector. It is primarily useful when used in conjunction with the `mdirt` function, though users may have other uses for it as well. See `expand.grid` for more flexible combination formats.

## Usage

```
thetaComb(theta, nfact, intercept = FALSE)
```

## Arguments

| | |
|---|---|
| `theta` | the vector from which all possible combinations should be obtained |
| `nfact` | the number of observations (and therefore the number of columns to return in the matrix of combinations) |
| `intercept` | logical; should a vector of 1's be appended to the first column of the result to include an intercept design component? Default is `FALSE` |

## Value

a matrix with all possible combinations

## Author(s)

Phil Chalmers <rphilip.chalmers@gmail.com>

## References

Chalmers, R. P. (2012). mirt: A Multidimensional Item Response Theory Package for the R Environment. *Journal of Statistical Software, 48*(6), 1-29. doi:10.18637/jss.v048.i06

## Examples

```
# all possible joint combinations for the vector -4 to 4
thetaComb(-4:4, 2)

# all possible binary combinations for four observations
thetaComb(c(0,1), 4)

# all possible binary combinations for four observations (with intercept)
thetaComb(c(0,1), 4, intercept=TRUE)
```

---

# traditional2mirt *Convert traditional IRT metric into slope-intercept form used in mirt*

---

## Description

This is a helper function for users who have previously available traditional/classical IRT parameters and want to know the equivalent slope-intercept translation used in `mirt`. Note that this function assumes that the supplied models are unidimensional by definition (i.e., will have only one slope/discrimination) and in the logistic metric (i.e., logistic-ogive scaling coefficient D=1). If there is no supported slope-intercept transformation available then the original vector of parameters will be returned by default.

## Usage

```
traditional2mirt(x, cls, ncat)
```

## Arguments

| | |
|---|---|
| `x` | a vector of parameters to transform |
| `cls` | the class or itemtype of the supplied model |
| `ncat` | the number of categories implied by the IRT model |

## Details

Supported class transformations for the `cls` input are:

**Rasch, 2PL, 3PL, 3PLu, 4PL** Form must be: (discrimination, difficulty, lower-bound, upper-bound)

**graded** Form must be: (discrimination, difficulty 1, difficulty 2, ..., difficulty k-1)

**gpcm** Form must be: (discrimination, difficulty 1, difficulty 2, ..., difficulty k-1)

**nominal** Form must be: (discrimination 1, discrimination 2, ..., discrimination k, difficulty 1, difficulty 2, ..., difficulty k)

## Value

a named vector of slope-intercept parameters (if supported)

## Examples

```
# classical 3PL model
vec <- c(a=1.5, b=-1, g=.1, u=1)
slopeint <- traditional2mirt(vec, '3PL', ncat=2)
slopeint

# classical graded model (four category)
vec <- c(a=1.5, b1=-1, b2=0, b3=1.5)
slopeint <- traditional2mirt(vec, 'graded', ncat=4)
slopeint

# classical generalize partial credit model (four category)
vec <- c(a=1.5, b1=-1, b2=0, b3=1.5)
slopeint <- traditional2mirt(vec, 'gpcm', ncat=4)
slopeint

# classical nominal model (4 category)
vec <- c(a1=.5, a2 = -1, a3=1, a4=-.5, d1=1, d2=-1, d3=-.5, d4=.5)
slopeint <- traditional2mirt(vec, 'nominal', ncat=4)
slopeint
```

---

# vcov-method *Extract parameter variance covariance matrix*

---

## Description

Extract parameter variance covariance matrix

## Usage

```
## S4 method for signature 'SingleGroupClass'
vcov(object)
```

### Arguments

| | |
|---|---|
| object | an object of class SingleGroupClass, MultipleGroupClass, or MixedClass |

### References

Chalmers, R. P. (2012). mirt: A Multidimensional Item Response Theory Package for the R Environment. *Journal of Statistical Software, 48*(6), 1-29. doi:10.18637/jss.v048.i06

### Examples

```
x <- mirt(Science, 1, SE=TRUE)
vcov(x)
```

---

## wald *Wald statistics for mirt models*

---

### Description

Compute a Wald test given an L vector or matrix of numeric contrasts. Requires that the model information matrix be computed (by passing SE = TRUE when estimating the model). Use wald(model) to observe how the information matrix columns are named, especially if the estimated model contains constrained parameters (e.g., 1PL).

### Usage

```
wald(object, L, C = NULL)
```

### Arguments

| | |
|---|---|
| object | estimated object from mirt, bfactor, multipleGroup, mixedmirt, or mdirt |
| L | a coefficient matrix with dimensions nconstrasts x npars.estimated, or a character vector giving the hypothesis in symbolic form (syntax format borrowed from the car package; see Details below). Omitting this value will return the column names of the information matrix used to identify the (potentially constrained) parameters |
| C | a constant vector of population parameters to be compared along side L, where length(C) == row(L). By default a vector of 0's is constructed. Note that when using the syntax input for L this argument is ignored<br>The following description is borrowed from car package documentation pertaining to the character vector input to the argument L: "The hypothesis matrix can be supplied as a numeric matrix (or vector), the rows of which specify linear combinations of the model coefficients, which are tested equal to the corresponding entries in the right-hand-side vector, which defaults to a vector of zeroes. |

Alternatively, the hypothesis can be specified symbolically as a character vector with one or more elements, each of which gives either a linear combination of coefficients, or a linear equation in the coefficients (i.e., with both a left and right side separated by an equals sign). Components of a linear expression or linear equation can consist of numeric constants, or numeric constants multiplying coefficient names (in which case the number precedes the coefficient, and may be separated from it by spaces or an asterisk); constants of 1 or -1 may be omitted. Spaces are always optional. Components are separated by plus or minus signs. Newlines or tabs in hypotheses will be treated as spaces. See the examples below."

## Author(s)

Phil Chalmers <rphilip.chalmers@gmail.com>

## References

Chalmers, R. P. (2012). mirt: A Multidimensional Item Response Theory Package for the R Environment. *Journal of Statistical Software, 48*(6), 1-29. doi:10.18637/jss.v048.i06

## Examples

```
# View parnumber index
data(LSAT7)
data <- expand.table(LSAT7)
mod <- mirt(data, 1, SE = TRUE)
coef(mod)

# see how the information matrix relates to estimated parameters, and how it lines up
#   with the parameter index
(infonames <- wald(mod))
index <- mod2values(mod)
index[index$est, ]

# second item slope equal to 0?
L <- matrix(0, 1, 10)
L[1,3] <- 1
wald(mod, L)

# same as above using character syntax input
infonames
wald(mod, "a1.5 = 0")

# simultaneously test equal factor slopes for item 1 and 2, and 4 and 5
L <- matrix(0, 2, 10)
L[1,1] <- L[2, 7] <- 1
L[1,3] <- L[2, 9] <- -1
L
wald(mod, L)

# Again, using more efficient syntax
```

```
infonames
wald(mod, c("a1.1 = a1.5", "a1.13 = a1.17"))

# log-Liklihood tests (requires estimating a new model)
cmodel <- 'theta = 1-5
           CONSTRAIN = (1,2, a1), (4,5, a1)'
mod2 <- mirt(data, cmodel)
# or, equivalently
#mod2 <- mirt(data, 1, constrain = list(c(1,5), c(13,17)))
anova(mod2, mod)

#####
# test equality of means in multi-group model:
#    H0: (mu1 - mu2) = (mu3 - mu4)

set.seed(12345)
a <- matrix(abs(rnorm(15,1,.3)), ncol=1)
d <- matrix(rnorm(15,0,.7),ncol=1)
itemtype <- rep('2PL', nrow(a))
N <- 500
dataset1 <- simdata(a, d, N, itemtype)
dataset2 <- simdata(a, d, N, itemtype, mu = .5)
dataset3 <- simdata(a, d, N, itemtype, mu = -1)
dataset4 <- simdata(a, d, N, itemtype, mu = -.5)
dat <- rbind(dataset1, dataset2, dataset3, dataset4)
group <- factor(rep(paste0('D', 1:4), each=N))
levels(group)
models <- 'F1 = 1-15'

# 3 means estimated
mod_free <- multipleGroup(dat, models, group = group, SE=TRUE,
                          invariance=c('slopes', 'intercepts', 'free_var','free_means'))
wald(mod_free) # obtain parameter names
# View(mod2values(mod_free))

# reference group mean = 0 by default
wald(mod_free, c("0 - MEAN_1.123 = MEAN_1.185 - MEAN_1.247"))
```

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