

# Package: PNAR (via r-universe)

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**Type** Package

**Title** Poisson Network Autoregressive Models

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**Description** Quasi likelihood-based methods for estimating linear and log-linear Poisson Network Autoregression models with  $p$  lags and covariates. Tools for testing the linearity versus several non-linear alternatives. Tools for simulation of multivariate count distributions, from linear and non-linear PNAR models, by using a specific copula construction. References include:  
Armillotta, M. and K. Fokianos (2023). ``Nonlinear network autoregression". Annals of Statistics, 51(6): 2526--2552. <doi:10.1214/23-AOS2345>.  
Armillotta, M. and K. Fokianos (2024). ``Count network autoregression". Journal of Time Series Analysis, 45(4): 584--612. <doi:10.1111/jtsa.12728>.  
Armillotta, M., Tsagris, M. and Fokianos, K. (2024). ``Inference for Network Count Time Series with the R Package PNAR". The R Journal, 15/4: 255--269. <doi:10.32614/RJ-2023-094>.

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 PNAR-package

*Poisson Network Autoregressive Models*


---

## Description

Quasi likelihood-based methods for estimating linear and log-linear Poisson Network Autoregression models with  $p$  lags and covariates. Tools for testing the linearity versus several non-linear alternatives. Tools for simulation of multivariate count distributions, from linear and non-linear PNAR models, by using a specific copula construction. References include:

Armiglotta, M. and K. Fokianos (2023). Nonlinear network autoregression. *Annals of Statistics*, 51(6): 2526–2552.

Armiglotta, M. and K. Fokianos (2024). Count network autoregression. *Journal of Time Series Analysis*, 45(4): 584–612.

Armiglotta, M., Tsagris, M. and Fokianos, K. (2024). Inference for Network Count Time Series with the R Package PNAR. *The R Journal*, 15/4: 255–269.

## Details

Package: PNAR  
Type: Package  
Version: 1.7  
Date: 2024-09-05  
License: GPL(>=2)

## Note

**Disclaimer:** Dr Mirko Armillotta and Konstantinos Fokianos wrote the initial functions. Dr Tsagris modified them, created the package and he is the maintainer.

We would like to acknowledge Manos Papadakis for his help with the "htest" class object and S3 methods (print() and summary() functions).

## Author(s)

Michail Tsagris, Mirko Armillotta and Konstantinos Fokianos.

## References

- Armillotta, M. and K. Fokianos (2024). Count network autoregression. *Journal of Time Series Analysis*, 45(4): 584–612.
- Armillotta, M. and K. Fokianos (2023). Nonlinear network autoregression. *Annals of Statistics*, 51(6): 2526–2552.
- Armillotta, M., Tsagris, M. and Fokianos, K. (2024). Inference for Network Count Time Series with the R Package PNAR. *The R Journal*, 15/4: 255–269.

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adja

*Generation of a network from the Stochastic Block Model*

---

## Description

This function generates a network from the Stochastic Block Model with  $K$  blocks.

## Usage

```
adja(N, K, alpha, directed = FALSE)
```

**Arguments**

N	The number of nodes on the network.
K	The number of blocks. Each block has dimension $N/K$ . $K$ should be chosen such that $N$ is divisible by $K$ .
alpha	The network density. A value in $[0, 1]$ defining the frequency of connections in the network.
directed	Logical scalar, whether to generate a directed network or not. If TRUE a directed network is generated.

**Details**

For each pair of nodes it performs a Bernoulli trial with values 1 "draw an edge", 0 "otherwise". The probabilities of these trials are bigger if the two nodes are in the same block, lower otherwise, and they are specified based on the number of nodes on the network  $N$  and network density *alpha*: Probability to draw an edge for a pair of nodes in the same block:  $\alpha * N^{-0.3}$ . Probability to draw an edge for a pair of nodes in different blocks:  $\alpha * N^{-1}$ .

**Value**

A row-normalized non-negative matrix describing the network. The main diagonal entries of the matrix are zeros, all the other entries are non-negative and the sum of elements over the rows equals one.

**Author(s)**

Mirko Armillotta, Michail Tsagris and Konstantinos Fokianos.

**References**

Faust, K. and S. Wasserman (1992). Blockmodels: Interpretation and evaluation. *Social Networks*, 14, 5–61.

**See Also**

[adja\\_gnp](#)

**Examples**

```
W <- adja(N = 20, K = 5, alpha = 0.1)
```

adja\_gnp

*Generation of a network from the Erdos-Renyi model***Description**

This function generates a network from the Erdos-Renyi model.

**Usage**

```
adja_gnp(N, alpha, directed = FALSE)
```

**Arguments**

N	The number of nodes on the network.
alpha	The network density. A value in $[0, 1]$ defining the frequency of connections in the network.
directed	Logical scalar, whether to generate a directed network. If TRUE a directed network is generated.

**Details**

For each pair of nodes it performs a Bernoulli trial with values 1 "draw an edge", 0 "otherwise". Each trial has the same probability of having an edge; this is equal to  $\alpha * N^{-0.3}$ , specified based on the number of nodes on the network  $N$  and the network density alpha.

**Value**

A row-normalized non-negative matrix describing the network. The main diagonal entries of the matrix are zeros, all the other entries are non-negative and the maximum sum of elements over the rows equals one.

**Author(s)**

Mirko Armillotta, Michail Tsagris and Konstantinos Fokianos.

**References**

Erdos, P. and A. Renyi (1959). On random graphs. *Publicationes Mathematicae*, 6, 290–297.

**See Also**

[adja](#)

**Examples**

```
W <- adja_gnp(N = 20, alpha= 0.1)
```

---

crime	<i>Chicago crime dataset</i>
-------	------------------------------

---

**Description**

Monthly number of burglaries on the south side of Chicago (552 blocks) during 2010-2015 (72 temporal observations).

**Usage**

```
crime
```

**Format**

A time series object ("ts" class) with multivariate time series, a matrix with 72 rows and 552 columns.

**Source**

Clark and Dixon (2021), available at <https://github.com/nick3703/Chicago-Data>.

**References**

Clark, N. J. and P. M. Dixon (2021). A class of spatially correlated self-exciting statistical models. *Spatial Statistics*, 43, 1–18.

**See Also**

[crime\\_W](#), [lin\\_estimnarpq](#), [log\\_lin\\_estimnarpq](#)

**Examples**

```
data(crime)
data(crime_W)
mod1 <- lin_estimnarpq( crime, crime_W, p = 1)
mod2 <- log_lin_estimnarpq( crime, crime_W, p = 1)
```

---

`crime_W`*Network matrix for Chicago crime dataset*

---

**Description**

Non-negative row-normalized adjacency matrix describing the network structure between Chicago census blocks.

**Usage**`crime_W`**Format**

A matrix with 552 rows and 552 columns.

**Source**

Clark and Dixon (2021), available at <https://github.com/nick3703/Chicago-Data>.

**References**

Clark, N. J. and P. M. Dixon (2021). A class of spatially correlated self-exciting statistical models. *Spatial Statistics*, 43, 1–18.

**See Also**

[crime](#), [lin\\_estimnarpq](#), [log\\_lin\\_estimnarpq](#)

**Examples**

```
data(crime)
data(crime_W)
mod1 <- lin_estimnarpq(crime, crime_W, p = 1)
mod2 <- log_lin_estimnarpq(crime, crime_W, p = 1)
```

---

`getN`*Count the number of events within a specified time*

---

**Description**

This function counts the number of events within a specified time.

**Usage**`getN(x, tt = 1)`

**Arguments**

x	A matrix of (positive) inter-event times.
tt	A positive time.

**Value**

The number of events within time `tt` (possibly 0), for each column of `x`.

**Author(s)**

Mirko Armillotta, Michail Tsagris and Konstantinos Fokianos.

**See Also**

[rcopula](#), [poisson.MODpq](#), [poisson.MODpq.log](#)

**Examples**

```
x <- rcopula(n = 100, N = 50, rho = 0.3)
getN(x)
```

---

global\_optimise\_LM\_stnarpq

*Optimization of the score test statistic for the ST-PNAR( $p$ ) model*

---

**Description**

Global optimization of the linearity test statistic for the Smooth Transition Poisson Network Autoregressive model of order  $p$  with  $q$  covariates (ST-PNAR( $p$ )) with respect to the nuisance scale parameter  $\gamma$ .

**Usage**

```
global_optimise_LM_stnarpq(gama_L = NULL, gama_U = NULL, len = 10, b, y, W,
p, d, Z = NULL, tol = 1e-9)
```

**Arguments**

gama_L	The lower value of the $\gamma$ values to consider. Use NULL if there is not information about its value. See the details for default computation.
gama_U	The upper value of the $\gamma$ values to consider. Use NULL if there is not information about its value. See the details for default computation.
len	The number of increments to consider for the $\gamma$ parameter.
b	The estimated parameters from the linear model, in the following order: (intercept, network parameters, autoregressive parameters, covariates). The length of the vector should be $2p + 1 + q$ , where $q$ denotes the number of covariates.



y	A $TT \times N$ time series object or a $TT \times N$ numerical matrix with the $N$ multi-variate count time series over $TT$ time periods.
W	The $N \times N$ row-normalized non-negative adjacency matrix describing the network. The main diagonal entries of the matrix should be zeros, all the other entries should be non-negative and the maximum sum of elements over the rows should equal one. The function row-normalizes the matrix if a non-normalized adjacency matrix is provided.
p	The number of lags in the model.
d	The lag parameter of non-linear variable (should be between 1 and $p$ ).
Z	An $N \times q$ matrix of covariates (one for each column), where $q$ is the number of covariates in the model. Note that they must be non-negative.
tol	Tolerance level for the optimizer.

### Details

The function optimizes the quasi score test statistic, under the null assumption of linearity, for testing linearity of Poisson Network Autoregressive model of order  $p$  against the following ST-PNAR( $p$ ) model, with respect to the unknown nuisance parameter ( $\gamma$ ). For each node of the network  $i = 1, \dots, N$  over the time sample  $t = 1, \dots, TT$

$$\lambda_{i,t} = \beta_0 + \sum_{h=1}^p (\beta_{1h} X_{i,t-h} + \beta_{2h} Y_{i,t-h} + \alpha_h e^{-\gamma X_{i,t-d}^2} X_{i,t-h}) + \sum_{l=1}^q \delta_l Z_{i,l}$$

where  $X_{i,t} = \sum_{j=1}^N W_{ij} Y_{j,t}$  is the network effect, i.e. the weighted average impact of node  $i$  connections, with the weights of the mean being  $W_{ij}$ , the single element of the network matrix  $W$ . The sequence  $\lambda_{i,t}$  is the expectation of  $Y_{i,t}$ , conditional to its past values.

The null hypothesis of the test is defined as  $H_0 : \alpha_1 = \dots = \alpha_p = 0$ , versus the alternative that at least one among  $\alpha_h$  is not 0. The test statistic has the form

$$LM(\gamma) = S'(\hat{\theta}, \gamma) \Sigma^{-1}(\hat{\theta}, \gamma) S(\hat{\theta}, \gamma)$$

where

$$S(\hat{\theta}, \gamma) = \sum_{t=1}^{TT} \sum_{i=1}^N \left( \frac{Y_{i,t}}{\lambda_{i,t}(\hat{\theta}, \gamma)} - 1 \right) \frac{\partial \lambda_{i,t}(\hat{\theta}, \gamma)}{\partial \alpha}$$

is the partition of the quasi score related to the vector of non-linear parameters  $\alpha = (\alpha_1, \dots, \alpha_p)$ , evaluated at the estimated parameters  $\hat{\theta}$  under the null assumption  $H_0$  (linear model) and  $\Sigma(\hat{\theta}, \gamma)$  is the variance of  $S(\hat{\theta}, \gamma)$ .

The optimization employs the Brent algorithm (Brent, 1973) applied in the interval from `gama_L` to `gama_U`. To be sure that the global optimum is found, the optimization is performed at `(len-1)` consecutive equidistant sub-intervals and then the maximum over them is taken as global optimum.

The values of `gama_L` and `gama_U` are computed internally as  $\text{gama\_L} = -\log(0.9)/X^2$  and  $\text{gama\_U} = -\log(0.1)/X^2$ , where  $X$  is the overall mean of  $X_{i,t}$  over the nodes  $i = 1, \dots, N$  and times  $t = 1, \dots, TT$ . Since the non-linear function  $e^{-\gamma X_{i,t-d}^2}$  ranges between 0 and 1, by considering  $X$  to be a representative value for the network mean, `gama_U` and `gama_L` would be the values of  $\gamma$  leading the non-linear switching function to be 0.1 and 0.9, respectively, so that in the optimization procedure the extremes of the function domain are excluded. Alternatively, their value can be supplied by the user. For details see Armillotta and Fokianos (2023, Sec. 4-5).

**Value**

A list including:

gama	The optimum value of the $\gamma$ parameter.
supLM	The value of the objective function at the optimum.
int	A vector with the extremes points of sub-intervals.

**Author(s)**

Mirko Armillotta, Michail Tsagris and Konstantinos Fokianos.

**References**

- Armillotta, M. and K. Fokianos (2023). Nonlinear network autoregression. *Annals of Statistics*, 51(6): 2526–2552.
- Armillotta, M. and K. Fokianos (2024). Count network autoregression. *Journal of Time Series Analysis*, 45(4): 584–612.
- Armillotta, M., Tsagris, M. and Fokianos, K. (2024). Inference for Network Count Time Series with the R Package PNAR. *The R Journal*, 15/4: 255–269.
- Brent, R. (1973) Algorithms for Minimization without Derivatives. Prentice-Hall, Englewood Cliffs N.J.

**See Also**

[score\\_test\\_stnarpq\\_j](#), [global\\_optimise\\_LM\\_tnarpq](#), [score\\_test\\_tnarpq\\_j](#)

**Examples**

```
data(crime)
data(crime_W)
mod1 <- lin_estimnarpq(crime, crime_W, p = 1)
b <- mod1$coefs[, 1]
global_optimise_LM_stnarpq(b = b, y = crime, W = crime_W, p = 1, d = 1)
```

---

global\_optimise\_LM\_tnarpq

*Optimization of the score test statistic for the T-PNAR(p) model*

---

**Description**

Global optimization of the linearity test statistic for the Threshold Poisson Network Autoregressive model of order  $p$  with  $q$  covariates (T-PNAR( $p$ )) with respect to the nuisance threshold parameter  $\gamma$ .

**Usage**

```
global_optimise_LM_tnarpq(gama_L = NULL, gama_U = NULL, len = 10, b, y, W,
p, d, Z = NULL, tol = 1e-9)
```

**Arguments**

gama_L	The lower value of the $\gamma$ values to consider. Use NULL if there is not information about its value.. See the details for default computation.
gama_U	The upper value of the $\gamma$ values to consider. Use NULL if there is not information about its value.. See the details for default computation.
len	The number of increments to consider for the $\gamma$ parameter.
b	The estimated parameters from the linear model, in the following order: (intercept, network parameters, autoregressive parameters, covariates). The dimension of the vector should be $2p + 1 + q$ , where $q$ denotes the number of covariates.
y	A $TT \times N$ time series object or a $TT \times N$ numerical matrix with the $N$ multivariate count time series over $TT$ time periods.
W	The $N \times N$ row-normalized non-negative adjacency matrix describing the network. The main diagonal entries of the matrix should be zeros, all the other entries should be non-negative and the maximum sum of elements over the rows should equal one. The function row-normalizes the matrix if a non-normalized adjacency matrix is provided.
p	The number of lags in the model.
d	The lag parameter of non-linear variable (should be between 1 and $p$ ).
Z	An $N \times q$ matrix of covariates (one for each column), where $q$ is the number of covariates in the model. Note that they must be non-negative.
tol	Tolerance level for the optimizer.

**Details**

The function optimizes the quasi score test statistic, under the null assumption of linearity, for testing linearity of Poisson Network Autoregressive model of order  $p$  against the following T-PNAR( $p$ ) model, with respect to the unknown nuisance parameter ( $\gamma$ ). For each node of the network  $i = 1, \dots, N$  over the time sample  $t = 1, \dots, TT$

$$\lambda_{i,t} = \beta_0 + \sum_{h=1}^p [\beta_{1h}X_{i,t-h} + \beta_{2h}Y_{i,t-h} + (\alpha_0 + \alpha_{1h}X_{i,t-h} + \alpha_{2h}Y_{i,t-h})I(X_{i,t-d} \leq \gamma)] + \sum_{l=1}^q \delta_l Z_{i,l}$$

where  $X_{i,t} = \sum_{j=1}^N W_{ij}Y_{j,t}$  is the network effect, i.e. the weighted average impact of node  $i$  connections, with the weights of the mean being  $W_{ij}$ , the single element of the network matrix  $W$ , and  $I()$  is the indicator function. The sequence  $\lambda_{i,t}$  is the expectation of  $Y_{i,t}$ , conditional to its past values.

The null hypothesis of the test is defined as  $H_0 : \alpha_0 = \alpha_{11} = \dots = \alpha_{2p} = 0$ , versus the alternative that at least one among  $\alpha_{s,h}$  is not 0, for  $s = 0, 1, 2$ . The test statistic has the form

$$LM(\gamma) = S'(\hat{\theta}, \gamma) \Sigma^{-1}(\hat{\theta}, \gamma) S(\hat{\theta}, \gamma)$$

where

$$S(\hat{\theta}, \gamma) = \sum_{t=1}^{TT} \sum_{i=1}^N \left( \frac{Y_{i,t}}{\lambda_{i,t}(\hat{\theta}, \gamma)} - 1 \right) \frac{\partial \lambda_{i,t}(\hat{\theta}, \gamma)}{\partial \alpha}$$

is the partition of the quasi score related to the vector of non-linear parameters  $\alpha = (\alpha_0, \dots, \alpha_{2p})$ , evaluated at the estimated parameters  $\hat{\theta}$  under the null assumption  $H_0$  (linear model) and  $\Sigma(\hat{\theta}, \gamma)$  is the variance of  $S(\hat{\theta}, \gamma)$ .

The optimization employs the Brent algorithm (Brent, 1973) applied in the interval from `gama_L` to `gama_U`. To be sure that the global optimum is found, the optimization is performed at `(len-1)` consecutive equidistant sub-intervals and then the maximum over them is taken as global optimum.

The values of `gama_L` and `gama_U` are computed internally as the mean over  $i = 1, \dots, N$  of 20% and 80% quantile of the empirical distribution of the network mean  $X_{i,t}$  for  $t = 1, \dots, TT$ . In this way the optimization is performed for values of  $\gamma$  such that the indicator function  $I(X_{i,t-d} \leq \gamma)$  is not always close to 0 or 1. Alternatively, their value can be supplied by the user. For details see Armillotta and Fokianos (2023, Sec. 4-5).

## Value

A list including:

<code>gama</code>	The optimum value of the $\gamma$ parameter.
<code>supLM</code>	The value of the objective function at the optimum.
<code>int</code>	A vector with the extremes points of sub-intervals.

## Author(s)

Mirko Armillotta, Michail Tsagris and Konstantinos Fokianos.

## References

- Armillotta, M. and K. Fokianos (2023). Nonlinear network autoregression. *Annals of Statistics*, 51(6): 2526–2552.
- Armillotta, M. and K. Fokianos (2024). Count network autoregression. *Journal of Time Series Analysis*, 45(4): 584–612.
- Armillotta, M., Tsagris, M. and Fokianos, K. (2024). Inference for Network Count Time Series with the R Package PNAR. *The R Journal*, 15/4: 255–269.
- Brent, R. (1973) Algorithms for Minimization without Derivatives. Prentice-Hall, Englewood Cliffs N.J.

## See Also

[score\\_test\\_tnarpq\\_j](#), [global\\_optimise\\_LM\\_stnarpq](#), [score\\_test\\_stnarpq\\_j](#)

**Examples**

```

data(crime)
data(crime_W)
mod1 <- lin_estimnarpq(crime, crime_W, p = 2)
b <- mod1$coefs[, 1]
global_optimise_LM_tnarpq(b = b, y = crime, W = crime_W, p = 2, d = 1)

```

---

lin_estimnarpq	<i>Estimation of the linear Poisson NAR(p) model model with p lags and q covariates (PNAR(p))</i>
----------------	---

---

**Description**

Estimation of the linear Poisson Network Autoregressive model of order  $p$  with  $q$  covariates (PNAR( $p$ )).

**Usage**

```

lin_estimnarpq(y, W, p, Z = NULL, uncon = FALSE, init = NULL,
xtol_rel = 1e-8, maxeval = 100)

```

**Arguments**

y	A $TT \times N$ time series object or a $TT \times N$ numerical matrix with the $N$ multivariate count time series over $TT$ time periods.
W	The $N \times N$ row-normalized non-negative adjacency matrix describing the network. The main diagonal entries of the matrix should be zeros, all the other entries should be non-negative and the maximum sum of elements over the rows should equal one. The function row-normalizes the matrix if a non-normalized adjacency matrix is provided.
p	The number of lags in the model.
Z	An $N \times q$ matrix of covariates (one for each column), where $q$ is the number of covariates in the model. Note that they must be non-negative.
uncon	logical, if TRUE an unconstrained optimization is run (default is FALSE).
init	A vector of starting values for the optimization algorithm. If this is NULL, the function computes them internally.
xtol_rel	The stopping tolerance of the optimization algorithm.
maxeval	The maximum number of evaluations the optimization algorithm will perform.

**Details**

This function performs constrained estimation of the linear Poisson NAR( $p$ ) model with  $q$  non-negative valued covariates, for each node of the network  $i = 1, \dots, N$  over the time sample  $t = 1, \dots, TT$ , defined as

$$\lambda_{i,t} = \beta_0 + \sum_{h=1}^p (\beta_{1h} X_{i,t-h} + \beta_{2h} Y_{i,t-h}) + \sum_{l=1}^q \delta_l Z_{i,l},$$

where  $X_{i,t} = \sum_{j=1}^N W_{ij} Y_{j,t}$  is the network effect, i.e. the weighted average impact of node  $i$  connections, with the weights of the mean being  $W_{ij}$ , the single element of the network matrix  $W$ . The sequence  $\lambda_{i,t}$  is the expectation of  $Y_{i,t}$ , conditional to its past values. The parameter  $\beta_0$  is the intercept of the model,  $\beta_{1h}$  are the network coefficients,  $\beta_{2h}$  are the autoregressive parameters, and  $\delta_l$  are the coefficients associated to the covariates  $Z_{i,l}$ .

The estimation of the parameters of the model is performed by Quasi Maximum Likelihood Estimation (QMLE), maximizing the following quasi log-likelihood

$$l(\theta) = \sum_{t=1}^{TT} \sum_{i=1}^N [Y_{i,t} \log \lambda_{i,t}(\theta) - \lambda_{i,t}(\theta)]$$

with respect to the vector of unknown parameters  $\theta$  described above. The coefficients are defined only in the non-negative real line.

By default, the optimization is constrained in the stationary region where  $\sum_{h=1}^p (\beta_{1h} + \beta_{2h}) < 1$ ; this can be removed by setting `uncons = TRUE`. However, the model estimates might be inconsistent if the estimated parameters lie outside the stationary region.

The ordinary least squares estimates are employed as starting values of the optimization procedure. Robust standard errors and z-tests are also returned.

## Value

A list with attribute class "PNAR" including:

<code>coefs</code>	A matrix with the estimated QMLE coefficients, their standard errors their Z-test statistics and the relevant p-values computed via the standard normal approximation.
<code>score</code>	The value of the quasi score function at the optimization point. It should be close to 0 if the optimization is successful.
<code>loglik</code>	The value of the maximized quasi log-likelihood.
<code>ic</code>	A vector with the Akaike information criterion (AIC), the Bayesian information criterion (BIC) and the Quasi information criterion (QIC).

Alternatively, these can be printed via the function `summary.PNAR`.

## Author(s)

Mirko Armillotta, Michail Tsagris and Konstantinos Fokianos.

## References

- Armillotta, M. and K. Fokianos (2023). Nonlinear network autoregression. *Annals of Statistics*, 51(6): 2526–2552.
- Armillotta, M. and K. Fokianos (2024). Count network autoregression. *Journal of Time Series Analysis*, 45(4): 584–612.
- Armillotta, M., Tsagris, M. and Fokianos, K. (2024). Inference for Network Count Time Series with the R Package PNAR. *The R Journal*, 15/4: 255–269..

**See Also**[log\\_lin\\_estimnarpq](#)**Examples**

```
data(crime)
data(crime_W)
mod1 <- lin_estimnarpq(crime, crime_W, p = 2)
summary(mod1)
```

---

lin_ic_plot	<i>Scatter plot of information criteria versus the number of lags in the linear Poisson NAR(<math>p</math>) model model with <math>p</math> lags and <math>q</math> covariates (PNAR(<math>p</math>))</i>
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---

**Description**

Scatter plot of information criteria versus the number of lags in the linear Poisson Network Autoregressive model of order  $p$  with  $q$  covariates (PNAR( $p$ )).

**Usage**

```
lin_ic_plot(y, W, p = 1:10, Z = NULL, unconstrained = FALSE, ic = "QIC")
```

**Arguments**

$y$	A $TT \times N$ time series object or a $TT \times N$ numerical matrix with the $N$ multivariate count time series over $TT$ time periods.
$W$	The $N \times N$ row-normalized non-negative adjacency matrix describing the network. The main diagonal entries of the matrix should be zeros, all the other entries should be non-negative and the maximum sum of elements over the rows should equal one. The function row-normalizes the matrix if a non-normalized adjacency matrix is provided.
$p$	A vector with integer numbers, the range of lags in the model, for which the AIC, BIC and QIC will be computed.
$Z$	An $N \times q$ matrix of covariates (one for each column), where $q$ is the number of covariates in the model. Note that they must be non-negative.
unconstrained	Logical, if TRUE an unconstrained optimization without stationarity constraints is performed (default is FALSE).
ic	The information criterion you want to plot, "QIC" (default value), "AIC" or "BIC".

**Details**

The function computes the AIC, BIC or QIC for a range of lag orders of the linear Poisson Network Autoregressive model of order  $p$  with  $q$  covariates (PNAR( $p$ )).

**Value**

A scatter plot with the lag order versus either QIC (default), AIC or BIC, and a vector with their values, for each lag order.

**Author(s)**

Mirko Armillotta, Michail Tsagris and Konstantinos Fokianos.

**References**

Armillotta, M. and K. Fokianos (2023). Nonlinear network autoregression. *Annals of Statistics*, 51(6): 2526–2552.

Armillotta, M. and K. Fokianos (2024). Count network autoregression. *Journal of Time Series Analysis*, 45(4): 584–612.

Armillotta, M., Tsagris, M. and Fokianos, K. (2024). Inference for Network Count Time Series with the R Package PNAR. *The R Journal*, 15/4: 255–269.

**See Also**

[lin\\_estimnarpq](#), [log\\_lin\\_ic\\_plot](#)

**Examples**

```
data(crime)
data(crime_W)
lin_ic_plot(crime, crime_W, p = 1:3)
```

---

lin_narpq_init	<i>Starting values for the linear Poisson NAR(<math>p</math>) model model with <math>p</math> lags and <math>q</math> covariates (PNAR(<math>p</math>))</i>
----------------	---

---

**Description**

Starting values for the linear Poisson Network Autoregressive model of order  $p$  with  $q$  covariates (PNAR( $p$ )).

**Usage**

```
lin_narpq_init(y, W, p, Z = NULL)
```



**Arguments**

y	A $TT \times N$ time series object or a $TT \times N$ numerical matrix with the $N$ multivariate count time series over $TT$ time periods.
W	The $N \times N$ row-normalized non-negative adjacency matrix describing the network. The main diagonal entries of the matrix should be zeros, all the other entries should be non-negative and the maximum sum of elements over the rows should equal one. The function row-normalizes the matrix if a non-normalized adjacency matrix is provided.
p	The number of lags in the model.
Z	An $N \times q$ matrix of covariates (one for each column), where $q$ is the number of covariates in the model. Note that they must be non-negative.

**Details**

The function computes starting values to be used in the function [lin\\_estimnarpq](#). These are simply the ordinary least squares estimators with a correction. If any of the the resulting coefficients is negative they become equal to 0.001.

**Value**

A vector with the initial values.

**Author(s)**

Mirko Armillotta, Michail Tsagris and Konstantinos Fokianos.

**References**

- Armillotta, M. and K. Fokianos (2023). Nonlinear network autoregression. *Annals of Statistics*, 51(6): 2526–2552.
- Armillotta, M. and K. Fokianos (2024). Count network autoregression. *Journal of Time Series Analysis*, 45(4): 584–612.
- Armillotta, M., Tsagris, M. and Fokianos, K. (2024). Inference for Network Count Time Series with the R Package PNAR. *The R Journal*, 15/4: 255–269.

**See Also**

[lin\\_estimnarpq](#)

**Examples**

```
data(crime)
data(crime_W)
x0 <- lin_narpq_init(crime, crime_W, p = 2)
```

---

log_lin_estimnarpq	<i>Estimation of the log-linear Poisson NAR(p) model with p lags and q covariates (log-PNAR(p))</i>
--------------------	---

---

### Description

Estimation of the log-linear Poisson Network Autoregressive model of order  $p$  with  $q$  covariates (log-PNAR( $p$ )).

### Usage

```
log_lin_estimnarpq(y, W, p, Z = NULL, unconstrained = FALSE, init = NULL,
xtol_rel = 1e-8, maxeval = 100)
```

### Arguments

y	A $TT \times N$ time series object or a $TT \times N$ numerical matrix with the $N$ multivariate count time series over $TT$ time periods.
W	The $N \times N$ row-normalized non-negative adjacency matrix describing the network. The main diagonal entries of the matrix should be zeros, all the other entries should be non-negative and the maximum sum of elements over the rows should equal one. The function row-normalizes the matrix if a non-normalized adjacency matrix is provided.
p	The number of lags in the model.
Z	An $N \times q$ matrix of covariates (one for each column), where $q$ is the number of covariates in the model.
unconstrained	logical, if TRUE an unconstrained optimization is performed (default is FALSE).
init	A vector of starting values for the optimization algorithm. If this is NULL, the function computes them internally.
xtol_rel	The stopping tolerance of the optimization algorithm.
maxeval	The maximum number of evaluations the optimization algorithm will perform.

### Details

This function performs a constrained estimation of the linear Poisson NAR( $p$ ) model with  $q$  non-negative valued covariates, for each node of the network  $i = 1, \dots, N$  over the time sample  $t = 1, \dots, TT$ , defined as

$$\nu_{i,t} = \beta_0 + \sum_{h=1}^p (\beta_{1h} X_{i,t-h} + \beta_{2h} Y_{i,t-h}) + \sum_{l=1}^q \delta_l Z_{i,l},$$

where  $X_{i,t} = \sum_{j=1}^N W_{ij} Y_{j,t}$  is the network effect, i.e. the weighted average impact of node  $i$  connections, with the weights of the mean being  $W_{ij}$ , the single element of the network matrix  $W$ . The sequence  $\nu_{i,t}$  is the log of the expectation of  $Y_{i,t}$ , conditional to its past values. The parameter  $\beta_0$  is

the intercept of the model,  $\beta_{1h}$  are the network coefficients,  $\beta_{2h}$  are the autoregressive parameters, and  $\delta_l$  are the coefficients associated to the covariates  $Z_{i,l}$ .

The estimation of the parameters of the model is performed by Quasi Maximum Likelihood Estimation (QMLE), maximizing the following quasi log-likelihood

$$l(\theta) = \sum_{t=1}^{TT} \sum_{i=1}^N \left[ Y_{i,t} \nu_{i,t}(\theta) - e^{\nu_{i,t}(\theta)} \right]$$

with respect to the vector of unknown parameters  $\theta$  described above.

By default, the optimization is constrained in the stationary region where  $\sum_{h=1}^p (|\beta_{1h}| + |\beta_{2h}|) < 1$ ; this can be removed by setting `uncons = TRUE`. However, the model estimates might be inconsistent if the estimated parameters lie outside the stationary region.

The ordinary least squares estimates are employed as starting values of the optimization procedure. Robust standard errors and z-tests are also returned.

## Value

A list with attribute class "PNAR" including:

<code>coefs</code>	A matrix with the estimated QMLE coefficients, their standard errors, their Z-test statistics and the relevant p-values computed via the standard normal approximation.
<code>score</code>	The value of the quasi score function at the optimization point. It should be close to 0 if the optimization is successful.
<code>loglik</code>	The value of the maximized quasi log-likelihood.
<code>ic</code>	A vector with the Akaike information criterion (AIC), the Bayesian information criterion (BIC) and the Quasi information criterion (QIC).

Alternatively, these can be printed via the function `summary.PNAR`.

## Author(s)

Mirko Armillotta, Michail Tsagris and Konstantinos Fokianos.

## References

- Armillotta, M. and K. Fokianos (2023). Nonlinear network autoregression. *Annals of Statistics*, 51(6): 2526–2552.
- Armillotta, M. and K. Fokianos (2024). Count network autoregression. *Journal of Time Series Analysis*, 45(4): 584–612.
- Armillotta, M., Tsagris, M. and Fokianos, K. (2024). Inference for Network Count Time Series with the R Package PNAR. *The R Journal*, 15/4: 255–269.

## See Also

[lin\\_estimnarpq](#)

**Examples**

```
data(crime)
data(crime_W)
mod1 <- log_lin_estimnarpq(crime, crime_W, p = 2)
summary(mod1)
```

---

log_lin_ic_plot	<i>Scatter plot of information criteria versus the number of lags in the log-linear Poisson NAR(<math>p</math>) model with <math>p</math> lags and <math>q</math> covariates (log-PNAR(<math>p</math>))</i>
-----------------	---

---

**Description**

Scatter plot of information criteria versus the number of lags in log-linear Poisson Network Autoregressive model of order  $p$  with  $q$  covariates (log-PNAR( $p$ )).

**Usage**

```
log_lin_ic_plot(y, W, p = 1:10, Z = NULL, uncons = FALSE, ic = "QIC")
```

**Arguments**

$y$	A $TT \times N$ time series object or a $TT \times N$ numerical matrix with the $N$ multivariate count time series over $TT$ time periods.
$W$	The $N \times N$ row-normalized non-negative adjacency matrix describing the network. The main diagonal entries of the matrix should be zeros, all the other entries should be non-negative and the maximum sum of elements over the rows should equal one. The function row-normalizes the matrix if a non-normalized adjacency matrix is provided.
$p$	A vector with integer numbers, the range of lags in the model, for which the AIC, BIC and QIC will be computed.
$Z$	An $N \times q$ matrix of covariates (one for each column), where $q$ is the number of covariates in the model. Note that they must be non-negative.
uncons	Logical, if TRUE an unconstrained optimization without stationarity constraints is performed (default is FALSE).
ic	The information criterion you want to plot, "QIC" (default value), "AIC" or "BIC".

**Details**

The function computes the AIC, BIC or QIC for a range of lag orders of the log-linear Poisson Network Autoregressive model of order  $p$  with  $q$  covariates (PNAR( $p$ )).

**Value**

A scatter plot with the lag order versus either QIC (default), AIC or BIC, and a vector with their values, for each lag order.

**Author(s)**

Mirko Armillotta, Michail Tsagris and Konstantinos Fokianos.

**References**

- Armillotta, M. and K. Fokianos (2023). Nonlinear network autoregression. *Annals of Statistics*, 51(6): 2526–2552.
- Armillotta, M. and K. Fokianos (2024). Count network autoregression. *Journal of Time Series Analysis*, 45(4): 584–612.
- Armillotta, M., Tsagris, M. and Fokianos, K. (2024). Inference for Network Count Time Series with the R Package PNAR. *The R Journal*, 15/4: 255–269.

**See Also**

[log\\_lin\\_estimnarpq](#), [lin\\_ic\\_plot](#)

**Examples**

```
data(crime)
data(crime_W)
log_lin_ic_plot(crime, crime_W, p = 1:3)
```

---

log_lin_narpq_init	<i>Starting values for the log-linear Poisson NAR(p) model with p lags and q covariates (log-PNAR(p))</i>
--------------------	---

---

**Description**

Starting values for the log-linear Poisson Network Autoregressive model of order  $p$  with  $q$  covariates (log-PNAR( $p$ )).

**Usage**

```
log_lin_narpq_init(y, W, p, Z = NULL)
```

**Arguments**

- |   |  |
|---|--|
| y | A $TT \times N$ time series object or a $TT \times N$ numerical matrix with the $N$ multivariate count time series over $TT$ time periods.   |
| W | The $N \times N$ row-normalized non-negative adjacency matrix describing the network. The main diagonal entries of the matrix should be zeros, all the other entries should be non-negative and the maximum sum of elements over the rows should equal one. The function row-normalizes the matrix if a non-normalized adjacency matrix is provided. |
| p | The number of lags in the model.   |
| Z | An $N \times q$ matrix of covariates (one for each column), where $q$ is the number of covariates in the model.  |

**Details**

This function computes initial values for the log-linear Poisson Network Autoregressive model of order  $p$  with  $q$  covariates (log-PNAR( $p$ )) with stationarity conditions. These initial values are simply the ordinary least squares estimators with a correction.

**Value**

A vector with the initial values.

**Author(s)**

Mirko Armillotta, Michail Tsagris and Konstantinos Fokianos.

**References**

Armillotta, M. and K. Fokianos (2023). Nonlinear network autoregression. *Annals of Statistics*, 51(6): 2526–2552.

Armillotta, M. and K. Fokianos (2024). Count network autoregression. *Journal of Time Series Analysis*, 45(4): 584–612.

Armillotta, M., Tsagris, M. and Fokianos, K. (2024). Inference for Network Count Time Series with the R Package PNAR. *The R Journal*, 15/4: 255–269.

**See Also**

[log\\_lin\\_estimnarpq](#)

**Examples**

```
data(crime)
data(crime_W)
mod1 <- log_lin_narpq_init(crime, crime_W, p = 2)
```

---

poisson.MODpq

*Generation of counts from a linear Poisson NAR( $p$ ) model with  $q$  covariates (PNAR( $p$ ))*

---

**Description**

Generation of multivariate count time series from a linear Poisson Network Autoregressive model of order  $p$  with  $q$  covariates (PNAR( $p$ )).

**Usage**

```
poisson.MODpq(b, W, p, Z = NULL, TT, N, copula = "gaussian",
  corrtype = "equicorrelation", rho, dof = 1)
```

### Arguments

b	The coefficients of the model, in the following order: (intercept, network parameters, autoregressive parameters, covariates). The dimension of the vector should be $2p + 1 + q$ , where $q$ denotes the number of covariates.
W	The $N \times N$ row-normalized non-negative adjacency matrix describing the network. The main diagonal entries of the matrix should be zeros, all the other entries should be non-negative and the maximum sum of elements over the rows should equal one. The function row-normalizes the matrix if a non-normalized adjacency matrix is provided.
p	The number of lags in the model.
Z	An $N \times q$ matrix of covariates (one for each column), where $q$ is the number of covariates in the model. Note that they must be non-negative.
TT	The temporal sample size.
N	The number of nodes on the network.
copula	Which copula function to use? The choices are "gaussian", "t", or "clayton".
rho	The value of the copula parameter ( $\rho$ ). A scalar in $[-1, 1]$ for elliptical copulas (Gaussian, t), a value greater than or equal to -1 for Clayton copula.
corrtype	Used only for elliptical copulas. The type of correlation matrix employed for the copula; it will either be the "equicorrelation" or "toeplitz". The "equicorrelation" option generates a correlation matrix where all the off-diagonal entries equal $\rho$ . The "toeplitz" option generates a correlation matrix whose generic off-diagonal $(i, j)$ -element is $\rho^{ i-j }$ .
dof	The degrees of freedom for Student's t copula.

### Details

This function generates counts from a linear Poisson NAR( $p$ ) model, where  $q$  non time-varying covariates are allowed as well. The counts are simulated from  $Y_t = N_t(\lambda_t)$ , where  $N_t$  is a sequence of  $N$ -dimensional IID Poisson count processes, with intensity 1, and whose structure of dependence is modelled through a copula construction  $C(\rho)$  on their associated exponential waiting times random variables. For details see Armillotta and Fokianos (2024, Sec. 2.1-2.2).

The sequence  $\lambda_{i,t}$  is the expectation of  $Y_{i,t}$ , conditional to its past values and it is generated by means of the following PNAR( $p$ ) model. For each node of the network  $i = 1, \dots, N$  over the time sample  $t = 1, \dots, TT$

$$\lambda_{i,t} = \beta_0 + \sum_{h=1}^p (\beta_{1h} X_{i,t-h} + \beta_{2h} Y_{i,t-h}) + \sum_{l=1}^q \delta_l Z_{i,l}$$

where  $X_{i,t} = \sum_{j=1}^N W_{ij} Y_{j,t}$  is the network effect, i.e. the weighted average impact of node  $i$  connections, with the weights of the mean being  $W_{ij}$ , the single element of the network matrix  $W$ . The parameter  $\beta_0$  is the intercept of the model,  $\beta_{1h}$  are the network coefficients,  $\beta_{2h}$  are the autoregressive parameters, and  $\delta_l$  are the coefficients associated to the covariates  $Z_{i,l}$ .

**Value**

A list including:

p2R	The Toeplitz correlation matrix, if employed in the copula or NULL else.
lambda	A $TT \times N$ time series object matrix of simulated Poisson means for $N$ time series over $TT$ .
y	A $TT \times N$ time series object matrix of simulated counts for $N$ time series over $TT$ .

**Author(s)**

Mirko Armillotta, Michail Tsagris and Konstantinos Fokianos.

**References**

- Armillotta, M. and K. Fokianos (2023). Nonlinear network autoregression. *Annals of Statistics*, 51(6): 2526–2552.
- Armillotta, M. and K. Fokianos (2024). Count network autoregression. *Journal of Time Series Analysis*, 45(4): 584–612.
- Armillotta, M., Tsagris, M. and Fokianos, K. (2024). Inference for Network Count Time Series with the R Package PNAR. *The R Journal*, 15/4: 255–269.
- Fokianos, K., Stove, B., Tjostheim, D., and P. Doukhan (2020). Multivariate count autoregression. *Bernoulli*, 26(1), 471–499.

**See Also**

[poisson.MODpq.log](#), [poisson.MODpq.nonlin](#), [poisson.MODpq.stnar](#), [poisson.MODpq.tnar](#)

**Examples**

```
W <- adja( N = 20, K = 5, alpha= 0.5)
y <- poisson.MODpq( b = c(0.5, 0.3, 0.2), W = W, p = 1, Z = NULL,
TT = 1000, N = 20, copula = "gaussian",
corrtype = "equicorrelation", rho = 0.5)$y
```

---

poisson.MODpq.log	<i>Generation of multivariate count time series from a log-linear Poisson NAR(p) model with q covariates (log-PNAR(p))</i>
-------------------	--

---

**Description**

Generation of counts from a log-linear Poisson Network Autoregressive model of order  $p$  with  $q$  covariates (log-PNAR( $p$ )).



**Usage**

```
poisson.MODpq.log(b, W, p, Z = NULL, TT, N, copula = "gaussian",
  corrtype = "equicorrelation", rho, dof = 1)
```

**Arguments**

b	The coefficients of the model, in the following order: (intercept, network parameters, autoregressive parameters, covariates). The dimension of the vector should be $2p + 1 + q$ , where $q$ denotes the number of covariates.
W	The $N \times N$ row-normalized non-negative adjacency matrix describing the network. The main diagonal entries of the matrix should be zeros, all the other entries should be non-negative and the maximum sum of elements over the rows should equal one. The function row-normalizes the matrix if a non-normalized adjacency matrix is provided.
p	The number of lags in the model.
Z	An $N \times q$ matrix of covariates (one for each column), where $q$ is the number of covariates in the model.
TT	The temporal sample size.
N	The number of nodes on the network.
copula	Which copula function to use? The "gaussian", "t", or "clayton".
rho	The value of the copula parameter ( $\rho$ ). A scalar in $[-1, 1]$ for elliptical copulas (Gaussian, t), a value greater or equal to -1 for Clayton copula.
corrtype	Used only for elliptical copulas. The type of correlation matrix employed for the copula; it will either be the "equicorrelation" or "toeplitz". The "equicorrelation" option generates a correlation matrix where all the off-diagonal entries equal $\rho$ . The "toeplitz" option generates a correlation matrix whose generic off-diagonal $(i, j)$ -element is $\rho^{ i-j }$ .
dof	The degrees of freedom for Student's t copula.

**Details**

This function generates counts from a log-linear Poisson NAR( $p$ ) model, where  $q$  non time-varying covariates are allowed as well. The counts are simulated from  $Y_t = N_t(e^{\nu_t})$ , where  $N_t$  is a sequence of  $N$ -dimensional IID Poisson count processes, with intensity 1, and whose structure of dependence is modelled through a copula construction  $C(\rho)$  on their associated exponential waiting times random variables. For details see Armillotta and Fokianos (2024, Sec. 2.1-2.2).

The sequence  $\nu_t$  is the log of the expectation of  $Y_t$ , conditional to its past values and it is generated by means of the following log-PNAR( $p$ ) model. For each node of the network  $i = 1, \dots, N$  over the time sample  $t = 1, \dots, TT$

$$\nu_{i,t} = \beta_0 + \sum_{h=1}^p (\beta_{1h} X_{i,t-h} + \beta_{2h} Y_{i,t-h}) + \sum_{l=1}^q \delta_l Z_{i,l}$$

where  $X_{i,t} = \sum_{j=1}^N W_{ij} Y_{j,t}$  is the network effect, i.e. the weighted average impact of node  $i$  connections, with the weights of the mean being  $W_{ij}$ , the single element of the network matrix  $W$ . The parameter  $\beta_0$  is the intercept of the model,  $\beta_{1h}$  are the network coefficients,  $\beta_{2h}$  are the autoregressive parameters, and  $\delta_l$  are the coefficients associated to the covariates  $Z_{i,l}$ .

**Value**

A list including:

p2R	The Toeplitz correlation matrix, if employed in the copula or NULL else.
log_lambda	A $TT \times N$ time series object matrix of simulated Poisson log-means for $N$ time series over $TT$ .
y	A $TT \times N$ time series object matrix of simulated counts for $N$ time series over $TT$ .

**Author(s)**

Mirko Armillotta, Michail Tsagris and Konstantinos Fokianos.

**References**

- Armillotta, M. and K. Fokianos (2023). Nonlinear network autoregression. *Annals of Statistics*, 51(6): 2526–2552.
- Armillotta, M. and K. Fokianos (2024). Count network autoregression. *Journal of Time Series Analysis*, 45(4): 584–612.
- Armillotta, M., Tsagris, M. and Fokianos, K. (2024). Inference for Network Count Time Series with the R Package PNAR. *The R Journal*, 15/4: 255–269.
- Fokianos, K., Stove, B., Tjostheim, D., and P. Doukhan (2020). Multivariate count autoregression. *Bernoulli*, 26(1), 471–499.

**See Also**

[poisson.MODpq](#), [poisson.MODpq.nonlin](#), [poisson.MODpq.stnar](#), [poisson.MODpq.tnar](#)

**Examples**

```
W <- adja( N = 20, K = 5, alpha= 0.5)
y <- poisson.MODpq.log( b = c(0.5, 0.3, 0.2), W = W, p = 1,
Z = NULL, TT = 1000, N = 20, copula = "gaussian",
corrtype = "equicorrelation", rho = 0.5)$y
```

---

poisson.MODpq.nonlin	<i>Generation of multivariate count time series from a non-linear Intercept Drift Poisson NAR(p) model with q covariates (ID-PNAR(p))</i>
----------------------	---

---

**Description**

Generation of counts from a non-linear Intercept Drift Poisson Network Autoregressive model of order  $p$  with  $q$  covariates (ID-PNAR( $p$ )).

**Usage**

```
poisson.MODpq.nonlin(b, W, gama, p, d, Z = NULL, TT, N, copula = "gaussian",
  corrtype = "equicorrelation", rho, dof = 1)
```

**Arguments**

b	The linear coefficients of the model, in the following order: (intercept, network parameters, autoregressive parameters, covariates). The dimension of the vector should be $2p + 1 + q$ , where $q$ denotes the number of covariates.
W	The $N \times N$ row-normalized non-negative adjacency matrix describing the network. The main diagonal entries of the matrix should be zeros, all the other entries should be non-negative and the maximum sum of elements over the rows should equal one. The function row-normalizes the matrix if a non-normalized adjacency matrix is provided.
gama	A scalar non-linear intercept drift parameter.
p	The number of lags in the model.
d	The lag parameter of non-linear variable (should be between 1 and $p$ ).
Z	An $N \times q$ matrix of covariates (one for each column), where $q$ is the number of covariates in the model. Note that they must be non-negative.
TT	The temporal sample size.
N	The number of nodes on the network.
copula	Which copula function to use? The "gaussian", "t", or "clayton".
rho	The value of the copula parameter ( $\rho$ ). A scalar in $[-1, 1]$ for elliptical copulas (Gaussian, t), a value greater than or equal to -1 for Clayton copula.
corrtype	Used only for elliptical copulas. The type of correlation matrix employed for the copula; it will either be the "equicorrelation" or "toeplitz". The "equicorrelation" option generates a correlation matrix where all the off-diagonal entries equal $\rho$ . The "toeplitz" option generates a correlation matrix whose generic off-diagonal $(i, j)$ -element is $\rho^{ i-j }$ .
dof	The degrees of freedom for Student's t copula.

**Details**

This function generates counts from a non-linear Intercept Drift Poisson NAR( $p$ ) model, where  $q$  non time-varying covariates are allowed as well. The counts are simulated from  $Y_t = N_t(\lambda_t)$ , where  $N_t$  is a sequence of  $N$ -dimensional IID Poisson count processes, with intensity 1, and whose structure of dependence is modelled through a copula construction  $C(\rho)$  on their associated exponential waiting times random variables. For details see Armillotta and Fokianos (2024, Sec. 2.1-2.2). The sequence  $\lambda_{i,t}$  is the expectation of  $Y_{i,t}$ , conditional to its past values and it is generated by means of the following ID-PNAR( $p$ ) model. For each node of the network  $i = 1, \dots, N$  over the time sample  $t = 1, \dots, TT$

$$\lambda_{i,t} = \frac{\beta_0}{(1 + X_{i,t-d})^\gamma} + \sum_{h=1}^p (\beta_{1h} X_{i,t-h} + \beta_{2h} Y_{i,t-h}) + \sum_{l=1}^q \delta_l Z_{i,l}$$

where  $X_{i,t} = \sum_{j=1}^N W_{ij} Y_{j,t}$  is the network effect, i.e. the weighted average impact of node  $i$  connections, with the weights of the mean being  $W_{ij}$ , the single element of the network matrix  $W$ .

The parameter  $\beta_0$  is the intercept of the model,  $\beta_{1h}$  are the network coefficients,  $\beta_{2h}$  are the autoregressive parameters,  $\gamma$  is the non-linear coefficient associated with the intercept drift, and  $\delta_l$  are the coefficients associated with the covariates  $Z_{i,l}$ . The coefficient  $d$  is considered as an extra parameter defining the lag of the network effect in the non-linear part of the model and is left to be set by the user. For details on ID-PNAR models see Armillotta and Fokianos (2023, Sec. 2).

## Value

A list including:

p2R	The Toeplitz correlation matrix, if employed in the copula or NULL else.
lambda	A $TT \times N$ time series object matrix of simulated Poisson means for $N$ time series over $TT$ .
y	A $TT \times N$ time series object matrix of simulated counts for $N$ time series over $TT$ .

## Author(s)

Mirko Armillotta, Michail Tsagris and Konstantinos Fokianos.

## References

- Armillotta, M. and K. Fokianos (2023). Nonlinear network autoregression. *Annals of Statistics*, 51(6): 2526–2552.
- Armillotta, M. and K. Fokianos (2024). Count network autoregression. *Journal of Time Series Analysis*, 45(4): 584–612.
- Armillotta, M., Tsagris, M. and Fokianos, K. (2024). Inference for Network Count Time Series with the R Package PNAR. *The R Journal*, 15/4: 255–269.

## See Also

[poisson.MODpq](#), [poisson.MODpq.log](#), [poisson.MODpq.stnar](#), [poisson.MODpq.tnar](#)

## Examples

```
W <- adja( N = 20, K = 5, alpha= 0.5)
y <- poisson.MODpq.nonlin( b = c(0.5, 0.3, 0.2), W = W, gama = 1, p = 1,
d = 1, Z = NULL, TT = 1000, N = 20, copula = "gaussian",
corrtype = "equicorrelation", rho = 0.5)$y
```

---

poisson.MODpq.stnar	<i>Generation of counts from a non-linear Smooth Transition Poisson NAR(<math>p</math>) model with <math>q</math> covariates (ST-PNAR(<math>p</math>))</i>
---------------------	--

---

## Description

Generation of multivariate count time series from a non-linear Smooth Transition Poisson Network Autoregressive model of order  $p$  with  $q$  covariates (ST-PNAR( $p$ )).

## Usage

```
poisson.MODpq.stnar(b, W, gama, a, p, d, Z = NULL, TT, N, copula = "gaussian",
  corrtype = "equicorrelation", rho, dof = 1)
```

## Arguments

b	The linear coefficients of the model, in the following order: (intercept, network parameters, autoregressive parameters, covariates). The dimension of the vector should be $2p + 1 + q$ , where $q$ denotes the number of covariates.
W	The $N \times N$ row-normalized non-negative adjacency matrix describing the network. The main diagonal entries of the matrix should be zeros, all the other entries should be non-negative and the maximum sum of elements over the rows should equal one. The function row-normalizes the matrix if a non-normalized adjacency matrix is provided.
gama	The scalar nuisance smoothing parameter.
a	Vector of non-linear parameters. The dimension of the vector should be $p$ .
p	The number of lags in the model.
d	The lag parameter of non-linear variable (should be between 1 and $p$ ).
Z	An $N \times q$ matrix of covariates (one for each column), where $q$ is the number of covariates in the model. Note that they must be non-negative.
TT	The temporal sample size.
N	The number of nodes on the network.
copula	Which copula function to use? The choices are "gaussian", "t", or "clayton".
rho	The value of the copula parameter ( $\rho$ ). A scalar in $[-1, 1]$ for elliptical copulas (Gaussian, t), a value greater than or equal to -1 for Clayton copula.
corrtype	Used only for elliptical copulas. The type of correlation matrix employed for the copula; it will either be the "equicorrelation" or "toeplitz". The "equicorrelation" option generates a correlation matrix where all the off-diagonal entries equal $\rho$ . The "toeplitz" option generates a correlation matrix whose generic off-diagonal $(i, j)$ -element is $\rho^{ i-j }$ .
dof	The degrees of freedom for Student's t copula.

## Details

This function generates counts from a non-linear Smooth Transition Poisson NAR( $p$ ) model, where  $q$  non time-varying covariates are allowed as well. The counts are simulated from  $Y_t = N_t(\lambda_t)$ , where  $N_t$  is a sequence of  $N$ -dimensional IID Poisson count processes, with intensity 1, and whose structure of dependence is modelled through a copula construction  $C(\rho)$  on their associated exponential waiting times random variables. For details see Armillotta and Fokianos (2024, Sec. 2.1-2.2).

The sequence  $\lambda_{i,t}$  is the expectation of  $Y_{i,t}$ , conditional to its past values and it is generated by means of the following ST-PNAR( $p$ ) model. For each node of the network  $i = 1, \dots, N$  over the time sample  $t = 1, \dots, TT$

$$\lambda_{i,t} = \beta_0 + \sum_{h=1}^p (\beta_{1h} X_{i,t-h} + \beta_{2h} Y_{i,t-h} + \alpha_h e^{-\gamma X_{i,t-h}^2} X_{i,t-h}) + \sum_{l=1}^q \delta_l Z_{i,l}$$

where  $X_{i,t} = \sum_{j=1}^N W_{ij} Y_{j,t}$  is the network effect, i.e. the weighted average impact of node  $i$  connections, with the weights of the mean being  $W_{ij}$ , the single element of the network matrix  $W$ .

The parameter  $\beta_0$  is the intercept of the model,  $\beta_{1h}$  are the network coefficients,  $\beta_{2h}$  are the autoregressive parameters,  $\alpha_h$  are the non-linear smooth transition parameters,  $\gamma$  is the nuisance smoothing parameter, and  $\delta_l$  are the coefficients associated to the covariates  $Z_{i,l}$ . The coefficient  $d$  is considered as an extra parameter defining the lag of the network effect in the non-linear part of the model and is left to be set by the user. For details on ST-PNAR models see Armillotta and Fokianos (2023, Sec. 2).

## Value

A list including:

p2R	The Toeplitz correlation matrix, if employed in the copula or NULL else.
lambda	A $TT \times N$ time series object matrix of simulated Poisson means for $N$ time series over $TT$ .
y	A $TT \times N$ time series object matrix of simulated counts for $N$ time series over $TT$ .

## Author(s)

Mirko Armillotta, Michail Tsagris and Konstantinos Fokianos.

## References

- Armillotta, M. and K. Fokianos (2023). Nonlinear network autoregression. *Annals of Statistics*, 51(6): 2526–2552.
- Armillotta, M. and K. Fokianos (2024). Count network autoregression. *Journal of Time Series Analysis*, 45(4): 584–612.
- Armillotta, M., Tsagris, M. and Fokianos, K. (2024). Inference for Network Count Time Series with the R Package PNAR. *The R Journal*, 15/4: 255–269.

**See Also**

[poisson.MODpq](#), [poisson.MODpq.log](#), [poisson.MODpq.nonlin](#), [poisson.MODpq.tnar](#)

**Examples**

```
W <- adja( N = 20, K = 5, alpha= 0.5)
y <- poisson.MODpq.stnar( b = c(0.5, 0.3, 0.2), W = W, gama = 0.2, a = 0.4,
p = 1, d = 1, Z = NULL, TT = 1000, N = 20, copula = "gaussian",
corrtype = "equicorrelation", rho = 0.5)$y
```

---

poisson.MODpq.tnar	<i>Generation of counts from a non-linear Threshold Poisson NAR(<math>p</math>) model with <math>q</math> covariates (T-PNAR(<math>p</math>))</i>
--------------------	---

---

**Description**

Generation of multivariate count time series from a non-linear Threshold Poisson network Autoregressive model of order  $p$  with  $q$  covariates (T-PNAR( $p$ )).

**Usage**

```
poisson.MODpq.tnar(b, W, gama, a, p, d, Z = NULL, TT, N, copula = "gaussian",
corrtype = "equicorrelation", rho, dof = 1)
```

**Arguments**

b	The linear coefficients of the model, in the following order: (intercept, network parameters, autoregressive parameters, covariates). The dimension of the vector should be $2p + 1 + q$ , where $q$ denotes the number of covariates.
W	The $N \times N$ row-normalized non-negative adjacency matrix describing the network. The main diagonal entries of the matrix should be zeros, all the other entries should be non-negative and the maximum sum of elements over the rows should equal one. The function row-normalizes the matrix if a non-normalized adjacency matrix is provided.
gama	The scalar nuisance threshold parameter.
a	Vector of non-linear parameters. The dimension of the vector should be $2p + 1$ .
p	The number of lags in the model.
d	The lag parameter of non-linear variable (should be between 1 and $p$ ).
Z	An $N \times q$ matrix of covariates (one for each column), where $q$ is the number of covariates in the model. Note that they must be non-negative.
TT	The temporal sample size.
N	The number of nodes on the network.
copula	Which copula function to use? The "gaussian", "t", or "clayton".

rho	The value of the copula parameter ( $\rho$ ). A scalar in $[-1, 1]$ for elliptical copulas (Gaussian, t), a value greater than or equal to -1 for Clayton copula.
corrtype	Used only for elliptical copulas. The type of correlation matrix employed for the copula; it will either be the "equicorrelation" or "toeplitz". The "equicorrelation" option generates a correlation matrix where all the off-diagonal entries equal $\rho$ . The "toeplitz" option generates a correlation matrix whose generic off-diagonal $(i, j)$ -element is $\rho^{ i-j }$ .
dof	The degrees of freedom for Student's t copula.

### Details

This function generates counts from a non-linear Threshold Poisson NAR( $p$ ) model, where  $q$  non time-varying covariates are allowed as well. The counts are simulated from  $Y_t = N_t(\lambda_t)$ , where  $N_t$  is a sequence of  $N$ -dimensional IID Poisson count processes, with intensity 1, and whose structure of dependence is modelled through a copula construction  $C(\rho)$  on their associated exponential waiting times random variables. For details see Armillotta and Fokianos (2024, Sec. 2.1-2.2).

The sequence  $\lambda_{i,t}$  is the expectation of  $Y_{i,t}$ , conditional to its past values and it is generated by means of the following T-PNAR( $p$ ) model. For each node of the network  $i = 1, \dots, N$  over the time sample  $t = 1, \dots, TT$

$$\lambda_{i,t} = \beta_0 + \sum_{h=1}^p [\beta_{1h} X_{i,t-h} + \beta_{2h} Y_{i,t-h} + (\alpha_0 + \alpha_{1h} X_{i,t-h} + \alpha_{2h} Y_{i,t-h}) I(X_{i,t-d} \leq \gamma)] + \sum_{l=1}^q \delta_l Z_{i,l}$$

where  $X_{i,t} = \sum_{j=1}^N W_{ij} Y_{j,t}$  is the network effect, i.e. the weighted average impact of node  $i$  connections, with the weights of the mean being  $W_{ij}$ , the single element of the network matrix  $W$ , and  $I(\cdot)$  is the indicator function.

The parameter  $\beta_0$  is the intercept of the model,  $\beta_{1h}$  are the network coefficients,  $\beta_{2h}$  are the autoregressive parameters, the  $\alpha$  vector of non-linear parameters is divided as follows:  $\alpha_0$  is the intercept,  $\alpha_{1h}$  are the network coefficients,  $\alpha_{2h}$  are the autoregressive parameters;  $\gamma$  is the nuisance threshold parameter, and  $\delta_l$  are the coefficients associated to the covariates  $Z_{i,l}$ . The coefficient  $d$  is considered as an extra parameter defining the lag of the network effect in the non-linear part of the model and is left to be set by the user. For details on T-PNAR models see Armillotta and Fokianos (2023, Sec. 2).

### Value

A list including:

p2R	The Toeplitz correlation matrix, if employed in the copula or NULL else.
lambda	A $TT \times N$ time series object matrix of simulated Poisson means for $N$ time series over $TT$ .
y	A $TT \times N$ time series object matrix of simulated counts for $N$ time series over $TT$ .

### Author(s)

Mirko Armillotta, Michail Tsagris and Konstantinos Fokianos.



References

Armiliotta, M. and K. Fokianos (2023). Nonlinear network autoregression. *Annals of Statistics*, 51(6): 2526–2552.

Armiliotta, M. and K. Fokianos (2024). Count network autoregression. *Journal of Time Series Analysis*, 45(4): 584–612.

Armiliotta, M., Tsagris, M. and Fokianos, K. (2024). Inference for Network Count Time Series with the R Package PNAR. *The R Journal*, 15/4: 255–269.

See Also

```
poisson.MODpq, poisson.MODpq.log, poisson.MODpq.nonlin, poisson.MODpq.stnar
```

Examples

```
W <- adja( N = 20, K = 5, alpha= 0.5)
y <- poisson.MODpq.tnar( b = c(0.5, 0.3, 0.2), W = W, gama = 1,
a = c(0.2, 0.2, 0.2), p = 1, d = 1, Z = NULL, TT = 1000, N = 20,
copula = "gaussian", corrtype = "equicorrelation", rho = 0.5)$y
```

---

rcopula	<i>Random number generation of copula functions</i>
---------	---

---

Description

Random number generation of copula functions.

Usage

```
rcopula(n, N, copula = "gaussian", corrtype = "equicorrelation",
rho, dof = 1, cholR = NULL)
```

Arguments

n	The number of random values to generate.
N	The number of variables for which random valeus will be generated.
copula	Which copula function to use? The "gaussian", "t", or "clayton".
rho	The the value of the copula parameter ( $\rho$ ). A scalar in $[-1, 1]$ for elliptical copulas (Gaussian, t), a value greater than or equal to -1 for Clayton copula.
corrtype	Used only for elliptical copulas. The type of correlation matrix employed for the copula; it will either be the "equicorrelation" or "toeplitz". The "equicorrelation" option generates a correlation matrix where all the off-diagonal entries equal $\rho$ . The "toeplitz" option generates a correlation matrix whose generic off-diagonal $(i, j)$ -element is $\rho^{ i-j }$ .
dof	The degrees of freedom for Student's t copula.
cholR	An alternative input for elliptical copulas, providing directly the Cholesky decomposition for a specific correlation matrix to be passed, otherwise leave it NULL.

**Details**

This function generates random copula values from Gaussian, Student's t, or Clayton copulas based on a single copula parameter and different correlation structures.

**Value**

An  $n \times N$  matrix with the simulated copula values.

**Author(s)**

Mirko Armillotta, Michail Tsagris and Konstantinos Fokianos.

**References**

Nelsen, Roger B. (1999). *An Introduction to Copulas*, Springer.

**See Also**

[getN](#), [poisson.MODpq](#), [poisson.MODpq.log](#)

**Examples**

```
u <- rcopula(n = 100, N = 50, rho = 0.3)
```

---

score\_test\_nonlinpq\_h0

*Linearity test against non-linear ID-PNAR( $p$ ) model*

---

**Description**

Quasi score test for testing linearity of Poisson Network Autoregressive model of order  $p$  against the non-linear Intercep Drift (ID) version (ID-PNAR( $p$ )).

**Usage**

```
score_test_nonlinpq_h0(b, y, W, p, d, Z = NULL)
```

**Arguments**

- |   |   |
|---|---|
| b | The estimated parameters from the linear PNAR model, in the following order: (intercept, network parameters, autoregressive parameters, covariates). The dimension of the vector should be $2p + 1 + q$ , where $q$ denotes the number of covariates. |
| y | A $TT \times N$ time series object or a $TT \times N$ numerical matrix with the $N$ multivariate count time series over $TT$ time periods.  |

W	The $N \times N$ row-normalized non-negative adjacency matrix describing the network. The main diagonal entries of the matrix should be zeros, all the other entries should be non-negative and the maximum sum of elements over the rows should equal one. The function row-normalizes the matrix if a non-normalized adjacency matrix is provided.
p	The number of lags in the model.
d	The lag parameter of non-linear variable (should be between 1 and $p$ ).
Z	An $N \times q$ matrix of covariates (one for each column), where $q$ is the number of covariates in the model. Note that they must be non-negative.

### Details

The function computes the quasi score test for testing linearity of Poisson Network Autoregressive model of order  $p$  against the following ID-PNAR( $p$ ) model. For each node of the network  $i = 1, \dots, N$  over the time sample  $t = 1, \dots, TT$

$$\lambda_{i,t} = \frac{\beta_0}{(1 + X_{i,t-d})^\gamma} + \sum_{h=1}^p (\beta_{1h} X_{i,t-h} + \beta_{2h} Y_{i,t-h}) + \sum_{l=1}^q \delta_l Z_{i,l}$$

where  $X_{i,t} = \sum_{j=1}^N W_{ij} Y_{j,t}$  is the network effect, i.e. the weighted average impact of node  $i$  connections, with the weights of the mean being  $W_{ij}$ , the single element of the network matrix  $W$ . The sequence  $\lambda_{i,t}$  is the expectation of  $Y_{i,t}$  conditional to its past values.

The null hypothesis of the test is defined as  $H_0 : \gamma = 0$ , versus the alternative  $H_1 : \gamma > 0$ . The test statistic has the form

$$LM = S'(\hat{\theta}) \Sigma^{-1}(\hat{\theta}) S(\hat{\theta}),$$

where

$$S(\hat{\theta}) = \sum_{t=1}^{TT} \sum_{i=1}^N \left( \frac{Y_{i,t}}{\lambda_{i,t}(\hat{\theta})} - 1 \right) \frac{\partial \lambda_{i,t}(\hat{\theta})}{\partial \gamma}$$

is the partition of the quasi score related to the non-linear parameter  $\gamma$ , evaluated at the estimated parameters  $\hat{\theta}$  under the null assumption  $H_0$  (linear model), and  $\Sigma(\hat{\theta})$  is the variance of  $S(\hat{\theta})$ . Under  $H_0$ , the test asymptotically follows the  $\chi^2$  distribution with 1 degree of freedom. For details see Armillotta and Fokianos (2023, Sec. 4).

### Value

A list with attribute class "htest" including:

statistic	The value of the $\chi^2$ test statistic.
parameter	The degrees of freedom of the $\chi^2$ distribution. This is always 1.
p.value	The p-value of the $\chi^2$ test statistic.
null.value	The value of the $\gamma$ parameter, which is equal to 0 under the null hypothesis.
alternative	The alternative hypothesis, $\gamma$ has to be greater than 0.
method	The name of the test.
data.name	Information on the arguments used.

Alternatively, these can be printed via the function [summary.nonlin](#).

**Author(s)**

Mirko Armillotta, Michail Tsagris and Konstantinos Fokianos.

**References**

Armillotta, M. and K. Fokianos (2023). Nonlinear network autoregression. *Annals of Statistics*, 51(6): 2526–2552.

Armillotta, M. and K. Fokianos (2024). Count network autoregression. *Journal of Time Series Analysis*, 45(4): 584–612.

Armillotta, M., Tsagris, M. and Fokianos, K. (2024). Inference for Network Count Time Series with the R Package PNAR. *The R Journal*, 15/4: 255–269.

**See Also**

[score\\_test\\_stnarpq\\_j](#), [score\\_test\\_tnarpq\\_j](#), [lin\\_estimnarpq](#)

**Examples**

```
data(crime)
data(crime_W)
mod1 <- lin_estimnarpq(crime, crime_W, p = 2)
ca <- mod1$coefs[, 1]
score_test_nonlinpq_h0(ca, crime, crime_W, p = 2, d = 1)
```

---

score\_test\_stnarpq\_DV    *Bound p-value for testing for smooth transition effects on PNAR( $p$ ) model*

---

**Description**

Computation of Davies bound p-value for the sup-type test for testing linearity of Poisson Network Autoregressive model of order  $p$  (PNAR( $p$ )) versus the non-linear Smooth Transition alternative (ST-PNAR( $p$ )).

**Usage**

```
score_test_stnarpq_DV(b, y, W, p, d, Z = NULL, gama_L = NULL,
  gama_U = NULL, len = 100)
```

**Arguments**

- |   |  |
|---|--|
| b | The estimated parameters from the linear model, in the following order: (intercept, network parameters, autoregressive parameters, covariates). The dimension of the vector should be $2p + 1 + q$ , where $q$ denotes the number of covariates. |
| y | A $TT \times N$ time series object or a $TT \times N$ numerical matrix with the $N$ multivariate count time series over $TT$ time periods.   |

W	The $N \times N$ row-normalized non-negative adjacency matrix describing the network. The main diagonal entries of the matrix should be zeros, all the other entries should be non-negative and the maximum sum of elements over the rows should equal one. The function row-normalizes the matrix if a non-normalized adjacency matrix is provided.
p	The number of lags in the model.
d	The lag parameter of non-linear variable (should be between 1 and $p$ ).
Z	An $N \times q$ matrix of covariates (one for each column), where $q$ is the number of covariates in the model. Note that they must be non-negative.
gama_L	The lower value of the nuisance parameter $\gamma$ to consider. Use NULL if there is not information about its value. See the details for default computation.
gama_U	The upper value of the nuisance parameter $\gamma$ to consider. Use NULL if there is not information about its value. See the details for default computation.
len	The length of the grid of values of $\gamma$ values to consider.

### Details

The function computes an upper-bound for the p-value of the sup-type test for testing linearity of Poisson Network Autoregressive model of order  $p$  (PNAR( $p$ )) versus the following Smooth Transition alternative (ST-PNAR( $p$ )). For each node of the network  $i = 1, \dots, N$  over the time sample  $t = 1, \dots, TT$

$$\lambda_{i,t} = \beta_0 + \sum_{h=1}^p (\beta_{1h} X_{i,t-h} + \beta_{2h} Y_{i,t-h} + \alpha_h e^{-\gamma X_{i,t-h}^2} X_{i,t-h}) + \sum_{l=1}^q \delta_l Z_{i,l}$$

where  $X_{i,t} = \sum_{j=1}^N W_{ij} Y_{j,t}$  is the network effect, i.e. the weighted average impact of node  $i$  connections, with the weights of the mean being  $W_{ij}$ , the single element of the network matrix  $W$ . The sequence  $\lambda_{i,t}$  is the expectation of  $Y_{i,t}$ , conditional to its past values.

The null hypothesis of the test is defined as  $H_0 : \alpha_1 = \dots = \alpha_p = 0$ , versus the alternative that at least one among  $\alpha_h$  is not 0. The test statistic has the form

$$LM(\gamma) = S'(\hat{\theta}, \gamma) \Sigma^{-1}(\hat{\theta}, \gamma) S(\hat{\theta}, \gamma),$$

where

$$S(\hat{\theta}, \gamma) = \sum_{t=1}^{TT} \sum_{i=1}^N \left( \frac{Y_{i,t}}{\lambda_{i,t}(\hat{\theta}, \gamma)} - 1 \right) \frac{\partial \lambda_{i,t}(\hat{\theta}, \gamma)}{\partial \alpha}$$

is the partition of the quasi score related to the vector of non-linear parameters  $\alpha = (\alpha_1, \dots, \alpha_p)$ , evaluated at the estimated parameters  $\hat{\theta}$  under the null assumption  $H_0$  (linear model), and  $\Sigma(\hat{\theta}, \gamma)$  is the variance of  $S(\hat{\theta}, \gamma)$ . Since the test statistic depends on an unknown nuisance parameter ( $\gamma$ ), the supremum of the statistic is considered in the test,  $\sup_{\gamma} LM(\gamma)$ . The function computes the bound of the p-value, suggested by Davies (1987), for the test statistic  $\sup_{\gamma} LM(\gamma)$ , with scalar nuisance parameter  $\gamma$ , as follows.

$$P(\chi_k^2 \geq M) + VM^{1/2(k-1)} \frac{e^{-M/2} 2^{-k/2}}{\Gamma(k/2)}$$

where  $M$  is the maximum of the test statistic  $LM(\gamma)$ , computed by the available sample, over a grid of values for the nuisance parameter  $\gamma_F = (\gamma_L, \gamma_1, \dots, \gamma_l, \gamma_U)$ ;  $k$  is the number of non-linear parameters tested. So the first summand of the bound is just the p-value of a chi-square test with  $k$  degrees of freedom. The second summand is a correction term depending on  $V$ , which is the approximated total variation computed as

$$V = |LM^{1/2}(\gamma_1) - LM^{1/2}(\gamma_L)| + |LM^{1/2}(\gamma_2) - LM^{1/2}(\gamma_1)| + \dots + |LM^{1/2}(\gamma_U) - LM^{1/2}(\gamma_l)|.$$

The feasible bound allows to approximate the p-values of the sup-type test in a straightforward way, by adding to the tail probability of a chi-square distribution a correction term which depends on the total variation of the process. For details see Armillotta and Fokianos (2023, Sec. 5).

The values of `gama_L` and `gama_U` are computed internally as  $\text{gama\_L} = -\log(0.9)/X^2$  and  $\text{gama\_U} = -\log(0.1)/X^2$ , where  $X$  is the overall mean of  $X_{i,t}$  over the nodes  $i = 1, \dots, N$  and times  $t = 1, \dots, TT$ . Since the non-linear function  $e^{-\gamma X_{i,t}^2}$  ranges between 0 and 1, by considering  $X$  to be a representative value for the network mean, `gama_U` and `gama_L` would be the values of  $\gamma$  leading the non-linear switching function to be 0.1 and 0.9, respectively, so that in the optimization procedure the extremes of the function domain are excluded. Alternatively, their values can be supplied by the user.

## Value

A list including:

DV	The Davies bound of p-values for sup test.
supLM	The value of the sup test statistic in the sample $y$ .

## Author(s)

Mirko Armillotta, Michail Tsagris and Konstantinos Fokianos.

## References

- Armillotta, M. and K. Fokianos (2023). Nonlinear network autoregression. *Annals of Statistics*, 51(6): 2526–2552.
- Armillotta, M. and K. Fokianos (2024). Count network autoregression. *Journal of Time Series Analysis*, 45(4): 584–612.
- Armillotta, M., Tsagris, M. and Fokianos, K. (2024). Inference for Network Count Time Series with the R Package PNAR. *The R Journal*, 15/4: 255–269.
- Davies, R. B. (1987). Hypothesis testing when a nuisance parameter is present only under the alternative. *Biometrika* 74, 33–43.

## See Also

[score\\_test\\_stnarpq\\_j](#), [global\\_optimise\\_LM\\_stnarpq](#)

**Examples**

```
data(crime)
data(crime_W)
mod1 <- lin_estimnarpq(crime, crime_W, p = 1)
ca <- mod1$coefs[, 1]
score_test_stnarpq_DV(ca, crime, crime_W, p = 1, d = 1)
```

---

score\_test\_stnarpq\_j    *Bootstrap test for smooth transition effects on PNAR(p) model*

---

**Description**

Computation of bootstrap p-value for the sup-type test for testing linearity of Poisson Network Autoregressive model of order  $p$  (PNAR( $p$ )) versus the non-linear Smooth Transition alternative (ST-PNAR( $p$ )).

**Usage**

```
score_test_stnarpq_j(supLM, b, y, W, p, d, Z = NULL, J = 499,
  gama_L = NULL, gama_U = NULL, tol = 1e-9, ncores = 1, seed = NULL)
```

**Arguments**

supLM	The optimized value of the test statistic. See the function <a href="#">global_optimise_LM_stnarpq</a> .
b	The estimated parameters from the linear model, in the following order: (intercept, network parameters, autoregressive parameters, covariates). The dimension of the vector should be $2p + 1 + q$ , where $q$ denotes the number of covariates.
y	A $TT \times N$ time series object or a $TT \times N$ numerical matrix with the $N$ multi-variate count time series over $TT$ time periods.
W	The $N \times N$ row-normalized non-negative adjacency matrix describing the network. The main diagonal entries of the matrix should be zeros, all the other entries should be non-negative and the maximum sum of elements over the rows should equal one. The function row-normalizes the matrix if a non-normalized adjacency matrix is provided.
p	The number of lags in the model.
d	The lag parameter of non-linear variable (should be between 1 and $p$ ).
Z	An $N \times q$ matrix of covariates (one for each column), where $q$ is the number of covariates in the model. Note that they must be non-negative.
J	The number of bootstrap samples to draw.
gama_L	The lower value of the nuisance parameter $\gamma$ to consider. Use NULL if there is not information about its value. See the details for default computation.
gama_U	The upper value of the nuisance parameter $\gamma$ to consider. Use NULL if there is not information about its value. See the details for default computation.

tol	Tolerance level for the optimizer.
ncores	Number of cores to use for parallel computing. By default the number of cores is set to 1 (no parallel computing). <b>Note:</b> If for some reason the parallel does not work then load the doParallel package yourselves.
seed	To replicate the results use a seed for the generator, an integer number.

## Details

The function computes a bootstrap p-value for the sup-type test for testing linearity of Poisson Network Autoregressive model of order  $p$  (PNAR( $p$ )) versus the following Smooth Transition alternative (ST-PNAR( $p$ )). For each node of the network  $i = 1, \dots, N$  over the time sample  $t = 1, \dots, TT$

$$\lambda_{i,t} = \beta_0 + \sum_{h=1}^p (\beta_{1h} X_{i,t-h} + \beta_{2h} Y_{i,t-h} + \alpha_h e^{-\gamma X_{i,t-h}^2} X_{i,t-h}) + \sum_{l=1}^q \delta_l Z_{i,l},$$

where  $X_{i,t} = \sum_{j=1}^N W_{ij} Y_{j,t}$  is the network effect, i.e. the weighted average impact of node  $i$  connections, with the weights of the mean being  $W_{ij}$ , the single element of the network matrix  $W$ . The sequence  $\lambda_{i,t}$  is the expectation of  $Y_{i,t}$ , conditional to its past values.

The null hypothesis of the test is defined as  $H_0 : \alpha_1 = \dots = \alpha_p = 0$ , versus the alternative that at least one among  $\alpha_h$  is not 0. The test statistic has the form

$$LM(\gamma) = S'(\hat{\theta}, \gamma) \Sigma^{-1}(\hat{\theta}, \gamma) S(\hat{\theta}, \gamma)$$

where

$$S(\hat{\theta}, \gamma) = \sum_{t=1}^{TT} \sum_{i=1}^N \left( \frac{Y_{i,t}}{\lambda_{i,t}(\hat{\theta}, \gamma)} - 1 \right) \frac{\partial \lambda_{i,t}(\hat{\theta}, \gamma)}{\partial \alpha}$$

is the partition of the quasi score related to the vector of non-linear parameters  $\alpha = (\alpha_1, \dots, \alpha_p)$ , evaluated at the estimated parameters  $\hat{\theta}$  under the null assumption  $H_0$  (linear model), and  $\Sigma(\hat{\theta}, \gamma)$  is the variance of  $S(\hat{\theta}, \gamma)$ .

Since the test statistic depends on an unknown nuisance parameter ( $\gamma$ ), the supremum of the statistic is considered in the test,  $\sup_{\gamma} LM(\gamma)$ . This value can be computed for the available sample by using the function `global_optimise_LM_stnarpq` and should be supplied here as an input `supLM`.

The function performs the bootstrap resampling of the test statistic  $\sup_{\gamma} LM(\gamma)$  by employing Gaussian perturbations of the score  $S(\hat{\theta}, \gamma)$ . For details see Armillotta and Fokianos (2023, Sec. 5).

The values of `gama_L` and `gama_U` are computed internally as `gama_L = -log(0.9)/X^2` and `gama_U = -log(0.1)/X^2`, where  $X$  is the overall mean of  $X_{i,t}$  over the nodes  $i = 1, \dots, N$  and times  $t = 1, \dots, TT$ . Since the non-linear function  $e^{-\gamma X_{i,t-h}^2}$  ranges between 0 and 1, by considering  $X$  to be a representative value for the network mean, `gama_U` and `gama_L` would be the values of  $\gamma$  leading the non-linear switching function to be 0.1 and 0.9, respectively, so that in the optimization procedure the extremes of the function domain are excluded. Alternatively, their value can be supplied by the user.

**Note:** For large datasets the function may require few minutes to run. Parallel computing is suggested to speed up the computations.



**Value**

A list including:

pJ	The bootstrap p-value of the sup test.
cpJ	The adjusted version of bootstrap p-value of the sup test.
gamaj	The optimal values of the $\gamma$ parameter for score test bootstrap replications.
supLMj	The values of perturbed test statistic at the optimum point gamaj.

**Author(s)**

Mirko Armillotta, Michail Tsagris and Konstantinos Fokianos.

**References**

- Armillotta, M. and K. Fokianos (2023). Nonlinear network autoregression. *Annals of Statistics*, 51(6): 2526–2552.
- Armillotta, M. and K. Fokianos (2024). Count network autoregression. *Journal of Time Series Analysis*, 45(4): 584–612.
- Armillotta, M., Tsagris, M. and Fokianos, K. (2024). Inference for Network Count Time Series with the R Package PNAR. *The R Journal*, 15/4: 255–269.

**See Also**

[score\\_test\\_stnarpq\\_DV](#), [global\\_optimise\\_LM\\_stnarpq](#), [score\\_test\\_tnarpq\\_j](#)

**Examples**

```
# load data
data(crime)
data(crime_W)
#estimate linear PNAR model
mod1 <- lin_estimnarpq(crime, crime_W, p = 2)
b <- mod1$coefs[, 1]

g <- global_optimise_LM_stnarpq(b = b, y = crime, W = crime_W, p = 2, d = 1)
supg <- g$supLM
score_test_stnarpq_j(supLM = supg, b = b, y = crime, W = crime_W, p = 2, d = 1, J = 5)
```

---

score\_test\_tnarpq\_j      *Bootstrap test for threshold effects on PNAR(p) model*

---

**Description**

Computation of bootstrap p-value for the sup-type test for testing linearity of Poisson Network Autoregressive model of order  $p$  (PNAR( $p$ )) versus the non-linear Threshold alternative (T-PNAR( $p$ )).

**Usage**

```
score_test_tnarpq_j(supLM, b, y, W, p, d, Z = NULL, J = 499,
  gama_L = NULL, gama_U = NULL, tol = 1e-9, ncores = 1, seed = NULL)
```

**Arguments**

supLM	The optimized value of the test statistic. See the function <a href="#">global_optimise_LM_tnarpq</a> .
b	The estimated parameters from the linear model, in the following order: (intercept, network parameters, autoregressive parameters, covariates). The dimension of the vector should be $2p + 1 + q$ , where $q$ denotes the number of covariates.
y	A $TT \times N$ time series object or a $TT \times N$ numerical matrix with the $N$ multivariate count time series over $TT$ time periods.
W	The $N \times N$ row-normalized non-negative adjacency matrix describing the network. The main diagonal entries of the matrix should be zeros, all the other entries should be non-negative and the maximum sum of elements over the rows should equal one. The function row-normalizes the matrix if a non-normalized adjacency matrix is provided.
p	The number of lags in the model.
d	The lag parameter of non-linear variable (should be between 1 and $p$ ).
Z	An $N \times q$ matrix of covariates (one for each column), where $q$ is the number of covariates in the model. Note that they must be non-negative.
J	The number of bootstrap samples to draw.
gama_L	The lower value of the nuisance parameter $\gamma$ to consider. Use NULL if there is not information about its value. See the details for default computation.
gama_U	The upper value of the nuisance parameter $\gamma$ to consider. Use NULL if there is not information about its value. See the details for default computation.
tol	Tolerance level for the optimizer.
ncores	Number of cores to use for parallel computing. By default the number of cores is set to 1 (no parallel computing). <b>Note:</b> If for some reason the parallel does not work then load the doParallel package yourselves.
seed	To replicate the results use a seed for the generator, an integer number.

**Details**

The function computes a bootstrap p-value for the sup-type test for testing linearity of Poisson Network Autoregressive model of order  $p$  (PNAR( $p$ )) versus the following Threshold alternative (T-PNAR( $p$ )). For each node of the network  $i = 1, \dots, N$  over the time sample  $t = 1, \dots, TT$

$$\lambda_{i,t} = \beta_0 + \sum_{h=1}^p [\beta_{1h}X_{i,t-h} + \beta_{2h}Y_{i,t-h} + (\alpha_0 + \alpha_{1h}X_{i,t-h} + \alpha_{2h}Y_{i,t-h})I(X_{i,t-d} \leq \gamma)] + \sum_{l=1}^q \delta_l Z_{i,l}$$

where  $X_{i,t} = \sum_{j=1}^N W_{ij}Y_{j,t}$  is the network effect, i.e. the weighted average impact of node  $i$  connections, with the weights of the mean being  $W_{ij}$ , the single element of the network matrix  $W$ . The sequence  $\lambda_{i,t}$  is the expectation of  $Y_{i,t}$ , conditional to its past values.

The null hypothesis of the test is defined as  $H_0 : \alpha_0 = \alpha_{11} = \dots = \alpha_{2p} = 0$ , versus the alternative that at least one among  $\alpha_{s,h}$  is not 0, for  $s = 0, 1, 2$ . The test statistic has the form

$$LM(\gamma) = S'(\hat{\theta}, \gamma) \Sigma^{-1}(\hat{\theta}, \gamma) S(\hat{\theta}, \gamma)$$

where

$$S(\hat{\theta}, \gamma) = \sum_{t=1}^{TT} \sum_{i=1}^N \left( \frac{Y_{i,t}}{\lambda_{i,t}(\hat{\theta}, \gamma)} - 1 \right) \frac{\partial \lambda_{i,t}(\hat{\theta}, \gamma)}{\partial \alpha}$$

is the partition of the quasi score related to the vector of non-linear parameters  $\alpha = (\alpha_0, \dots, \alpha_{2p})$ , evaluated at the estimated parameters  $\hat{\theta}$  under the null assumption  $H_0$  (linear model), and  $\Sigma(\hat{\theta}, \gamma)$  is the variance of  $S(\hat{\theta}, \gamma)$ .

Since the test statistic depends on an unknown nuisance parameter ( $\gamma$ ), the supremum of the statistic is considered in the test,  $\sup_{\gamma} LM(\gamma)$ . This value can be computed for the available sample by using the function `global_optimise_LM_tnarpq` and should be supplied here as an input `supLM`.

The function performs the bootstrap resampling of the test statistic  $\sup_{\gamma} LM(\gamma)$  by employing Gaussian perturbations of the score  $S(\hat{\theta}, \gamma)$ . For details see Armillotta and Fokianos (2023, Sec. 5).

The values of `gama_L` and `gama_U` are computed internally as the mean over  $i = 1, \dots, N$  of 20% and 80% quantiles of the empirical distribution of the network mean  $X_{i,t}$  for  $t = 1, \dots, TT$ . In this way the optimization is performed for values of  $\gamma$  such that the indicator function  $I(X_{i,t-d} \leq \gamma)$  is not always close to 0 or 1. Alternatively, their value can be supplied by the user. For details see Armillotta and Fokianos (2023, Sec. 4-5).

**Note:** For large datasets the function may require few minutes to run. Parallel computing is suggested to speed up the computations.

## Value

A list including:

<code>pJ</code>	The bootstrap p-value of the sup test.
<code>cpJ</code>	The adjusted version of bootstrap p-value of the sup test.
<code>gama_j</code>	The optimal values of the $\gamma$ parameter for score test bootstrap replications.
<code>supLMj</code>	The values of perturbed test statistic at the optimum point <code>gama_j</code> .

## Author(s)

Mirko Armillotta, Michail Tsagris and Konstantinos Fokianos.

## References

- Armillotta, M. and K. Fokianos (2023). Nonlinear network autoregression. *Annals of Statistics*, 51(6): 2526–2552.
- Armillotta, M. and K. Fokianos (2024). Count network autoregression. *Journal of Time Series Analysis*, 45(4): 584–612.
- Armillotta, M., Tsagris, M. and Fokianos, K. (2024). Inference for Network Count Time Series with the R Package PNAR. *The R Journal*, 15/4: 255–269.

**See Also**

[global\\_optimise\\_LM\\_tnarpq](#), [global\\_optimise\\_LM\\_stnarpq](#), [score\\_test\\_stnarpq\\_j](#)

**Examples**

```
# load data
data(crime)
data(crime_W)
#estimate linear PNAR model
mod1 <- lin_estimnarpq(crime, crime_W, p = 2)
b <- mod1$coefs[, 1]

g <- global_optimise_LM_tnarpq(b = b, y = crime, W = crime_W, p = 2, d = 1)
supg <- g$supLM
score_test_tnarpq_j(supLM = supg, b = b, y = crime, W = crime_W, p = 2, d = 1, J = 5)
```

---

summary.DV

*S3 methods for extracting the results of the bound p-value for testing for smooth transition effects on PNAR(p) model*

---

**Description**

S3 methods for extracting the results of the bound p-value for testing for smooth transition effects on PNAR( $p$ ) model.

**Usage**

```
## S3 method for class 'DV'
summary(object, ...)
## S3 method for class 'summary.DV'
print(x, ...)
## S3 method for class 'DV'
print(x, ...)
```

**Arguments**

object	An object containing the results of the function <a href="#">score_test_stnarpq_DV</a> .
x	An object containing the results of the function <a href="#">score_test_stnarpq_DV</a> .
...	Extra arguments the user can pass.

**Details**

The functions print the output of the bound p-value for testing for smooth transition effects on PNAR( $p$ ) model.

**Value**

The functions print the results of the function [score\\_test\\_stnarpq\\_DV](#).

**Author(s)**

Mirko Armillotta, Michail Tsagris and Konstantinos Fokianos.

**References**

Armillotta, M., Tsagris, M. and Fokianos, K. (2024). Inference for Network Count Time Series with the R Package PNAR. *The R Journal*, 15/4: 255–269.

Davies, R. B. (1987). Hypothesis testing when a nuisance parameter is present only under the alternative. *Biometrika* 74, 33–43.

**See Also**

[score\\_test\\_stnarpq\\_DV](#)

**Examples**

```
data(crime)
data(crime_W)
mod1 <- lin_estimnarpq(crime, crime_W, p = 1)
ca <- mod1$coefs[, 1]
a <- score_test_stnarpq_DV(ca, crime, crime_W, p = 1, d = 1)
print(a)
summary(a)
```

---

summary.nonlin

*S3 methods for extracting the results of the non-linear hypothesis test*


---

**Description**

S3 methods for extracting the results of the non-linear hypothesis test.

**Usage**

```
## S3 method for class 'nonlin'
summary(object, ...)
## S3 method for class 'summary.nonlin'
print(x, ...)
## S3 method for class 'nonlin'
print(x, ...)
```

**Arguments**

object	An object containing the results of the function <a href="#">score_test_nonlinpq_h0</a> .
x	An object containing the results of the function <a href="#">score_test_nonlinpq_h0</a> .
...	Extra arguments the user can pass.

**Details**

The functions print the output of the non-linear hypothesis test.

**Value**

The functions print the results of the function `score_test_nonlinpq_h0`.

**Author(s)**

Mirko Armillotta, Michail Tsagris and Konstantinos Fokianos.

**References**

Armillotta, M., Tsagris, M. and Fokianos, K. (2024). Inference for Network Count Time Series with the R Package PNAR. *The R Journal*, 15/4: 255–269.

**See Also**

`score_test_nonlinpq_h0`

**Examples**

```
data(crime)
data(crime_W)
mod1 <- lin_estimnapq(crime, crime_W, p = 2)
ca <- mod1$coefs[, 1]
a <- score_test_nonlinpq_h0(ca, crime, crime_W, p = 2, d = 1)
print(a)
summary(a)
```

---

summary.PNAR

*S3 methods for extracting the results of the estimation functions*


---

**Description**

S3 methods for extracting the results of the estimation functions.

**Usage**

```
## S3 method for class 'PNAR'
summary(object, ...)
## S3 method for class 'summary.PNAR'
print(x, ...)
## S3 method for class 'PNAR'
print(x, ...)
```

**Arguments**

object	An object containing the results of the estimation function <a href="#">lin_estimnarpq</a> or <a href="#">log_lin_estimnarpq</a> .
x	An object containing the results of the estimation function <a href="#">lin_estimnarpq</a> or <a href="#">log_lin_estimnarpq</a> .
...	Extra arguments the user can pass.

**Details**

These functions print the output of the estimation functions.

**Value**

The print.PNAR() function prints the coefficients of the model. The summary.PNAR() function prints the output in the lm() style.

**Author(s)**

Mirko Armillotta, Michail Tsagris and Konstantinos Fokianos.

**References**

Armillotta, M., Tsagris, M. and Fokianos, K. (2024). Inference for Network Count Time Series with the R Package PNAR. *The R Journal*, 15/4: 255–269.

**See Also**

[log\\_lin\\_estimnarpq](#)

**Examples**

```
data(crime)
data(crime_W)
mod1 <- lin_estimnarpq(crime, crime_W, p = 2)
mod1
print(mod1)
summary(mod1)
```

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